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Article

Anticoagulation with osocimab in patients with kidney failure undergoing hemodialysis: a randomized phase 2 trial

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Individuals with kidney failure undergoing hemodialysis are at elevated risk for thromboembolic events. Factor (F) XI, which is in the intrinsic pathway of coagulation, is emerging as an attractive target for new anticoagulants that may be safer than existing agents. Osocimab-an inhibitory FXIa antibody-is a potential treatment option for such patients. We conducted a phase 2b, double-blind, placebo-controlled trial, in which 704 participants (448 male, 256 female) with kidney failure undergoing hemodialysis were randomized to receive lower- or higher-dose osocimab or placebo. In total, 686 participants (436 male, 250 female) received treatment for ≤18 months (planned minimal treatment period of 6 months). The co-primary outcomes were clinically relevant bleeding (a composite of major and clinically relevant nonmajor bleeding) and a composite of the incidence of moderate, severe or serious adverse events. Clinically relevant bleeding occurred in 16/232 (6.9%) and 11/224 (4.9%) participants who received lower- and higher-dose osocimab, respectively, and in 18/230 participants (7.8%) who received a placebo. For the composite adverse event endpoint, incidences were 51%, 47% and 43% in the lower-dose osocimab, higher-dose osocimab and placebo groups, respectively. These results suggest that osocimab is associated with a low risk of bleeding and is generally well tolerated in this population; findings that require confirmation in larger trials. Clinical Trials. gov identifier, NCT04523220.

Chronic kidney disease, which affects almost 10% of the global population, is a major cause of morbidity and mortality¹. Diabetes and hypertension are common causes of kidney failure². Most individuals with kidney failure who are managed with hemodialysis are at risk of major adverse vascular events such as myocardial infarction, stroke and other thromboembolic events^{2,3}. Although long-term anticoagulation therapy has the potential to prevent these complications, its use remains understudied because people with kidney failure are also at

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Fig. 1 | **Study design.** Overview of the study design. Participants with kidney failure undergoing hemodialysis were randomized to receive lower- or higher-dose osocimab or placebo for \leq 18 months. The follow-up period ended 5 months after the last study intervention and 4 months after the end of the main or extension treatment period.

increased risk of major bleeding⁴. Around one in seven such individuals will experience a major bleed within 3 years of dialysis initiation⁵, and the bleeding risk is further increased, up to tenfold, when people with kidney failure requiring hemodialysis are treated with warfarin⁶. Furthermore, a recent trial comparing warfarin with apixaban for stroke prevention in patients with atrial fibrillation on regular hemodialysis reported comparably high rates of major bleeding (9.7% versus 8.5%)⁷. Therefore, with limited evidence supporting the use of warfarin or direct oral anticoagulants in this patient population, there remains an unmet need for safer anticoagulants for the prevention of thromboembolic events.

Inhibitors of factor XI or its activated form, factor XIa, may be safer than the currently available anticoagulants because factor XI seems to be more important for thrombosis than for hemostasis⁸. Thus, individuals with reduced factor XI levels are at lower risk for thrombosis than those with normal levels⁹, but rarely experience serious bleeding⁸. Conversely, high levels of factor XI increase the risk for thrombosis^{8,10,11}. Therefore, factor XI has emerged as a target for potentially safer anticoagulants and may be a particularly attractive target in individuals with kidney failure requiring regular dialysis³.

Osocimab is a long-acting, fully human inhibitory antibody directed against factor XIa^{12,13}. In a phase 2 dose-finding study in individuals undergoing total knee arthroplasty, a single intravenous dose of osocimab administered before surgery was superior, or after surgery was noninferior, to enoxaparin for prevention of postoperative venous thromboembolism¹⁴. Furthermore, osocimab was associated with low rates of clinically relevant bleeding, the composite of major and clinically relevant nonmajor bleeding¹⁴. To address the unmet need for safer anticoagulants for patients with end-stage kidney disease, we conducted the CONVERT trial (ClinicalTrials.gov identifier, NCT04523220) to compare the rates of clinically relevant bleeding with osocimab and placebo in participants with kidney failure who were undergoing hemodialysis (Fig. 1).

Results

Patient disposition

From August 2020 to April 2021, a total of 704 participants from 147 sites in 19 countries were randomized. After the exclusion of 18 participants who did not receive the study drug, 686 individuals started the main 6-month treatment period and were included in the analyses (Fig. 2). The main treatment period was completed by 199 participants

in the lower-dose osocimab group (84.7%), 194 participants in the higher-dose osocimab group (82.9%) and 206 participants in the placebo group (87.7%). The extended treatment period was completed by 178 participants in the lower-dose osocimab group (75.5%), 174 participants in the higher-dose osocimab group (74.4%) and 176 participants in the placebo group (74.9%). The median previous duration on maintenance dialysis was 4.0 years (interquartile range, 2.0–7.2 years). Overall, 249 participants (36.3%) had known atherosclerotic disease and 290 (42.3%) were taking aspirin. The baseline characteristics and median duration of study treatment were similar across groups (Table 1).

Primary outcomes

Clinically relevant bleeding (the composite of major and clinically relevant nonmajor bleeding) occurred in 16 of 232 participants (6.9%) who received lower-dose osocimab, 11 of 224 (4.9%) who received higher-dose osocimab and 18 of 230 (7.8%) who received placebo (Table 2). Major bleeding occurred in three (1.3%), two (0.9%) and seven (3.0%) participants in the lower- and higher-dose osocimab and placebo groups, respectively. In those over 75 years of age, clinically relevant bleeding per 100 person-years was 20 (90% confidence interval (Cl), 7–39) events for lower-dose osocimab, 8 (90% Cl, 1–18) events for higher-dose osocimab and 23 (90% Cl, 8–44) events for placebo. Corresponding values for individuals on low-dose aspirin were 12 (90% Cl, 6–20), 7 (90% Cl, 3–13) and 16 (90% Cl, 9–25) events per 100 person-years, respectively. Cause-specific hazard ratios and subdistribution hazard ratios for the primary outcomes are provided in Extended Data Tables 1–4.

The incidences of moderate, severe or serious adverse events were 51% (n = 118), 47% (n = 106) and 43% (n = 99) in the lower-dose osocimab, higher-dose osocimab and placebo groups, respectively (Table 2).

Secondary outcomes

Prespecified secondary outcomes were osocimab plasma levels, prolongation of the activated partial thromboplastin time and inhibition of factor XIa. There were dose-dependent changes in these variables and their timecourses were comparable from the first to the last measured dose (Fig. 3).

Safety

Injection-site reactions occurred in 14, 17 and 1 participant in the lower-dose osocimab, higher-dose osocimab and placebo groups,



Fig. 2 | Patient disposition. Summary of patient flow in the phase 2b CONVERT trial.

respectively, resulting in study drug discontinuation by two participants from each osocimab group and no participants in the placebo group. Further details regarding the incidence of treatment-related serious adverse events are provided in Extended Data Table 5.

In total, 26 participants received lower- or higher-dose osocimab (n = 13 each) within 30 days preceding major surgery or intervention (Extended Data Table 6). The predicted median factor XIa inhibition levels at the time of surgery or intervention were 58% and 69% with the lower and higher osocimab doses, respectively, and no bleeding events were reported during or within 2 weeks of surgery or intervention. Of those randomized to placebo or randomized to osocimab but who did not receive any treatment, 21 participants underwent major surgery or intervention. There was one procedure-related, clinically relevant nonmajor bleed after dialysis catheter removal in the placebo group.

Exploratory outcomes

For the prespecified exploratory outcomes, major adverse vascular events occurred in 3 of 232 (1.3%) participants who received lower-dose osocimab, 6 of 224 (2.7%) who received higher-dose osocimab and 7 of 230 (3.0%) who received placebo. In participants with known atherosclerotic disease, major adverse vascular events occurred in 4 of 167 (2.4%) individuals who received osocimab and in 6 of 82 (7.3%) who received placebo. No participant experienced symptomatic venous thromboembolism.

Clotting of the dialysis circuit was scored in each participant at every study visit (0, no clot; 1, trace of clot; 2, intermediate between 1 and 3; and 3, fully clotted system necessitating interruption of hemodialysis session). Dialysis circuit clotting scores of 2 or 3 at any visit were reported in 29.3% and 27.2% of participants in the lower- and higher-dose osocimab groups, respectively, and in 41.3% of those in the placebo group. The relative risk of moderate-to-complete clotting at one or more visits was significantly lower in both osocimab groups than in the placebo group (lower dose versus placebo, 0.71 (95% Cl, 0.54–0.93; P = 0.0085); higher dose versus placebo, 0.66 (95% Cl, 0.49–0.87; P = 0.0021)).

Sensitivity analyses

Aalen–Johansen estimates of the cumulative incidences (that is, the expected proportions of participants with an outcome over time, taking competing risks into account) for the primary outcomes and the exploratory efficacy outcome are provided in Extended Data Figs. 1–3.

For the primary outcome of clinically relevant bleeding, the cumulative incidence risks over the main treatment period were 4.3% (90% CI, 2.48–6.90) for lower-dose osocimab, 3.57% (90% CI, 1.91–6.04) for higher-dose osocimab and 6.09% (90% CI, 3.84–9.04) for placebo. For the second primary outcome of the composite of moderate, severe or serious adverse events over the main treatment period, the cumulative incidence risks were 38.37% (90% CI, 33.10-43.60) for lower-dose osocimab, 32.17% (90% CI, 27.16-37.28) for placebo.

For the exploratory efficacy outcome of the incidence of arteriovenous fistula or graft thrombosis over the main treatment period, the cumulative incidence risks were 1.7% (90% CI, 0.70-3.61) for lower-dose osocimab, 2.68% (90% CI, 1.29-4.91) for higher-dose osocimab and 3.91% (90% CI, 2.18-6.43) for placebo.

Discussion

Individuals with kidney failure requiring hemodialysis are at increased risk of bleeding^{3,4}. The results of the primary outcomes of this phase 2b trial suggest that, compared with placebo, osocimab does not increase

Table 1 | Baseline demographics and clinical characteristics of study participants

Characteristic	Lower-dose osocimab (n=232)	Higher-dose osocimab (n=224)	Placebo (n=230)
Age, median (range) (years)	61 (28–91)	61 (25–90)	60 (24–90)
Age group, n (%)			
<60 years	104 (44.8)	103 (46.0)	108 (47.0)
60–75 years	98 (42.2)	84 (37.5)	94 (40.9)
>75 years	30 (12.9)	37 (16.5)	28 (12.2)
Male sex, n (%)	143 (61.6)	143 (63.8)	150 (65.2)
Body mass index ^a , <i>n</i> (%)			
25-30kgm ⁻²	62 (26.7)	87 (38.8)	93 (40.4)
≥30 kg m ⁻²	74 (31.9)	65 (29.0)	62 (27.0)
White, <i>n</i> (%)	191 (82.3)	185 (82.6)	185 (80.4)
Geographic region, n (%)			
Western Europe	36 (15.5)	37 (16.5)	35 (15.2)
Eastern Europe	133 (57.3)	121 (54.0)	123 (53.4)
Asia Pacific	21 (9.0)	19 (8.4)	19 (8.2)
North America	39 (16.8)	37 (16.5)	37 (16.1)
Australia and Israel	3 (1.2)	10 (4.4)	16 (6.9)
Dialysis duration, median (interquartile range) years	4.05 (2.0–7.2)	4.0 (2.0–7.5)	3.85 (1.8–7.0)
Dialysis access, n (%)			
Fistula	191 (82.3)	180 (80.4)	189 (82.2)
Graft	22 (9.5)	20 (8.9)	13 (5.7)
Catheter	19 (8.2)	24 (10.7)	28 (12.2)
Heparin use during dialysis, n (%)	221 (95.3)	218 (97.3)	218 (94.8)
Etiology of kidney disease, n (%)			
Diabetes	65 (28.0)	53 (23.6)	59 (25.6)
Hypertension	51 (22.0)	67 (30.0)	50 (21.7)
Glomerulonephritis	32 (13.8)	30 (13.4)	33 (14.3)
Pyelonephritis	11 (4.7)	10 (4.5)	10 (4.3)
Polycystic kidney disease	23 (9.9)	23 (10.3)	21 (9.1)
Other	50 (21.5)	41 (18.3)	57 (24.8)
Diabetes, n (%)	90 (38.8)	82 (36.6)	87 (37.8)
Hypertension, n (%)	213 (91.8)	212 (94.6)	215 (93.5)
Coronary heart disease, n (%)	56 (24.1)	61 (27.2)	47 (20.4)
Myocardial infarction, n (%)	17 (7.3)	15 (6.7)	11 (4.8)
Peripheral artery disease, n (%)	18 (7.8)	23 (10.3)	22 (9.6)
Stroke, n (%)	12 (5.1)	18 (7.7)	18 (7.7)
Atrial fibrillation, n (%)	15 (6.5)	17 (7.6)	14 (6.1)
Previous cardiovascular event ^b , <i>n</i> (%)	37 (15.9)	43 (19.2)	40 (17.4)
Atherosclerosis°, n (%)	81 (34.9)	86 (38.4)	82 (35.7)
History of venous thromboembolism, <i>n</i> (%)	8 (3.4)	10 (4.5)	10 (4.3)
Platelet count, mean (s.d.)×10 ⁹ l ⁻¹	209 (62)	209 (56)	212 (58)
Hemoglobin, median (interquartile range) (mg dl-1)	10.9 (10.2–11.8)	11.1 (10.4–11.8)	11.0 (10.4–11.7)
Low-dose aspirin, n (%)	97 (41.8)	95 (42.4)	98 (42.6)
Duration of study drug administration, median (interquartile range) (months)	9.1 (7.0–11.7)	9.2 (7.1–11.7)	9.0 (7.1–11.4)
Length of study, n (%)			
≤6 months	37 (15.9)	33 (14.7)	26 (11.3)
>6-9months	76 (32.8)	73 (32.6)	89(38.7)
>9 months	119 (51.3)	118 (52.7)	115 (50.0)
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Percentages may not total 100 because of rounding. ^aBody mass index is the weight in kilograms divided by the square of the height in meters. ^bPrevious cardiovascular events include stroke, transient ischemic attack, myocardial infarction, deep vein thrombosis and pulmonary embolism. ^cAtherosclerosis was defined as a history of ischemic stroke, transient ischemic attack, unstable angina, myocardial infarction, peripheral artery disease or aortic aneurysm.

Table 2 | Safety and efficacy outcomes

	Lower-dose osocimab (n=232)	Higher-dose osocimab (n=224)	Placebo (n=230)
Clinically relevant bleeding (primary outcome)			
n (%)	16 (6.9)	11 (4.9)	18 (7.8)
Events (90% CI) per 100 patient-years	9.7 (6.1–14.1)	6.7 (3.7–10.3)	10.8 (7.0–15.3)
Composite of moderate, severe or serious adverse events (primary o	utcome)		
n (%)	118 (50.9)	106 (47.3)	99 (43.0)
Major bleeding			
n (%)	3 (1.3)	2 (0.9)	7 (3.0)
Events (90% CI) per 100 patient-years	1.8 (0.5–3.8)	1.2 (0.2–2.9)	4.2 (2.0–7.1)
Clinically relevant nonmajor bleeding			
n (%)	13 (5.6)	9 (4.0)	11 (4.8)
Events (90% CI) per 100 patient-years	7.9 (4.7–11.8)	5.5 (2.9–8.8)	6.6 (3.7–10.2)
Types of major bleeding, n (%)			
Gastrointestinal	1 (0.4)	0	1 (0.4)
Urogenital (kidney or bladder)	0	0	2 (0.9)
Intracranial	0	0	1 (0.4)
Eye (intraocular or retinal)	1 (0.4)	0	2 (0.9)
Skin (any vascular access site)	0	2 (0.9)	0
Respiratory (pulmonary)	1 (0.4)	0	0
Procedural	0	0	1 (0.4)
Types of clinically relevant nonmajor bleeding, n (%)			
Gastrointestinal	1 (0.4)	1 (0.4)	0
Epistaxis	2 (0.9)	1 (0.4)	2 (0.9)
Urogenital	3 (1.3)	2 (0.9)	1 (0.4)
Skin	4 (1.7)	2 (0.9)	2 (0.9)
Vascular access site	2 (0.9)	4 (1.8)	6 (2.6)
Conjunctival	1 (0.4)	0	0
Major adverse vascular events ^a			
n (%)	3 (1.3)	6 (2.7)	7 (3.0)
Events (90% CI) per 100 patient-years	1.7 (0.5–3.7)	3.6 (1.6–6.3)	4.1 (1.9–6.9)
Type of event, n (%)			
Myocardial infarction	2 (0.9)	4 (1.8)	4 (1.7)
Ischemic stroke	1 (0.4)	1 (0.4)	2 (0.9)
Major amputation	0	0	1 (0.4)
Systemic embolism	0	1 (0.4)	0
Atherosclerotic subgroup ^a			
n/n (%)	2/81 (2.5)	2/86 (2.3)	6/82 (7.3)
Events (90% CI) per 100 patient-years	3.7 (0.7–8.7)	3.3 (0.6–7.8)	11.0 (4.8–19.2)
Dialysis circuit clotting, n (%)			
Score of 2 or 3	68 (29.3)	61 (27.2)	95 (41.3)
Score of 3	4 (1.7)	4 (1.8)	10 (4.3)
Access thrombosis, n (%)	6 (2.6)	8 (3.6)	11 (4.8)
Serious adverse events, n (%)			
Serious adverse event	70 (30.2)	64 (28.6)	63 (27.4)
Serious adverse event leading to discontinuation of study drug	8 (3.4)	10 (4.5)	13 (5.7)
All-cause death			
n (%)	12 (5.2%)	9 (4.0%)	11 (4.8%)
Events (90% CI) per 100 patient-years	7.0 (4.0–10.6)	5.4 (2.8–8.6)	6.4 (3.6–9.8)
Injection-site reactions, n (%)	14 (6.0)	17 (7.6)	1 (0.4)
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Incidences are reported by the number of participants having the specific event up to 30 days after the end of study treatment for that individual. ^aVascular death (due to myocardial infarction, stroke, pulmonary or systemic embolism), nonfatal myocardial infarction or stroke, major amputation for vascular etiology, acute limb ischemia and symptomatic venous thromboembolism or incidence of thrombosis of arteriovenous fistulas or grafts.





in the boxes; boxes indicate 25th and 75th percentiles; the vertical lines extend to a maximum distance of 1.5 interquartile ranges; values outside of this range are plotted separately. Analyses are based on the pharmacodynamic analysis set, n = 686 (lower-dose osocimab, n = 232; higher-dose osocimab, n = 224; placebo, n = 230).

the risk of clinically relevant bleeding or the risk of moderate, severe or serious adverse events. No bleeding events were reported in the small number of osocimab-treated individuals who underwent major surgery or intervention, including the 13 participants who underwent kidney transplantation. Therefore, osocimab seems to be associated with a low rate of clinically relevant bleeding in individuals requiring hemodialysis, which is similar to what was previously reported with osocimab in subjects undergoing elective knee arthroplasty¹⁴. These findings were observed despite the fact that osocimab prolongs the activated partial thromboplastin time in a dose-dependent manner.

Low event rates precluded exploratory analysis of efficacy regarding the incidence of major adverse vascular events. However, the exploratory analysis of dialysis circuit clotting, which was assessed in all participants, revealed that both doses of osocimab were associated with a significant reduction in the risk of moderate-to-complete dialysis circuit clotting compared with placebo, providing proof of concept that osocimab has antithrombotic effects beyond those of heparin¹⁵. Extracorporeal circuits activate factor XII and trigger clotting via the intrinsic pathway¹⁶. By inhibiting factor XIa, which is activated by factor XIIa, osocimab seems to be able to attenuate this process to a greater extent than heparin.

At present, there are no safe and effective anticoagulants for the prevention of thromboembolic events in individuals with kidney failure requiring hemodialysis. Prospective and observational studies have reported little or no benefit and increased bleeding in those receiving oral anticoagulants^{17,18}. Two randomized trials comparing vitamin K antagonists with apixaban in individuals with kidney failure undergoing hemodialysis who had atrial fibrillation were able to recruit only 154 and 97 participants, respectively, and were therefore underpowered to yield definitive results^{7,19}. The current findings and those of a trial comparing fesomersen—an antisense oligonucleotide that reduces the hepatic synthesis of factor XI—with placebo in individuals with end-stage kidney disease requiring hemodialysis²⁰, raise the possibility that factor XI inhibitors may be safer than currently available anticoagulants.

Like the above-mentioned apixaban trials⁷¹⁹, the composite of major and clinically relevant nonmajor bleeding was a primary outcome of the current study. Osocimab is an experimental agent and, as such, adverse events were also included as a primary safety outcome. Stroke and systemic embolism were key secondary outcomes in the apixaban trials because the patients had atrial fibrillation, whereas most of the patients enrolled in the present study did not have atrial fibrillation. Given that the incidences of stroke and systemic embolism are lower in those without atrial fibrillation, a broader range of thromboembolic events was included as exploratory outcomes. Furthermore, because osocimab was compared with a placebo in this trial, the only patients with atrial fibrillation who were eligible for enrollment were those who had been deemed unsuitable for anticoagulation therapy.

Some methodological aspects of our trial require comment. The strengths of the study include the double-blind trial design, the extended duration of treatment and the consistent findings with lower and higher doses of osocimab, which render it unlikely that the reported safety findings reflect a play of chance. Nonetheless, because of the modest sample size, additional studies are needed to assess the safety of osocimab. Finally, although end-stage kidney disease is more prevalent in older than in younger individuals, the prevalence of end-stage kidney disease in those aged 45 to 64 years has increased by 56% over the past 20 years¹⁵. This may explain why the mean age of participants in the present study was 61 years. Additional studies are needed to assess the safety of osocimab more thoroughly in older subjects.

In summary, compared with placebo, osocimab did not increase the risk of clinically relevant bleeding in individuals with kidney failure requiring hemodialysis. Appropriately powered phase 3 trials are needed to determine whether osocimab reduces the risk of thromboembolic events in this vulnerable and understudied population.

Online content

Any methods, additional references, Nature Portfolio reporting summaries, source data, extended data, supplementary information, acknowledgements, peer review information; details of author contributions and competing interests; and statements of data and code availability are available at https://doi.org/10.1038/s41591-023-02794-7.

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CONVERT Investigators

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A full list of members and their affiliations appears in the Supplementary Information.

Methods

Additional information about authors, study sites and investigators, and outcome definitions is provided in the Supplementary Appendix.

Study design and oversight

This phase 2b, randomized, double-blind, parallel-group trial compared two subcutaneous osocimab dosing regimens with a placebo in individuals with kidney failure requiring hemodialysis (Fig. 1). This multicenter study enrolled participants globally, including in North America, Europe, Asia and Australia.

An academic Steering Committee, in collaboration with the sponsor (Bayer), was responsible for the design and oversight of the study. The sponsor was responsible for data collection, maintenance and analysis. An Institutional Review Board at each participating center approved the protocol and all participants provided informed consent. This study was conducted in accordance with the consensus ethical principles derived from international guidelines including the Declaration of Helsinki and Council for International Organizations of Medical Sciences (CIOMS) International Ethical Guidelines, applicable ICH Good Clinical Practice Guidelines and all applicable laws and regulations.

The Steering Committee was blinded to treatment assignment as were the members of a Central Independent Adjudication Committee, who adjudicated all deaths, suspected bleeds, and cardiovascular or thromboembolic events. An independent Data and Safety Monitoring Committee periodically reviewed trial outcomes and adverse events with the support of an independent Statistical Analysis Center.

The protocol and accompanying documents are available with the full text of this article online.

Participants

Individuals with kidney failure were eligible for inclusion if they were 18 years of age or older and were undergoing hemodialysis (at least three times per week for a minimum of 9 h per week) and stable for at least 3 months. Patients with atrial fibrillation who were not considered to be candidates for therapeutic anticoagulation by their treating physicians were eligible for inclusion.

Participants were eligible for inclusion if all the following criteria were met:

- (1) ≥ 18 years of age
- (2) End-stage kidney disease undergoing hemodialysis (including hemodiafiltration) for ≥3 months and stable, in the view of the investigator
- (3) Body weight \geq 50 kg
- (4) Men and women were eligible; contraceptive use should be consistent with local regulations regarding the methods of contraception for those participating in clinical studies
- (5) Capable of providing signed informed consent

Participants were not eligible for inclusion if any of the following criteria applied:

- (1) Recent (<6 months before screening) clinically significant bleeding
- (2) Hemoglobin $< 9.0 \text{ g dl}^{-1}$
- (3) Platelet count $<100 \times 10^9 l^{-1}$
- (4) Activated partial thromboplastin time or prothrombin time above the upper limit of normal
- (5) Hepatic disease associated with alanine aminotransferase over three times the upper limit of normal, or total bilirubin over two times the upper limit of normal with direct bilirubin over 20% of the total
- (6) Sustained uncontrolled hypertension (diastolic blood pressure ≥100 mmHg and/or systolic blood pressure ≥180 mmHg)

- (7) Known intracranial neoplasm, arteriovenous malformation or aneurysm
- (8) Known bleeding disorders
- (9) Recent (<3 months before screening) thromboembolic event
- (10) Recent (<3 months before screening) major surgery or scheduled major surgery during study participation
- (11) Scheduled living donor renal transplant during study participation
- (12) Persistent heart failure, as classified by the New York Heart Association classification of III or higher
- (13) Receiving antiplatelet therapy, except acetylsalicylic acid ≤150 mg per day
- (14) Receiving anticoagulation in therapeutic doses, other than standard anticoagulation during the hemodialysis procedure
- (15) Life expectancy <6 months
- (16) Active malignancy requiring treatment during study participation (except nonmelanoma skin cancer or cervical carcinoma in situ)
- (17) Known hypersensitivity to the investigational drug or to inactive constituents of the study drug
- (18) Participation in another clinical study with an investigational medicinal product within 30 days or within five half-lives of such, whichever is longer, before randomization and during the study
- (19) Any other conditions, which, in the opinion of the investigator or sponsor, would render the individual unsuitable for inclusion

Randomization and study treatment

Participants were randomized between lower-dose osocimab and lower-dose placebo in a ratio of 2:1, and higher-dose osocimab and higher-dose placebo in a ratio of 2:1. Different placebos were needed because a larger volume was administered in the higher-dose group than in the lower-dose group. Participants were centrally assigned using an interactive web-response system and covariate-adaptive randomization. Covariates included geographical region, age, previous major adverse cardiovascular event, dialysis access via catheter, low-dose aspirin use, diabetes and atrial fibrillation. After randomization, there was a 1-week pretreatment period where participants underwent three hemodialysis sessions to establish a baseline for adverse events without treatment. Participants then received osocimab in a lower- or higher-dose regimen or placebo for 6 months in the main treatment period followed by an extension period of up to 12 months. or until the last study participant had completed their 6-month main treatment period. The lower-dose osocimab regimen consisted of a 105 mg loading dose followed by a monthly maintenance dose of 52.5 mg, and the higher-dose regimen consisted of a 210 mg loading dose followed by a monthly maintenance dose of 105 mg. A matching placebo was provided in identical-appearing vials. Treatments were administered subcutaneously in the abdomen no more than 1 h before dialysis. Heparin was administered as per usual during the dialysis sessions. The maximum duration of treatment was 18 months, and the trial was stopped when the last randomized participant completed 6 months of treatment.

Study outcomes

Primary outcomes. The primary outcomes were (1) clinically relevant bleeding, namely the composite of major and clinically relevant non-major bleeding, and (2) the composite of moderate or severe adverse events and serious adverse events. Bleeding was classified as major if it was overt and associated with a decrease in hemoglobin of 2 g dl⁻¹ or more; necessitated transfusion of two or more units of blood; occurred in a critical area or organ; or contributed to death. Overt bleeding not meeting these criteria, but that necessitated medical examination or intervention, or had clinical consequences, was classified as clinically relevant nonmajor bleeding. If neither set of criteria was met, bleeding was classified as minor.

Prespecified secondary outcomes were assessments of the change from baseline in osocimab concentration and key pharmacodynamic parameters, namely activated partial thromboplastin time and inhibition of factor XIa. Osocimab concentrations were measured by immunoassay, activated partial thromboplastin times were measured using C.K. Prest–a kaolin activator (Diagnostica Stago)–and factor XIa inhibition was quantified using a proprietary fluorogenic assay. Assays were conducted after in vitro neutralization of heparin to eliminate the potential effects of heparin in these assays.

Exploratory outcomes. Prespecified exploratory outcomes included the incidence of major adverse vascular events, the composite of vascular death due to myocardial infarction, stroke or pulmonary or systemic embolism; nonfatal myocardial infarction or stroke; major amputation of vascular etiology; acute limb ischemia and symptomatic venous thromboembolism.

Additionally, the incidence of arteriovenous fistula or graft thrombosis, and clotting of the dialysis circuit was assessed at every study visit as a prespecified exploratory outcome. Dialysis circuit clotting in the filter and air trap was assessed at the end of the hemodialysis procedure by study personnel blinded to treatment allocation. Clotting scores were assigned using a visual scoring system (0, no clot; 1, trace of clot; 2, intermediate between 1 and 3; and 3, fully clotted system necessitating interruption of hemodialysis session).

Statistical analysis

The major objectives of this study were to document the incidences of major and clinically relevant nonmajor bleeding for assessment of safety and arterial and venous thrombotic events for exploration of efficacy. There were no formal hypotheses. All analyses were descriptive and there was no formal a priori sample size calculation. Based on historic operational experience, of approximately 600 participants assigned randomly to study intervention, 555 participants would be expected to complete the main treatment period (assuming a true incidence rate of 15 participants lost to follow up per 100 participant years across all groups). Data were analyzed using SAS base v.9.4 (SAS/STAT v.14.3).

The efficacy and safety analyses were performed in the safety analysis set, defined as all randomized participants who received at least one dose of study medication. Safety and efficacy outcomes included all events that occurred during study treatment and up to 30 days after the last study drug administration. Outcomes were described by incidence proportions and cause-specific incidence rates per 100 person-years and their 90% CIs. Cause-specific hazard ratios were estimated with Cox proportional hazards model, subdistribution hazard ratios were estimated by Fine-Gray subdistribution hazards model. Because of the exploratory nature of these analyses, no multiplicity adjustment was performed. The number of participants who had a clotting score of 2 or 3 (moderate-to-complete dialysis circuit clotting) at one or more visits during the overall treatment period was compared between treatment groups by calculating relative risks and 95% CIs, where CIs were determined with the SAS procedure PROC FREQ by inverting two separate one-sided exact tests that were based on the score statistic (post hoc analysis; Farrington-Manning score statistic)^{21,22}.

Sensitivity analyses

The cumulative incidence functions for the event-of-interest as well as the associated competing events together with the corresponding CIs were estimated for each treatment arm using Aalen–Johansen estimators. The competing events for the primary endpoints were death and premature discontinuation of exposure to assigned treatment. The cumulative incidences were estimated for time-to-event endpoints by Aalen–Johansen estimators with the competing event. The differences of the Aalen–Johansen estimators between high-dose osocimab and placebo and low-dose osocimab and placebo are presented with 90% CIs.

Reporting summary

Further information on research design is available in the Nature Portfolio Reporting Summary linked to this article.

Data availability

Availability of the data underlying this publication will be determined according to Bayer's commitment to the EFPIA/PhRMA 'Principles for responsible clinical trial data sharing.' This pertains to the scope, timepoint and process of data access. As such, Bayer commits to sharing clinical trial data at the patient and study level upon request from qualified research personnel, as well as protocols from clinical trials for medicines and indications approved in the United States and European Union (necessary for conducting legitimate research). This applies to data on new medicines and indications that have been approved by the European Union and United States regulatory agencies on or after 1 January 2014. Interested researchers can use www.vivli.org to request access to anonymized patient-level data and supporting documents from clinical studies to conduct further research that can help advance medical science or improve patient care. Information about the Bayer criteria for listing studies and other relevant information is provided in the member section of the portal. Data access to anonymized patient-level data, protocols and clinical study reports will be granted after approval by an independent scientific review panel. Data will be made available within 6 months after signing the Data Use Agreement to researchers who provide a methodologically sound proposal. Bayer is not involved in the decisions made by the independent review panel. Bayer will take all necessary measures to ensure that patient privacy is safeguarded.

References

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Author contributions

J.I.W. and L.B.T. drafted the initial manuscript. J.H. and L.B.T. made substantial contributions to the acquisition of data and Á.F.P. made substantial contributions to the analysis of data. All authors had full access to the study data and contributed to the critical review and revision of the manuscript and had final responsibility for the decision to submit it for publication.

Competing interests

L.B.T., J.H., Á.F.P. and D.K. are employees of Bayer; J.I.W. holds the Canada Research Chair (TierI) in Thrombosis and the Heart and Stroke Foundation J.F. Mustard Chair in Cardiovascular Research at McMaster University, and has served as a consultant and received honoraria from Alnylam, Alexion, Anthos, Bayer, Boehringer Ingelheim, Bristol Myers Squibb, Daiichi Sankyo, Janssen, Merck, PhaseBio, Regeneron, Servier and VarmX; J.F. reports consultancy or speaker honoraria from Alnylam, AstraZeneca, Bayer, Boehringer Ingelheim, Calliditas, Chinook, Novartis, Omeros, Travere, VeraTx

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and Vifor; he also serves on data safety monitoring boards of trials by Novo Nordisk and Visterra. J.F. is an associate editor of Kidney International; K.A.A.F. reports grant funding from Bayer and AstraZeneca and consulting for Bayer, Medscape, Baim Institute for Clinical Research and the Thrombosis Research Institute; D.L.B. reports the following relationships: Advisory Boards of Angiowave, Bayer, Boehringer Ingelheim, Cardax, CellProthera, Cereno Scientific, Elsevier Practice Update Cardiology, High Enroll, Janssen, Level Ex, McKinsey, Medscape Cardiology, Merck, MyoKardia (now a subsidiary of Bristol Myers Squibb), NirvaMed, Novo Nordisk, PhaseBio, PLx Pharma, Regado Biosciences and Stasys; Board of Directors of Angiowave (stock options), Boston VA Research Institute, Bristol Myers Squibb (stock), DRS.LINQ (stock options), High Enroll (stock), Society of Cardiovascular Patient Care and TobeSoft: Chair: Inaugural Chair, American Heart Association Quality Oversight Committee; Consultant: Broadview Ventures; Data Monitoring Committees: Acesion Pharma, Assistance Publique-Hôpitaux de Paris, Baim Institute for Clinical Research (formerly Harvard Clinical Research Institute, for the PORTICO trial, funded by St. Jude Medical, now Abbott), Boston Scientific (Chair, PEITHO trial), Cleveland Clinic (including for the ExCEED trial, funded by Edwards), Contego Medical (Chair, PERFORMANCE 2), Duke Clinical Research Institute, Mayo Clinic, Mount Sinai School of Medicine (for the ENVISAGE trial, funded by Daiichi Sankyo; for the ABILITY-DM trial, funded by Concept Medical), Novartis, Population Health Research Institute; Rutgers University (for the National Institutes of Health (NIH)-funded MINT Trial); Honoraria: American College of Cardiology (Senior Associate Editor, Clinical Trials and News, ACC.org; Chair, ACC Accreditation Oversight Committee), Arnold and Porter law firm (work related to Sanofi/Bristol Myers Squibb clopidogrel litigation), Baim Institute for Clinical Research (formerly Harvard Clinical Research Institute; RE-DUAL PCI clinical trial steering committee funded by Boehringer Ingelheim; AEGIS-II executive committee funded by CSL Behring), Belvoir Publications (Editor in Chief, Harvard Heart Letter), Canadian Medical and Surgical Knowledge Translation Research Group (clinical trial steering committees), Cowen and Company, Duke Clinical Research Institute (clinical trial steering committees, including for the PRONOUNCE trial, funded by Ferring Pharmaceuticals), HMP Global (Editor in Chief, Journal of Invasive Cardiology), Journal of the American College of Cardiology (Guest Editor; Associate Editor), K2P (co-chair, interdisciplinary curriculum), Level Ex, Medtelligence/ReachMD (CME steering committees), MJH Life Sciences, Oakstone CME (Course Director, Comprehensive Review of Interventional Cardiology), Piper Sandler, Population Health Research Institute (for the COMPASS operations committee, publications committee, steering committee, and US national co-leader, funded by Bayer), Slack Publications (Chief Medical Editor, Cardiology Today's Intervention), Society of Cardiovascular Patient Care (Secretary/Treasurer), WebMD (CME steering committees), Wiley (steering committee); Other: Clinical Cardiology (Deputy Editor), NCDR-ACTION Registry Steering

Committee (Chair), VA CART Research and Publications Committee (Chair): Patent: Sotagliflozin (named on a patent for sotagliflozin assigned to Brigham and Women's Hospital who assigned to Lexicon; neither I nor Brigham and Women's Hospital receive any income from this patent); Research Funding: Abbott, Acesion Pharma, Afimmune, Aker Biomarine, Amarin, Amgen, AstraZeneca, Bayer, Beren, Boehringer Ingelheim, Boston Scientific, Bristol Myers Squibb, Cardax, CellProthera, Cereno Scientific, Chiesi, CinCor, CSL Behring, Eisai, Ethicon, Faraday Pharmaceuticals, Ferring Pharmaceuticals, Forest Laboratories, Fractyl, Garmin, HLS Therapeutics, Idorsia, Ironwood, Ischemix, Janssen, Javelin, Lexicon, Lilly, Medtronic, Merck, Moderna, MyoKardia (now a subsidiary of Bristol Myers Squibb), NirvaMed, Novartis, Novo Nordisk, Owkin, Pfizer, PhaseBio, PLx Pharma, Recardio, Regeneron, Reid Hoffman Foundation, Roche, Sanofi, Stasys, Synaptic, The Medicines Company, Youngene and 89Bio; Royalties: Elsevier (Editor, Braunwald's Heart Disease); Site Co-Investigator: Abbott, Biotronik, Boston Scientific, CSI, Endotronix, St. Jude Medical (now Abbott), Philips, SpectraWAVE, Svelte, Vascular Solutions; Trustee: American College of Cardiology; Unfunded Research: FlowCo, Takeda; R.T. has received royalties from Thermo Fisher Scientific, and consulting fees from the US Food and Drug Administration and Fresenius Medical Care; he has served in a data safety committee for Alnylam, and scientific advisory committees for Aggamin, Bayer, Comanche Biopharma and Hero; and he is on the board of directors for CAMP4, and is an equity holder with Aggamin, CAMP4, Comanche Biopharma and Hero; W.C.W. holds the Gordon A. Cain Chair in Nephrology at Baylor College of Medicine, and has served as a consultant and has received honoraria from Akebia/Otsuka, Ardelyx, AstraZeneca, Bayer, Boehringer Ingelheim/Lilly, GlaxoSmithKline, Merck Sharp & Dohme/Merck, Pharmacosmos, Reata, Unicycive and Zydus.

Additional information

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Extended Data Fig. 1 | Cumulative incidence risk (Aalen–Johansen) of the time-to-first treatment-emergent composite of major and clinically relevant nonmajor bleeding (primary outcome; safety analysis set). Cumulative incidence (%) = Aalen–Johansen estimates of the cumulative probability for an event, calculated as 100 x (1 minus the Aalen–Johansen estimate of the survival

function). Overall study period: Only treatment-emergent events that occurred during the overall study period are considered. Treatment-emergent events are those that occurred after the start of the study treatment and until the last dose of the study treatment plus 30 days.

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Extended Data Fig. 2 | Cumulative incidence risk (Aalen–Johansen) of the time-to-first treatment-emergent moderate, severe, or serious adverse event (primary outcome; safety analysis set). Cumulative incidence (%) = Aalen– Johansen estimates of the cumulative probability for an event, calculated as 100 x (1 minus the Aalen–Johansen estimate of the survival function). Overall study period: Only treatment-emergent events that occurred during the overall study period are considered. Treatment-emergent events are those that occurred after the start of the study treatment and until the last dose of the study treatment plus 30 days.

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Extended Data Fig. 3 | Cumulative incidence risk (Aalen–Johansen) of timeto-first arteriovenous fistula or graft thrombosis (exploratory outcome; safety analysis set). Cumulative incidence (%) = Aalen–Johansen estimates of the cumulative probability for an event, calculated as 100 x (1 minus the Aalen–Johansen estimate of the survival function). Overall study period: Only treatment-emergent events that occurred during the overall study period are considered. Treatment-emergent events are those that occurred after the start of the study treatment and until the last dose of the study treatment plus 30 days.

Extended Data Table 1 | Cause-specific hazard ratios and subdistribution hazard ratios for treatment-emergent primary outcome (composite of major and clinically relevant nonmajor bleeding) and competing events (safety analysis set)

	Cox proportional hazard regression model			nodel	Fine-Gray model		
	Cause-specific hazard ratio (95% CI)	Relative hazard reduction*	90% CI	P value (logrank test)	Subdistribution hazard ratio (95% CI)	90% CI	P value (Gray's test)
Lower-dose osocimab vs. placebo							
Composite of major and clinically	0.90 (0.46–1.77)	9.9%	(0.51 - 1.58)	0.760	0.88 (0.45-1.72)	(0.50–1.54)	0.708
relevant nonmajor bleeding events							
Competing event 1: Death due to	1.12 (0.48–2.64)		(0.55 - 2.30)	0.793	1.12 (0.48-2.62)	(0.55 - 2.28)	0.796
other reasons							
Competing event 2: Premature	0.94 (0.60–1.48)		(0.64–1.37)	0.773	0.93 (0.59–1.46)	(0.64–1.36)	0.763
discontinuation of study medication							
Higher-dose osocimab vs. placebo							
Composite of major and clinically	0.62 (0.29–1.31)	38.1%	(0.33–1.16)	0.206	0.61 (0.29–1.29)	(0.33–1.15)	0.196
relevant nonmajor bleeding events							
Competing event 1: Death due to	0.90 (0.37-2.22)		(0.42 - 1.92)	0.825	0.92 (0.38-2.27	(0.44–1.96)	0.861
other reasons							
Competing event 2: Premature	1.01 (0.65–1.58)		(0.70–1.47)	0.961	1.04 (0.67–1.61)	(0.72 - 1.50)	0.877
discontinuation of study medication							

*Relative hazard reduction = 100 x (1 - hazard ratio)%

The Cox model and the Fine-Gray model were fitted with the treatment group as a factor. P values are two-sided with the treatment group as a fixed factor.

Overall study period: only treatment-emergent events that occurred during the overall study period are considered. Treatment-emergent events are those that occurred after the start of the study treatment and up until the last dose of study treatment plus 30 days.

Extended Data Table 2 | Cause-specific hazard ratios and subdistribution hazard ratios for primary outcome (composite of major bleeding and clinically relevant nonmajor bleeding) and competing events (safety analysis set)

	Cox pro	Cox proportional hazard regression model				Fine-Gray model		
	Cause-specific hazard ratio (95% CI)	Relative hazard reduction*	90% CI	P value (logrank test)	Subdistribution hazard ratio (95% CI)	90% CI	P value (Gray's test)	
Lower-dose osocimab vs. placebo								
Composite of major and clinically	1.00 (0.57-1.77)	-0.5%	(0.62 - 1.62)	0.987	0.98 (0.56-1.72)	(0.61-1.57)	0.940	
relevant nonmajor bleeding events								
Competing event 1: Death due to	1.39 (0.70–2.77)		(0.78 - 2.48)	0.350	1.39 (0.70–2.76)	(0.78 - 2.48)	0.346	
other reasons								
Higher-dose osocimab vs. placebo								
Composite of major and clinically	0.84 (0.47-1.53)	15.7%	(0.51 - 1.39)	0.571	0.84 (0.46-1.51)	(0.51 - 1.38)	0.559	
relevant nonmajor bleeding events								
Competing event 1: Death due to	1.09 (0.53-2.26)		(0.59–2.01)	0.819	1.11 (0.54–2.29)	(0.60-2.04)	0.779	
other reasons								

*Relative hazard reduction = 100 x (1 - hazard ratio)%

The Cox model and the Fine-Gray model were fitted with the treatment group as a factor. P values are two-sided with the treatment group as a fixed factor.

Events that occurred during the full study period consisting of the main treatment, extension treatment and post-treatment follow-up periods are considered.

Extended Data Table 3 | Cause-specific hazard ratios and subdistribution hazard ratios for the treatment-emergent composite of moderate, severe, or serious adverse events (primary outcome) and competing events (safety analysis set)

	Cox pro	Cox proportional hazard regression model			Fine-Gray model		
	Cause-specific hazard ratio (95% CI)	Relative hazard reduction*	90% CI	P value (logrank test)	Subdistribution hazard ratio (95% CI)	90% CI	P value (Gray's test)
Lower-dose osocimab vs. placebo							
Composite of moderate, severe, or	1.25 (0.96–1.64)	-25.2%	(1.00–1.57)	0.098	1.26 (0.97–1.64)	(1.01–1.57)	0.089
serious adverse events							
Competing event 1: Death due to	NA						
other reasons							
Competing event 2: Premature	0.78 (0.39–1.57)		(0.44 - 1.40)	0.488	0.73 (0.37-1.45)	(0.41 - 1.30)	0.371
discontinuation of study medication							
Higher-dose osocimab vs. placebo							
Composite of moderate, severe, or	1.14 (0.87–1.50)	-13.9%	(0.91–1.43)	0.350	1.14 (0.87–1.50)	(0.91–1.44)	0.338
serious adverse events							
Competing event 1: Death due to	NA						
other reasons							
Competing event 2: Premature	0.85 (0.43-1.67)		(0.48 - 1.50)	0.632	0.81 (0.41-1.58)	(0.46 - 1.42)	0.525
discontinuation of study medication							

*Relative hazard reduction = 100 x (1 - hazard ratio)%

The Cox model and the Fine-Gray model were fitted with the treatment group as a factor. *P* values are two-sided with the treatment group as a fixed factor. Overall study period: only treatment-emergent events that occurred during the overall study period are considered. Treatment-emergent events are those that occurred after the start of the study treatment and until the last dose of study treatment plus 30 days. NA, not available

Extended Data Table 4 | Cause-specific hazard ratios and subdistribution hazard ratios for treatment-emergent major bleeding events, clinically relevant nonmajor bleeding events, and competing events (safety analysis set)

		Cox PH regress	sion model		Fine-Gray model		
	Cause-specific hazard ratio (95% CI)	Relative hazard reduction*	90% CI	P value (logrank test)	Subdistribution hazard ratio (95% CI)	90% CI	P value (Gray's test)
Lower-dose osocimab vs. placebo							
Major bleeding events	0.29 (0.06–1.40)	70.9%	(0.08–1.09)	0.101	0.28 (0.06–1.35)	(0.08–1.05)	0.091
Competing event 1: Death due to other reasons	2.30 (0.71–7.47)		(0.86-6.18)	0.154	2.27 (0.70–7.33)	(0.85–6.07)	0.161
Competing event 2: Premature discontinuation of study medication	1.41 (0.77–2.58)		(0.85–2.34)	0.266	1.40 (0.77–2.55)	(0.84–2.32)	0.274
Clinically relevant nonmajor bleeding events	1.16 (0.42–3.19)	-15.6%	(0.49–2.71)	0.779	1.13 (0.41–3.12)	(0.48–2.65)	0.809
Competing event 1: Death due to other reasons	1.84 (0.62–5.49)		(0.74-4.61)	0.267	1.82 (0.61–5.39)	(0.73-4.53)	0.278
Competing event 2: Premature discontinuation of study medication	1.22 (0.67–2.20)		(0.74–2.00)	0.522	1.20 (0.66–2.16)	(0.73–1.96)	0.550
Higher-dose osocimab vs. placebo							
Major bleeding events	0.15 (0.02–1.20)	85.2%	(0.03–0.86)	0.038	0.14 (0.02–1.17)	(0.03–0.83)	0.035
Competing event 1: Death due to other reasons	1.29 (0.35–4.81)		(0.43-3.90)	0.701	1.30 (0.35–4.82)	(0.43–3.90)	0.696
Competing event 2: Premature discontinuation of study medication	1.38 (0.75–2.54)		(0.83–2.31)	0.300	1.39 (0.75–2.54)	(0.83–2.31)	0.291
Clinically relevant nonmajor bleeding	1.03 (0.36–2.95)	-3.4%	(0.43–2.49)	0.950	1.02 (0.36–2.91)	(0.42–2.46)	0.968
Competing event 1: Death due to other reasons	1.04 (0.30–3.59)		(0.37–2.94)	0.953	1.04 (0.30–3.58)	(0.37–2.94)	0.952
Competing event 2: Premature discontinuation of study medication	1.24 (0.69–2.25)		(0.76–2.05)	0.470	1.24 (0.69–2.23)	(0.75–2.03)	0.478

*Relative hazard reduction = 100 x (1 – hazard ratio)%

The Cox model and the Fine-Gray model were fitted with the treatment group as a factor. P values are two-sided with the treatment group as a fixed factor.

Main treatment period: only treatment-emergent events that occurred during the main treatment period are considered. The main treatment period starts with the loading dose at visit 5 and ends with the end-of-main-treatment-period visit after 6 months. Treatment-emergent events are those that occurred after the start of the study treatment and until the last dose of study treatment plus 30 days.

Extended Data Table 5 | Summary of Adverse Events

	Lower-dose	Higher-dose	Placebo
	Osocimab	Osocimab	(N=230)
	(N=232)	(N=224)	
Adverse events — n (%)			
Mild	66 (28.4)	81 (36.2)	74 (32.2)
Moderate	77 (33.2)	71 (31.7)	60 (26.1)
Severe	36 (15.5)	28 (12.5)	35 (15.2)
Adverse events related to treatment — n (%)			
Mild	24 (10.3)	24 (10.7)	16 (7.0)
Moderate	4 (1.7)	9 (4.0)	7 (3.0)
Severe	1 (0.4)	2 (0.9)	0
Serious adverse events related to treatment — n (%)	1 (0.4)	2 (0.9)	2 (0.9)

	Lower-dose	Higher-dose	Placebo or no study
	Osocimab	Osocimab	treatment received
Patients — n	13	13	21
Procedures — n	14 (5 major)	16 (6 major)	25
Factor XIa inhibition at surgery — median (range) $\%$	58 (42–73)	69 (54–92)	n.a.
Factor XIa inhibition at major surgery — median (range) $\%$	57 (47–71)	77 (65–92)	n.a.
Clinically relevant nonmajor bleeding [*] — n	0	0	1

Extended Data Table 6 | Surgical Interventions Within One Month After Administration of Study Drug, Extent of Factor XI Inhibiton, and Associated Bleeding Outcomes*

n.a. denotes not available.

*Clinically relevant nonmajor bleeding occurring during or within 2 weeks from surgery/intervention.

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n/a	Cor	firmed
		The exact sample size (n) for each experimental group/condition, given as a discrete number and unit of measurement
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		The statistical test(s) used AND whether they are one- or two-sided Only common tests should be described solely by name; describe more complex techniques in the Methods section.
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		For null hypothesis testing, the test statistic (e.g. <i>F</i> , <i>t</i> , <i>r</i>) with confidence intervals, effect sizes, degrees of freedom and <i>P</i> value noted Give <i>P</i> values as exact values whenever suitable.
\checkmark		For Bayesian analysis, information on the choice of priors and Markov chain Monte Carlo settings
\checkmark		For hierarchical and complex designs, identification of the appropriate level for tests and full reporting of outcomes
\checkmark		Estimates of effect sizes (e.g. Cohen's d, Pearson's r), indicating how they were calculated
		Our web collection on statistics for biologists contains articles on many of the points above.

Software and code

Policy information about availability of computer code

Data collection Data were collected using the electronic Case Report Forms RAVE by Medidata.

Data analysis	Data were analysed using SAS base 9.4 - SAS/STAT 14.3. Only standard methods were used for the analysis as described in the
	paper.

For manuscripts utilizing custom algorithms or software that are central to the research but not yet described in published literature, software must be made available to editors and reviewers. We strongly encourage code deposition in a community repository (e.g. GitHub). See the Nature Portfolio guidelines for submitting code & software for further information.

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Availability of the data underlying this publication will be determined according to Bayer's commitment to the EFPIA/PhRMA "Principles for responsible clinical trial data sharing". This pertains to the scope, timepoint and process of data access. As such, Bayer commits to sharing clinical trial data upon request from qualified research personnel at the patient and study level, as well as protocols from clinical trials for medicines and indications approved in the United States (US) and European Union (EU) necessary for conducting legitimate research. This applies to data on new medicines and indications that have been approved by the EU and US regulatory agencies on or after January 01, 2014.

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Research involving human participants, their data, or biological material

Policy information about studies with <u>human participants or human data</u>. See also policy information about <u>sex, gender (identity/presentation)</u>, <u>and sexual orientation</u> and <u>race, ethnicity and racism</u>.

Reporting on sex and gender	Data from females and males were aggregated and not reported separately
	Baseline demographic data present ethnicity (% white) and geographical region.
Reporting on race, ethnicity, or other socially relevant	Lower-dose osocimab: n = 232; median (range) age, 61.0 (28–91) years; sex (male, %), 61.6 Higher-dose osocimab: n = 224; median (range) age, 61.0 (25–90) years; sex (male, %), 63.8 Placebo: n = 230; median (range) age, 60.0 years; sex (male, %), 65.2
Population characteristics	Participants were identified and recruited according to standard industry guidelines and regulations. Research staff at study sites reviewed their rosters of patie with kidney disease undergoing hemodialysis and approached those who met the inclusion criteria to determine their interest in participation in the trial. Interes patients without exclusion criteria provided signed informed consent prior to performing any study specific testing or procedures. After screening, patients with either randomized or identified as screen failures. Recruitment materials for the study included pocket reference cards and posters for study site personnel and
Recruitment	patient welcome booklet. Recruited patients reflected the age, sex, and racial makeup of the rosters of patients at the various dialysis centers, which may have limited the diversity of those included.
Ethics oversight	An Institutional Review Board at each participating center approved the protocol and all participants provided informed consent. This study was conducted in accordance with the consensus ethical principles derived from international guidelines including the Declaration of Helsinki and Council for International Organizations of Medical Sciences (CIOMS) International Ethical Guidelines, applicable ICH Good Clinical Practice Guidelines and all applicable laws and
Note that full information on the appro	$_{ m >V}$ regulations. A list of all IEC and IRBs is provided in supplementary material.

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Please select the one below that is the best fit for your research. If you are not sure, read the appropriate sections before making your selection.

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For a reference copy of the document with all sections, see <u>nature.com/documents/nr-reporting-summary-flat.pdf</u>

Life sciences study design

 \Box

All studies must disclose on these points even when the disclosure is negative.

Sample size	A total of 704 participants from 147 sites in 19 countries were randomized.
Data exclusions	18 participants who did not receive study drug were excluded.
Replication	All pharmacokinetic and pharmacodynamic assays were performed in duplicate at multiple time points.
Randomization	Participants were centrally assigned in a 1:1:1 ratio to lower- or higher-dose osocimab, or to placebo using an interactive web-response system and covariate-adaptive randomization.
Blinding	This was a double-blind trial. The Steering Committee was blinded to treatment assignment as were the members of a Central Independent Adjudication Committee, who adjudicated all deaths, suspected bleeding, and cardiovascular or thromboembolic events.

Reporting for specific materials, systems and methods

We require information from authors about some types of materials, experimental systems and methods used in many studies. Here, indicate whether each material, system or method listed is relevant to your study. If you are not sure if a list item applies to your research, read the appropriate section before selecting a response.

Materials & experimental systems		Me	Methods	
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\bigtriangledown	Eukaryotic cell lines	\checkmark	Flow cytometry	
\checkmark	Palaeontology and archaeology	\checkmark	MRI-based neuroimaging	
\checkmark	Animals and other organisms			
	🛛 Clinical dataDual use research of concern			
\checkmark	Plants			

Note that full information on the approval of the study protocol must also be provided in the manuscript.

Clinical data

Policy information about <u>clinical studies</u>

All manuscripts should comply with the ICMJE guidelines for publication of clinical research and a completed CONSORT checklist must be included with all submissions.

Clinical trial registration	NCT04523220
Study protocol	Study protocol and SAP provided
	The energy was reasonable for data collection, maintenance, and evaluate. Data was collected with all straight
Data collection	Case Report Forms (RAVE by Medidata).
Outcomes	The primary outcomes were (1) clinically relevant bleeding, namely the composite of major and clinically relevant nonmajor bleeding, and (2) the composite of moderate or severe adverse events and serious adverse events.
	Bleeding was classified as major if it was overt and associated with a decrease in hemoglobin of 2 g/dL or more;
	necessitated transfusion of two or more units of blood, occurred in a critical area or organ; or contributed to death
	Overt blooding not monthly those oritorial but that popossitated medical evamination or intervention, or bed divised
	Over bleeding not meeting these chiena, but that necessitated medical examination of intervention, of had clinical
	consequences, was classified as clinically relevant nonmajor bleeding. If heither set of criteria was met, bleeding was
	classified as minor.
	Prespecified Ssecondary outcomes were assessments of the change from baseline in osocimah concentration and
	key pharmaged visioning parameters, named accelerated partial thrombon lacting time and inhibition of factor. Vis
	Osocimad concentrations were measured by immunoassay, activated partial thromboplastin times were measured
	using C.K. Prest®, a kaolin activator (Diagnostica Stago, France), and factor XIa inhibition was guantified using a
	proprietary fluorogenic assay. Assays were conducted after in vitro neutralization of heparin to eliminate the potential
	offorte of bonarin in those accourt
	eneus of nepanin in dese assays.
	Prespectied Eexploratory outcomes included the incidence of major adverse vascular events, the composite of
	vascular death due to myocardial infarction, stroke, or pulmonary or systemic embolism; nonfatal myocardial
	infarction or stroke: major amputation of vascular etiology: acute limb ischemia, and symptomatic venous
	thromboembolism
	Additionally the incidence of arteriovaneus fietule or graft thrombasis, and eletting of the dislusis sirguit was
	Additionally, the incidence of antenovenous listula of grant thrombosis, and dotting of the dialysis circuit was
	assessed at every study visit as a prespecified exploratory outcome. Dialysis circuit clotting in the filter and air trap
	was assessed at the end of the hemodialysis procedure by study personnel blinded to treatment allocation. S.
	Clotting scores were assigned using a visual scoring system (0, no clot: 1, trace of clot: 2, intermediate between 1
	and 3: and 3 fully clotted system accessitating interview of homodalysis session)
	and 5, and 5, runy dotted system necessitating interruption of heriodialysis session).

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