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OPEN Preterm births prevalence during the COVID-19 pandemic in Brazil: results from the national database

Charles M'poca Charles^{1,2}, Luiz Alves Souza Neto^{3,4}, Camila Ferreira Soares⁵, Tacildo Souza Araújo³, Cristiano Torezzan⁶, Everton Emanuel Campos Lima^{5,6}, Aline Munezero¹, Luis Bahamondes^{1,7}, Renato Teixeira Souza¹, Maria Laura Costa¹, José Guilherme Cecatti¹ & Rodolfo Carvalho Pacagnella¹

The SARS-CoV-2 (COVID-19) pandemic impacted the health systems between and within countries, and in the course of the pandemic sexual and reproductive health services were the most disrupted. Findings from high-income settings have reported significant changes in preterm birth prevalence during the pandemic period. To understand the possible effects of the COVID-19 pandemic on preterm birth numbers at the Brazilian national level. We compare the number of preterm deliveries during the COVID-19 pandemic period (2020 and 2021) with previous years. We conducted a populationbased cross-sectional study taking the period from January 2017 to December 2021 to account. We use individual-level live births data from the Brazilian Live Birth Information System (SINASC), and we estimate the odds ratio (OR) of preterm deliveries using propensity score weighting analysis in Brazil and its regions. During the study period (from 2017 to 2021), about 2.7 million live births were recorded per year, and the missing value for gestational age at delivery was less than 1.5%. The preterm birth prevalence slightly increased during the COVID-19 pandemic compared to the pre-pandemic period (11.32% in 2021 vs 11.09% in 2019, p-value < 0.0001). After adjusting for sociodemographic variables, the OR of preterm births in Brazil has significantly increased, 4% in 2020 (OR: 1.04 [1.03–1.05] 95% CI, p-value < 0.001), and 2% in 2021(OR: 1.02 [1.01–1.03] 95% CI, p-value < 0.001), compared to 2019. At the regional level, the preterm birth pattern in the South, Southeast and Northeast regions show a similar pattern. The highest odds ratio was observed in the South region (2020 vs 2019, OR: 1.07 [1.05–1.10] 95% CI; 2021 vs 2019, OR: 1.03 [1.01–1.06] 95% CI). However, we also observed a significant reduction in the ORs of preterm births in the northern region during the COVID-19 pandemic (2020 vs 2019, OR: 0.96 [0.94–0.98] 95% CI) and (2021 vs 2019, OR: 0.97 [0.95–0.99] 95% CI). Our analysis shows that the pandemic has increased regional variation in the number of preterm births in Brazil in 2020 and 2021 compared to the pre-pandemic years.

The COVID-19 pandemic impacted Latin American countries as hard as in more developed locations. Due to the sanitary crises, we saw in many countries disrupting health systems, and as consequence, mortality exceeded its usual numbers and life expectancy reductions in many countries¹. In addition to controversial health policies, conflicting messages and long-time central government resistance to implementing population mobility restrictions², Brazil was one of the most affected countries by COVID-19 worldwide. Pregnant women were also a risk group, as maternal mortality skyrocketed during the pandemic³.

Maternal mortality is also an important proxy for the quality of country's health services. Another obstetric condition that is sensitive to suboptimal clinical care is preterm birth⁴. Studying preterm birth is important

¹Department of Obstetrics and Gynecology, University of Campinas, Campinas, SP, Brazil. ²Provincial Health Administration, DPS Manica, Chimoio, Mozambique. ³Institute of Mathematics, Statistics and Scientific Computing (IMEEC), University of Campinas, Campinas, SP, Brazil. ⁴School of Applied Sciences (FCA), University of Campinas, Campinas, SP, Brazil. ⁵College of Philosophy and Human Sciences (IFCH), University of Campinas, Campinas, SP, Brazil. ⁶Center for Population Studies (NEPO), University of Campinas, Campinas, SP, Brazil. ⁷Campinas Reproductive Health Research Center (CEMICAMP), Campinas, Brasil. ^{III}email: rodolfop@unicamp.br

because it is the primary cause of neonatal deaths, and its prevalence is rising in most low- and middle-income countries, despite many efforts to revert it^{5,6}. While several risk factors have been well-established, the key factor responsible for preterm deliveries remains unknown in half of the cases⁷. Since the onset of the pandemic, several studies have identified an association between COVID-19 infection and adverse perinatal outcomes, such as stillbirths and premature deliveries^{8–11}. These findings are also contradictory because some analyses indicate that the number of preterm deliveries increased during the pandemic, while other studies suggested a reduction in such types of pregnancies¹¹.

One of the most important underlying mechanisms for preterm birth is the inflammatory condition. The systemic inflammation may trigger cervical effacement and uterine contraction through increasing prostaglandins¹². The SARS-CoV-2 infection is a systemic inflammatory disease; therefore, we may argue that it could lead to preterm birth. For example, among women with SARS-CoV-2 pneumonia, empirical evidence shows an increased preterm birth rate¹³. However, the infection itself may not represent the whole mechanism related to preterm delivery.

In addition, we may also argue that changes in individual behavior are associated with lockdown and other population restrictions policies, implemented to mitigate SARS-CoV-2 dissemination, and that may have influenced to some extent the number of preterm births. As an example, an Australian study showed a lower risk of preterm birth in pregnant women during lockdowns in comparison to those born before the pandemic¹⁴. Other studies also indicate a decrease in preterm birth rates^{15,16}, although the same empirical evidence was not corroborated elsewhere that fail to identify differences in the number of preterm pregnancies¹⁷.

Despite inconclusive findings, there is a consensus that the COVID-19 pandemic period brought many challenges to the country's health systems, and there is still scarce information about its real consequences on perinatal health while considering low- and middle-income countries. Therefore, in this study, we aim to assess the changes in preterm birth counts in Brazil and its regions, by comparing the number of preterm deliveries during the pandemic (2020 and 2021) and pre-pandemic periods (2017, 2018 and 2019).

Methods

We performed extensive use of the publicly available microdata of live birth, collected by the Brazilian Ministry of Health, and launched by the Brazilian Live Birth Information System (SINASC in Portuguese)¹⁸. The SINASC is an e-birth registration system developed by the Department of Informatics of the National Unified Health System (DATASUS). This system was implemented in 1990. The data are collected routinely immediately after each birth through a standardized document (declaration of live births), which was updated in 2010 to ensure a better quality of the information recorded¹⁹. The updated version included many important variables for the study of preterm birth, such as sociodemographic and obstetric variables.

The data was downloaded (as of August 12, 2022), and updated (as of April 15, 2023) from http://svs.aids. gov.br/dantps/cgiae/sinasc/, and we consider all live births equal or superior to 22 weeks, from January 2017 to December 2021. We extracted individual-level data regarding gestational age at birth, maternal age, marital status, ethnicity, schooling (as a proxy of women's income), parity, gravidity, mode of delivery, region and federal state of residence, number of living children, number of antenatal care (ANC) visits, and newborn's weight and sex. These variables are available in the SINASC for each birth, and they are highly associated with preterm delivery. We did not exclude multiple pregnancies and neonates with congenital anomalies for the analysis. All categorical variables were converted to binary dummies by using the one-hot encoding procedure. Less than 1.5% of the data had missing or unknown information. Notwithstanding, SINASC data quality has recently shown enormous improvement. Of course, while considering more disaggregated geographical levels, the data may still need some adjustments and corrections. For this study, we work with Brazil and great regions, and that reduces significantly defective concerns such as under-registration of birth counts; as in Brazil and its regions, the rate of underreported data is generally less than 1%, except in the North and Northeast regions where the rate is about^{20,21} 1.7%. In addition, Castanheira and Kohler considered inadequate to apply any correction method to birth registrations, given the recent fertility dynamics in the country²². Lima et al.²³ also show that recent SINASC information does not require data corrections at lesser disaggregated levels, such as Federal States and great regions. However, we acknowledge that the unprecedented burden on the health system during the pandemic may have influenced the data quality. The study protocol was published elsewhere²⁴.

To reduce the influence of past trends in prematurity prevalence, we restrict our preterm birth analysis to pairwise years comparison, initiating from 2017 until 2021. We did not include data before 2017 to avoid the influence of the Zika virus outbreak (between 2015 and 2016) on birth counts and overall fertility²⁵. We created four stacked datasets (2017–2018; 2018–2019; 2019–2020; 2019–2021) and we added, for each dataset, two dichotomous variables of interest: one to indicate whether the birth was preterm (y=1) and (y=0) otherwise, and another measure indicates the period, i.e. the current year in the dataset (z=0) vs. the following year (z=1). This last variable is useful for identifying the control group (preterm births occurring in years before COVID-19) and the treatment group (preterm deliveries occurring during the pandemic).

Statistic model. Our analysis was based on a quasi-experimental approach using a Propensity Score Weighting (PSW) method^{26,27}. PSW was designed to control for selection bias in non-experimental studies, for which it is desirable to assess the average effect of some variable that emulates a control/treatment process. Propensity scores are used to match untreated versus treated individuals, understanding that there is a probability of these last being exposed to a certain stimulus or intervention²⁸.

As the first step, a multiple logistic regression analysis was used to fit the binary control variable (z) as a function of the mother's and obstetrics' characteristics: age, ethnicity/skin colour, schooling, parity, mode of delivery, number of previous children, marital status, number of antenatal care visits and new-born weight. With the regression estimates, we extracted a vector (e) that gives the probability of treatment assignment to a random individual conditioned to a given set of covariates (x), i.e. e(x) = P(z=1|x).

The vector (e) is called the Propensity Score, and it was used to control for selection bias and to derive the weights of a second regression model. The control was made by pruning samples corresponding to the tails of the Propensity Score vector, to keep only samples that can be considered comparable to each other. Figure 1 shows the kernel density estimate (KDE) plot for the Propensity Score referring to births in Brazil in the years 2019 and 2020, before (a) and after (b) a 10% pruning of each tail. In this example, 80% of the original dataset was selected for the final phase of the analysis.

The set of weights was estimated as follows: for the individuals in the treatment group, w = 1/e(x), and for the individuals in the control group w = 1/(1-e(x)). As a final step, we estimate a new regression, fitting the outcome of interest (preterm birth) controlled by the covariates and using the propensity scores as weights.

Results

About 2.7 million live births were recorded annually from 2017 to 2021¹⁸. In Table 1, we present the percentage share of clinical and socioeconomic mother's characteristics, comparing the last three years of our analysis, 2019 until 2021.

Overall, in Brazil, the preterm birth counts were around 11%, and this number did not change much compared to pre-pandemic years, in this case, 2019. Also, in terms of the mother's characteristics, we did not identify considerable changes in the last three years of our analysis.

In Table 2, we show the results of the multiple regression analysis using PSW for Brazil, by pairwise year comparison. We were interested to see if the pandemic (treatment period, 2020, and 2021) somehow affected the chances of preterm birth counts in the country. Our results show that the Odds Ratio (ORs) of preterm births in Brazil has increased by 4% in 2020 (95% CI 1.03–1.05), and 2% (95% CI 1.01–1.03) in 2021, compared to 2019 after controlling for other sociodemographic variables.

In addition, the pairwise comparison for the period 2017 to 2019 shows small or non-significant changes in the ORs of preterm births, and the odds ratios of preterm pregnancies fluctuated between values of below and above 1%. This means that during the pandemic the chances of preterm deliveries have increased somewhat to two and four per cent.

In Fig. 2, we bring the odds ratios for Brazil and its regions. These estimates are based on complete models, controlled by the same variables described in Table 2. Across regions, the odds ratios of preterm births showed a small decline or even stalled values between the pre-pandemic periods of 2017–2019, seen especially in the South and Midwest regions of Brazil. However, while we consider the pandemic period effect (2019 vs. 2020, and 2019 vs. 2021), the chances of preterm pregnancies increased again. In the Southeast and the less developed Brazilian region of the Northeast, for example, there was a small decrease in the odds ratios between 2017 and 2019, but during the pandemic period, the chances of preterm births increased even more, especially in the Northeastern part of Brazil. The Northern region was the only location that had a reduction in the odds ratios of preterm births during the COVID-19 pandemic period (2020 and 2021). These results may also indicate that the effect of the pandemic on the prevalence of preterm births was uneven across subnational areas of the country.

Our finding showed a significant change in caesarean delivery rate during the pandemic period compared to the previous period (OR 1.09 [1.08–1.10] in 2020, and OR 1.10 [1.09–1.11] in 2021), Table 2. The analysis of the mode of delivery by gestational age, for the pairwise comparison of 2019 and 2020, showed a trend of increasing caesarean delivery in all gestational age groups. Moreover, preterm babies had a higher risk of being delivered by caesarean in 2020 and 2021 compared to the previous years. This pattern was also observed in the South, Southeast, and Northeast regions, Figs. 3 and 4.

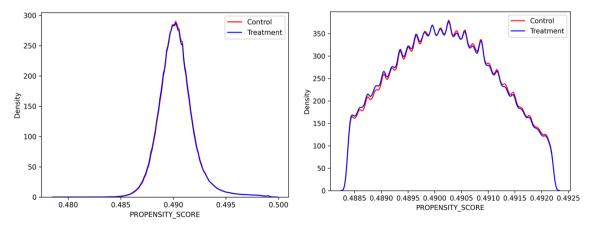


Figure 1. Example of Kernel density estimate (KDE) plot for the Propensity Score referring to births in Brazil in the years 2019 and 2020. *Source* Brazilian Live Birth Information System (SINASC) (2023).

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Brown 55.96 1,594,267 57.06 1,533,251 Indigenous 0.93 26,373 0.91 25,741 Not stated/Unknown 2.63 74,987 3 70,881 Mother's schooling 0 1,533,251 1,000 1,533,251 Mother's schooling 0 1,4987 3 70,881 Mother's schooling 0 15.36 431,144 8 to 17 years 61.36 1,748,186 62.23 1,698,877 12 and more 21.27 606,145 21.12 589,807 Not stated/unknown 1.15 32,752 1.31 29,178 Mother's marital status 52.39 1,492,765 50.34 1,374,363 Widow 0.16 4,693 0.17 4,603 Divorced 1.36 38,748 1.45 39,619 Not stated/Unknown 0.95 26,942 1.02 27,806 Type of pregnancy 1 57,846 1,467 0.05 1,262 No	6.81	181,875
Indigenous 0.93 26,373 0.91 25,741 Not stated/Unknown 2.63 74,987 3 70,881 Mother's schooling 74,987 3 70,881 Mother's schooling 16.22 462,063 15.36 431,144 8 to 11 years 61.36 1,748,186 62.23 1,698,877 12 and more 21.27 606,145 21.12 589,807 Not stated/unknown 1.15 32,752 1.31 29,178 Mother's marital status 52.39 1,492,765 50.34 1,374,363 Widow 0.16 4,693 0.17 4,603 Divorced 1.36 38,748 1.45 39,619 Not stated/Unknown 0.95 26,942 1.02 27,806 Type of pregnancy 51.36 2,112 568,636 Twin 2.13 60,61 2.11 57,846 Triplet and more 0.05 1,467 0.05 1,262	0.45	12,106
Not stated/Unknown 2.63 74,987 3 70,881 Mother's schooling 0 74,987 3 70,881 Mother's schooling 0 1 74,987 3 70,881 Mother's schooling 0 1 74,987 3 70,881 0 to 7 years 16.22 462,063 15.36 431,144 8 to 11 years 61.36 1,748,186 62.23 1,698,877 12 and more 21.27 606,145 21.12 589,807 Not stated/unknown 1.15 32,752 1.31 29,178 Mother's marital status Single 45.14 1,285,998 47.02 1,283,754 Married/Cohabit 52.39 1,492,765 50.34 1,374,363 Widow 0.16 4,693 0.17 4,603 Divorced 1.36 38,748 1.45 39,619 Not stated/Unknown 0.95 26,942 1.02 27,806 Type of pregnancy Single 97.76 <t< td=""><td>56.75</td><td>1,516,269</td></t<>	56.75	1,516,269
Not stated/Unknown 2.63 74,987 3 70,881 Mother's schooling 0 74,987 3 70,881 Mother's schooling 0 15.36 431,144 8 to 11 years 61.36 1,748,186 62.23 1,698,877 12 and more 21.27 606,145 21.12 589,807 Not stated/unknown 1.15 32,752 1.31 29,178 Mother's marital status 52.39 1,492,765 50.34 1,374,363 Widow 0.16 4,693 0.17 4,603 Divorced 1.36 38,748 1.45 39,619 Not stated/Unknown 0.95 26,942 1.02 27,806 Type of pregnancy 1.36 38,748 1.45 39,619 Not stated/Unknown 0.95 26,942 1.02 27,806 Type of pregnancy 1.36 38,748 1.45 39,619 Not stated/Unknown 0.95 1,467 0.05 1,262 Noit stated/Unk	1.06	28,216
Mother's schooling Id.22 462,063 15.36 431,144 8 to 11 years 61.36 1,748,186 62.23 1,698,877 12 and more 21.27 606,145 21.12 589,807 Not stated/unknown 1.15 32,752 1.31 29,178 Mother's marital status 53,752 1.31 29,178 Mother's marital status 1,283,754 1,492,765 50.34 1,374,363 Midow 0.16 4,693 0.17 4,603 Divorced 1.36 38,748 1.45 39,619 Not stated/Unknown 0.95 26,942 1.02 27,806 Type of pregnancy 50.34 1,374,363 14 1,285,200 97.74 2,668,636 Twin 2.13 60,61 2.11 57,846 17 Triplet and more 0.05 1,467 0.05 1,262 Not stated/Unknown 0.07 1,869 0.08 2,401	2.47	65,923
0 to 7 years 16.22 462,063 15.36 431,144 8 to 11 years 61.36 1,748,186 62.23 1,698,877 12 and more 21.27 606,145 21.12 589,807 Not stated/unknown 1.15 32,752 1.31 29,178 Mother's marital status		
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12 and more 21.27 606,145 21.12 589,807 Not stated/unknown 1.15 32,752 1.31 29,178 Mother's marital status	62.62	1,673,570
Not stated/unknown 1.15 32,752 1.31 29,178 Mother's marital status 1,283,754 1,374,363 Married/Cohabit 52.39 1,492,765 50.34 1,374,363 Widow 0.16 4,693 0.17 4,603 Divorced 1.36 38,748 1.45 39,619 Not stated/Unknown 0.95 26,942 1.02 27,806 Type of pregnancy 2,785,200 97.74 2,668,636 Twin 2.13 60,61 2.11 57,846 Triplet and more 0.05 1,467 0.05 1,262 Not stated/Unknown 0.07 1,869 0.08 2,401	21.48	583,779
Mother's marital status 45.14 1,285,998 47.02 1,283,754 Married/Cohabit 52.39 1,492,765 50.34 1,374,363 Widow 0.16 4,693 0.17 4,603 Divorced 1.36 38,748 1.45 39,619 Not stated/Unknown 0.95 26,942 1.02 27,806 Type of pregnancy Single 97.76 2,785,200 97.74 2,668,636 Twin 2.13 60,61 2.11 57,846 Triplet and more 0.07 1,869 0.08 2,401 Number of antenatal visits 50.70 1,262 1.262	1.30	34,898
Single 45.14 1,285,998 47.02 1,283,754 Married/Cohabit 52.39 1,492,765 50.34 1,374,363 Widow 0.16 4,693 0.17 4,603 Divorced 1.36 38,748 1.45 39,619 Not stated/Unknown 0.95 26,942 1.02 27,806 Type of pregnancy 2,785,200 97.74 2,668,636 Twin 2.13 60,61 2.11 57,846 Triplet and more 0.05 1,467 0.05 1,262 Not stated/Unknown 0.07 1,869 0.08 2,401		
Married/Cohabit 52.39 1,492,765 50.34 1,374,363 Widow 0.16 4,693 0.17 4,603 Divorced 1.36 38,748 1.45 39,619 Not stated/Unknown 0.95 26,942 1.02 27,806 Type of pregnancy 50,611 2.11 57,846 Twin 2.13 60,611 2.11 57,846 Triplet and more 0.05 1,467 0.05 1,262 Not stated/Unknown 0.07 1,869 0.08 2,401	48.39	1,292,963
Widow 0.16 4,693 0.17 4,603 Divorced 1.36 38,748 1.45 39,619 Not stated/Unknown 0.95 26,942 1.02 27,806 Type of pregnancy 5 2,785,200 97.74 2,668,636 Twin 2.13 60,61 2.11 57,846 Triplet and more 0.05 1,467 0.05 1,262 Not stated/Unknown 0.07 1,869 0.08 2,401	48.76	1,302,820
Divorced 1.36 38,748 1.45 39,619 Not stated/Unknown 0.95 26,942 1.02 27,806 Type of pregnancy 2,785,200 97.74 2,668,636 Twin 2.13 60,61 2.11 57,846 Triplet and more 0.05 1,467 0.05 1,262 Not stated/Unknown 0.07 1,869 0.08 2,401	0.19	4,978
Not stated/Unknown 0.95 26,942 1.02 27,806 Type of pregnancy	1.48	39,576
Type of pregnancy 97.76 2,785,200 97.74 2,668,636 Twin 2.13 60,61 2.11 57,846 Triplet and more 0.05 1,467 0.05 1,262 Not stated/Unknown 0.07 1,869 0.08 2,401	1.19	31,709
Number of antenatal visits 97.76 2,785,200 97.74 2,668,636 Twin 2.13 60,61 2.11 57,846 Triplet and more 0.05 1,467 0.05 1,262 Not stated/Unknown 0.07 1,869 0.08 2,401		,. 07
Twin 2.13 60,61 2.11 57,846 Triplet and more 0.05 1,467 0.05 1,262 Not stated/Unknown 0.07 1,869 0.08 2,401 Number of antenatal visits	97.72	2,611,194
Triplet and more 0.05 1,467 0.05 1,262 Not stated/Unknown 0.07 1,869 0.08 2,401 Number of antenatal visits	2.14	57,061
Not stated/Unknown 0.07 1,869 0.08 2,401 Number of antenatal visits <t< td=""><td>0.05</td><td>1,319</td></t<>	0.05	1,319
Number of antenatal visits	0.03	2,472
	0.09	2,772
1.52 45,400 1.75 47,276	1.04	49.085
1 to 3 5 35 152 492 6 04 164 042	1.84	49,085
1 to 3 5.35 152,483 6.04 164,943 4 to 6 20.26 577,17 20.70 565,211	5.34	142,687
	19.15	511,652
7 and more 72.43 2,063,669 71.01 1,938,920 Not stated/Unknown 0.44 12,418 0.50 13,795	73.14 0.54	1,954,282 14,430

Table 1. Descriptive statistics for Brazil live births 2019, 2020 and 2021. Source Brazilian Live BirthInformation System (SINASC) (2023).

	Odds ratio (95% CI)					
	2017-2018	2018-2019	2019-2020	2019-2021		
Intercept	157.97*** (152.54-163.60)	31.54*** (29.90-33.27)	33.12*** (31.39-34.95)	35.41*** (33.56-37.36)		
Year	1.01** (1.00-1.02)	1.00 (0.99–1.1)	1.04*** (1.03-1.05)	1.02*** (1.01-1.03)		
Weight	0.99*** (0.99-0.99)	0.99*** (0.99-0.99)	0.99*** (0.99-0.99)	0.99*** (0.99-0.99)		
Mother's age						
19-34	REF	REF	REF	REF		
<19	1.10*** (1.11-1.12)	1.20*** (1.19-1.21)	1.19*** (1.18-1.21)	1.20*** (1.19-1.22)		
>34	1.24*** (1.23-1.25)	1.23*** (1.22-1.24)	1.26*** (1.25-1.27)	1.26*** (1.25-1.27)		
Multiparous	1.11*** (1.10-1.12)	1.17*** (1.16-1.18)	1.19*** (1.18-1.20)	1.20*** (1.19-1.21)		
Caesarean delivery	1.08*** (1.07-1.09)	1.08*** (1.07-1.09)	1.09*** (1.08-1.10)	1.10***(1.09-1.11)		
Sex						
Female	Ref	Ref	Ref	Ref		
Unknown	1.74*** (1.40-2.18)	2.66*** (1.93-3.67)	2.79*** (2.03-3.84)	2.55*** (1.83-3.55)		
Male	1.38*** (1.37-1.39)	1.38*** (1.37-1.39)	1.39*** (1.38-1.40)	1.40*** (1.39-1.41)		
Race/colour				1		
White	Ref	Ref	Ref	Ref		
Black	0.90*** (0.88-0.92)	0.90*** (0.89-0.91)	0.90*** (0.88-0.91)	0.89*** (0.88-0.91)		
Asian	0.90 (0.85-0.95)	0.94 (0.89-0.98)	0.99 (0.94-1.05)	0.97 (0.92-1.03)		
Brown	0.98*** (0.97-0.99)	0.99 (0.98-1.00)	1.03*** (1.02-1.04)	1.04*** (1.03-1.05)		
Indigenous	1.05*** (1.01-1.10)	1.10*** (1.07-1.14)	1.50*** (1.44-1.54)	1.60*** (1.57-1.63)		
Mother's schooling			·			
12 and more	ref	ref	ref	ref		
8 to 11 years	0.96 (0.95–0.97)	1.04*** (1.03-1.05)	1.05*** (1.04-1.06)	1.05***(1.03-1.06)		
0 to 7 years	1.00 (0.99–1.01)	1.14*** (1.13–1.16)	1.15*** (1.13–1.16)	1.14*** (1.12-1.16)		
Mother's marital statu	S		•			
Single	0.95*** (0.94-0.96)	0.97*** (0.96-0.98)	0.93*** (0.92-0.94)	0.95*** (0.94-0.96)		
Married/Cohabit	ref	ref	ref	ref		
Widow	1.00 (0.92–1.98)	0.99 (0.89–1.09)	0.96 (0.88-1.06)	1.05 (0.96-1.15)		
Type of pregnancy			•			
Single	ref	ref	ref	ref		
Twin	2.91*** (2.86-2.96)	3.44*** (2.36-3.52)	3.44*** (3.36-3.52)	3.49*** (3.42-3.58)		
Triplet and more	11.70*** (9.55-14.34)	4.36*** (2.68-7.09)	4.56*** (2.86-7.28)	5.24*** (3.29-8.34)		
Number of antenatal	visits					
None	ref	ref	ref	ref		
1 to 3	1.52*** (1.48-1.57)	1.59*** (1.53-1.64)	1.50*** (1.45-1.55)	1.44*** (1.39-1.49)		
4 to 6	1.26*** (1.23-1.30)	1.32*** (1.28-1.37)	1.28*** (1.24-1.33)	1.24***(1.20-1.29)		
7 and more	0.66*** (0.64-0.67)	0.70*** (0.68-0.72)	0.70*** (0.67-0.72)	0.67*** (0.65-0.69)		

Table 2. Logistic regression analysis using Propensity Score Weighting for preterm birth in Brazil 2017–2021.Source Brazilian Live Birth Information System (SINASC) (2023). Significance level $p < 0.05^*$, $p < 0.01^{**}$ and $p < 0.001^{***}$.

Discussion

Using Ministry of Health data, we assessed the odds ratios of preterm births in Brazil and its regions, before and during the pandemic. Our results indicated that, during the pandemic years of 2020 and 2021, preterm births have significantly increased as compared to pre-pandemic periods. This increase was not homogeneous across the country, and in certain regions, the pandemic has disrupted previous decline patterns or even accelerated the past trend of preterm deliveries growing; as observed in Northeastern, Southeastern, and Southern regions.

The ethnicity (indigenous women), low level of education, low number of antenatal care visits, and multiparity, extreme maternal ages, were associated with an increased risk of PTB. These data are similar to the findings of other population-based studies^{29,30}.

Finding from individual studies and systematic review has suggested a global reduction in ANC clinic visits, maternity healthcare-seeking, and unscheduled care visits³¹; the same pattern was also observed in Brazil, where the quality of ANC was low (only 35.8% of the study participants had adequate), In addition, the risk of inad-equate ANC was higher among pregnant women with black/brown skin colour and multiparous when compared to their contra part³². These factors potentially contributed to worsening pregnancy outcomes (including the preterm birth rate), even for married/cohabiting women.

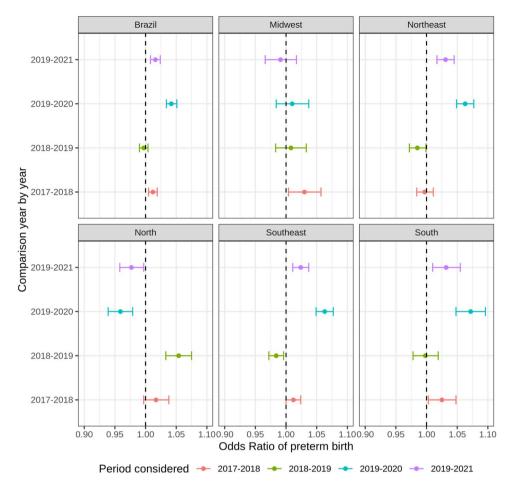
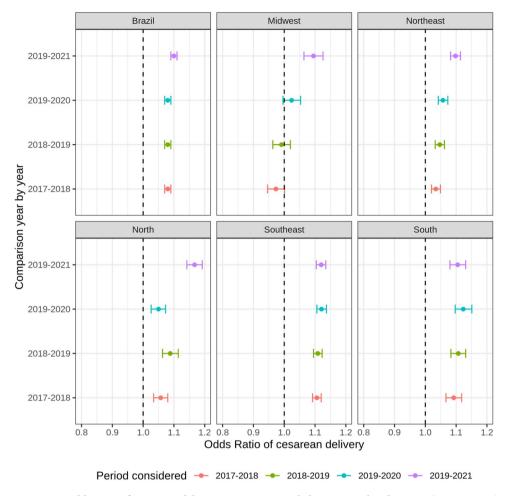


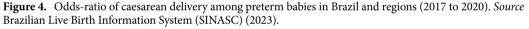
Figure 2. Odds-ratios of preterm birth for Brazil and its regions 2017 to 2021. *Source* Brazilian Live Birth Information System (SINASC) (2023).



Figure 3. Caesarean delivery rate for categories of gestational ages in Brazil (2017 to 2021). *Source* Brazilian Live Birth Information System (SINASC) (2023).

During the COVID-19 access to the Internet and DICT (Digital Information and Communication Technologies) was heterogeneous within the Brazilian regions, and municipality, public and private health systems. And, to the best of our knowledge, data regarding the coverage of virtual or remote antenatal care were not available in the database [SINASC], and the ANC visits are not desegregated by the mode of consultation [remote vs inperson]), therefore, we have not considered this variable in our analysis.





The pandemic brought the attention of health experts and demographers that took the time to understand how COVID-19 could affect birth counts and, for instance, the chances of preterm deliveries in the country. Brazil is a country that suffered excessive mortality due to the COVID-19 pandemic^{2,33}, as well as health facilities also were stressed by the high number of COVID-19 cases, and many services could not be properly provided by health units³⁴. This exogenous sanitary problem might have also affected women's antenatal care, especially among those that require more attention from public health services, i.e., mothers from low socioeconomic strata. Uncertainty and economic restrictions caused by the pandemic context may also play an important role in reproduction³⁵, and compromise pregnancy and antenatal care in Brazil.

Among Brazilian regions, the Northeast requires special attention because this is a region marked by historically lower socioeconomic development that could be in turn associated with restricted health services access^{36,37}, and the lack of strategy to mitigate the impact of the pandemic at different governmental levels^{3,33}. Notwithstanding, the COVID-19 pandemic brought an enormous burden to Brazil's Northern and Northeastern regions and revealed a sudden disruption of health care services^{38,39}. These setbacks might in turn affect the preterm birth rate.

Our findings differ from other studies that indicated a reduction in preterm deliveries during the COVID-19 pandemic^{10,14,40–44}. This could be partly explained by the measures applied to face the pandemic, which was uncoordinatedly implemented in Brazil³⁶. Regional inequality in health services access and the slow degree of responsiveness of the Brazilian National Health System could have played a role in the unequal pandemic effects on preterm births across Brazilian regions. As previous studies indicate, less sub-national inequality is seen in high-income countries, recognized by strict lockdown policies and with developed health services according to the needs posed by the pandemic. Moreover, Brazil had more severe cases; one out of seven maternities had intensive unit beds, therefore resulting in the phase three delay—concerning receiving proper diagnosis and timely treatment^{1,45,46}.

However, our findings suggested a different pattern of preterm birth rate in the Northern region. In the Northern region of Brazil, more than two-thirds of pregnant women did not attend antenatal care, and higher excess mortality (especially in Manaus city), which might have caused severe perinatal outcomes (miscarriage and fetal death)^{47,48}.

Our study suggested an increased rate of caesarean delivery among preterm babies in 2020, and 2021 compared to previous periods. Therefore, we may speculate that the increased risk of PTB in 2020, and 2021 may be related to non-spontaneous (provider-initiated) preterm birth^{49,50}.

It is important to mention that we concentrate our analysis and interpretations on the year's effect only (comparison between control versus treatment, or pre-pandemic vs. pandemic period), and we do not get into detail about the other control variables, despite the models have shown important differences in preterm pregnancies among distinguished demographic and socioeconomic groups.

This study has some strengths and limitations. Our data covers the entire population of live births in Brazil, with information at the individual level²⁰. The analysis of the different geographic regions allowed us to picture preterm birth developments in a country recognized for its regional inequality. The main limitation is related to the study design, which does not allow us to infer causality but only refers to the association between the pandemic and preterm births. We also did not assess the direct impact of COVID-19 on the occurrence of preterm births, and we considered the years 2020 and 2021 as risk factors that caused changes (from social, economic, and epidemiological order) brought by the pandemic onset. Likewise, our model did not include all variables associated with preterm birth, for example, human development index, availability and access to health services before and during the COVID-19 pandemic, cigarette smoking, BMI, maternal income, unemployment, maternal underlying medical conditions, and maternal infection (vector-borne diseases, urinary tract, genital, and respiratory infection [including COVID-19]). We did not assess the prevalence of fetal deaths and the abortion rates. But we recognize that these outcomes could have increased in situations of reduced access to adequate health services, impacting Brazil's birth rates.

Although we did not see an expressive increase of preterm births, we still argue that the disruption of sexual and reproductive health services may have influenced pregnancy outcomes. Therefore, monitoring the preterm birth rate might be an essential strategy for assessing the quality of maternal and perinatal care and might help providers and policymakers to develop strategies to mitigate the problem.

Data availability

The datasets analysed during the current study are publicly available from the Brazilian Live Birth Information System (SINASC) http://svs.aids.gov.br/dantps/cgiae/sinasc/.

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Author contributions

R.C.P., C.M.C., L.B., J.G.C., C.T. and R.T.S. conceived the study. C.M.C., A.M., L.A.N., T.A., and C.F.S. extracted the data from the SINASC. R.C.P., C.T., E.E.C.L., L.A.N., and C.F.S. conducted the statistical analysis, and collaborate in writing the manuscript first draft. R.C.P., L.B., M.L.C., C.M.C., and C.F.S. collaborated to statistical analysis. R.C.P., C.T., E.E.C.L., L.A.N., E.E.C.L., and C.M.C. wrote the first Draft. All authors critically reviewed the manuscript. All the authors read and approved the final version of the manuscript.

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Competing interests

The authors declare no competing interests.

Additional information

Correspondence and requests for materials should be addressed to R.C.P.

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