

OPEN A Mendelian Randomization Study of Plasma Homocysteine and **Multiple Myeloma**

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Observational studies have demonstrated an association between elevated homocysteine (Hcy) level and risk of multiple myeloma (MM). However, it remains unclear whether this relationship is causal. We conducted a Mendelian randomization (MR) study to evaluate whether genetically increased Hcy level influences the risk of MM. We used the methylenetetrahydrofolate reductase (MTHFR) C677T polymorphism as an instrumental variable, which affects the plasma Hcy levels. Estimate of its effect on plasma Hcy level was based on a recent genome-wide meta-analysis of 44,147 individuals, while estimate of its effect on MM risk was obtained through meta-analysis of case-control studies with 2,092 cases and 4,954 controls. By combining these two estimates, we found that per one standard-deviation (SD) increase in natural log-transformed plasma Hcy levels conferred a 2.67-fold increase in risk for MM (95% confidence interval (CI): 1.12–6.38; $P = 2.7 \times 10^{-2}$). Our study suggests that elevated Hcy levels are causally associated with an increased risk of developing MM. Whether Hcy-lowering therapy can prevent MM merits further investigation in long-term randomized controlled trials (RCTs).

Multiple myeloma (MM) is a malignant disease characterized by proliferation of clonal plasma cells in the bone marrow and typically accompanied by the secretion of monoclonal immunoglobulins that are detectable in the plasma or urine, causing anemia, pathological fracture, and the clinical symptoms of renal failure. Among other plasma cell dyscrasias, such as Waldenström's macroglobulinaemia and primary amyloidosis, multiple myeloma is the second most frequent haematological malignancy with an age-adjusted incidence of six per 100,000 per year in the USA and Europe^{1,2}. However, the cause of the MM is still largely unknown. Risk factors such as age, positive family history, smoking, alcohol consumption, ionizing radiation, industrial occupation, and obesity have been reported for the development of MM^{3,4}. Because treatments for MM are limited, the best approach to reduce mortality and morbidity is primary prevention through modification of acquired risk factors.

Homocysteine (Hcy), a sulfur-containing amino acid, is formed in the demethylation of the essential amino acid methionine⁵. Previous studies have demonstrated that high plasma Hcy concentration is an independent risk factor for cerebrovascular, peripheral vascular, and cardiovascular disease^{6,7}. Furthermore, observational studies showed that plasma levels of Hcy as well as its metabolizing factors were associated with the overall risk of cancer⁸⁻¹¹. However, Hcy level is also related to smoking status, blood pressure, and social class. Thus, this relation could be subject to residual confounding, reverse-causality bias, or both¹¹. In view of these, we conducted a Mendelian randomization analysis to assess whether elevated Hcy level is causally associated within creased risk of MM.

In the absence of evidence from high-quality randomized controlled trials (RCTs), the principles of Mendelian randomization (MR) can be applied to strengthen or refute the causality of biomarkers in disease etiology. MR is a study design in which genetic variants are served as instrumental variables for estimating the unconfounded effect of an exposure (for example, Hcy) on a disease (for example, stroke)¹¹. This approach, which is conceptually similar to a RCT, is based on the principle that genetic variants are randomly allocated at meiosis, and consequently these genetic variants are independent of many factors that bias observational studies, such as confounding and

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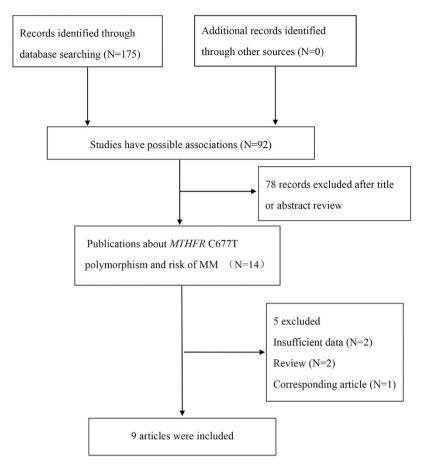


Figure 1. Flow diagram of study selection. The terms "N" in the boxes represent the number of corresponding studies.

reverse causation. MR methods have been used previously to investigate the influence of type 2 diabetes(T2D) and fasting glucose (FG) on coronary heart disease(CHD) risk, which supporting a causal relationship between T2D and CHD¹². MR methods may be of particular relevance for understanding the etiology of MM since the date of disease onset is often poorly recognized clinically and MR studies assess the effect of lifetime exposures.

MR analyses using methylenetetrahydrofolate reductase (*MTHFR*) C677T polymorphism as an instrument variable have been carried out in the past^{11,13–15}. The researchers provided evidence from MR that plasma Hcy level is causally related to stroke, schizophrenia, T2D and offspring birth weight. In the present study, we provide evidence on the presence, direction, and magnitude of a causal effect of plasma Hcy levels on MM risk by performing a MR study based on the *MTHFR* C677T polymorphism as an instrumental variable.

Trial sequential analysis (TSA) is an approach that provides the required information size in meta-analyses. Under the premise of not expanding the type I error, we can earlier draw a conclusion using TSA, which can terminate defects of invalid test and save medical resources, compared to the traditional meta-analysis. And using TSA can make our result of meta-analysis be more stable and reliable ¹⁶.

Results

Study selection and characteristics. The process of literature retrieval and exclusion was shown in Fig. 1. The literature search identified a total of 92 potential articles related to *MTHFR* C677T polymorphism and risk of MM. 78 articles were excluded after reading abstract and title. Then, 5 articles were excluded once again due to insufficient data, review or corresponding article. Finally, 9 studies with a total of 2,092 cases and 4,954 controls, were included in our meta-analysis^{17–25}. The main characteristics of included studies were shown in Table 1. Among those 9 studies, 6 studies were carried out in Caucasians, while 3 were in Asian populations. Most of those included studies used RFLP-PCR to test the genotype of *MTHFR* C677T polymorphism, and a small part of which were genotyped by Microarray or TaqMan assay. Genotypes distribution in the controls of all included studies were in agreement with HWE (Hardy Weinberg Equilibrium).

Association of *MTHFR* **C677T polymorphism with risk of multiple myeloma.** The main results of the current meta-analysis and heterogeneity were summarized in Table 2. In conclusion, we found significant association between MTHFR C677T polymorphism and MM susceptibility under allele model (T vs. C, OR = 1.17, 95% CI = 1.02–1.34, P = 0.03). However, no remarkable association was observed between MTHFR C677T polymorphism and MM risk under other four genetic models (TT vs. CC, OR = 1.16, 95% CI = 0.98–1.37,

				Genotyping	Control	Sample size	Cases			Controls			HWE (P
Study	Year	Country	Ethnicity	methods	Source	(case/control)	CC	CT	TT	CC	CT	TT	value)
Jiang ¹⁸	2014	China	Asia	Microarray	НВ	30/157	9	11	10	72	66	19	0.52
Martino ¹⁷	2014	Mixed	Caucasian	TaqMan assay	НВ, РВ	1264/1797	554	525	185	767	787	243	0.07
Lima ¹⁹	2008	Brazil	Caucasian	PCR-RFLP	PB	123/188	52	57	14	92	79	17	0.99
Moon ²⁰	2007	Korea	Asia	TaqMan assay	PB	196/434	57	103	36	144	196	94	0.08
Kim ²¹	2007	Korea	Asia	PCR-RFLP/ Real-time PCR	РВ	173/1700	58	80	35	540	863	297	0.13
Chiusolo ²²	2006	Italy	Caucasian	PCR-RFLP	PB	100/100	31	44	25	36	45	19	0.46
Lincz ²³	2003	Australia	Caucasian	PCR-RFLP	PB	90/299	38	44	8	145	133	21	0.20
González-Fraile ²⁴	2002	Spain	Caucasian	PCR-RFLP	PB	90/79	31	48	11	38	32	9	0.57
González Ordóñez ²⁵	2000	Spain	Caucasian	PCR-RFLP	Unknown	26/200	5	17	4	92	88	20	0.88

Table 1. Main characteristics of studies included in the meta-analysis. PB, Population-based; HB, Hospital-based; PCR, polymerase chain reaction; RFLP, restriction fragment length polymorphism; HWE, Hardy-Weinberg equilibrium.

Genetic models	OR (95% CI)	P value	Model	I ² value (%)
T vs. C	1.17 (1.02, 1.34)	0.03	Random	47
TT vs. CC	1.16 (0.98, 1.37)	0.08	Fixed	29
TC vs. CC	1.18 (0.96, 1.45)	0.12	Random	46
TT + TC vs. CC	1.22 (1.00, 1.48)	0.05	Random	47
TT vs. TC + CC	1.13 (0.98, 1.32)	0.10	Fixed	20

Table 2. Meta-analysis of the association between MTHFR C677T polymorphism and MM risk.

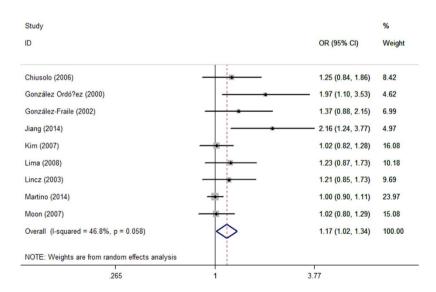


Figure 2. Forest plots of the risk of MM associated with the *MTHFR* C677T polymorphism under allele **model** (T vs. C). The solid diamonds and horizontal lines correspond to the study-specific ORs and 95% CIs. The gray areas reflect the study-specific weight. The hollow diamonds represent the pooled ORs and 95% CIs of the overall population. The vertical solid lines show the OR of 1 and the vertical dashed lines indicate the corresponding pooled OR.

P = 0.08; TC vs. CC, OR = 1.18, 95% CI = 0.96-1.45, P = 0.12; TT + TC vs. CC, OR = 1.22, 95% CI = 1.00-1.48, P = 0.05; TT vs. TC + CC, OR = 1.13, 95% CI = 0.98-1.32, P = 0.10) (Table 2, Fig. 2).

Using the trial sequential analysis (TSA), the required information size is 2823 subjects to demonstrate the issue. Until now, the cumulative Z-curve crossed the trial sequential monitoring boundary and the required information size has been reached, confirming that *MTHFR* C677T polymorphism is associated with increased risk of MM and further relevant trials are unnecessary (Fig. 3). The trial sequential analysis adjusted 95% confidence interval was 0.99 to 1.38.

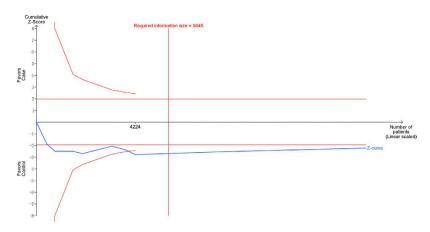


Figure 3. Trial sequential analysis of 9 studies reporting MTHFR C677T polymorphism. The required information size was calculated using $\alpha=0.05$ (two sided), $\beta=0.20$ (power 80%), $D^2=67.0\%$, a relative risk increase of 16.81% and an event proportion of 38.02% in the control arm. The blue cumulative Z-curve was constructed using a random effects model.

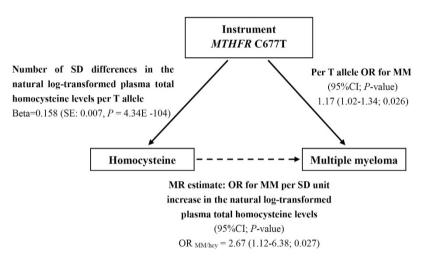


Figure 4. Schematic representation of the Mendelian randomization design. The risk estimate for the association between *MTHFR* C677T polymorphism and MM risk was obtained from the present meta-analysis. The effect of *MTHFR* C677T polymorphism on the SD change in natural log-transformed plasma Hcy levels was obtained from a recent meta-analysis of genome-wide association studies. SE = standard error, SD = standard deviation.

Mendelian randomization analysis for the association of *MTHFR* **C677T polymorphism with multiple myeloma risk.** In order to assess the association between genetically increased plasma Hcy level and risk of multiple myeloma, we performed a MR analysis. By combining two pooled estimates, OR $_{\text{MM/per T-allele}}$ from a meta-analysis of 9 case-control studies and beta $_{\text{hcy/per T-allele}}$ from a meta-analysis of genome-wide association studies by van Meurs and colleagues, we observed that each 1-SD increase in natural-log-transformed plasma Hcy level was associated with a 2.67-fold increased risk of MM (95% CI: 1.12–6.38; $P = 2.7 \times 10^{-2}$) (Fig. 4). Using the online sample size and power calculator, we had 96.5% power to detect the causal odds ratio.

Sensitivity analysis and publication bias. The leave-one-out sensitivity analysis showed that no single study qualitatively altered the pooled ORs, indicating the reliability and stability of our results. Begg's funnel plot and Egger's test were performed to evaluate the potential publication bias of literatures. The shape of the funnel plot showed no evidence of obvious asymmetry (Fig. 5). The Egger's test result did not support the existence of publication bias (TT ν s. CT + CC, t = 1.53, P = 0.169).

Discussion

Our MR study demonstrated that a genetic increase in natural log-transformed plasma Hcy by 1 SD was associated with a 2.67-fold increased risk of MM, providing strong evidence in support of a causal role of Hcy on MM susceptibility. Since genetic effects on Hcy levels represent differences that generally persist throughout adult life, the estimate of our MR study reflects an effect of Hcy over the course of a lifetime. Meanwhile, our findings

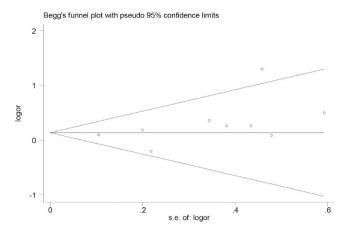


Figure 5. Begg's funnel plot for the *MTHFR* **C677T polymorphism and MM risk** (TT *vs.* CT + CC). Each circle represents a separate study for the indicated association. Logor = natural logarithm of OR, s.e. = standard error.

are consistent with previous evidence from observational studies that plasma higher Hcy levels influence risk of $MM^{8,9}$. To our knowledge, this report is the first to provide evidence for putative causal nature of the association between plasma Hcy and MM.

As we all know, MTHFR is a key enzymes of the methylation reaction. MTHFR converts 5,10-methylenetetrahydrofolate into 5-methyltetrahydrofolate and this reaction provide a methyl for Hcy into methionine in the catalyzed reaction by methionine synthase (MTR) and MTR requires vitamin B12 as a coenzyme²⁶. A previous study reported that homozygous mutation of MTHFR C677T reduced by approximately 70% of the mean enzyme activity and the heterozygous mutation of MTHFR C677T reduced by approximately 35% of the mean MTHFR activity²⁷. There is no doubt that MTHFR C677T polymorphism was related to elevated plasma Hcy levels and lower folate level²⁸⁻³¹, which was consistent with recent GWAS meta-analysis³².

Hcy, a well-known cardiovascular risk factor, involves in one-carbon methyl group-transfer metabolism³³. The mechanism has been considered critical for Hcy metabolism in carcinogenesis in terms of DNA synthesis, repair and methylation^{34,35}. Previous studies have demonstrated that high plasma level of Hcy was associated with risk of a wide range of cancers, such as breast cancer, lung cancer, colorectal cancer and cervical cancer^{36–41}. Although there have been no substantial and enough studies to strengthen the view that Hyper-Hcy levels were associated with MM risk, our MR analysis convincingly indicated elevated Hcy was causally associated with increased risk of MM. Therefore, it is necessary to conduct great scale RCTs to assess Hcy-lowering for the treatment and prevention of MM.

Our research has advantages in itself. Firstly, with the application of Mendelian randomization, the effect estimates are closer to the real situation and we are able to overcome potential confounding and reverse causation that may bias evaluations from observational studies. Secondly, Our MR analysis described the association of a lifetime of exposure to Hcy-increasing allele in the general population, whereas observational studies merely provide insights from intervention for shorter periods in individuals at risk. Lastly, the present data from the largest GWAS meta-analysis for Hcy level (44,147 individuals) and from the current meta-analysis for MM risk (2,092 cases and 4,954 controls) have enabled us to more precisely examine our study hypothesis than if we used individual-level data from a small study^{32,42}.

A few limitations of our study should be considered. First of all, MR estimates which utilize instrumental variables accounting for little variance in a trait tend to be biased towards the null⁴³. In this study, we used only one genetic variant as the instrumental variable that influences the plasma Hcy levels. Secondly, it seems difficult for us to exclude the pleiotropy of *MTHFR* C677T polymorphism since data on other clinical parameters across C677T genotypes are rarely provided from most qualified articles, requiring further confirmation. Thirdly, canalization, the process by which compensatory feedback mechanisms attenuate the phenotypic consequences of genetic variation, has been extensively investigated in the circumstance of MR^{44–46}. Although compensatory feedback interactions tend to bias results towards the null, the presence of this mechanism would not alter the statistical significance or direction of the effects we found through MR. Finally, considering the differences in minor allele frequencies between populations and other demographic characteristics in the included studies, it is hard to ignore an impact of population stratification.

At last, we have to emphasize this point about the impact of epigenetics on Mendelian randomization as a result of epigenetics gaining recognition as an independent field of study within the last decades. An increasing number of reports suggest that random distribution of epigenetic changes (e.g. gene expression) at conception should be considered on the assumption of MR analysis⁴⁷. Because that, some researchers raised "Two-step epigenetic Mendelian randomization" for establishing the causal role of epigenetic processes in pathways to disease^{48,49}.

In conclusion, our analysis provides a puissant evidence for a causal role of increased plasma Hcy levels in the etiology of MM. These findings may provide a new insight for further investigating the potential pathogenesis of MM and therapeutic target by decreasing the plasm Hcy levels to prevent the onset and progression of MM. Nevertheless, substantial long-term RCTs assessing the effect of Hcy-lowering on the risk of MM should be carried out in future.

Methods

Data on gene association with multiple myeloma risk. To estimate the association of the MTHFR C677T polymorphism with multiple myeloma risk, we performed a meta-analysis of case-control studies. We conducted a comprehensive search in Pubmed, Embase, Web of science databases for all eligible studies (updated to Sep 30th, 2015) by two authors independently using the following strategies. Key words or terms used for searching were: "methylenetetrahydrofolate reductase" or "MTHFR", "multiple myeloma", and "polymorphism" or "variation" or "variant" or "mutation" or "genotype" or "allele" or "SNP", without any restriction on the language. Reference lists of relevant articles were reviewed manually to look for additional studies. For inclusion, studies had to meet the following criteria: (1) evaluation for the association between MTHFR C677T polymorphism and multiple myeloma; (2) studies were designed as the case-control type; (3) genotype frequencies for both cases and controls were available. Studies were excluded if: (1) no detailed genotype frequency; and (2) case reports, family-based studies, abstracts, editorials and review articles. When multiple literatures reported the same population, only the most recent one with the largest sample sets was selected for this meta-analysis. Two authors selected the articles independently according to the above criteria. Any uncertainty regarding the eligibility was adjudged by further joint inspection of the publications.

The following data were independently extracted by two investigators from each eligible article according to a fixed protocol: first author's name, publication year, country and ethnicity of population, genotyping methods, source of control, number of cases and controls, genotype distributions in cases and controls and the HWE in controls (*P* value). If these were not possible, the authors of the publications were contacted via E-mail for more detailed data.

Data on Gene Association with Hcy. Estimate of the effect sizes of the *MTHFR* C677Tpolymorphism on the plasma Hcy levels was based on the findings of a recent GWAS meta-analysis³². The meta-analysis included data from a total of 44,147 white individuals of European ancestry derived from 10 GWAS on Hcy levels.

Statistical Analysis. *Meta-analysis.* Hardy-Weinberg equilibrium (HWE) of genotypes distribution in the control group was checked by the χ^2 -test and P < 0.05 was considered as significant disequilibrium. Studies with controls not in HWE were subjected to a sensitivity analysis. The pooled odds ratios (ORs) with their 95% confidence intervals (95% CIs) were calculated to evaluate the strength of the association between *MTHFR* C677T polymorphism and multiple myeloma risk based on different genetic models: allele model (T vs. C), homozygous model (TT vs. CC), heterozygous model (CT vs. CC), dominant model (TT + CT vs. CC), and recessive model (TT vs. CT + CC). Statistical heterogeneity between eligible studies was evaluated by using the Cochran's Q statistic and I² test⁵⁰. P < 0.1 and I² exceeding 50% indicated substantial heterogeneity across studies, then a random-effects model was chosen to perform meta-analysis, otherwise, the fixed-effects model was selected. Begg's funnel plot and Egger's regression test were used to search for publication bias and a P value > 0.05 suggested no significant publication bias have been detected⁵¹.

Trial sequential analysis. Meta-analyses may result in type I errors owing to an increased risk of random error when sparse data are analysed and due to repeated significance testing when a cumulative meta-analysis is updated with new trials $^{52-54}$. We therefore challenged the meta-analyses with the application of trial sequential analysis. Trial sequential analysis is similar to interim analysis in a single trial where the monitoring boundaries are used to decide whether the P value is sufficiently small to show the anticipated effect and whether the trial should be terminated early. In the same manner, trial sequential monitoring boundaries can be applied to meta-analyses $^{16,54-56}$. TSA depends on the quantification of the required information size. We calculated a diversity-adjusted (D^2) required information size, since the heterogeneity adjustment with I^2 underestimates the required information size I^2 ? TSA was performed with the intention to maintain an overall I^2 risk of a type I error and a power of I^2 so I^2 risk of a type I error and a power of I^2 so I^2 risk of a type I error and a power of I^2 risk of a type I error and I^2 risk of a type I err

When the cumulative Z-curve crosses the trial sequential monitoring boundary, a sufficient level of evidence may have been reached and further trials are unnecessary. If the Z-curve does not cross any of the boundaries and the required information size has not been reached, evidence to reach a conclusion is insufficient⁵⁸. We used software Trial Sequential Analysis (version 0.9, http://www.ctu.dk/tsa/) and provided the 95% confidence intervals adjusted for sparse data or repetitive testing, which we describe as the TSA adjusted 95% confidence intervals.

Mendelian randomization estimates. We calculated a MR estimate of the effect of the plasma Hcy levels on the risk of multiple myeloma (OR $_{\rm MM/Hcy}$) as log OR $_{\rm MM/Hcy}$ = (log OR $_{\rm MM/per\,T-allele}$)/beta $_{\rm Hcy/per\,T-allele}$, as in previous studies 59,60 . Log OR $_{\rm MM/Hcy}$ is the (log) increase of multiple myeloma risk by SD unit increase in the natural log-transformed plasma Hcy (MR estimate). Log OR $_{\rm MM/per\,T-allele}$ is the (log) increase in multiple myeloma risk per allele (gene-multiple myeloma association). Beta $_{\rm Hcy/per\,T-allele}$ is the number of SD differences in the natural log-transformed plasma Hcy levels per allele (SD/allele) (gene-Hcy association). The standard error of the MR estimate was derived using the Delta method 61 . Using an online sample size and power calculator for Mendelian randomization with a binary outcome (http://spark.rstudio.com/sb452/power/) 62 , we estimate the power, considering sample size, case proportion, odds ratio per SD change in the natural log-transformed plasma Hcy, 0.05 type I error rate, and assuming the variance in Hcy level explained by MTHFR C677T polymorphism is $R^2 = 0.01$. All P values were two sided. All above statistical analyses were performed using STATA software version 12.0 (STATA Corporation, College Station, TX, USA).

References

- 1. Landgren, O. et al. Risk of monoclonal gammopathy of undetermined significance (MGUS) and subsequent multiple myeloma among African American and white veterans in the United States. Blood 107, 904–906, doi: 10.1182/blood-2005-08-3449 (2006).
- 2. Röllig, C., Knop, S. & Bornhäuser, M. Multiple myeloma. The Lancet 385, 2197-2208 (2015).
- Wallin, A. & Larsson, S. C. Body mass index and risk of multiple myeloma: a meta-analysis of prospective studies. Eur J Cancer 47, 1606–1615, doi: 10.1016/j.ejca.2011.01.020 (2011).
- 4. Lope, V. et al. Occupation, exposure to chemicals, sensitizing agents, and risk of multiple myeloma in Sweden. Cancer Epidemiol Biomarkers Prev 17, 3123–3127, doi: 10.1158/1055-9965.epi-08-0343 (2008).
- Liu, Y. et al. Individual and Joint Associations of Methylenetetrahydrofolate Reductase C677T Genotype and Plasma Homocysteine With Dyslipidemia in a Chinese Population With Hypertension. Clin Appl Thromb Hemost, doi: 10.1177/1076029615609686 (2015).
- Refsum, H., Ueland, P. M., Nygard, O. & Vollset, S. E. Homocysteine and cardiovascular disease. Annu Rev Med 49, 31–62, doi: 10.1146/annurev.med.49.1.31 (1998).
- 7. Dinavahi, R. & Falkner, B. Relationship of homocysteine with cardiovascular disease and blood pressure. *Journal of clinical hypertension (Greenwich, Conn.)* **6,** 494–498; quiz 499–500 (2004).
- 8. Wang, X., Zhu, W. & Tian, P. Detection and clinical significance of blood homocysteine β2 -microglobulin and erythrocyte sedimentation rate in multiple myeloma. *Acta Academiae Medicinae Neimongol* 35, 174–178 (2013) (in chinese).
- 9. Gu, X. Clinical research on Hcy and FA levels of patients with hematologic malignancies. Occup and Health 29, 2415–2416 (2013) (in chinese).
- 10. Xie, S. Z. et al. Association between the MTHFR C677T polymorphism and risk of cancer: evidence from 446 case-control studies. *Tumour Biol*, doi: 10.1007/s13277-015-3648-z (2015).
- Casas, J. P., Bautista, L. E., Smeeth, L., Sharma, P. & Hingorani, A. D. Homocysteine and stroke: evidence on a causal link from mendelian randomisation. *Lancet* 365, 224–232, doi: 10.1016/s0140-6736(05)17742-3 (2005).
- Ahmad, O. S. et al. A Mendelian randomization study of the effect of type-2 diabetes on coronary heart disease. Nature communications 6, 7060, doi: 10.1038/ncomms8060 (2015).
- 13. Numata, S. et al. Evaluation of an association between plasma total homocysteine and schizophrenia by a Mendelian randomization analysis. BMC Med Genet 16, 54, doi: 10.1186/s12881-015-0197-7 (2015).
- 14. Yajnik, C. S. *et al.* Maternal homocysteine in pregnancy and offspring birthweight: epidemiological associations and Mendelian
- randomization analysis. *Int J Epidemiol* **43**, 1487–1497, doi: 10.1093/ije/dyu132 (2014).

 15. Huang, T., Ren, J., Huang, J. & Li, D. Association of homocysteine with type 2 diabetes: a meta-analysis implementing Mendelian
- randomization approach. *BMC Genomics* **14**, 867, doi: 10.1186/1471-2164-14-867 (2013).

 16. Brok, J., Thorlund, K., Gluud, C. & Wetterslev, J. Trial sequential analysis reveals insufficient information size and potentially false positive results in many meta-analyses. *J Clin Epidemiol* **61**, 763–769, doi: 10.1016/j.jclinepi.2007.10.007 (2008).
- 17. Martino, A. et al. Genetic variants and multiple myeloma risk: IMMEnSE validation of the best reported associations—an extensive replication of the associations from the candidate gene era. Cancer Epidemiol Biomarkers Prev 23, 670–674, doi: 10.1158/1055-9965. epi-13-1115 (2014).
- 18. Jiang, N. *et al.* The relationship between methylenetetrahydrofolate reductase polymorphism and hematological malignancy. *Clin Lab* **60**, 767–774 (2014).
- 19. Lima, C. S. *et al.* Polymorphisms of methylenetetrahydrofolate reductase (MTHFR), methionine synthase (MTR), methionine synthase reductase (MTRR), and thymidylate synthase (TYMS) in multiple myeloma risk. *Leuk Res* **32**, 401–405, doi: 10.1016/j. leukres.2007.06.001 (2008).
- 20. Moon, H. W. et al. MTHFR 677CC/1298CC genotypes are highly associated with chronic myelogenous leukemia: a case-control study in Korea. Leuk Res 31, 1213–1217, doi: 10.1016/j.leukres.2006.10.016 (2007).
- 21. Kim, H. N. et al. Polymorphisms involved in the folate metabolizing pathway and risk of multiple myeloma. Am J Hematol 82, 798–801, doi: 10.1002/ajh.20967 (2007).
- Chiusolo, P. et al. Analysis of MTHFR polymorphisms and P16 methylation and their correlation with clinical-biological features of multiple myeloma. Ann Hematol 85, 474–477, doi: 10.1007/s00277-006-0097-1 (2006).
- 23. Lincz, L. F. et al. Methionine synthase genetic polymorphism MS A2756G alters susceptibility to follicular but not diffuse large B-cell non-Hodgkin's lymphoma or multiple myeloma. Br J Haematol 120, 1051–1054 (2003).
- 24. González-Fraile, M. I. *et al.* Methylenetetrahydrofolate reductase genotype does not play a role in multiple myeloma pathogenesis. *British Journal of Haematology* 117, 890–892 (2002).
- 25. González Ordóñez, A. J. *et al.* Normal frequencies of the C677T genotypes on the methylenetetrahydrofolate reductase (MTHFR) gene among lymphoproliferative disorders but not in multiple myeloma. *Leukemia and Lymphoma* **39**, 607–612 (2000).
- 26. Martinez-Frias, M. L. et al. Maternal polymorphisms 677C-T and 1298A-C of MTHFR, and 66A-G MTRR genes: is there any relationship between polymorphisms of the folate pathway, maternal homocysteine levels, and the risk for having a child with Down syndrome? Am J Med Genet A 140, 987–997, doi: 10.1002/ajmg.a.31203 (2006).
- Frosst, P. et al. A candidate genetic risk factor for vascular disease: a common mutation in methylenetetrahydrofolate reductase. Nat Genet 10, 111–113, doi: 10.1038/ng0595-111 (1995).
- 28. Li, W. X., Lv, W. W., Dai, S. X., Pan, M. L. & Huang, J. F. Joint associations of folate, homocysteine and MTHFR, MTR and MTRR gene polymorphisms with dyslipidemia in a Chinese hypertensive population: a cross-sectional study. *Lipids Health Dis* 14, 101, doi: 10.1186/s12944-015-0099-x (2015).
- 29. Husemoen, L. L. et al. MTHFR C677T genotype and cardiovascular risk in a general population without mandatory folic acid fortification. Eur J Nutr 53, 1549–1559, doi: 10.1007/s00394-014-0659-2 (2014).
- 30. Yakub, M. et al. Polymorphisms in MTHFR, MS and CBS genes and homocysteine levels in a Pakistani population. PLoS One 7, e33222, doi: 10.1371/journal.pone.0033222 (2012).
- 31. Barbosa, P. R. et al. Association between decreased vitamin levels and MTHFR, MTR and MTRR gene polymorphisms as determinants for elevated total homocysteine concentrations in pregnant women. Eur J Clin Nutr 62, 1010–1021, doi: 10.1038/si.eicn.1602810 (2008).
- 32. van Meurs, J. B. et al. Common genetic loci influencing plasma homocysteine concentrations and their effect on risk of coronary artery disease. Am J Clin Nutr 98, 668–676, doi: 10.3945/ajcn.112.044545 (2013).
- 33. Selhub, J. Public health significance of elevated homocysteine. Food Nutr Bull 29, S116-125 (2008).
- 34. Ferroni, P. et al. Determinants of homocysteine levels in colorectal and breast cancer patients. Anticancer Res 29, 4131-4138 (2009).
- 35. Ulrich, C. M. *et al.* Mathematical modeling of folate metabolism: predicted effects of genetic polymorphisms on mechanisms and biomarkers relevant to carcinogenesis. *Cancer Epidemiol Biomarkers Prev* 17, 1822–1831, doi: 10.1158/1055-9965.epi-07-2937 (2008).
- 36. Tastekin, D. et al. Plasma homocysteine, folate and vitamin B12 levels in patients with lung cancer. Exp Oncol 37, 218-222 (2015).
- 37. Singh, N., Maturu, V. N. & Behera, D. Total Plasma Homocysteine Level Assessment and Timing of Folate/B12 Supplementation Prior to Initiation of Pemetrexed-Based Chemotherapy for Nonsquamous Non-Small Cell Lung Cancer Patients: An Irrelevant Investigation, an Unnecessary Delay, or Both? Oncologist 20, e21, doi: 10.1634/theoncologist.2015-0040 (2015).
- 38. Sen, S. et al. Role of cystathionine beta-synthase in human breast Cancer. Free Radic Biol Med 86, 228–238, doi: 10.1016/j. freeradbiomed.2015.05.024 (2015).

- 39. Pooja, S. et al. MTHFR 677C > T polymorphism and the risk of breast cancer: evidence from an original study and pooled data for 28031 cases and 31880 controls. PLoS One 10, e0120654, doi: 10.1371/journal.pone.0120654 (2015).
- 40. Cho, S. H. et al. 3'-UTR Polymorphisms in the MiRNA Machinery Genes DROSHA, DICER1, RAN, and XPO5 Are Associated with Colorectal Cancer Risk in a Korean Population. PLoS One 10, e0131125, doi: 10.1371/journal.pone.0131125 (2015).
- 41. Pathak, S. et al. Serum one-carbon metabolites and risk of cervical cancer. Nutr Cancer 66, 818-824, doi: 10.1080/01635581. 2014.916318 (2014).
- 42. Ma, L. M., Ruan, L. H. & Yang, H. P. Meta-analysis of the association of MTHFR polymorphisms with multiple myeloma risk. *Scientific Reports* 5, doi: 1073510.1038/srep10735 (2015).
- 43. Fewell, Z., Davey Smith, G. & Sterne, J. A. The impact of residual and unmeasured confounding in epidemiologic studies: a simulation study. *Am J Epidemiol* **166**, 646–655, doi: 10.1093/aje/kwm165 (2007).
- 44. Lawlor, D. A., Harbord, R. M., Sterne, J. A., Timpson, N. & Davey Smith, G. Mendelian randomization: using genes as instruments for making causal inferences in epidemiology. *Stat Med* 27, 1133–1163, doi: 10.1002/sim.3034 (2008).
- 45. Smith, G. D. & Ebrahim, S. 'Mendelian randomization': can genetic epidemiology contribute to understanding environmental determinants of disease? *Int J Epidemiol* 32, 1–22 (2003).
- 46. Mokry, L. E., Ahmad, O., Forgetta, V., Thanassoulis, G. & Richards, J. B. Mendelian randomisation applied to drug development in cardiovascular disease: a review. *J Med Genet* **52**, 71–79, doi: 10.1136/jmedgenet-2014-102438 (2015).
- 47. Ogbuanu, I. U., Zhang, H. & Karmaus, W. Can we apply the Mendelian randomization methodology without considering epigenetic effects? *Emerg Themes Epidemiol* **6**, 3, doi: 10.1186/1742-7622-6-3 (2009).
- 48. Allard, C. et al. Mendelian randomization supports causality between maternal hyperglycemia and epigenetic regulation of leptin gene in newborns. *Epigenetics* 10, 342–351, doi: 10.1080/15592294.2015.1029700 (2015).
- Relton, C. L. & Davey Smith, G. Two-step epigenetic Mendelian randomization: a strategy for establishing the causal role of epigenetic processes in pathways to disease. Int J Epidemiol 41, 161–176, doi: 10.1093/ije/dyr233 (2012).
- 50. Higgins, J. P. & Thompson, S. G. Quantifying heterogeneity in a meta-analysis. Stat Med 21, 1539–1558, doi: 10.1002/sim.1186 (2002).
- 51. Egger, M., Davey Smith, G., Schneider, M. & Minder, C. Bias in meta-analysis detected by a simple, graphical test. *BMJ* **315**, 629–634 (1997).
- 52. Brok, J., Thorlund, K., Wetterslev, J. & Gluud, C. Apparently conclusive meta-analyses may be inconclusive—Trial sequential analysis adjustment of random error risk due to repetitive testing of accumulating data in apparently conclusive neonatal meta-analyses. *Int J Epidemiol* 38, 287–298, doi: 10.1093/ije/dyn188 (2009).
- 53. Turner, R. M., Bird, S. M. & Higgins, J. P. The impact of study size on meta-analyses: examination of underpowered studies in Cochrane reviews. *PLoS One* **8**, e59202, doi: 10.1371/journal.pone.0059202 (2013).
- 54. Wettersley, J., Thorlund, K., Brok, J. & Gluud, C. Trial sequential analysis may establish when firm evidence is reached in cumulative meta-analysis. *J Clin Epidemiol* 61, 64–75, doi: 10.1016/j.jclinepi.2007.03.013 (2008).
- 55. Higgins, J. P., Whitehead, A. & Simmonds, M. Sequential methods for random-effects meta-analysis. *Stat Med* **30**, 903–921, doi: 10.1002/sim.4088 (2011).
- 56. Thorlund, K. et al. The number of patients and events required to limit the risk of overestimation of intervention effects in meta-analysis--a simulation study. PLoS One 6, e25491, doi: 10.1371/journal.pone.0025491 (2011).
- 57. Wetterslev, J., Thorlund, K., Brok, J. & Gluud, C. Estimating required information size by quantifying diversity in random-effects model meta-analyses. *BMC Med Res Methodol* **9**, 86, doi: 10.1186/1471-2288-9-86 (2009).
- 58. Holst, L. B., Petersen, M. W., Haase, N., Perner, A. & Wetterslev, J. Restrictive versus liberal transfusion strategy for red blood cell transfusion: systematic review of randomised trials with meta-analysis and trial sequential analysis. *BMJ* 350, h1354 (2015).
- 59. Nishi, A. et al. Meta-analyses of blood homocysteine levels for gender and genetic association studies of the MTHFR C677T polymorphism in schizophrenia. Schizophr Bull 40, 1154–1163, doi: 10.1093/schbul/sbt154 (2014).
- Pichler, I. et al. Serum iron levels and the risk of Parkinson disease: a Mendelian randomization study. PLoS Med 10, e1001462, doi: 10.1371/journal.pmed.1001462 (2013).
- 61. Greco, M. F., Minelli, C., Sheehan, N. A. & Thompson, J. R. Detecting pleiotropy in Mendelian randomisation studies with summary data and a continuous outcome. *Stat Med* 34, 2926–2940, doi: 10.1002/sim.6522 (2015).
- 62. Burgess, S. Sample size and power calculations in Mendelian randomization with a single instrumental variable and a binary outcome. *Int J Epidemiol* 43, 922–929, doi: 10.1093/ije/dyu005 (2014).

Author Contributions

Y.X. and D.-J.H. conceived this work and interpreted the data. X.-H.L. and D.-J.H. revised the manuscript. Y.X. and Z.-Q.H. wrote the paper. Y.X., Z.-Q.H. and Z.-M.T. prepared tables 1 and 2 and figures 1–5. All authors reviewed this manuscript and approved the final draft.

Additional Information

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