REVIEW ARTICLE



Effects of maternal folate and vitamin B12 on gestational diabetes mellitus: a dose-response meta-analysis of observational studies

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To comprehensively estimate the association of gestational diabetes mellitus (GDM) risk with maternal red blood cell (RBC) folate, plasma/serum folate, dose and duration of folic acid supplement (FAS) intake and vitamin B12 separately. PubMed, Web of science, CNKI, and Wanfang Databases were searched through March 26, 2021. We synthesized data using random-effects model meta-analysis in Stata 12.0. Sensitivity, subgroup and dose-response analyses were also performed. The certainty of evidence was evaluated using the Grading of Recommendations, Assessment, Development and Evaluations (GRADE). Twenty six datasets from thirteen eligible observational studies were included in the study. We found a significant increase of GDM risk with the highest versus lowest category of RBC folate (OR = 1.96, 95% Cl: 1.48–2.61, I2 = 0.0%, moderate-certainty evidence) and plasma/serum folate (OR = 1.23, 1.02–1.48, I2 = 57.8%, low-certainty evidence). The dose-response analysis revealed that each 200 ng/ml increase in RBC folate was significantly associated with 8% higher GDM risk. No significant association between dose of FAS intake and GDM risk was found with very low cetainty. Meanwhile, longer duration (\geq 3 months) of FAS conferred 56% significant higher GDM risk (OR = 1.56, 1.02–2.39, very low certainty evidence). No significant association of GDM risk with highest plasma/serum B12 was observed compared to lowest B12 (OR = 0.77, 0.58–1.02, very low-certainty evidence). Moderate-certainty evidence suggests that higher RBC folate appears to significantly increase GDM risk. Higher plasma/serum folate may increase GDM risk but with low certainty. Further well-designed trials or prospective studies are needed.

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INTRODUCTION

Gestational diabetes mellitus (GDM) is a condition in which glucose intolerance is of variable severity with onset or first detected during pregnancy resulting in varying degrees of hyperglycemia [1]. The global prevalence of GDM has been in climbing trend and estimated to be 16.9% worldwide in women (20–49 years) varying with population characteristics and the diagnostic criteria used [2, 3]. The pathogenesis of GDM may include genetic susceptibility, exacerbated insulin resistance, chronic inflammatory response, oxidative stress and other aspects [4]. GDM is of relevance to elevated adverse perinatal complications and cardiometabolic risk such as obesity, type 2 diabetes hypertension, hyperlipidemia and lifetime cardiovascular disease in GDM women and their offspring [5–7].

A number of factors associated with GDM risk have been implicated, including obesity, excess gestational weight gain, family history of diabetes, physical inactivity, advanced maternal age and genetic factors [8]. It is a critical and urgent issue to identify more modifiable risk factors and promote GDM prevention from a public health and maternal and child health perspective.

Folate and vitamin B12 are water-soluble vitamins and could not be synthesized by the body itself and must be obtained from food. In the daily diet, folate is mainly from green leafy vegetables, fruits, egg yolks, and legumes, while vitamin B12 comes from animal products. Folate and vitamin B12 are key cofactors in one carbon metabolic pathway, which gets involved in DNA methylation and cell metabolism via generating precursors of nucleotide biosynthesis and methyl groups for methylation reactions [9, 10]. With the rising of maternal blood volume and renal blood flow, requirement of folate and vitamin B12 is dramatically elevated during pregnancy and the risk of inadequate intake also arises [11].

Since the early 1990s, folic acid supplement (FAS) has been recommended globally prior to and during the first trimester of pregnancy to prevent neural tube defects (NTDs) [12]. As many women continue to take folic acid supplements beyond the first trimester, the combined effects of supplementation and fortification have resulted in substantial increases in folate concentrations among women of reproductive age [13]. However, the GDM prevalence has been rising continuously. Emerging epidemiological studies reveal that an imbalance of high folate status or intake and low vitamin B12 status during pregnancy predisposes women

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to elevated GDM risk and their offspring to insulin resistance and adiposity and low birthweight [10, 14]. Chen et al. [15] reported that higher maternal RBC folate and vitamin B12 levels in early pregnancy are significantly associated with GDM risk, while the balance of folate/vitamin B12 is not significantly associated with GDM. Another two studies also showed that higher maternal RBC folate concentrations increase the GDM risk in Chinese population [16, 17]. However, there are inconsistent findings regarding the association of folate status and risk of GDM. Those results are difficult to compare due to the varied dosage and duration of FAS and tested in serum or plasma [18-23]. In addition, to our knowledge, only one systematic review included two studies for quantitative analysis and examined the influence of vitamin B12 insufficiency on GDM, however, the studies had small sample sizes and did not exclude the potential confounders [24]. Moreover, the relationship between vitamin B12 and GDM is still unclear and recent studies have conflicting results [15, 25, 26].

Therefore, we conducted a meta-analysis with the following aims: (1) to systematically benchmark current knowledge of the associations of RBC and plasma/serum folate, dose and duration of FAS, vitamin B12 with GDM; (2) to further explore the potential dose-response relationship if sufficient data available; and (3) to assess the robustness of the findings by performing sensitivity analyses and subgroup analyses based on study characteristics.

METHODS

We have used the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) checklist [27].

Search strategy

We comprehensively searched relevant literatures in PubMed, Web of Science, China National Knowledge Infrastructure (CNKI), and Wanfang Databases from inception to March 26, 2021. The following search string was used in PubMed: ("folic acid" [MeSH Terms] OR (folic AND acid) OR folate OR (vitamin AND (B-9 OR B9))) OR ((vitamin AND (B-12 OR B12)) OR cobalamin OR cobalamin [MeSH Terms])) AND ("diabetes, gestational" [MeSH Terms] OR (gestational AND diabetes)). The search string for Web of Science: ((ALL = folate OR ALL = "folic acid" OR TS = "folic acid" OR (ALL = folic AND ALL = acid)) OR (ALL = vitamin AND (ALL = B-9 OR ALL = B9)) OR (ALL = vitamin AND (ALL = B-12 OR ALL = B12)) OR ALL = cobalamin) AND (ALL = gestational AND ALL = diabetes). Chinese keywords were used in China National Knowledge Infrastructure (CNKI), and Wanfang Databases, if the articles were in Chinese. Additional articles were also manually searched in bibliographies of retrieved articles and reviews.

Study selection

Inclusion criteria were as following: (1) Reported intake of folate status or vitamin B12 as exposure and in at least two categories; (2) Provided sufficient data for calculation of their effect on GDM. (3) Cohort, case-control and cross-sectional studies were eligible.

Studies were excluded if: (1) Exposure of folate status or vitamin B12 was not evaluated or not directly linked to GDM; (2) Studies were not conducted in population. (3) Reviews, guideline, comment, case reports; (4) Studies were not written in English or Chinese; (5) Duplicate studies or those based on the overlapping data sets, of which those with the largest sample size and most detailed results were included preferentially.

NoteExpress version 3.0.4.6732 was used to merge retrieved citations and eliminate duplications.

Data extraction

Two independent investigators (NN Li, JC Jiang) extracted the following information from eligible studies: tittle, first author, year of publication, study design, country, number of participants/GDM, mean age and or range, exposure of folate status or B12,

criteria of GDM assessment and date, effect size, and covariates adjusted. If a study reported several estimates with adjustment for different confounders, results were included in the analyses for the one adjusting for the largest number of covariates. If any data were not available, we attempted to contact the authors for help. Any disagreement were resolved by reaching a consensus with a third independent reviewer (LL Guo).

Quality assessment

The methodological quality of the enrolled cohort/case-control studies were assessed using Newcastle–Ottawa Scale (NOS) [28], and the cross sectional studies using Agency for Healthcare Research and Quality (USA) (AHRQ) [29]. The maximum score of NOS and AHRQ were 9 and 11, respectively. For NOS, score of \leq 6, 7–8, and 9 represent low, medium, and high quality, respectively. For AHRQ, score of \leq 3, 4–7, and 8–11 represent low, medium, and high quality, respectively.

We also use the Grading of Recommendations Assessment, Development and Evaluation (GRADE) approach to rate the certainty of evidence [30]. The evidence were assessed in five domains for downgrade (risk of bias, imprecision, inconsistency, indirectness, publication bias) and three domains for upgrade (large magnitude of effect, dose-response gradient, all residual confounding would decrease magnitude of effect). We rated the certainty of the evidence as "very low," "low," "moderate," or "high" and prepared evidence profiles.

Statistical analysis

We used odds ratio (OR) with 95% confidence intervals (CI) to pool the association of each exposure with GDM. For the highest versus the lowest category meta-analysis, we adopted the DerSimonian and Laird (D&L) random effects model (if P for heterogeneity <0.05) and Mantel-Haenszel fix effects model (if P for heterogeneity >0.05) to calculate the adjusted ORs and 95% CI [31] using metan command in Stata. Meanwhile, publication bias was detected by funnel plot and evaluated using Egger's asymmetry test. Heterogeneity among different studies was tested by Cochrane chi-squared test and calculated by I-squared statistics (variation in effect size attributable to heterogeneity). I² of 25% or less was considered to be low, 26-50% to be moderate, 51-75% to be high, and 76% or more to be very high heterogeneity. To test the potential influence of each study on pooled effect size, "leaveone-out" sensitivity analyses were performed to evaluate the reliability of the results by excluding a single study at a time using metaninf command in STATA. Subgroup analyses were also conducted by number of participants, study design, geographical region, adjusting for history of DM and adjusting for intakes of other nutrients or not. As advised by the guidelines, if the number of pooled studies for specified exposure was not limited (less than ten) or of high heterogeneity, we would perform meta-regression analysis [32, 33].

We conducted dose-response analyses for the association of GDM with increased level of each exposure using a generalized least squares trend estimation [34], according to the methods recommended by Greenland and Longnecker. We extracted the mean level of exposure, the number of participants and cases, and the ORs with 95% CI in each category. If neither median nor mean values were reported, we used the midpoint of the range. If the lowest category was open-ended, the lowest boundary was set to zero. If the highest category was open-ended, we considered the width of that category to be the same as the width of the adjacent category [35]. We examined the potential nonlinear association by modeling each exposure levels using restricted cubic splines with three knots [36]. Non-linearity test was conducted using Waldtype test by testing the null hypothesis that the coefficient of the second spline was equal to zero.

All procedures of meta-analyses were carried out by using the relevant publicly available command in Stata version 12.0

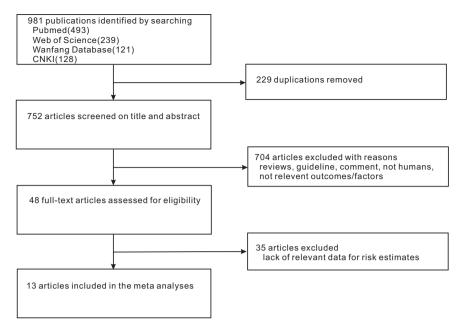


Fig. 1 Flow chart of literature search for eligible articles in the meta-analyses of maternal folate, vitamin B12 and GDM risk.

(StataCorp, College Station, TX, USA) with a two-sided $\it P < 0.05$ as statistically significant.

RESULTS

Study selection and characteristics

A detailed identification of eligible studies is outlined in Fig. 1. Of the 981 publications retrieved from PubMed, Web of Science, CNKI, and Wanfang Databases, We identified 48 articles by removing the replications and screening on the tittle and abstracts. After reviewing full-text articles, we eventually included 13 eligible studies in present meta-analyses, consisting of ten prospective cohort studies, one case control study and two cross sectional studies.

The characteristics of the 13 studies are presented in Table 1. These 13 individual publications contained 26 eligible datasets including 3 for RBC folate, 8 for plasma folate, 4 for FAS dose, 3 for FAS duration and 8 for plasma/serum vitamin B12. The number of participants involved in 13 studies ranged from 326 to 14553, with mean number of 2420. The incidence of GDM ranged from 3.25% to 18.31% in ten cohorts. The prevalence of GDM in two cross sectional studies were 17.96% and 22.17%. International Association of the Diabetes and Pregnancy Study Groups (IADPSG) criteria was used as the main method of GDM assessment by 7 of 13 eligible studies. The adjusted ORs were calculated by comparing the highest vs. the lowest level of each exposure adjusting for covariates such as maternal age or BMI plus other factors.

Based on the NOS and AHRQ, All studies were rated of medium to high quality, of which five studies were scored 9, six studies were scored 8 and two studies were scored 7 (Table 1).

RBC folate and GDM risk

As presented in the forest plot (Fig. 2), three eligible studies with no heterogeneity ($l^2=0.0\%$, P=0.401) were pooled and a OR of 1.96 (95% Cl: 1.48–2.61) was obtained for the risk of GDM comparing the highest versus the lowest category of RBC folate. Owing to dose-response gradient, we rated the certainty of evidence as moderate. The details of the evidence profile for GRADE rating were listed in Supplementary Table 1.

Furthermore, the sensitivity analyses also consistently revealed a significant positive association persisted with the exclusion of each study in turn and confirmed the stability of the results (pooled

estimates ranged between 1.91 and 2.34) (Supplementary Fig. 1a). As suggested in subgroup analyses (Table 2), all with no heterogeneity, increased risk of GDM in highest category of RBC folate was seen in subgroups of >1000 participants, <1000 participants, adjusting for history of DM or not, adjusting for intakes of other nutrients or not (pooled effect estimates ranged between 1.72 and 2.47). In addition, no significant differences were found in the risk of GDM in studies that assessed >1000 participants vs. <1000 participants, adjusting for history of DM vs. not, adjusting for intakes of other nutrients vs. not (all $P_{\rm between}$ > 0.05).

Plasma/serum folate and GDM risk

For the highest vs. the lowest category of plasma/serum folate, eight datasets from six studies with high heterogeneity ($I^2 = 57.8\%$, P = 0.020) were included in the random-effects meta-analysis (Fig. 2). A significantly positive association of plasma/serum folate with GDM risk were also observed (OR = 1.23 (95% CI: 1.02–1.48). The certainty of evidence was low (Supplementary Table 1).

Similarly, we evaluated the fluctuation of the overall risk of GDM regarding folate status by excluding one study at a time. When excluding the three datasets in turn [15, 23, 25], the association of GDM risk and plasma/serum folate turned to be insignificant, with wide confidence bounds (Supplementary Fig. 1b).

In the subgroup analyses (Table 2), we observed significantly increased risk of GDM with higher plasma/serum folate in the groups of studies with number of participants >1000, or with estimates adjusting for diabetes mellitus history, or with estimates adjusting for intakes of other nutrients, or the Asian population, or the cross-sectional studies. Test for heterogeneity between subgroups may be invalid due to moderate heterogeneity observed in one or more subgroups.

Dose and duration of FAS intake and GDM risk

Due to serious imprecision and inconsistency, very low certainty of evidence showed that women with higher intake of FAS demonstrated a greater risk of GDM but not statistically significant (OR = 1.50, 95% Cl: 0.72–3.11), with very high heterogeneity ($l^2 = 87.4\%$, P < 0.001). Regarding duration of FAS, as suggested by the meta-analysis of three datasets from two studies with moderate heterogeneity ($l^2 = 43.7\%$, P = 0.170), a longer duration (≥ 3 months) of FAS intake was associated with higher risk of GDM (OR = 1.56,

Quality score* 6 ∞ ∞ ∞ 6 œ ∞ Age, preconceptional BMI, gestational Age, preconceptional BMI, family history of weight gain, snoring, gravidity and parity ethnicity, primiparity, weight gain at OGTT and other of diabetes, previous Maternal age, parity, parity, family history energy, glycemic load, saturated fat, total fiber, and origin, smoking, gestation of bloods, occurrence of GDM, plasma B6 and B12. Adjusted covariables alcohol use, dietary prepregnancy BMI, history of diabetes, prepregnancy BMI, prepregnancy BMI, cigarette smoking, smoking, drinking, Maternal age, ethnicity (Chinese, education, income, Age, parity, ethnic diabetes, smoking exposure, and diabetes, physical Age, ethnicity, education, parity, monthly income, smoking, alcohol Age, race, parity, family history of supplement use drinking status. gestational BMI Age, education, fetal sex, family intake, physical intakes of total Malay, Indian), employment, enrollment. and BMI at serum B12, heme iron activity, activity, history week gestation 26–28 weeks' gestation 24–28 weeks' gestation 26–28 weeks' gestation 24–28 weeks' gestation 24–28 weeks' gestation 24-28 weeks' Period for GDM assessment Around 28gestation Ϋ́ The modified WHO 1999 criteria **GDM assessment** guidelines (1998) The WHO 1999 criteria management Self-report IADPSG IADPSG IADPSG IADPSG GDM Characteristics of 13 eligible studies included in the meta-analyses of maternal folate, vitamin B12 and GDM risk. From prepregnancy through midpregnancy 24-28 weeks' gestation 19-24 weeks' gestation 26-28 weeks' gestation 9-13 weeks' gestation Period for exposure measurement 26.9 ± 5.3 weeks' Prepregnancy First trimester gestation Serum folate, Serum B12 Plasma folate, Plasma B12 folate, FAS, Serum B12 **3BC** folate, RBC folate Exposure Serum FAS FAS FAS GDM:29.01 ± 3.15 non-GDM:27.89 ± Age (mean ± SD, range; years) 30.3 ± 5.88 30.2 ± 4.0 29.4 ± 4.5 25-44 18-46 26.7 3.18 ¥ 1938/249 2282/392 Women/ cases 1058/180 4353/374 344/143 913/164 14553/ 824 406/90 Study design CCS CSS CSS S S S S S Singapore Country China China China China China NSA ¥ Mengying Li, 2019 First author, year Kaipeng Xie, 2019 Shuying Li, 2019 Beibei Zhu, 2016 Nithya Sukumar, 2016 Jun S. Lai, 2017 Qian Li, 2019 Xiaotian Chen, 2021

Table 1.

Table 1. continued	Q Q									
First author, year	Country	Study design	Women/ cases	Age (mean ± SD, range; years)	Exposure	Period for exposure measurement	GDM assessment	Period for GDM assessment	Adjusted covariables	Quality score*
					Serum folate, Serum B12				prepregnancy BMI, and family history of diabetes. Mutual adjustments were then made for serum folate and vitamin B12.	
Pengju Liu, 2020	China	Ŋ	366/67	GDM:30.5 ± 4.0 non-GDM:28.9 ± 3.5	RBC folate	Before 12 weeks' gestation	IADPSG	24–28 weeks' gestation	Age, physical activity, BMI, parity, family history of diabetes, use of folic acid supplements, HOMA-IR, C-reactive protein, hemoglobin, vitamin B12, and serum homocysteine.	٥
Tanja Jankovic- Karasoulos, 2020	Australia	SS	3196/104	GDM: 28.9±5.2 non-GDM:27.9± 5.9	Serum folate, Serum B12	15±1 weeks' gestation	The new WHO 2016 criteria	Around 26- week gestation	Maternal age, BMI and smoking status at 15 weeks	8
Krishnaveni, 2009	India	೮	774/49	23	Plasma folate, Plasma B12	Before 30 weeks' gestation	Carpenter- Coustan criteria	30±2 weeks' gestation	Age, BMI, socioeconomic status, religion, parity and family history of diabetes	7
Gang Cheng, 2019	China	Ŋ	950/97	<25;46, 25-29;449; 30-34;335; 35:120	FAS duration	During prepregnancy and pregnancy	IADPSG	24–28 weeks' gestation	Age, education, household income, postpartum BMI, parity, passive parity, passive smoking, family history of diabetes, iron supplementation, multivitamin supplementation, gestational dietary intake, and alcohol use	σ
Lili Huang, 2019	China	S	326/33	28.4 ± 3.15	FAS duration	16–18 weeks' gestation	IADPSG	24–28 weeks' gestation	Age, pre-pregnancy BMI, education level, parity, passive smoking, monthly income, energy and overall nutrients intake, TG, HDL-c, LDL-c concentrations	Q

GDM gestational diabetes mellitus, SD standard deviation, CS cohort study, CCS case control study, CSS cross sectional study, NA not available, RBC red blood cell, FAS folic acid supplement, IADPSG International Association of the Diabetes and Pregnancy Study Groups, BMI body mass index, OGTT oral glucose tolerance test, TG Triglyceride, HDL high-density lipoprotein, LDL low-density lipoprotein.
*CS and CCS were assessed using Newcastle-Ottawa Scale (NOS), and CSS using Agency for Healthcare Research and Quality (USA) (AHRQ).

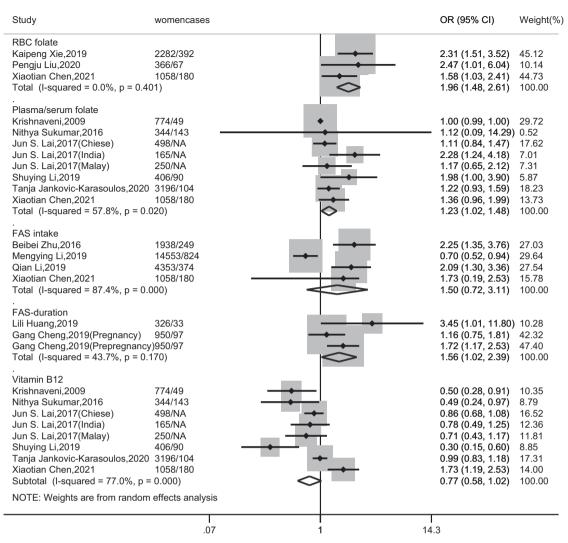


Fig. 2 Forest plot of meta-analyses evaluating association of gestational diabetes mellitus (GDM) risk with maternal red blood cell (RBC) and plasma/serum folate, dose and duration of folic acid supplement (FAS) and plasma/serum vitamin B12 (highest vs. lowest categories).

95% CI: 1.02–2.39) (Fig. 2). Because of poor consistency, the certainty of evidence was categorized as very low (Supplementary Table 1).

Of note, excluding the study conducted in the America by Li et al. [19] removed the observed statistical heterogeneity ($I^2=0.0\%$, P=0.928), and we found the highest category of FAS intake had a significantly 2.13 times increased risk of GDM compared to lowest category of FAS (95% CI: 1.52–2.98) (Supplementary Fig. 1c).

Furthermore, the subgroup analyses indicated that the result reported by Li et al. (OR = 0.70, 95% Cl: 0.52–0.94) was from western country, and evaluated the relationship between FAS intake during prepregnancy and self-reported GDM, while the other three studies with no heterogeneity (pooled OR = 2.13, 95% Cl: 1.52–2.98) were based on Chinese population, the data of FAS intake were collected during pregnancy and the diagnoses of GDM was all in line with IADPSG criteria, which may explain the inconsistent estimates (Table 2) among these studies. By contrast, test for heterogeneity between subgroups of adjusting for history of DM or intakes of other nutrients may be invalid due to considerable heterogeneity observed in one or more subgroups.

Plasma/serum vitamin B12 and GDM risk

With respect to plasma/serum vitamin B12 level, the randomeffects meta-analysis pooled eight datasets from six studies with very high heterogeneity ($l^2 = 77.0\%$, P < 0.001). No significant association of GDM risk with the highest B12 was observed compared to the lowest B12 (OR = 0.77, 95% Cl: 0.58–1.02) (Fig. 2) with very low certainty (Supplementary Table 1), the downgrade was mainly due to serious imprecision and inconsistency.

However, the sensitivity analysis by excluding the study by Chen [15] attenuated the heterogeneity ($I^2 = 66.0\%$, P = 0.007) and strengthened this effect, rendering it statistically significant (OR = 0.70, 95% CI: 0.54–0.90). When excluding the other seven datasets in turn, the association of GDM risk and vitamin B12 appeared to be inconclusive (Supplementary Fig. 1e).

In the subgroup analyses (Table 2), we observed significant decreased risk of GDM with higher plasma/serum vitamin B12 in the groups of studies with lower number of participants (<1000), or with estimates adjusting for intakes of other nutrients, or the case control or cross sectional studies.

Dose-response analyses

First, we performed nonlinear dose-response association analyses of GDM risk with RBC folate and FAS using a restricted cubic spine model. The studies and data points on the dose response curve were presented in the Supplementary Table 2. We observed a positive nonlinear association of GDM risk with RBC folate ($P_{\text{nonlinearity}} = 0.0004$; Fig. 3) and FAS intake ($P_{\text{nonlinearity}} = 0.0002$; Fig. 4). Within 500 ng/ml, higher RBC folate was associated

Characteristics and covariates	2	OR (95% CI)	- D*	#	2	OR (95% CI)	2 D*	#
	RBC folate			Detween	Plasma/serum folate			Detween
Overall	ĸ	1.96 (1.48, 2.61)	0.00%, 0.401		8	1.23 (1.02, 1.48)	57.80%, 0.02	
Number of participants				0.592				0.032
>1000	2	1.91 (1.32, 2.77)	0.00%, 0.638		2	1.27 (1.02, 1.57)	0.00%, 0.638	
<1000	1	2.47 (1.01, 6.04)			9	1.24 (0.96, 1.59)	57.40%, 0.038	
Study design				ı				0.059
S	3	1.96 (1.48, 2.61)	0.00%, 0.401		33	1.12 (0.92, 1.35)	58.7, 0.089	
CCS	0				_	1.12 (0.09, 14.11)	1	
CSS	0				4	1.45 (1.01, 2.10)	50.8, 0.107	
Geographical region				1				0.146
Asian	8	1.96 (1.48, 2.61)	0.00%, 0.401		9	1.26 (1.00, 1.59)	65.50%, 0.013	
Western	0				2	1.22 (0.93, 1.59)	0.00%, 0.948	
Adjusting for history of DM				0.308				0.146
Yes	2	1.72 (1.17, 2.52)	0.00%, 0.374		9	1.26 (1.00, 1.59)	65.50%, 0.013	
No	1	2.31 (1.51, 3.52)	1		2	1.22 (0.93, 1.59)	0.00%, 0.948	
Adjusting for intakes of other nutrients				0.592				0.018
Yes	1	2.47 (1.01, 6.04)	•		5	1.42 (1.03, 1.98)	34.60%, 0.02	
No	2	1.91 (1.32, 2.77)	0.00%, 0.638		3	1.12 (0.92, 1.35)	58.70%, 0.089	
	FAS intake				FAS duration			
Overall	4	1.50 (0.72, 3.11)	87.40%, <0.001	1	3	1.56 (1.03, 2.39)	43.70%, 0.17	
Number of participants				1				
>1000	4	1.50 (0.72, 3.11)	87.40%, <0.001		0			
<1000	0				3	1.56 (1.03, 2.39)	43.70%, 0.17	
Study design				1				
CS	4	1.50 (0.72, 3.11)	87.40%, <0.001		3	1.56 (1.03, 2.39)	43.70%, 0.170	
Region				<0.001				
Asian	3	2.13 (1.52, 2.98)	0.00%, 0.928		3	1.56 (1.03, 2.39)	43.70%, 0.170	
Western	1	0.70 (0.52, 0.94)	,		0			
Adjusting for history of DM				0.004				0.178
Yes	æ	1.29 (0.54, 3.12)	87.10%, <0.001		2	1.43 (0.98, 2.10)	42.50%, 0.187	
No	1	2.25 (1.35, 3.76)	1		1	3.45 (1.01, 11.8)	1	
Adjusting for intakes of other nutrients				0.003				
Yes	2	0.19 (0.41, 3.47)	93.20%, <0.001		3	1.56 (1.03, 2.39)	43.70%, 0.170	
ON	2	2.17 (1.35, 3.50)	0.00%, 0.711		0			
Period for exposure measurement				<0.001				<0.001
Prepregnancy	-	0.70 (0.52, 0.94)	¥		1	1.72 (1.17, 2.53)	¥	
Pregnancy	33	2.13 (1.52, 2.98)	0.00%, 0.928		1	1.16 (0.75, 1.81)	¥	
Prepregnancy and early pregnancy	0				1	3.45 (1.01, 11.8)	· ·	
Plasma/serum Vitamin B12								
Overall	8	0.77 (0.58, 1.02)	77.00%, <0.001	•				
Number of participants				<0.001				
>1000	2	1.28 (0.74, 2.20)	85.50%, 0.009					

Table 2. continued								
Characteristics and covariates	2	OR (95% CI)	l², p* P	$P_{ m between}^{\ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ $	N	OR (95% CI)	l², p*	P _{between}
<1000	9	0.63 (0.47, 0.84)	55.10%, 0.049					
Study design			0	0.011				
CS	3	0.99 (0.58, 1.69)	84.9%, 0.001					
CCS	1	0.49 (0.24, 0.97)						
CSS	4	0.68 (0.47, 0.97)	62.90%, 0.044					
Region			0	0.401				
Asian	9	0.75 (0.50, 1.12)	80.80%, <0.001					
Western	2	0.76 (0.39, 1.48)	72.70%, 0.056					
Adjusting for history of DM			0	0.401				
Yes	9	0.75 (0.50, 1.12)	80.80%, <0.001					
No	2	0.76 (0.39, 1.48)	72.70%, 0.056					
Adjusting for intakes of other nutrients			0	9000				
Yes	2	0.65 (0.47, 0.90)	58.10%, 0.049					
ON	3	0.99 (0.58, 1.69)	84.90%, 0.001					
GDM gestational diabetes mellitus, DM diabetes mellitus, FAS folic acid supplement, CS cohort study, CCS case control study, CSS cross sectional study, OR odds ratio, CI confidence intervals.	tes mellitus, FAS fo	olic acid supplement, CS	cohort study, CCS case	control stud	y, CSS cross sectional study,	OR odds ratio, C/ confide	ence intervals.	

*P-heterogeneity within subgroups using a random-effects model.

*P-heterogeneity between subgroups using a fixed-effects model.

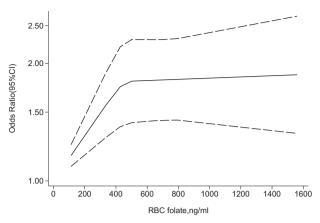


Fig. 3 Dose response association between maternal red blood cell (RBC) folate and gestational diabetes mellitus (GDM) risk.

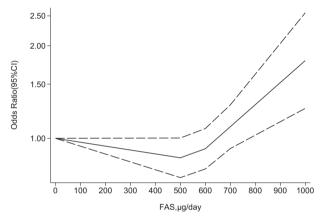


Fig. 4 Dose response association between folic acid supplement (FAS) intake and gestational diabetes mellitus (GDM) risk.

with higher GDM risk. When RBC folate >500 ng/ml, then GDM risk appeared not to increase (Fig. 3). The GDM risk appeared to be the lowest in the point of 400 μ g/day FAS intake, and when the intake of FAS was up to around 800 μ g/day or above, a significant increased risk of GDM began to be observed (Fig. 4).

Then we conducted linear dose-response analyses. The two-stage fixed-effects dose-response model provided the evidence of a linear association of GDM risk with RBC folate, and we found that each 200 ng/ml increase in RBC folate was related to 8% higher risk of GDM (OR = 1.08, 95% Cl: 1.03–1.13, n=3 studies). However, we did not find the linear dose-response association of GDM risk with FAS intake (OR = 1.04, 95% Cl: 0.94–1.15, P=0.457, n=3 studies, for each 100 µg/day increment in FAS intake).

In addition, we did not conduct dose-response analyses due to a lack of sufficient data regarding each category of plasma/serum folate, duration of FAS and plasma/serum vitamin B12.

Publication bias

For each meta-analysis including less than ten studies, the minimum number of studies suggested for use of funnel plot [32, 37], the power of publication bias assessment may be too low to distinguish true bias from chance. As presented in Supplementary Fig. 2, there was no evidence of publication bias for RBC folate (Egger's test P=0.745, Begg's test P=0.602), FAS intake (Egger's test P=0.367, Begg's test P=0.497), FAS duration (Egger's test P=0.585, Begg's test P=0.602). The evidence of publication bias was inconclusive for plasma/serum folate (Egger's test P=0.009, Begg's test P=0.458) and plasma/serum vitamin B12 (Egger's test P=0.139, Begg's test P=0.013).

DISCUSSIONS

We systematically conducted the meta-analyses to evaluate the association of GDM risk with one-carbon metabolism-related B vitamins (RBC and plasma/serum folate, dose and duration of FAS intake, vitamin B12) based on thirteen observational studies with 31459 participants and 2766 GDM cases. We found a 96% and 23% significant increase risk of GDM with the highest versus the lowest category of RBC folate and plasma/serum folate, with moderate and low certainty, respectively. The dose-response analysis also revealed that each 200 ng/ml increase in RBC folate was associated with 8% significant higher risk of GDM. No significant increased risk of GDM was observed in the highest vs. the lowest category of FAS intake. A longer duration (≥3 months) of FAS intake was significantly associated with 56% higher risk of GDM. However, due to the very low certainty evidence—downgraded for serious inconsistency, it is uncertain of this effect. In addition, no statistically significant association of GDM risk with the highest plasma/serum B12 level was observed in comparison with the lowest B12.

Because estimating natural dietary intake of folate is difficult and imprecise, RBC and plasma/serum folate, as the biomarkers, are more useful and reliable for evaluating the folate nutritional status. Generally in accordance with the results of a previous meta-analysis [16] without distinguishing specified folate status, we observed similar patterns of increasing GDM risk in relation to excess RBC and plasma/serum folate, respectively. Intriguingly, Chen also reported that the decrease in serum folate from early to midgestation was negatively associated with GDM risk (aOR = 0.95, 95% CI:0.90–0.99), which further reinforced the detrimental effect of excess folate on GDM [15].

As the biomarker of folate status, plasma/serum folate is an indicator of recent folate intake and is dramatically affected by FAS [38]. By comparison RBC folate responds slower to changes in folate intake and represents the long-term folate status [38]. But RBC and plasma/serum folate concentrations are highly correlated. Furthermore, their correlation was reported to be modified by BMI and genotype and substantially by low plasma vitamin B-12 [39]. Previous studies in various populations have consistently shown that NTD risk decreased with increasing RBC folate concentration in relation to dietary folate intake levels. How and to what extent FAS has impact on maternal RBC and plasma/serum folate concentrations still remains uncertain. An ancillary study within the Folic Acid Clinical Trial (FACT) has found that high-dose FAS in early pregnancy increased maternal serum folate but not RBC folate concentrations, suggesting tissue saturation [40]. Notably, a randomized controlled trial in Malaysia indicated that weekly ironfolic acid (IFA) supplements containing 2.8 mg folic acid increased RBC folate and had a lower risk of NTD more than those containing 0.4 mg folate [41]. So it is an interesting and important problem to identify the mutual relation among daily folic acid intake, RBC and plasma/serum folate concentrations, further assess and establish optimal RBC and plasma/serum folate for NTD prevention and recommend the daily total folic acid intake in various population [42]. However, the exact cut-offs to classify clinical and subclinical deficiency or excess remain debated. Therefore, based on the results of our meta-analyses, we anticipate that optimal RBC and plasma/serum folate should be evaluated and figured out in consideration of the potential risk of GDM as well as other adverse effects.

In contrast with the positive associations of GDM risk with RBC and plasma/serum folate, the adverse effect of higher FAS intake on GDM was observed but not statistically significant. However, when we performed the sensitivity and subgroup analysis by excluding the study conducted by Li et al. [19], heterogeneity of the other three Chinese cohort studies disappeared and a significant 2.13 times increased risk of GDM was obtained with the highest vs. the lowest category of FAS intake. The study excluded was based on western population and focused on the prepregnancy intake of FAS and self-reported GDM, which may

account for the heterogeneity. In addition, inconsistent with the results of Li et al., Zhu reminded that an increased risk of GDM was not apparent for women using FAS before pregnancy alone [18]. To add, we did not question the adequate intake of FAS, with the reason that we only observed a significant increased risk of GDM when the intake of FAS was up to around $800\,\mu\text{g}/\text{day}$ or above as revealed by the non-linear dose-response analysis.

Concomitantly, we found that the significant positive relationship between duration of FAS and GDM risk. However, only two eligible studies both based on Chinese cohorts were included in the analysis and one of them [21] investigated FAS duration before and after pregnancy separately and the other [22] focused on FAS duration before pregnancy alone or in the second trimester alone or the total duration from prepregnancy to early pregnancy separately, which made the results mixed. Notably, In combination of dose and duration of FAS, Qian Li et al. [20] found that FAS above 800 µg/day for long duration (continuous at least 4 weeks prepregnancy and continued for at least 16 weeks during pregnancy before OGTT) was related to the highest GDM risk compared with nonusers in a Chinese population, indicating an additive effect exists. In addition to GDM risk, increasing evidence was reported that high-dose FAS intake during pregnancy may contribute to the rise in gestational hypertension [43, 44] and in adverse outcomes of the offspring, such as allergic disease and autism spectrum disorder [45, 46]. Thus, this highlights the importance of systematically confirming the potential adverse effects of excess dose and duration of FAS intake with differentiating between prepregnancy and other trimester periods separately in various population with larger sample and optimizing the FAS strategy in further studies.

Similar significant shift of risk estimates also occurred in regard to plasma/serum vitamin B12. After excluding the study by Chen et al. [15] in the sensitivity analysis, we observed a significant 30% decrease in GDM risk in the highest categories of plasma/serum B12 compared with the lowest B12, along with reduced heterogeneity. It is reported that elevated maternal vitamin B12 requirements and low serum vitamin B12 concentrations associate with physiological changes during pregnancy [47]. High serum vitamin B12 levels in some disorders such as liver disease may mask functional B12 deficiency [48]. We assumed the discrepancy of the GDM-B12 relationship may be partially explained by the different physiological status and the time points of plasma/serum vitamin B12 measurement between the studies. Furthermore, when excluding the study without adjusting for intakes of other nutrient, the protective effect of plasma/serum B12 on GDM risk turned to be significant. We speculated that the effect of plasma/ serum B12 on GDM may be confounded by other factors such as energy and overall nutrients intakes, iron, folate, multivitamin and other supplements. Of note, The highest ORs of GDM were observed among women with combined vitamin B12 insufficiency and high folate concentration [23, 25]. Furthermore, the synergistic adverse effect of imbalance of plasma/serum folate and vitamin B12 (higher folate/B12) on GDM risk was also reported to be further elevated by older maternal age and higher prepregnancy BMI with significant additive interaction, and the effect of folate on GDM could be confounded by serum vitamin B12 levels [25]. The modified effect of age may be explained by the age-related decline in vitamin B12 absorption despite adequate intake in older adults [10]. Although there are not enough data on the interactions between folate and vitamin B12 imbalance and effect modifiers on the risk of GDM, the findings give a promising direction to understand its mechanism and have potential implications for antenatal supplement recommendations. Further studies are needed to confirm the relationship between GDM risk and the imbalance of folate and vitamin B12 from prepregnancy through pregnancy and effect modifiers such as age, BMI and the presence of other nutrients.

Mechanism of the effect

The underlying mechanism of the possible adverse effect of excess folate on GDM remains to be elucidated. The imbalance between folate and vitamin B12 may be one of the possible explanations. Folate and vitamin B12 are key cofactors in one carbon metabolic pathway. One carbon pathway generates precursors for nucleotide biosynthesis and methyl groups for methylation reactions [10]. The deficiency of folate and vitamin B12 may lead to disruption of DNA synthesis, cellular inflammation, adiposity dysfunction with increased lipogenesis and homocysteine levels, which might lead to glucose intolerance and anemia [49]. Furthermore, sufficient folate status could mask the vitamin B12 deficiency, exacerbate the deleterious effect and clinical manifestations of vitamin B12 deficiency with increasing dose of folate and might participate in the pathogenesis of GDM through worsening insulin resistance [50]. Recent study also revealed that low vitamin B12 levels in pregnancy plays a role in epigenetic regulation by altering adipose-derived circulating miRNAs during adipocyte differentiation and resulted in an adverse insulin resistance phenotype [51]. In addition, high folic acid intake and excess unmetabolized plasma folate may get involved in the developing of GDM by reducing natural killer cell cytotoxicity and immune function [52, 53].

Strengths and limitations

The systematical meta-analyses have several strengths. The majority of the included studies were high-quality cohort studies and the pooled analyses were based on risk estimates adjusted for confounding covariates, which strengthen the conclusion. Moreover, we examined strictly methodological quality of the evidence using NOS and AHRQ and assessed the certainty of the evidence via GRADE tool. In addition, we conducted sensitivity analysis, subgroup analysis and dose-response analysis to shed light on the true associations with consideration of potential heterogeneity. In terms of RBC folate, the consistent results of the multiple subgroup analyses and dose-response gradient reinforce its effect on GDM risk and upgrade the certainty of evidence.

Inevitably, we acknowledge several limitations. First, the relatively limited number of eligible studies for each exposure (<10 studies) may make it difficult and less power to safely detect potential bias or sources of heterogeneity. We attempted to address this issue by using random-effect model and subgroup analyses. Secondly, comprehensive dose-response analyses were not conducted as planned due to a lack of sufficient data regarding each category of plasma/serum folate, duration of FAS and plasma/serum vitamin B12. Thirdly, as to the effect of dose and duration of FAS and plasma/serum vitamin B12 on GDM risk, considering serious imprecision, inconsistency, and unexplained heterogeneity, the certainty of evidence remain very low. The results should be interpreted with caution.

So further high-quality longitudinal studies or randomized controlled trials are warranted to clarify the true independent relationships. Furthermore, future studies should measure FAS intake, vitamin B12 and their biomarkers at different pregnancy stages, and adjust the confounding factors such as intakes of other nutrient, age, BMI, history of diabetes mellitus and other covariates in various population.

CONCLUSIONS

To conclude, moderate-certainty evidence of observational studies indicates that higher RBC folate appears to be significantly associated with an increasing risk of GDM. The effect across multiple prespecified subgroups and a dose response effect reinforce this conclusion. Higher plasma/serum folate may be significantly associated with increased GDM risk at low certainty. For other comparisons, the certainty of evidence was rated as "very low", which means that the associations of GDM risk with dose or

duration of FAS, maternal plasma/serum vitamin B12 remain unclear. More well-designed studies are needed to confirm the associations.

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AUTHOR CONTRIBUTIONS

NL developed the idea and study design; NL and JJ carried out the literature searching, data-extracting and critical appraisal; NL, JJ and LG performed the statistical analyses and interpretation; NL drafted and edited the manuscript; JJ and LG gave advice on extensive revision. All authors have read and approved the submission of final manuscript.

COMPETING INTERESTS

The authors declare no competing interests.

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