scientific reports



OPEN Gender differences in the association between body mass index and recent suicide attempts in Chinese patients with first-episode drug-naïve major depressive disorder

Junjun Liu^{1,2,3,6}, Zhe Li^{2,6}, Fengnan Jia^{2,3,6}, Hsinsung Yuan¹, Yue Zhou^{2,4}, Xingzhi Xia^{2,4}, Ruchang Yang^{2,3}, Yuxuan Wu^{2,3}, Xiaobin Zhang², Gang Ye², Xiangdong Du^{2,3™} & Xiangyang Zhang^{5⊠}

Controversial evidence exists on the relationship between body mass index (BMI) and suicide attempts (SA) in patients with major depressive disorder (MDD). This cross-sectional study aimed to explore the association between BMI and SA in first-episode drug-naïve (FEDN) MDD patients in China. The study was conducted from 2016 to 2018 in Taiyuan, China. Univariate and multivariate logistic regression analyzed the BMI-SA association, with subgroup analysis for gender. Threshold effects were examined using two-piecewise regression. In males, BMI was significantly associated with SA (OR = 0.84, 95% CI 0.74-0.94, P=0.003) after full adjustment, but not in females (OR = 0.97, 95% CI 0.89-1.06, P = 0.541). The interaction with gender was significant (P for interaction < 0.05). Smoothing plots revealed an L-shaped BMI-SA relationship in both genders, with BMI inflection points at 27.3 kg/m² in males and 21.4 kg/m² in females. Below the inflection points, BMI is negatively associated with SA in males (OR = 0.75, 95% CI 0.66 - 0.86, P < 0.001) and females (OR = 0.48, 95% CI 0.32 - 0.72, P < 0.001). Above the inflection points, no association existed for both genders (all P > 0.05). Results showed an L-shaped nonlinear BMI-SA relationship in FEDN MDD patients but differing BMI inflection points between genders, thus contributing to effective prevention programs for suicide.

Major depressive disorder (MDD) is a common serious mental disorder characterized by persistent depressed mood, loss of interest and motivation, sleep disturbance, changes in appetite, feelings of worthlessness or guilt, and even recurrent suicidal thoughts and behaviors¹. The World Health Organization (WHO) ranked MDD as the third leading contributor to global disease burden in 2017, with predictions that it will become second by 2030². At its most severe, depression can lead to suicide, which caused over 700,000 deaths worldwide in 2019, representing approximately 1.3% of all deaths³. A previous study reported that 63% of suicide decedents had MDD at the time of death⁴, with a lifetime prevalence of suicide attempts (SA) in people with MDD as high as 23.7%⁵. Moreover, past SA strongly predicts completed suicides and is associated with poorer long-term out-

The causes of SA in patients with MDD are complex, involving numerous biological, psychological, social, and environmental factors⁷. The epidemiology of suicide in depressed patients varies widely, influenced by factors including earlier onset⁸, male gender⁹, depressive severity¹⁰, past suicide attempts (SA)¹¹, drug or alcohol

¹Nanjing Meishan Hospital, Nanjing 210041, People's Republic of China. ²Department of Psychiatry, Suzhou Guangji Hospital, The Affiliated Guangji Hospital of Soochow University, Suzhou 215137, People's Republic of China. ³Medical College of Soochow University, Suzhou 215137, People's Republic of China. ⁴Xuzhou Medical University, Xuzhou 221004, People's Republic of China. ⁵CAS Key Laboratory of Mental Health, Institute of Psychology, Chinese Academy of Sciences, Beijing 100000, People's Republic of China. ⁶These authors contributed equally: Junjun Liu, Zhe Li and Fengnan Jia. [™]email: xiangdong-du@163.com; zhangxy@psych.ac.cn

abuse or dependence¹², physical illness¹³, and physical inactivity¹⁴. In addition, some objective indicators, such as body mass index (BMI), an indicator of an unhealthy lifestyle, nutrition, physical activity, an unbalanced intake of energy, and insulin resistance, have also been the focus of researchers¹⁵. Factors affecting BMI may also influence neurodevelopment and the neuroendocrine system, which can affect the risk of developing mental illness¹⁶. However, the relationship between BMI and adult SA in patients of different genders is still unclear, with controversial research results. Some studies found low BMI associated with increased suicide risk in men, while high BMI correlated with suicidal ideation in women^{17–19}. One study showed increased BMI in women linked to depression and suicidal thoughts, while lower BMI in men tied to depression, attempts, and ideation²⁰. Another study found that obesity was associated with a higher suicide risk in women but not in men²¹. Inconsistencies across studies may stem from differing ethnicities, populations, hospitalization histories, antidepressant use, or other factors, making it difficult to fully understand or compare the results.

First-episode drug-naïve (FEND) MDD patients were defined as having a first symptomatic episode but no prior antidepressant or other medication use²². Patients with FEND MDD provide a unique opportunity to examine the relationship between SA and BMI in different genders without lifestyle or pharmacotherapy confounders like disease course or medical comorbidities. However, to our knowledge, no studies have specifically examined the association between BMI and SA in patients with FEND MDD, especially in different sex groups. Therefore, this cross-sectional study aimed to assess the relationship between BMI and SA in a relatively large sample of FEND MDD patients in China, stratified by gender.

Materials and methods Study population

The study protocol adhered to the Declaration of Helsinki and was approved by the ethics committee of the First Hospital of Shanxi Medical University (No. 2016-Y27), with all participants providing written informed consent. All research was performed in accordance with relevant guidelines and regulations.

This cross-sectional study was conducted at the First Hospital of Shanxi Medical University (Taiyuan, Shanxi Province, China) from September 2016 to December 2018. Inclusion criteria were: (1) first-episode patients with MDD without medication; (2) age between 18 and 60 years; and (3) clinical diagnosis of MDD in accordance with the structured clinical interview of the Diagnostic and Statistical Manual of Mental Disorders, Fourth Edition (DSMIV). Exclusion criteria included the following: (1) serious physical illness such as cancer, persistent infection, epilepsy, brain injury, stroke, and diabetes (n=9); (2) pregnant or breastfeeding (n=10); (3) alcohol or drug abuse (n=9); (4) severe personality disorder or other mental disorders that affect BMI, such as eating disorders (n=15); (5) refusal to participate (n=21); (7) unable to be interviewed due to an acute clinical condition (n=5); and (8) other unknown reasons (n=9) (Fig. 1). A total of 1796 patients were recruited and diagnosed by two experienced psychiatrists, with 78 excluded, leaving 1,718 participants (588 males and 1130 females).

Socio-demographic characteristics and anthropometric data

Sociodemographic and clinical data were collected via a structured questionnaire, including age, sex, education, marital status, age of onset, and duration of illness. Body weight, height, systolic blood pressure (SBP), and diastolic blood pressure (DBP) were measured using standard calibrated protocols. Body mass index (BMI) was calculated as weight (kg) divided by height squared (m²).

Suicide attempt history was assessed through a face-to-face interview with the question "Have you ever attempted suicide in your lifetime?", derived from the WHO/EURO multicenter study²³. Patients with FEND MDD who responded "yes" to this inquiry were regarded as having attempted suicide. Then, we probed more into the frequency, mode, and precise dates of suicide attempts. Totally, 346 MDD patients were reported to have SA within the previous month. Of which, one patient attempted suicide four times, two patients three times, 26 patients twice, and 317 patients once. Participants were stratified into suicide attempt and non-attempt groups by gender based on their responses.

Clinical assessment

Two qualified psychiatrists used the Chinese version of the Modified Structured Clinical Interview for DSM-IV-TR criteria (SCID-I/P) to provide a consensus diagnosis of MDD for each patient.

The 17-item Hamilton Depression (HAMD-17) Chinese version was used to assess depression severity. It contains 17 items, with 8 scored on a 5-point scale from 0 (absent) to 4 (severe) and 9 scored from 0 (absent) to 2 (symptom-specific severity). Total scores range from 0 to 52, with higher scores indicating more severe depression. The HAMD-17 has demonstrated strong validity and reliability in China and has been widely used in previous research²⁴.

The Hamilton Anxiety Rating Scale (HAMA) Chinese version was used to evaluate anxiety symptom severity²⁵. It consists of 14 symptom-defined items scored on a 5-point Likert scale from 0 (absence) to 4 (severe). Total scores range from 0 to 56, encompassing both somatic anxiety (e.g., physical discomfort) and psychogenic anxiety (e.g., mental agitation, psychological distress).

Two certified psychiatrists with over 5 years of clinical experience received pre-study training on the use of the rating scales. Inter-observer reliability for HAMA and HAMD total scores post-training remained above 0.8 with repeated assessments. Additionally, they were unaware of the patients' clinical conditions.

Blood samples

Blood samples were taken from each participant between the hours of 6:00 and 9:00 am after an overnight fast. On the same day, all samples were submitted to the hospital's laboratory facility and measured before 11:00 a.m. Using an automated clinical ARCHITECT Immulite 2000 SR analyzer, the following hormones were measured:

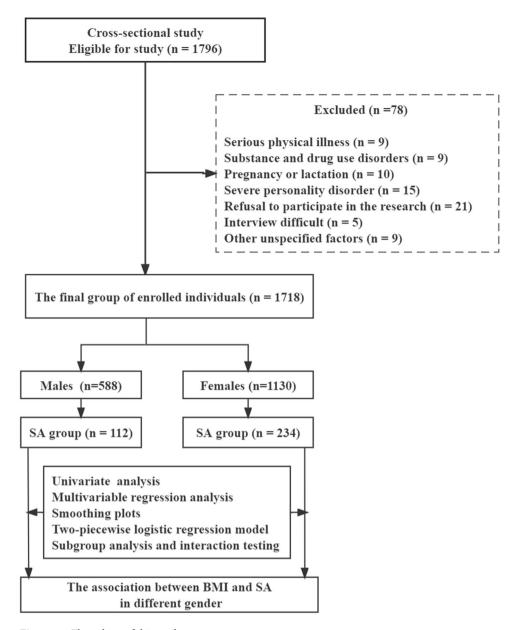


Figure 1. Flow chart of this study.

free triiodothyronine (FT3), free thyroxine (FT4), thyroid stimulating hormone (TSH), antithyroglobulin (TgAb), and thyroid peroxidase antibody (TPOAb) (Abbott, Longford, Ireland). An auto-analyzer was used to measure the levels of fasting blood sugar and the lipid profile, comprising low-density lipoprotein cholesterol (LDL-C), high-density lipoprotein cholesterol (HDL-C), cholesterol (TC), and triglyceride (TG) (ARCHITECT c8000 system, Abbott Laboratories, Illinois, USA).

Statistical analysis

Continuous variables were assessed for normality using the Kolmogorov–Smirnov test. Normally distributed variables were summarized as mean ± standard deviation (SD), and non-normally distributed variables as median and interquartile range (IQR). Categorical variables were expressed as frequencies and percentages (%). Characteristic differences between male and female patients were examined using independent t-tests and Mann–Whitney U tests for continuous data and chi-square tests for categorical data. The linear and nonlinear relationships between SA and BMI were evaluated using multivariate linear regression models and two-piece piecewise regression models, adjusting for age at onset, length of illness, education, HAMA, HAMD, TSH, TGAb, TPOAb, FBG, TC, TG, HDL-c, TG, LDL-c, systolic pressure, and diastolic pressure in men and women. Stratified and interaction analyses were further conducted to explore the potential modifier and interaction effects on the SA-BMI association. Variables with a variance inflation factor (VIF) > 5 were excluded from the final model to avoid multicollinearity.

Covariates were also screened and included if they significantly altered the BMI–SA estimate by > 10% or were associated with SA $(P < 0.10)^{26}$. Log-likelihood ratio tests compared differences between regression models.

The statistical software programs R 4.3.0 (http://www.r-project.org, The R Foundation) and EmpowerStats (http://www.empowerstats.com, X&Y Solution, Inc., Boston, Massachusetts, USA) were used for all analyses. Statistical significance was defined as a two-tailed *p*-value < 0.05.

Results

Characteristics of the participants

A total of 1,718 FEDN MDD patients met the criteria for enrollment and were included in this study, with 588 males and 1130 females, ages 34.87 ± 12.43 years (range: 18-60 years), and a mean disease duration of 5.00 (3.00, 8.00) months. The overall mean level of BMI was $(24.37 \pm 1.92 \text{ kg/m}^2)$, $(24.41 \pm 2.04 \text{ kg/m}^2)$ among men, and $(24.35 \pm 1.86 \text{ kg/m}^2)$ among women. The prevalence of SA in male and female patients was 19.05% (112/588) and 20.71% (234/1130), respectively (Table 1).

Female patients had higher levels of age, age at onset, systolic pressure, and were more likely to be married than male patients (all P < 0.05), but their duration of illness (months), HAMD, HAMA, TSH (uIU/ml), TGAb (IU/l), TPOAb (IU/l), FT3 (pmol/l), FT4 (pmol/l), FBG (mmol/l), TC (mmol/l), HDL-c (mmol/l), TG (mmol/l), LDL-c (mmol/l), BMI (kg/m²), and DBP (mmHg) were not significantly different (all P > 0.05). After Bonferroni corrections, there was no significant difference in education between different genders (P = 0.009, Bonferroni corrected P < 0.05/6 = 0.0083).

Univariate logistic analysis of SA in males and females

The univariate logistic regression analysis was performed by taking SA as the dependent variable (Y = 1) and using the clinical data and blood indicators from different gender groups as independent variables (Table 2). The results showed that HAMD, HAMA, TSH (uIU/ml), TGAb (IU/l), TPOAb (IU/l), FBG (mmol/l), TC (mmol/l), HDL-c (mmol/l), BMI (kg/m2), systolic pressure (mmHg), and diastolic pressure (mmHg) were all significantly related to SA in male patients (P all < 0.05). In female patients, the association between

| Characteristics | Total (n = 1718) | Males (n = 588) | Females (n = 1130) | P-value |
|------------------------------|----------------------|----------------------|----------------------|---------|
| Age (years) | 34.87 ± 12.43 | 33.12 ± 12.22 | 35.78 ± 12.45 | < 0.001 |
| Age at onset (years) | 34.66 ± 12.31 | 32.91 ± 12.09 | 35.57 ± 12.34 | < 0.001 |
| Duration of illness (months) | 5.00 (3.00, 8.00) | 5.00 (3.00, 8.00) | 5.00 (3.00, 8.00) | 0.363 |
| HAMD | 30.30 ± 2.94 | 30.28 ± 2.95 | 30.30 ± 2.94 | 0.883 |
| HAMA | 20.80 ± 3.47 | 20.66 ± 3.38 | 20.87 ± 3.52 | 0.231 |
| TSH (uIU/ml) | 5.07 ± 2.56 | 4.95 ± 2.48 | 5.13 ± 2.60 | 0.157 |
| TGAb (IU/l) | 21.46 (14.43, 43.62) | 21.19 (14.23, 37.30) | 21.80 (14.46, 46.81) | 0.355 |
| TPOAb (IU/l) | 17.43 (12.32, 34.61) | 16.81 (12.48, 32.86) | 18.20 (12.24, 35.27) | 0.590 |
| FT3 (pmol/l) | 4.91 ± 0.72 | 4.95 ± 0.72 | 4.88 ± 0.72 | 0.055 |
| FT4 (pmol/l) | 16.70 ± 3.10 | 16.72 ± 3.15 | 16.69 ± 3.07 | 0.838 |
| FBG (mmol/l) | 5.40 ± 0.65 | 5.37 ± 0.64 | 5.41 ± 0.65 | 0.229 |
| TC (mmol/l) | 5.25 ± 1.11 | 5.22 ± 1.06 | 5.26 ± 1.13 | 0.429 |
| HDL-c (mmol/l) | 1.22 ± 0.29 | 1.22 ± 0.30 | 1.22 ± 0.28 | 0.629 |
| TG (mmol/l) | 2.17 ± 0.99 | 2.14 ± 1.00 | 2.18 ± 0.98 | 0.497 |
| LDL-c (mmol/l) | 2.98 ± 0.86 | 2.96 ± 0.85 | 3.00 ± 0.87 | 0.428 |
| BMI (kg/m²) | 24.37 ± 1.92 | 24.41 ± 2.04 | 24.35 ± 1.86 | 0.539 |
| Systolic pressure (mmHg) | 119.48 ± 10.91 | 118.76 ± 10.40 | 119.86 ± 11.16 | 0.047 |
| Diastolic pressure (mmHg) | 75.95 ± 6.74 | 75.90 ± 6.89 | 75.98 ± 6.67 | 0.820 |
| Suicide attempts | 346(20.14%) | 112 (19.05%) | 234 (20.71%) | 0.416 |
| Marital status | | | | 0.007 |
| Single | 502 (29.22%) | 196 (33.33%) | 306 (27.08%) | |
| Marriage | 1216 (70.78%) | 392 (66.67%) | 824 (72.92%) | |
| Education | | | | 0.009 |
| Junior high school | 413 (24.04%) | 125 (21.26%) | 288 (25.49%) | |
| Senior high school | 760 (44.24%) | 271 (46.09%) | 489 (43.27%) | |
| College | 449 (26.14%) | 170 (28.91%) | 279 (24.69%) | |
| Postgraduate | 96 (5.59%) | 22 (3.74%) | 74 (6.55%) | |

Table 1. Comparison of socio-demographic and clinical characteristics between male and female patients with FEDN MDD. FEDN, first-episode drug-naïve; MDD, major depressive disorder; HAMD, 17-item Hamilton Rating Scale for Depression; HAMA, 14-item Hamilton Anxiety Rating Scale; TSH, thyroid-stimulating hormone; TGAb, thyroglobulin antibody; TPOAb, thyroid peroxidase antibody; FT3, free triiodothyronine; FT4, free thyroxine; FBG, fasting blood glucose; TC, total cholesterol; HDL-c, high-density lipoprotein cholesterol; TG, triglyceride; LDL-c, low-density lipoprotein cholesterol; BMI, body mass index.

| | Males (n = 588) | | Females (n = 1130) | | |
|------------------------------|-------------------|----------|--------------------|---------|--|
| Covariate | OR (95% CI) | P-value | OR (95% CI) | P-value | |
| Age (years) | 1.01 (1.00, 1.03) | 0.156 | 1.01 (1.00, 1.02) | 0.139 | |
| Age at onset (years) | 1.01 (1.00, 1.03) | 0.161 | 1.01 (1.00, 1.02) | 0.144 | |
| Duration of illness (months) | 1.03 (0.99, 1.08) | 0.103 | 1.03 (1.00, 1.06) | 0.031 | |
| HAMD | 1.31 (1.21, 1.41) | < 0.001 | 1.39 (1.31, 1.48) | < 0.001 | |
| HAMA | 1.34 (1.25, 1.43) | < 0.001 | 1.37 (1.31, 1.44) | < 0.001 | |
| TSH (uIU/ml) | 1.40 (1.27, 1.54) | < 0.001 | 1.38 (1.29, 1.46) | < 0.001 | |
| TGAb (IU/l) | 1.00 (1.00, 1.00) | 0.003 | 1.00 (1.00, 1.00) | < 0.001 | |
| TPOAb (IU/l) | 1.00 (1.00, 1.00) | < 0.001 | 1.00 (1.00, 1.00) | < 0.001 | |
| FT3 (pmol/l) | 0.84 (0.63, 1.11) | 0.225 | 1.13 (0.92, 1.38) | 0.233 | |
| FT4 (pmol/l) | 1.01 (0.95, 1.08) | 0.771 | 0.99 (0.94, 1.04) | 0.672 | |
| FBG (mmol/l) | 2.04 (1.48, 2.80) | < 0.001 | 1.60 (1.29, 1.99) | < 0.001 | |
| TC (mmol/l) | 1.87 (1.52, 2.30) | < 0.001 | 1.67 (1.46, 1.91) | < 0.001 | |
| HDL-c (mmol/l) | 0.29 (0.14, 0.60) | < 0.001 | 0.27 (0.16, 0.46) | < 0.001 | |
| TG (mmol/l) | 1.12 (0.92, 1.36) | 0.276 | 1.20 (1.04, 1.38) | 0.013 | |
| LDL-c (mmol/l) | 1.60 (1.25, 2.04) | < 0.001 | 1.39 (1.18, 1.64) | < 0.001 | |
| BMI (kg/m²) | 0.89 (0.80, 0.99) | 0.031 | 1.04 (0.96, 1.13) | 0.299 | |
| Systolic pressure (mmHg) | 1.04 (1.02, 1.07) | < 0.0001 | 1.06 (1.05, 1.08) | < 0.001 | |
| Diastolic pressure (mmHg) | 1.05 (1.02, 1.08) | < 0.001 | 1.09 (1.06, 1.11) | < 0.001 | |
| Marital status | | | | | |
| Single | 1.0 (reference) | | 1.0 (reference) | | |
| Marriage | 1.46 (0.93, 2.32) | 0.104 | 0.96 (0.69, 1.32) | 0.787 | |
| Education | | | | | |
| Junior high school | 1.0 (reference) | | 1.0 (reference) | | |
| Senior high school | 0.93 (0.54, 1.60) | 0.790 | 0.65 (0.46, 0.91) | 0.014 | |
| College | 1.01 (0.56, 1.82) | 0.964 | 0.58 (0.39, 0.87) | 0.009 | |
| Postgraduate | 1.58 (0.56, 4.46) | 0.389 | 0.90 (0.50, 1.62) | 0.718 | |

Table 2. Univariate analysis for Suicide attempts in male and female patients with FEDN MDD. CI, confidence interval. OR, odds ratio.

duration of illness, HAMD, HAMA, TSH (uIU/ml), TGAb (IU/l), TPOAb (IU/l), FBG (mmol/l), TC (mmol/l), HDL-c (mmol/l), TG (mmol/l), systolic pressure (mmHg), diastolic pressure (mmHg), and SA was significant (all P < 0.05). Compared with junior high school patients, senior high school and college patients were negatively associated with SA (all P < 0.05).

Linear relationship between SA and BMI in males and females

Multivariable logistic regression analysis was used to evaluate the linear association between BMI and SA by adjusting the covariates. The model with non-adjusted covariates equaled univariate logistic regression analysis. The minimally adjusted covariates (Adjust I) included age at onset, duration of illness, and education, and the fully adjusted covariates (Adjust II) included age at onset, duration of illness, education, HAMA, HAMD, TSH, TGAb, TPOAb, FBG, TC, TG, HDL-c, TG, LDL-c, systolic pressure, and diastolic pressure (Table 3). For male patients, an increase in BMI elevated the risk of SA in all three models (all P < 0.05), with an OR = 0.84 (95% CI 0.74–0.94, P = 0.003) in the fully-adjusted model. For female patients, the association between BMI and SA was not significant in all three models (all P > 0.05). The interaction effect between different genders was significant (all P for interaction < 0.05), indicating that the relationship between BMI and SA was affected by gender.

For sensitivity analysis, we converted BMI from a continuous variable to a categorical variable. Based on guidelines from the Working Group on Obesity in China (WGOC)¹⁶, participants were classified into BMI categories of underweight ($<18.5 \text{ kg/m}^2$), normal weight ($18.5-23.9 \text{ kg/m}^2$), overweight ($24-27.9 \text{ kg/m}^2$), or obesity ($\ge 28 \text{ kg/m}^2$). Compared to those with normal weight, overweight males exhibited reduced odds of SA (OR = 0.38, 95% CI 0.22-0.64, P<0.001), whereas underweight females showed increased odds of SA (OR = 7.85, 95% CI 1.47-41.84, P=0.016). However, the small sample size of just 7 underweight females limits the statistical power; thus, the elevated SA risk associated with being underweight in females should be interpreted cautiously.

Nonlinear relationship between SA and BMI in males and females

An L-shaped relationship between SA and BMI was observed for both male and female patients by a piecewise regression model, and the estimated inflection point was 27.3 kg/m^2 for males and 21.4 kg/m^2 for females (Fig. 2 and Table 4). For men, once the BMI was lower than 27.3 kg/m^2 , a significantly negative association between BMI and SA was found (OR = 0.75, 95% CI 0.66-0.86, P < 0.001), while no significant BMI–SA association was detected (OR = 1.71, 95% CI 0.93-3.15, P = 0.087) if the BMI was higher than 27.3 kg/m^2 . In women, the risk

| Variable | | Non-Adjusted | | Adjusted I | | Adjusted II | |
|-----------------------------|------|--------------------------|---------|--------------------------|---------|--------------------------|---------|
| | n | OR (95% CI) | P-value | OR (95% CI) | P-value | OR (95% CI) | P-value |
| Males | | <u>'</u> | | ' | | ' | ' |
| BMI (kg/m²) | 558 | 0.89 (0.80, 0.99) | 0.031 | 0.88 (0.79, 0.98) | 0.020 | 0.84 (0.74, 0.94) | 0.003 |
| BMI category | | | | | | | |
| <18.5 kg/m ² | 3 | 6,846,260.28 (0.00, Inf) | 0.975 | 7,748,200.55 (0.00, Inf) | 0.975 | 8,879,381.13 (0.00, Inf) | 0.984 |
| 18.5-23.9 kg/m ² | 237 | 1 (reference) | | 1 (reference) | | 1 (reference) | |
| 24- 27.9 kg/m ² | 320 | 0.56 (0.36, 0.86) | 0.008 | 0.53 (0.34, 0.82) | 0.005 | 0.38 (0.22, 0.64) | < 0.001 |
| ≥28 kg/m ² | 28 | 0.88 (0.34, 2.28) | 0.795 | 0.85 (0.32, 2.26) | 0.744 | 0.44 (0.14, 1.37) | 0.158 |
| Females | | | | | | | |
| BMI (kg/m²) | 1130 | 1.04 (0.96, 1.13) | 0.299 | 1.04 (0.96, 1.12) | 0.326 | 0.97 (0.89, 1.06) | 0.541 |
| BMI category | | | | | | | |
| <18.5 kg/m ² | 7 | 5.90 (1.30, 26.88) | 0.022 | 4.92 (1.06, 22.72) | 0.041 | 7.85 (1.47, 41.84) | 0.016 |
| 18.5-23.9 kg/m ² | 445 | 1 (reference) | | 1 (reference) | | 1 (reference) | |
| 24-27.9 kg/m ² | 642 | 1.21 (0.89, 1.64) | 0.216 | 1.18 (0.86, 1.60) | 0.301 | 1.01 (0.70, 1.45) | 0.963 |
| ≥ 28 kg/m ² | 36 | 1.70 (0.79, 3.67) | 0.174 | 1.68 (0.77, 3.64) | 0.190 | 1.26 (0.51, 3.11) | 0.611 |
| P for interaction | | 0.018 | | 0.019 | | 0.029 | |

Table 3. Relationship between BMI and SA in different models. BMI, body mass index; CI, confidence interval. Model I adjusted for age at onset, duration of illness, education. Model II adjusted for age at onset, duration of illness, education, HAMA, HAMD, TSH, TGAb, TPOAb, FBG, TC, TG, HDL-c, LDL-c, systolic pressure, diastolic pressure.

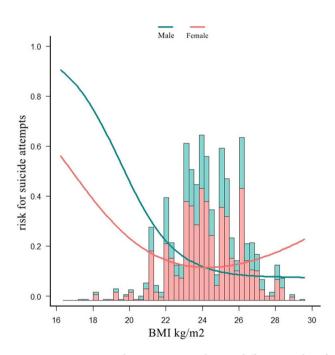


Figure 2. Association between BMI and SA in different gender of patients with FEDN MDD. A non-linear association between BMI and SA was found in the multivariable logistic regression model in both male and female patients (all *P* for non-linearity < 0.05). All adjusted for age at onset, duration of illness, education, HAMA, HAMD, TSH, TGAb, TPOAb, FBG, TC, TG, HDL-c, TG, LDL-c, systolic pressure, diastolic pressure.

of SA was negatively associated with BMI (OR = 0.48, 95% CI 0.32–0.72, P < 0.001) if the BMI was lower than 21.4 kg/m², but no significant BMI–SA association (OR = 1.06, 95% CI 0.96–1.16, P = 0.255) was detected when the BMI was higher than 21.4 kg/m². The dose–response relationship was agreed upon with the threshold effect analysis, and P was < 0.001 for the log-likelihood ratio test (Table 4) in both gender groups, indicating a nonlinear association between BMI and SA.

| Inflection point of BMI (kg/m²) | OR | 95% CI | P-value |
|---------------------------------|------|-----------|---------|
| Males | | | |
| <27.3 | 0.75 | 0.66-0.86 | < 0.001 |
| > = 27.3 | 1.71 | 0.93-3.15 | 0.087 |
| Log likelihood ratio test | | | < 0.001 |
| Females | | | |
| <21.4 | 0.48 | 0.32-0.72 | < 0.001 |
| >=21.4 | 1.06 | 0.96-1.16 | 0.255 |
| Log likelihood ratio test | | | < 0.001 |

Table 4. Threshold effect of body mass index and Suicide attempts using piecewise logistic regression model in different sex. All adjusted for age at onset, duration of illness, education, HAMA, HAMD, TSH, TGAb, TPOAb, FBG, TC, TG, HDL-c, LDL-c, systolic pressure, diastolic pressure.

Exploration of modifier and interaction effects on the SA-BMI association

We explored potential modifier or interaction effects from sex, education, and marital status, and found that only sex might be a potential interaction factor for the SA-BMI association (*P* for interaction = 0.029) (Fig. 3).

To elucidate the relationship between BMI and SA risk in specific demographic groups, male and female patients were stratified by age (<20 years, 20-29 years, 30-39 years, >=40 years) for subgroup analyses (Table 5). Among males, BMI exhibited a negative correlation with SA risk in the 20-29 years age group (OR = 0.51, 95% CI 0.35-0.74, P<0.001), with a marginally significant interaction test (P for interaction = 0.013). For females, BMI also showed a negative correlation with SA risk, but specifically in the <20 years age group (OR = 0.63, 95% CI 0.41-0.96, P=0.033), and the interaction test was statistically significant (P for interaction = 0.071).

Discussion

To our knowledge, this is the first study to examine the nonlinear association between BMI and SA in a relatively large sample of FEDN MDD patients, stratified by gender. Our study found that this relationship was significantly moderated by gender. A L-shaped nonlinear correlation between BMI and SA was observed in both genders, with a threshold effect at inflection points of 27.3 kg/m² in males and 21.4 kg/m² in females. BMI below the inflection point was negatively associated with SA in both genders, while no significant association was found above the inflection point.

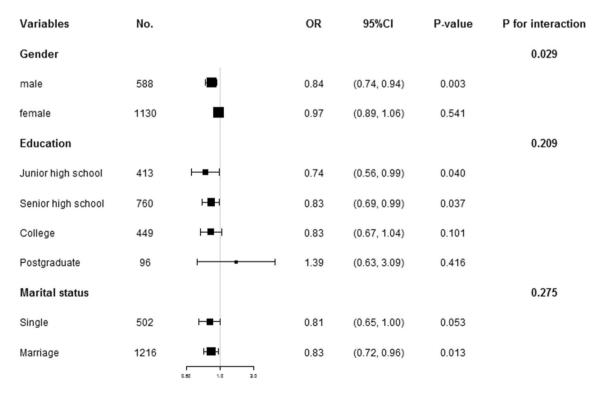


Figure 3. Stratified and interaction analyses of the association between BMI and SA. The OR (95% CI) was derived from the Logistic regression model. (Age at onset, duration of illness, HAMA, HAMD, TSH, TGAb, TPOAb, FBG, TC, TG, HDL-c, TG, LDL-c, SBP and DBP were adjusted).

| | | | 1 | |
|-------------------|-----|------|-----------|---------|
| | n | OR | 95% CI | P-value |
| Males (years) | | | , | |
| < 20 | 98 | 1.13 | 0.87-1.48 | 0.353 |
| 20-29 | 172 | 0.51 | 0.35-0.74 | < 0.001 |
| 30-39 | 130 | 0.85 | 0.61-1.19 | 0.348 |
| > = 40 | 188 | 0.80 | 0.63-1.03 | 0.084 |
| P for interaction | | | | 0.013 |
| Females (years) | | | | |
| < 20 | 131 | 0.63 | 0.41-0.96 | 0.033 |
| 20-29 | 288 | 1.03 | 0.83-1.26 | 0.805 |
| 30-39 | 237 | 1.05 | 0.84-1.31 | 0.688 |
| > = 40 | 474 | 0.96 | 0.83-1.12 | 0.635 |
| P for interaction | | | | 0.071 |

Table 5. Subgroup analysis in different age groups between BMI and SA. All adjusted for age at onset, duration of illness, education, HAMA, HAMD, TSH, TGAb, TPOAb, FBG, TC, TG, HDL-c, LDL-c, systolic pressure, diastolic pressure.

Accumulating evidence suggests that underweight is a common comorbidity of SA, with the risk increasing as BMI decreases²⁷. In a meta-analysis, Perera et al.²⁸ found that among adult populations (>= 18 years old), underweight people had the highest risk of committing suicide compared to those of normal weight, but obesity and overweight people were substantially related to a lower risk of suicide. In large Swedish cohort research, Batty et al.29 showed an increased risk of SA in underweight men, and they also found that low weight was inversely correlated with SA within the normal BMI range. Individuals with a low BMI are predisposed to diminished physical functioning, which in turn is associated with emotional distress and more severe depressive symptoms¹⁶. A study conducted in the United States found that among men with MDD, a lower BMI was correlated with increased depressive symptom severity³⁰. A recent systematic review by Amiri et al. explored the association between obesity and suicide over the past decade and showed an inverse association between overweight and obesity and SA and suicide mortality, while the association between overweight and obesity was positively associated with suicidal ideation³¹. Research by Carpenter et al.³² showed that a lower BMI or different weight status was associated with the probability of past-year SA and suicidal ideation in a sample of the general adult male population in the United States. This phenomenon, known as the obesity paradox, is supported by mounting evidence that a higher BMI may be protective in some patient populations³². Our findings provide credence to the obesity paradox in FEDN MDD patients. Samaan et al.³³ found that the adiposity and obesity-associated (FTO) gene rs9939609 obesity-risk variant was paradoxically inversely associated with the risk of depression at the observational level, supporting the idea that the relationship between obesity and MDD can be explained, at least in part, by shared genetic factors. Adipose tissue is now recognized not only as an energy reservoir but also as a dynamic endocrine organ that secretes many adipocytokines with high metabolic activity. These cytokines are associated with metabolic anomalies and neurohormonal activity, including sympathetic nervous system activation, renin-angiotensin-aldosterone system activation, and leptin regulation³⁴. Leptin levels may be linked to a higher risk of suicidal behavior, according to a recent meta-analysis³⁵. Westling et al.³⁶ reported that leptin levels in cerebrospinal fluid (CSF) were involved in neuroendocrine dysfunction in MDD patients who made SA. In addition to biological factors, current sociocultural influences on body image, particularly social media, further propagate an attentional focus on modifying specific aspects of one's physique (e.g., enhancing muscularity, decreasing body fat, altering shape)³⁷. This hyper-focused scrutiny of certain body regions can potentiate anxiety and distress surrounding particular features among adult women. It is hypothesized that such societal pressures, coupled with ones own motivations for thinness, may engender psychological distress in individuals at weight extremes compared to population norms³⁸. Meanwhile, in this study, a similar L-shaped non-linear relationship between BMI and SA was determined for male or female patients, and there is no incremental advantage when BMI is greater than the threshold point.

However, evidence on the BMI–SA association is mixed. Some research has shown no association between BMI and SA³⁹, while others have reported a positive association. Zhang et al.⁴⁰ discovered that among 104,907 adolescents aged 12–15 years in 45 low- and middle-income countries (LMICs), being overweight and obese were significantly associated with a higher risk of SA. According to a recent review by Heneghan et al., Americans who are obese are more likely to commit suicide⁴¹. Individuals with bipolar disorder showed a similar association, with obese patients being almost twice as likely to have attempted suicide in the past as patients of normal weight⁴². These findings that are inconsistent with the results of this study may be due to the following factors: (1) Previous studies only focused on the linear relationship between SA and BMI, while our results examine the non-linear relationship for both men and women. (2) The participants in this study are special since the study population was made up of FEDN MDD patients, as opposed to the general community or individuals who had received depression treatment in many other studies. (3) The cultural and ethnic differences (e.g., lifestyle, beliefs, values, genetic factors, etc.) between the West and the East may have an impact on the prevalence of obesity, depression, and SA in the population⁴³. For instance, the well-known Chinese proverb "laugh and gain fat" refers to the relationship between enjoyment and obesity. The Chinese consider that being overweight is a

sign of good fortune because lucky people can eat more and put on more weight. However, discrimination is frequently associated with being overweight in Western nations. (4) The study population's low average BMI of $24.37 \pm 1.92 \text{ kg/m}^2$ and the limited sample size of obese participants (BMI > = 28 kg/m^2 , 64 in total) may have some bearing on the study's findings.

Sex differences are observed in both clinical and epidemiological features of MDD. According to studies, women are more likely than men to experience MDD⁴⁴. Females with MDD are more likely to exhibit increased anxiety, along with higher rates of comorbid anxiety disorders, whereas males with MDD are more likely to show increased psychomotor agitation and to have comorbid substance abuse disorders⁴⁵. Males with MDD are more likely to succeed in their attempts at suicide than females, despite the fact that females with MDD are more likely to attempt suicide⁴⁶. This study found that although there was an L-shaped nonlinear correlation between BMI and SA in both male and female patients with depression, there were differences in inflection points, with 27.3 kg/m² in males and 21.4 kg/m² in females, which has not been reported in previous studies. Although there were gender differences in the relationship between BMI and SA in patients with FEDN MDD, the mechanisms behind these observations remain unclear, and several studies indicate that differences in sex hormone levels may partially explain the effects. Given that depressive disorders and autoimmune diseases are more common in females, sex hormones appear to be a likely factor in sex differences in cytokine-associated depression⁴⁷. It is possible that sex-specific hormonal substrates contributed to the observed sex differences in cytokine-associated depressed mood at the neurochemical level⁴⁸. Estrogen therapy shows efficacy for perimenopausal depression, supporting interactions between serotonin and gonadal hormones underlying sex differences in the antidepressant response⁴⁹⁻⁵¹. Additionally, estrogen has also been proposed to potentially help explain sex-related differences in depression. Nevertheless, there is evidence to suggest that estrogen may influence mood through regulating various aspects of noradrenergic, serotonergic, GABAergic, and dopaminergic transmission⁵². The sexually dimorphic features in the cortico-limbic-striatal neural system may correspond to the sex differences in clinical observations of MDD and are associated with mood and emotion regulation dysfunction in MDD⁵³. Sex-related patterns of abnormalities within the cortico-limbic-striatal nervous system, such as predominant prefrontal-limbic abnormalities in MDD females versus predominant prefrontal-striatal abnormalities in MDD males, suggest that differences in neural circuitry may mediate gender differences in the clinical presentation of MDD and potential targets for gender-differentiated treatment of the disorder⁵⁴. Finally, because of the different contexts in which men and women develop as children and adolescents, there may be additional cognitive and psychological elements that influence gender disparities⁵². During prenatal and adolescent development, sex hormones may influence a number of behaviors, including psychological changes during puberty (e.g., increased risk-taking behavior in early adolescence, gender differences in internalizing symptoms, i.e., girls have more anxiety and self-esteem problems than boys), gender-typed activity interests (e.g., career choices), gender identity, sexual orientation, cognitive abilities, and behavioral problems⁵⁵. Previous research on cognitive development and the effects of gender coping styles has shown differences in boys' and girls' preferences for coping strategies⁵⁶. In addition, a longitudinal study of depressive cognitions in children and adolescents found that some cognitive differences in girls, including perceptions of appearance and self-worth, were more stable and idiosyncratic than those of boys before gender differences in depression emerged⁵⁷. Female patients in this study had a higher age and age of disease onset than males, which may have influenced the relationship between SA and BMI to some extent, although we have controlled for this in the multifactorial regression model. In the future, it will be necessary to match relevant factors, such as age, to further validate the results of this study.

This study has several limitations. First, patients were recruited from one hospital in Taiyuan, China, and were ethnically Han Chinese. Confirmation in diverse populations is needed. Second, the cross-sectional design cannot determine causal relationships between BMI and SA. Future longitudinal studies should investigate this causal association. Third, SA were identified using interviews and medical records rather than structured diagnostics, without details on severity or ideation. Specific suicide assessments are needed in future work. Fourth, the participants in this study were MDD patients in their first episode who had not received any medication for their depressive symptoms. Even though there was a benefit to selecting first-episode and drug-naïve patients, we cannot entirely exclude patients whose diagnosis may have changed to bipolar disorder since the original depressed episode may have paralleled the MDD episode. Finally, unmeasured confounders like smoking, alcohol use, personality, income, social factors, and biological variables could influence the BMI-suicide attempt relationship. Future research should account for additional confounders to elucidate the underlying mechanisms.

Conclusions

In conclusion, this study showed an L-shaped nonlinear association between BMI and SA in Chinese FEDN MDD patients for both men and women, with inflection points of 27.3 kg/m^2 for men and 21.4 kg/m^2 for women, respectively. The findings provided academics with an exciting psychopathology viewpoint on how to reduce suicide risk in FEDN MDD patients and prevent suicide. Furthermore, because of the study's limitations, such as its cross-sectional design, absence of a structured evaluation instrument, and inability to collect some relevant data, the findings of this investigation should be regarded as preliminary and verified in a follow-up study employing a longitudinal design.

Data availability

The datasets used and/or analysed during the current study available from the corresponding author on reasonable request.

Received: 28 March 2023; Accepted: 20 September 2023

Published online: 25 September 2023

References

- 1. Malhi, G. S. & Mann, J. J. Depression. Lancet 392(10161), 2299-2312. https://doi.org/10.1016/S0140-6736(18)31948-2 (2018).
- 2. World Health Organization. Depression and Other Common Mental Disorders: Global Health Estimates (World Health Organization (WHO), 2017).
- 3. Singh, R. et al. Newspaper reporting of suicide in Nepal: Quality assessment against World Health Organization media guidelines. Health Sci. Rep. 5(2), e547. https://doi.org/10.1002/hsr2.547 (2022).
- 4. Harwood, D., Hawton, K., Hope, T. & Jacoby, R. Psychiatric disorder and personality factors associated with suicide in older people: A descriptive and case-control study. *Int. J. Geriatr. Psychiatry* 16(2), 155–165. https://doi.org/10.1002/1099-1166(200102)16:2% 3c155::aid-gps289%3e3.0.co;2-0 (2001).
- Dong, M. et al. Prevalence of suicidal behaviors in patients with major depressive disorder in China: A comprehensive metaanalysis. J. Affect. Disord. 225, 32–39. https://doi.org/10.1016/j.jad.2017.07.043 (2018).
- Goldman-Mellor, S. J. et al. Suicide attempt in young people: A signal for long-term health care and social needs. JAMA Psychiatry 71(2), 119–127. https://doi.org/10.1001/jamapsychiatry.2013.2803 (2014).
- 7. Ye, G. et al. Suicide attempt rate and the risk factors in young, first-episode and drug-naïve Chinese Han patients with major depressive disorder. BMC Psychiatry 22(1), 612. https://doi.org/10.1186/s12888-022-04254-x (2022).
- 8. Cha, C. B. et al. Annual research review: Suicide among youth epidemiology, (potential) etiology, and treatment. J. Child. Psychol. Psychiatry 59(4), 460–482. https://doi.org/10.1111/jcpp.12831 (2018).
- 9. Hawton, K., I Comabella, C. C., Haw, C. & Saunders, K. Risk factors for suicide in individuals with depression: A systematic review. J. Affect. Disord. 147(1–3), 17–28. https://doi.org/10.1016/j.jad.2013.01.004 (2013).
- Brådvik, L., Mattisson, C., Bogren, M. & Nettelbladt, P. Long-term suicide risk of depression in the Lundby cohort 1947–1997– severity and gender. Acta Psychiatry Scand. 117(3), 185–191. https://doi.org/10.1111/j.1600-0447.2007.01136.x (2008).
- 11. Holma, K. M. et al. Incidence and predictors of suicide attempts in DSM-IV major depressive disorder: A 5-year prospective study. Am. J. Psychiatry 167(7), 801–808. https://doi.org/10.1176/appi.ajp.2010.09050627 (2010).
- 12. Perez, J., Beale, E., Overholser, J., Athey, A. & Stockmeier, C. Depression and alcohol use disorders as precursors to death by suicide. Death Stud. 46(3), 619–627. https://doi.org/10.1080/07481187.2020.1745954 (2022).
- Bergman Levy, T., Barak, Y., Sigler, M. & Aizenberg, D. Suicide attempts and burden of physical illness among depressed elderly inpatients. Arch. Gerontol. Geriatr. 52(1), 115–117. https://doi.org/10.1016/j.archger.2010.02.012 (2011).
- 14. Rönnqvist, I., Nilsson, F. K. & Nordenskjöld, A. Electroconvulsive therapy and the risk of suicide in hospitalized patients with major depressive disorder. *JAMA Netw. Open* 4(7), e2116589. https://doi.org/10.1001/jamanetworkopen.2021.16589 (2021).
- 15. Liu, J. et al. Association between triglyceride glucose index and suicide attempts in patients with first-episode drug-naïve major depressive disorder. Front. Psychiatry 14, 1231524. https://doi.org/10.3389/fpsyt.2023.1231524 (2023).
- Liu, J. et al. Association between body mass index and suicide attempts in Chinese patients of a hospital in Shanxi district with first-episode drug-naïve major depressive disorder. J. Affect. Disord. 339, 377–383. https://doi.org/10.1016/j.jad.2023.06.064 (2023).
- 17. Dong, C., Li, W. D., Li, D. & Price, R. A. Extreme obesity is associated with attempted suicides: Results from a family study. *Int. J. Obes. (Lond).* 30(2), 388–390. https://doi.org/10.1038/sj.ijo.0803119 (2006).
- MacLean, J., Kinley, D. J., Jacobi, F., Bolton, J. M. & Sareen, J. The relationship between physical conditions and suicidal behavior among those with mood disorders. J. Affect. Disord. 130(1-2), 245-250. https://doi.org/10.1016/j.jad.2010.10.028 (2011).
- Mather, A. A., Cox, B. J., Enns, M. W. & Sareen, J. Associations of obesity with psychiatric disorders and suicidal behaviors in a nationally representative sample. *J. Psychosom. Res.* 66(4), 277–285. https://doi.org/10.1016/j.jpsychores.2008.09.008 (2009).
- Carpenter, K. M., Hasin, D. S., Allison, D. B. & Faith, M. S. Relationships between obesity and DSM-IV major depressive disorder, suicide ideation, and suicide attempts: Results from a general population study. Am. J. Public Health. 90(2), 251–257. https://doi.org/10.2105/aiph.90.2.251 (2000).
- 21. Branco, J. C. et al. Association between obesity and suicide in woman, but not in man: A population-based study of young adults. Psychol. Health Med. 22(3), 275–281. https://doi.org/10.1080/13548506.2016.1164870 (2017).
- Peng, P. et al. Association between thyroid hormones and comorbid psychotic symptoms in patients with first-episode and drugnaïve major depressive disorder. Psychiatry Res. 320, 115052. https://doi.org/10.1016/j.psychres.2023.115052 (2023).
- 23. Platt, S. et al. Parasuicide in Europe: The WHO/EURO multicentre study on parasuicide. I. Introduction and preliminary analysis for 1989. Acta Psychiatry Scand. 5(2), 97–104. https://doi.org/10.1111/j.1600-0447.1992.tb01451.x (1992).
- 24. Lin, J. *et al.* Validation of the Chinese version of the Hamilton rating scale for depression in adults with epilepsy. *Epilepsy. Behav.* **89**, 148–152. https://doi.org/10.1016/j.yebeh.2018.10.009 (2018).
- Yang, W. et al. Prevalence and clinical profiles of comorbid anxiety in first episode and drug naïve patients with major depressive disorder. J. Affect. Disord. 257, 200–206. https://doi.org/10.1016/j.jad.2019.06.052 (2019).
- 26. Jaddoe, V. W. et al. First trimester fetal growth restriction and cardiovascular risk factors in school age children: population based cohort study. BMJ 348, g14. https://doi.org/10.1136/bmj.g14 (2014).
- 27. Klinitzke, G., Steinig, J., Blüher, M., Kersting, A. & Wagner, B. Obesity and suicide risk in adults–a systematic review. J. Affect. Disord. 145(3), 277–284. https://doi.org/10.1016/j.jad.2012.07.010 (2013).
- 28. Perera, S. et al. Body mass index is an important predictor for suicide: Results from a systematic review and meta-analysis. Suicide Life Threat Behav. 46(6), 697–736. https://doi.org/10.1111/sltb.12244 (2016).
- 29. Batty, G. D., Whitley, E., Kivimäki, M., Tynelius, P. & Rasmussen, F. Body mass index and attempted suicide: Cohort study of 1,133,019 Swedish men. *Am. J. Epidemiol.* 172(8), 890–899. https://doi.org/10.1093/aje/kwq274 (2010).
- Dreimüller, N. et al. Body mass index (BMI) in major depressive disorder and its effects on depressive symptomatology and antidepressant response. J. Affect. Disord. 256, 524–531. https://doi.org/10.1016/j.jad.2019.06.067 (2019).
- Amiri, S. & Behnezhad, S. Body mass index and risk of suicide: A systematic review and meta-analysis. J. Affect. Disord. 238, 615–625. https://doi.org/10.1016/j.jad.2018.05.028 (2018).
- 32. Lavie, C. J., De Schutter, A. & Milani, R. V. Healthy obese versus unhealthy lean: The obesity paradox. *Nat. Rev. Endocrinol.* 11(1), 55–62. https://doi.org/10.1038/nrendo.2014.165 (2015).
- 33. Samaan Z, Lee YK, Gerstein HC, Engert JC, Bosch J, Mohan V, Diaz R, Yusuf S, Anand SS, Meyre D EpiDREAM Genetics Investigators. Obesity genes and risk of major depressive disorder in a multiethnic population: a cross-sectional study. *J. Clin. Psychiatry* 76(12), e1611–e1618. https://doi.org/10.4088/JCP.14m09720 (2015).
- 34. Sun, D., Li, W., Zhang, H., Li, Y. & Zhang, Q. Inverted U-shaped relationship between body mass index and multivessel lesions in Chinese patients with myocardial infarction: A cross-sectional study. *J. Int. Med. Res.* 48(7), 300060520932820. https://doi.org/10.1177/0300060520932820 (2020).
- 35. González-Castro, T. B. et al. Evaluation of leptin levels in serum as a biomarker for suicide behavior: systematic review and meta-analysis. Int. J. Neurosci. 131(1), 49–55. https://doi.org/10.1080/00207454.2020.1733558 (2021).
- 36. Westling, S., Ahrén, B., Träskman-Bendz, L. & Westrin, A. Low CSF leptin in female suicide attempters with major depression. J. Affect. Disord. 81(1), 41–48. https://doi.org/10.1016/j.jad.2003.07.002 (2004).
- 37. Kim, H. et al. Association of body mass index with suicide behaviors, perceived stress, and life dissatisfaction in the korean general population. Psychiatry Investig. 15(3), 272–278. https://doi.org/10.30773/pi.2017.06.28 (2018).
- 38. Jackson, S. E., Johnson, F., Croker, H. & Wardle, J. Weight perceptions in a population sample of English adolescents: Cause for celebration or concern? *Int. J. Obes. (Lond.)* **39**(10), 1488–1493. https://doi.org/10.1038/ijo.2015.126 (2015).

- Austin, A. E., van den Heuvel, C. & Byard, R. W. Body mass index and suicide. Am. J. Forensic Med. Pathol. 35(2), 145–147. https://doi.org/10.1097/PAF.000000000000094 (2014).
- 40. Zhang, M. Z. et al. Body mass index and the risk of suicidal ideation and suicide attempt among youth in 45 low-and middle-income countries. J. Affect. Disord. 298(Pt A), 357–363. https://doi.org/10.1016/j.jad.2021.11.018 (2022).
- 41. Heneghan, H. M., Heinberg, L., Windover, A., Rogula, T. & Schauer, P. R. Weighing the evidence for an association between obesity and suicide risk. Surg. Obes. Relat. Dis. 8(1), 98–107. https://doi.org/10.1016/j.soard.2011.10.007 (2012).
- 42. Gomes, F. A. et al. Obesity is associated with previous suicide attempts in bipolar disorder. Acta Neuropsychiatry 22(2), 63–67. https://doi.org/10.1111/j.1601-5215.2010.00452.x (2010).
- 43. Yu, M., Shi, Y., Gu, L. & Wang, W. "Jolly fat" or "sad fat": A systematic review and meta-analysis of the association between obesity and depression among community-dwelling older adults. *Aging Ment. Health.* 26(1), 13–25. https://doi.org/10.1080/13607863. 2020.1857687 (2022).
- 44. Kessler, R. C., Chiu, W. T., Demler, O., Merikangas, K. R. & Walters, E. E. Prevalence, severity, and comorbidity of 12-month DSM-IV disorders in the National Comorbidity Survey Replication. *Arch. Gen. Psychiatry* 62(6), 617–627. https://doi.org/10.1001/archpsyc.62.6.617 (2005).
- Marcus, S. M. et al. Gender differences in depression: Findings from the STAR*D study. J. Affect. Disord. 87(2-3), 141-150. https://doi.org/10.1016/j.jad.2004.09.008 (2005).
- Oquendo, M. A. et al. Ethnic and sex differences in suicide rates relative to major depression in the United States. Am. J. Psychiatry 158(10), 1652–1658. https://doi.org/10.1176/appi.ajp.158.10.1652 (2001).
- Parker, G. & Brotchie, H. Gender differences in depression. Int. Rev. Psychiatry 22(5), 429–436. https://doi.org/10.3109/09540261. 2010.492391 (2010).
- Eisenberger, N. I., Inagaki, T. K., Rameson, L. T., Mashal, N. M. & Irwin, M. R. An fMRI study of cytokine-induced depressed mood and social pain: The role of sex differences. *Neuroimage* 47(3), 881–890. https://doi.org/10.1016/j.neuroimage.2009.04.040 (2009).
- Soares, C. N., Almeida, O. P., Joffe, H. & Cohen, L. S. Efficacy of estradiol for the treatment of depressive disorders in perimenopausal women: A double-blind, randomized, placebo-controlled trial. Arch. Gen. Psychiatry 58(6), 529–534. https://doi.org/10. 1001/archpsyc.58.6.529 (2001).
- Young, E. A. et al. Sex differences in response to citalopram: A STAR*D report. J Psychiatry Res. 43(5), 503–511. https://doi.org/ 10.1016/j.jpsychires.2008.07.002 (2009).
- 51. Gartlehner, G., Chapman, A., Strobelberger, M. & Thaler, K. Differences in efficacy and safety of pharmaceutical treatments between men and women: an umbrella review. *PLoS One* 5(7), e11895. https://doi.org/10.1371/journal.pone.0011895 (2010).
- 52. Berlanga, C. & Flores-Ramos, M. Different gender response to serotonergic and noradrenergic antidepressants. A comparative study of the efficacy of citalopram and reboxetine. *J. Affect. Disord.* **95**(1), 119–123. https://doi.org/10.1016/j.jad.2006.04.029 (2006).
- Marchand, W. R. Cortico-basal ganglia circuitry: A review of key research and implications for functional connectivity studies of mood and anxiety disorders. Brain Struct. Funct. 215(2), 73–96. https://doi.org/10.1007/s00429-010-0280-y (2010).
- Kong, L. et al. Sex differences of gray matter morphology in cortico-limbic-striatal neural system in major depressive disorder. J. Psychiatry Res. 47(6), 733–739. https://doi.org/10.1016/j.jpsychires.2013.02.003 (2013).
- 55. Sandberg, D. E. & Gardner, M. Differences/disorders of sex development: Medical conditions at the intersection of sex and gender. Annu. Rev. Clin. Psychol. 18, 201–231. https://doi.org/10.1146/annurey-clinpsy-081219-101412 (2022).
- Ebbinkhuijsen, M., Bevelander, K. E., Buijzen, M. & Kleemans, M. Children's emotions after exposure to news: Investigating chat conversations with peers as a coping strategy. J. Youth Adolesc. 50(7), 1424–1436. https://doi.org/10.1007/s10964-021-01408-0
- Kujawa, A. J. et al. Attentional biases for emotional faces in young children of mothers with chronic or recurrent depression. J. Abnorm. Child. Psychol. 39(1), 125–135. https://doi.org/10.1007/s10802-010-9438-6 (2011).

Acknowledgements

The authors thank the First Clinical Medical College, Shanxi Medical University for the supports.

Author contributions

Study Design: X.Z., X.D., J.L.. Investigation: J.L., Z.L., F.J., H.Y., Y.Z., X.X., R.Y., Y.W. Analysis and interpretation of data: J.L., X.Z., G.Y., Z.L. Drafting of the manuscript: J.L., Z.L, F.J. Critical revision of the manuscript: X.D., X.Z.. Approval of the final version for publication: J.L., Z.L., F.J., X.D., X.Z..

Funding

This work was supported by the Suzhou Gusu Health Talents Scientific Research Project (Nos. GSWS2021053, GSWS2019070), Key Diagnosis and treatment Program of Suzhou (LCZX202016), the Suzhou clinical Medical Center for mood disorders (Szlcyxzx202109), the Medical Science and Technology Development Foundation, Nanjing Department of Health (Nos. YKK21216, YKK20184, YKK22264). The funding sources of this study had no role in study design, data collection and analysis, decision to publish, or preparation of the article.

Competing interests

The authors declare no competing interests.

Additional information

Correspondence and requests for materials should be addressed to X.D. or X.Z.

Reprints and permissions information is available at www.nature.com/reprints.

Publisher's note Springer Nature remains neutral with regard to jurisdictional claims in published maps and institutional affiliations.

Open Access This article is licensed under a Creative Commons Attribution 4.0 International License, which permits use, sharing, adaptation, distribution and reproduction in any medium or format, as long as you give appropriate credit to the original author(s) and the source, provide a link to the Creative Commons licence, and indicate if changes were made. The images or other third party material in this article are included in the article's Creative Commons licence, unless indicated otherwise in a credit line to the material. If material is not included in the article's Creative Commons licence and your intended use is not permitted by statutory regulation or exceeds the permitted use, you will need to obtain permission directly from the copyright holder. To view a copy of this licence, visit http://creativecommons.org/licenses/by/4.0/.

© The Author(s) 2023