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OPEN Histidine-rich glycoprotein as a prognostic biomarker for sepsis

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Various biomarkers have been proposed for sepsis: however, only a few become the standard. We previously reported that plasma histidine-rich qlycoprotein (HRG) levels decreased in septic mice, and supplemental infusion of HRG improved survival in mice model of sepsis. Moreover, our previous clinical study demonstrated that HRG levels in septic patients were lower than those in noninfective systemic inflammatory response syndrome patients, and it could be a biomarker for sepsis. In this study, we focused on septic patients and assessed the differences in HRG levels between the nonsurvivors and survivors. We studied ICU patients newly diagnosed with sepsis. Blood samples were collected within 24 h of ICU admission, and HRG levels were determined using an enzyme-linked immunosorbent assay. Ninety-nine septic patients from 11 institutes in Japan were included. HRG levels were significantly lower in non-survivors (n = 16) than in survivors (n = 83) (median, 15.1 [interquartile ranges, 12.7–16.6] vs. 30.6 [22.1–39.6] µg/ml; p < 0.01). Survival analysis revealed that HRG levels were associated with mortality (hazard ratio 0.79, p < 0.01), and the Harrell C-index (predictive power) for HRG was 0.90. These results suggested that HRG could be a novel prognostic biomarker for sepsis.

Abbreviations

APACHE Acute physiology and chronic evaluation

Confidence interval

ELISA Enzyme-linked immunosorbent assay

HR Hazard ratio

HRG Histidine-rich glycoprotein **ICU** Intensive care unit

IQR Interquartile ranges **PCT** Procalcitonin

SIRS Systemic inflammatory response syndrome

SOFA Sequential organ failure assessment

Sepsis is one of the leading causes of death worldwide^{1,2}. However, all clinical trials conducted for the development of therapeutic drugs for sepsis for decades have failed. One of the problems hindering an early initiation of treatment for sepsis is unavailability of a good biomarker to diagnose and evaluate severity of sepsis³. Many markers and scoring systems have been studied to predict the severity and mortality in septic patients. Procalcitonin (PCT) is commonly used to assist in the diagnosis of acute infection in clinical settings^{4,5} and is somewhat considered as a prognostic biomarker 6.7. However, a single biomarker cannot sufficiently determine the prognosis^{3,7,8}. Thus, new biomarkers available to estimate sepsis severity and enable earlier treatment initiation

Histidine-rich glycoprotein (HRG) is a 75 kDa plasma glycoprotein produced in the liver and present at concentration of approximately 100 µg/ml⁹. It is considered to be involved in many functions in biological

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Variable	Total	Sepsis	Severe sepsis	Septic shock
N	99	10	43	46
Age, years	72.0 (64.0-78.0)	75.0 (68.3–80.5)	72.0 (64.0-80.0)	71.0 (60.8–76.3)
Male sex, n	69 (69.7%)	8 (80.0%)	28 (65.1%)	33 (71.7%)
28-day death, days	16 (16.2%)	0	8 (18.6%)	8 (17.4%)
ICU stay, days	10.0 (5.0–16.0)	7.5 (3.8–15.5)	11.0 (5.0–18.0)	9.5 (5.0–16.0)
APACHE II score	25.0 (21.0-31.0)	20.0 (13.5–21.8)	24.0 (22.0-30.0)	28.0 (21.5-34.5)
SOFA score	11.0 (8.0-13.0)	4.5 (2.0-6.0)	9.0 (7.0-12.0)	13.0 (11.0-16.0)
Source of infection, n (death)	99 (16)			
Lung	20 (4)	3	7 (2)	10 (2)
Gastrointestinal	19 (2)	3	10 (1)	6 (1)
Hepatic	1	0	0	1
Gallbladder	8	0	3	5
Urinary	14 (1)	1	6 (1)	7
Bone/soft tissue	13 (2)	2	5 (1)	6(1)
Others	24 (7)	1	12 (3)	11 (4)
Ventilation, n	53 (53.5%)	5 (50.0%)	23 (53.5%)	25 (54.3%)
Inotropes (Day 1)	54 (54.5%)	1 (10.0%)	7 (16.3%)	46 (100%)
Blood purification		•	•	•
Chronic dialysis	4 (4.0%)	0	1 (2.3%)	3 (6.5%)
Renal replacement therapy	29 (29.3%)	0	15 (34.9%)	14 (30.4%)
Polymyxin B hemoperfusion	7 (7.1%)	0	3 (7.0%)	4 (8.7%)
Liver failure	2 (2.0%)	0	2 (4.7%)	0
AIDS	0	0	0	0
Hematologic malignancies	6 (6.1%)	1 (10.0%)	1 (2.3%)	4 (8.7%)

Table 1. Patient characteristics. Expressed as median (interquartile range). *APACHE* acute physiology and chronic evaluation, *SOFA* sequential organ failure assessment, *AIDS* acquired immunodeficiency syndrome.

systems, such as coagulation, immune response, and modulation of angiogenesis^{9,10}. In particular, HRG is suggested to play an essential role in host defense mechanisms^{10–12}. Recently, we reported that plasma HRG levels rapidly decreased in mice with sepsis¹³. Moreover, supplemental HRG infusion significantly improved the sepsis survival rate in mice, while the knockdown of HRG levels exacerbated mortality¹³. Since HRG strongly induced the spherical shape in human neutrophils, suppressed the neutrophil adhesion to vascular endothelial cells, maintained passage through microcapillaries, and reduced ROS production, these effects may contribute to the beneficial effects of HRG detected in septic animals¹³. Furthermore, HRG has been recently reported to protect the vascular endothelial cells from LPS-induced disorder and increased permeability¹⁴. Our animal study as well as in vitro experiments suggested that plasma HRG might be a useful biomarker of sepsis and that supplemental therapy with HRG may provide a novel strategy for the treatment of sepsis. In fact, our previous clinical study on 70 patients with systemic inflammatory response syndrome (SIRS) showed that HRG levels of septic patients were significantly lower than those of noninfective SIRS patients and that HRG levels were significantly associated with mortality within the SIRS population¹⁵. However, whether HRG could predict the prognosis of septic patients remains unclear.

This study included septic patients according to Sepsis-2 definition ¹⁶ and evaluate the survival rate of septic patients classified according to plasma HRG levels. HRG levels were also compared with clinical parameters such as PCT in septic patients.

Results

Patients. Patients were prospectively enrolled from October 2014 to September 2016 in 11 Japanese hospitals: 2 university and 9 general hospitals. Written informed consent was obtained from 101 patients. Two were excluded because of lack of 28-day survival data, and finally, 99 patients were analyzed. The patient characteristics are shown in Table 1 and Supplementary Tables S1 and S2. The median age was 72.0 (interquartile ranges [IQR], 64.0–78.0) years and 70% were males (69 males and 30 females). Their median acute physiology and chronic evaluation (APACHE) II and sequential organ failure assessment (SOFA) scores were 25.0 (IQR, 21.0–31.0) and 11.0 (IQR, 8.0–13.0), respectively. The 28-day mortality rate was 16.2% (16/99 cases). A total of 46 patients (46.5%) were diagnosed with septic shock, 43 (43.4%) with severe sepsis, and 10 (10.1%) with sepsis according to the Sepsis-2 definition. The 28-day mortality rates in patients with septic shock, severe sepsis, and sepsis were 17.4% (8/46 patients), 18.6% (8/43 patients), 0% (0/10 patients), respectively.

Plasma HRG and PCT levels on the first day of ICU admission. The median HRG level of 99 septic patients was 26.49 (IQR, 19.55–38.40) μg/ml in this study.

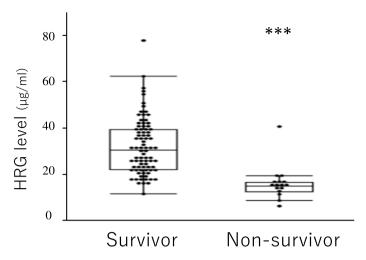


Figure 1. Plasma HRG levels in septic patients on the first day of ICU admission. We compared HRG levels between survivors (n = 83) and non-survivors (n = 16). A box-and-whisker plot showing median, 25th, and 75th percentiles. The bars represent the 5th and 95th percentiles. Circles represent HRG levels in patients. ***P < 0.01. *HRG* histidine-rich glycoprotein.

	Univariate analysis			Adjusted with APACHE II score	
Variables	HR (95% CI)	P	Harrell C-index	Adjusted HR (95% CI)	P
HRG	0.79 (0.71-0.87)	< 0.01	0.90	0.81 (0.72-0.89)	< 0.01
PCT	0.9993 (0.99-1.00)	0.81	0.54	NA	NA
APACHE II score	1.12 (1.06-1.20)	< 0.01	0.72	NA	NA
SOFA score	1.19 (1.04-1.37)	0.010	0.68	NA	NA

Table 2. Associations between each variable and mortality. We used Cox's proportional hazard model to evaluate associations between each variable and mortality. *HRG* histidine-rich glycoprotein, *PCT* procalcitonin, *APACHE* Acute Physiology and Chronic Evaluation, *SOFA* Sequential Organ Failure Assessment, *HR* hazard ratio, *Adjusted HR* hazard ratio adjusted according to APACHE II score.

Figure 1 shows the primary outcome results. Comparing the non-survivors (n = 16) and survivors (n = 83), HRG levels in the former were significantly lower than those in the latter (median, 15.1 [IQR, 12.7–16.6] vs. 30.6 [22.1–39.6] μ g/ml; Mann–Whitney test, P < 0.01). PCT levels in non-survivors and survivors were 9.23 (IQR, 3.30–100) and 31.8 (11.0–84.2) ng/ml, respectively, but without statistical differences (P = 0.42).

HRG levels in patients with septic shock, severe sepsis, and sepsis were 25.5 (IQR, 17.6–37.6), 30.8 (19.6–40.9), and 28.9 (22.2–37.8) μ g/ml, respectively, but without statistical differences (Kruskal–Wallis test, P = 0.48) (Supplementary Figure S1).

There were no significant correlations between HRG levels and other parameters, such as white blood cell count (Spearman's rho, 0.071, P = 0.48), C-reactive protein (-0.0018, P = 0.99), PCT (0.089, P = 0.38), platelet number (0.039, P = 0.70), and fibrinogen (0.066, P = 0.55).

Survival analysis. Table 2 shows associations between plasma levels of each marker and mortality. Higher first-day HRG levels were significantly associated with lower risk of mortality. The hazard ratio (HR) was 0.79 (95% confidence interval [CI], 0.71-0.87, P < 0.01). Even after adjusting with the APACHE II score, the golden standard in evaluating the patient's severity, the HRG level remained an independent prognostic factor (adjusted HR, 0.81; 95% CI, 0.72-0.89; P < 0.01). PCT levels were not statistically associated with mortality (P = 0.81). The Harrell C-index for mortality was as follows: HRG, 0.90 (95% CI, 0.81-0.99); PCT, 0.54 (0.37-0.71); APACHE II score, 0.72 (0.58-0.86); and SOFA score, 0.68 (0.56-0.81).

When patients were divided into four subgroups according to quartiles of HRG level, Kaplan–Meier curves (Fig. 2A) showed that the mortality in the Q1 group (lowest HRG subgroup) was significantly higher than that in others (log-rank test, Bonferroni corrected P < 0.01).

Supplementary Table S3 shows that the sensitivity and specificity of HRG levels associated with mortality at the cut-off level of $20.0 \mu g/ml$ was 0.94 and 0.84, respectively. Thus, when patients were divided into high and low HRG groups according to this cut-off level, Kaplan–Meier curves (Fig. 2B) showed that the mortality in the latter group was significantly higher than that in the former (HR, 57.2; 95% CI, 11.5-1034; P < 0.01).

We analyzed whether HRG was associated with mortality, regardless of the severity of organ dysfunction. Supplementary Figure S2 shows the associations between HRG level and mortality in subgroups stratified with

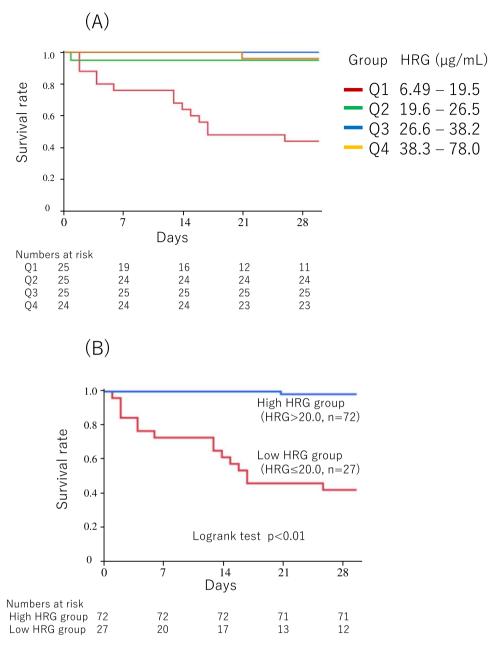


Figure 2. Kaplan–Meier survival curves. **(A)** Patients were divided into four subgroups according to quartiles of HRG level; Q1, 6.49–19.5 μ g/ml, n = 25; Q2, 19.6–26.5 μ g/ml, n = 25; Q3, 26.6–38.2 μ g/ml, n = 25; Q4, 38.3–78.0 μ g/ml, n = 24. The mortality of the Q1 group was significantly higher than others (log-rank test, Bonferroni corrected P < 0.01). **(B)** Patients were divided into high and low HRG groups according to the cut-off level of 20.0 μ g/ml. At the cut-off level of 20.0 μ g/ml, the sensitivity and specificity of HRG levels associated with mortality was 0.94 and 0.84, respectively. *HRG* histidine-rich glycoprotein.

APACHE II score and each parameter of SOFA score: shock, respiratory dysfunction, liver dysfunction, renal dysfunction, and platelet number. HRG levels were shown to be significantly associated with mortality in all subgroups.

Discussion

In this multicenter prospective study, first-day HRG levels in non-survivors were demonstrated to be significantly lower than those in survivors among septic patients. HRG levels were also shown to be significantly associated with mortality, and determination of HRG provided a good prognostic estimation as a biomarker for sepsis. Moreover, these properties could also be found in patients with organ failure, such as severe hepatic or renal failure.

In our previous investigation, 70 patients with SIRS, including 20 with sepsis, were examined. We showed that HRG levels in septic patients were significantly lower than those in noninfective SIRS patients and that

HRG levels were significantly associated with mortality within the SIRS population¹⁵. However, the association of HRG with sepsis survival was difficult to determine due to the limited sample size¹⁷. In the current study, we focused on sepsis and investigated a new cohort of 99 septic patients from 11 institutes, and we could obtain similar results. Thus, these results were reproducible in independent cohorts.

In this study, first-day HRG levels were shown to be associated with mortality. Harrell C-index, the predictive power index, for mortality was 0.90 for HRG. This C-index was superior to the C-indices of APACHE II (0.72) and SOFA scores (0.68), the golden standards in evaluating the sepsis severity. The SOFA score is included in the new definition of sepsis (Sepsis-3)¹⁸, and previous studies demonstrated that SOFA score is a good indicator of prognosis in ICU-admitted critically ill patients, especially those with severe sepsis¹⁹⁻²¹. Conversely, Liu et al. reported in their meta-analysis that PCT was associated with mortality in septic patients and can moderately predict sepsis mortality⁶. However, the prognostic accuracy of PCT has been reported to be inadequate for clinical use, especially in patients with renal failure^{22,23}. In the current study, the C-index for PCT was 0.54 and inferior to that for HRG. Therefore, we suggest that HRG might be superior to APACHE II score, SOFA score, and PCT as a prognostic biomarker.

First-day $\dot{H}RG$ levels were found to distinguish survivors from non-survivors. Kaplan–Meier curves (Fig. 2A) clearly showed that only Q1 group, patients with first-day HRG levels of < 19.5 μ g/ml, had high mortality rate of \geq 50% and that almost all patients in other groups survived. This unique property would facilitate the determination of death threshold when using HRG as a biomarker. In other words, HRG was thought to identify patients who really need intensive treatments. Intensive and timely care for patients with first-day HRG levels of < 19.5 μ g/ml may possibly reduce sepsis mortality. In addition, our previous animal study suggested that supplemental infusion of HRG might provide a novel strategy for the treatment of sepsis 13, indicating that HRG might serve as a therapeutic as well as a biomarker.

Moreover, when patients were divided into subgroups according to the degree of organ dysfunctions, reduced HRG levels were shown to be significantly associated with mortality in all subgroups, including hepatic and renal dysfunction (Supplementary Figure S2). Current sepsis biomarkers have problems in evaluating patients with organ failure⁷. There are some cases in which PCT cannot work well, especially in patients with renal failure^{7,23}. Our data suggested that HRG had sufficient ability in evaluating the severity of septic patients in these settings, including hepatic and renal failures. These results strongly suggested that plasma HRG would be a useful prognostic biomarker for sepsis.

This study had some limitations. First, the old definition of sepsis (Sepsis-2) was used because this study was started before the Sepsis-3 definition was established. Old data could affect interpretation of our results. In this study, all patients had suspicious infection and SOFA score of 2 or higher, however it is unclear whether it was increase of > 2 score. Therefore, another study should be conducted using the Sepsis-3 definition, and we have already started the study. Second, although we focused on sepsis survival, we studied 99 septic patients, including 16 non-survivors. To reinforce our claim that HRG could be a prognostic marker for sepsis, a larger study should be conducted. Third, we had no data on exact number of eligible patients and the incidence rate of sepsis in 11 hospitals. Selection bias could affect the results. Fourth, only one-time HRG levels were assessed on day 1 of ICU, and we had no data on the time course and the exact timing of blood sampling. Timing of sampling in consideration of treatment, and time-dependent changes in HRG levels would be more valuable and may reflect treatment effects. Whether HRG-guided therapy is possible with time course data may be evaluated. Fifth, we had no data regarding patient comorbidities and the pathogen. Further investigations on the association of HRG with patient characteristics and with the causes of sepsis will be conducted. Sixth, because of the lack of exact data on the cause of death, it is possible that patients with withdrawal from treatment were included, which could affect the results. Seventh, for measuring HRG, we used our in-house ELISA, which established in our previously published study. This may affect the results of this study.

In conclusion, in this study, first-day plasma HRG levels in non-survivors were found to be significantly lower than those in survivors among septic patients. HRG were also found to have a good prognostic accuracy as a biomarker for sepsis. Therefore, we suggest that HRG might be a good biomarker in diagnosing sepsis, evaluating sepsis severity, and predicting patient outcomes. Studies with larger cohorts are needed to confirm our findings.

Methods

Study design. This was a multicenter, prospective, and observational investigation approved by the Institutional Review Board of any institution involved, including the Okayama University Graduate School of Medicine, Dentistry, and Pharmaceutical Sciences. It was registered at the UMIN Clinical Trials Registry (UMIN000017651, registered 22 May 2015—retrospectively registered, principal investigator's name is Hiroshi Morimatsu). This study was performed in accordance with the ethical standards laid down in the Declaration of Helsinki and its later amendments. This observational study was also reported following the Strengthening the Reporting of Observational Studies in Epidemiology guidelines²⁴.

Patients and data collection. Patients who were newly diagnosed with sepsis according to Sepsis-2¹⁶ were prospectively enrolled in the study. The inclusion criteria were patients admitted to the ICU and the occurrence of sepsis. The exclusion criteria were age younger than 20 years, pregnancy, or failure to obtain consent.

Clinical and laboratory data were obtained daily while in the ICU. Initial SOFA score and APACHE II score were calculated using clinical parameters and blood test results. Patients were classified according to the previous definitions of sepsis (Sepsis-2), the guidelines of the American College of Chest Physicians/Society of Critical Care Medicine and the International Surviving Sepsis Campaign Guidelines Committee 16. A follow-up at 28 days was performed to determine survivors and non-survivors.

Measurement methods. After obtaining written informed consent for participation from the patients or their relatives, blood samples were collected in tubes containing K_2 -EDTA at a convenient timing within 24 h of sepsis diagnosis, regardless of the treatment. The samples were processed and immediately frozen in each hospital, then transported to Okayama University, and stored at -80 °C for later analysis.

HRG levels were determined using a modified quantitative sandwich enzyme-linked immunosorbent assay (ELISA) as described previously¹⁵. In brief, we used a rat monoclonal antibody against human HRG (made inhouse, number 75-14) as the capture antibody and horseradish peroxidase-conjugated nickel-nitrilotriacetic acid (Ni-NTA HRP conjugate; Qiagen, Venlo, Netherland) for detection instead of secondary antibody. Plasma samples were diluted 1:50 in phosphate-buffered saline (PBS) containing 1% bovine serum albumin (BSA) for measurement. A standard curve is established using serial dilutions of known amounts of purified HRG (made in-house). Each plasma sample was tested in duplicate, and assays were repeated twice independently.

Outcomes. The primary outcome was HRG levels in survivors and non-survivors. The secondary outcome was the association between HRG levels and mortality.

Statistical analysis. The statistical approach was mostly designed a priori. Only the analysis provided in Supplementary Figure S2 arose out of the data exploration process and was therefore designed post hoc.

The median, interquartile ranges (IQR, 25th–75th percentiles), and box-whisker plot were used to summarize the variables. The ICU first-day markers were compared using the Mann–Whitney/Kruskal–Wallis tests between/among subgroups, and the statistical significance level was set as 0.05 (two-sided). The contribution of each marker to the 28-day mortality was assessed using Cox's proportional hazard model adjusted with APACHE II score, also especially in case of HRG, stratified with each parameter of SOFA score. The hazard ratio (HR) and 95% confidence interval (CI) were estimated. The cumulative survival rate was estimated and tested using the Kaplan–Meier method and the log-rank test with Bonferroni correction for four subgroups according to quartiles of HRG level, and using the Kaplan–Meier method and the log-rank test for subgroups defined by the cut-off value of HRG specified with the logistic regression receiver operating characteristic analysis; the HR was calculated using the Cox's proportional hazard model. Spearman rank correlation coefficient was calculated to assess correlations between HRG and other variables. All analyses were performed using the JMP Pro 12 software (SAS Institute Inc., Chicago, IL), except for the Harrell C-index that was performed using the STATA 12 software (SAS Institute Inc.).

Data availability

The datasets generated and analysed during the current study are available from the corresponding author on reasonable request.

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References

- 1. Rhee, C. et al. Incidence and trends of sepsis in US hospitals using clinical vs claims data, 2009–2014. JAMA 318, 1241–1249 (2017).
- Fleischmann, C. et al. Assessment of global incidence and mortality of hospital-treated sepsis current estimates and limitations. Am. J. Respir. Crit. Care Med. 193, 259–272 (2016).
- 3. Wu, C. C. et al. Comparison of diagnostic accuracy in sepsis between presepsin, procalcitonin, and C-reactive protein: A systematic review and meta-analysis. Ann. Intensive Care 7, 91 (2017).
- 4. Bloos, F. & Reinhart, K. Rapid diagnosis of sepsis. Virulence 5, 154-160 (2014).
- 5. Rhodes, A. *et al.* Surviving sepsis campaign: International guidelines for management of sepsis and septic shock: 2016. *Intensive Care Med.* **43**, 304–377 (2017).
- Liu, D. et al. Prognostic value of procalcitonin in adult patients with sepsis: A systematic review and meta-analysis. PLoS ONE 10, e0129450 (2015).
- 7. Yang, H. S. et al. Prognostic value of presepsin in adult patients with sepsis: Systematic review and meta-analysis. PLoS ONE 13, e0191486 (2018).
- 8. Parlato, M. et al. Circulating biomarkers may be unable to detect infection at the early phase of sepsis in ICU patients: The CAP-TAIN prospective multicenter cohort study. Intensive Care Med. 44, 1061–1070 (2018).
- 9. Poon, I. K. H., Patel, K. K., Davis, D. S., Parish, C. R. & Hulett, M. D. Histidine-rich glycoprotein: The Swiss Army knife of mammalian plasma. *Blood* 117, 2093–2101 (2011).
- Blank, M. & Shoenfeld, Y. Histidine-rich glycoprotein modulation of immune/autoimmune, vascular, and coagulation systems. Clin. Rev. Allergy Immunol. 34, 307–312 (2008).
- 11. Shannon, O. et al. Histidine-rich glycoprotein promotes bacterial entrapment in clots and decreases mortality in a mouse model of sepsis. Blood 116, 2365–2372 (2010).
- 12. Nishimura, Y. *et al.* Histidine-rich glycoprotein augments natural killer cell function by modulating PD-1 expression via CLEC-1B. *Pharmacol. Res. Perspect.* 7, e00481 (2019).
- 13. Wake, H. et al. Histidine-rich glycoprotein prevents septic lethality through regulation of immunothrombosis and inflammation. EBioMedicine 9, 180–194 (2016).
- 14. Gao, S. et al. Histidine-rich glycoprotein ameliorates endothelial barrier dysfunction through regulation of NF-kB and MAP signal pathway. Br. J. Pharmacol. 176, 2808–2824 (2019).
- Kuroda, K. et al. Decrease in histidine-rich glycoprotein as a novel biomarker to predict sepsis among systemic inflammatory response syndrome. Crit. Care Med. 46, 570–576 (2018).
- Levy, M. M. et al. 2001 SCCM/ESICM/ACCP/ATS/SIS International sepsis definitions conference. Crit. Care Med. 31, 1250–1256 (2003).
- 17. Nishibori, M., Wake, H. & Morimatsu, H. Histidine-rich glycoprotein as an excellent biomarker for sepsis and beyond. *Crit. Care* 22, 209 (2018).
- 18. Singer, M. et al. The third international consensus definitions for sepsis and septic shock (Sepsis-3). JAMA 315, 801-810 (2016).

- 19. Yousef, A. A. & Suliman, G. A. The predictive prognostic values of serum TNF-α in comparison to SOFA score monitoring in critically ill patients. *Biomed. Res. Int.* **2013**, 258029 (2013).
- Jones, A. E., Trzeciak, S. & Kline, J. A. The sequential organ failure assessment score for predicting outcome in patients with severe sepsis and evidence of hypoperfusion at the time of emergency department presentation. Crit. Care Med. 37, 1649–1654 (2009).
- Vincent, J. L. et al. Use of the SOFA score to assess the incidence of organ dysfunction/failure in intensive care units: Results of a multicenter, prospective study. Crit. Care Med. 26, 1793–1800 (1998).
- 22. Henriquez-Camacho, C. & Losa, J. Biomarkers for sepsis. Biomed. Res. Int. 2014, 547818 (2014).
- Grace, E. & Turner, R. M. Use of procalcitonin in patients with various degrees of chronic kidney disease including renal replacement therapy. Clin. Infect. Dis. 59, 1761–1767 (2014).
- 24. von Elm, E. et al. Strengthening the reporting of observational studies in epidemiology (STROBE) statement: Guidelines for reporting observational studies. BMJ 335, 806–808 (2007).

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Author contributions

K.K., H.W., M.N., and H.M. contributed to the study concept and design. K.K., K.I., Y.M., H.M. and the ORION investigators participated in data collection. K.K., N.K., H.W., S.M., and M.N. did the laboratory work. All authors participated in data analysis and interpretation. K.K., M.Y., and H.M. contributed to the statistical analysis. K.K., N.K., H.W., M.N., and H.M. contributed to drafting of the report. M.N. and H.M. were responsible for obtaining funding. M.N. and H.M. supervised the study.

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Competing interests

The authors declare no competing interests.

Additional information

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