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# No evidence for differential gene expression in major depressive disorder PBMCs, but robust evidence of elevated biological ageing

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## Abstract

The increasingly compelling data supporting the involvement of immunobiological mechanisms in Major Depressive Disorder (MDD) might provide some explanation for the variance in this heterogeneous condition. Peripheral blood measures of cytokines and chemokines constitute the bulk of evidence, with consistent meta-analytic data implicating raised proinflammatory cytokines such as IL6, IL1 $\beta$  and TNF. Among the potential mechanisms linking immunobiological changes to affective neurobiology is the accelerated biological ageing seen in MDD, particularly via the senescence associated secretory phenotype (SASP). However, the cellular source of immunobiological markers remains unclear. Pre-clinical evidence suggests a role for peripheral blood mononuclear cells (PBMC), thus here we aimed to explore the transcriptomic profile using RNA sequencing in PBMCs in a clinical sample of people with various levels of depression and treatment response comparing it with that in healthy controls (HCs). There were three groups with major depressive disorder (MDD): treatment-resistant ( $n = 94$ ), treatment-responsive ( $n = 47$ ) and untreated ( $n = 46$ ). Healthy controls numbered 44. Using PBMCs gene expression analysis was conducted using RNAseq to a depth of 54.5 million reads. Differential gene expression analysis was performed using DESeq2. The data showed no robust signal differentiating MDD and HCs. There was, however, significant evidence of elevated biological ageing in MDD vs HC. Biological ageing was evident in these data as a transcriptional signature of 888 age-associated genes (adjusted  $p < 0.05$ , absolute  $\log_2\text{fold} > 0.6$ ) that also correlated strongly with chronological age (spearman correlation coefficient of 0.72). Future work should expand clinical sample sizes and reduce clinical heterogeneity. Exploration of RNA-seq signatures in other leukocyte populations and single cell RNA sequencing may help uncover more subtle differences. However, currently the subtlety of any PBMC signature mitigates against its convincing use as a diagnostic or predictive biomarker.

## Introduction

Major Depressive Disorder (MDD) remains one of the most aetiologically opaque of human disorders, yet one that continues to exert a powerfully negative toll on human health – physical as well as mental. MDD is both heterogeneous in its phenotypic expression and complex in its genetic and physiological correlates. Among the latter there are increasingly compelling data supporting the involvement of immunobiology in MDD. However,

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We pay tribute to our Janssen colleague, the late Dr Peter de Boer. Peter was a man of great knowledge, integrity, kindness and humour. And we were lucky enough to have him at the heart of the NIMA consortium from the outset. He will be sorely missed.

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the mechanisms underpinning this relationship remain unclear. Peripheral blood measures constitute the bulk of evidence with consistent meta-analytic data implicating raised proinflammatory cytokines. The most comprehensive genome-wide association study (GWAS) to date on MDD used 7 major cohorts and identified 44 independent loci and 153 genes<sup>1</sup>. Forty-five of these were in the extended major histocompatibility complex (MHC), which is central to acquired immunity and to leukocyte interactions.

Whole-transcriptome studies offer another variant of genome-wide search for disease-related mechanisms by measuring mRNA expression levels of each gene in a relevant tissue. RNA sequencing (RNA-seq) uses next-generation sequencing to provide a quantitation of RNA or gene expression. Recent studies have used this method in MDD. One of the largest examined a total sample of 922 people (463 with MDD and 459 health controls) and sequenced RNA from whole blood<sup>2</sup>. A relatively small number of genes were found to be associated with MDD (29) at a very relaxed false discovery rate (FDR) threshold of 0.25. With the more customary and restrictive FDR threshold of 0.05, no significant genes were found. They also showed modest enrichment for the IFN  $\alpha/\beta$  pathway, which included three significant genes at FDR < 0.25.

A number of potential mechanisms have linked immunobiological changes to affective neurobiology. Among these is the accelerated biological ageing seen in MDD. Immune cell senescence has a well-documented effect on both epigenome and transcriptome<sup>3</sup>. MDD has also been linked to the senescence associated secretory phenotype (SASP), a dynamic secretory molecular pathway indicative of cellular senescence<sup>4</sup>. This speaks to a more elaborate biology linking cell biology, transcriptome and inflammatory proteins produced by the cell.

The cellular source of immunobiological markers in depression remains a key unanswered question. PBMCs are a key source of peripheral cytokines and pre-clinical models have suggested some PBMC subsets can enter the brain and contribute to onset of sickness behaviour in the context of stress. Monocytes recruited to the brain have been linked to behavioural changes associated with anxiety and with direct effects on neuronal dendritic spine remodelling linked to cognitive deficits<sup>5–7</sup>. Similarly, CD4+ T cells have been linked to stress-related behavioural changes via mitochondrial fission leading to xanthine upregulation and subsequent oligodendrocyte proliferation in the amygdala<sup>8</sup>.

Given the weight of the preclinical evidence suggesting a role for PBMCs, we aimed to explore the transcriptomic profile using RNA-seq in PBMCs in a clinical sample of people with various levels of depression and treatment response and compare with that in healthy controls.

We aimed to answer the following research questions.

- (1) Is there evidence of differential gene expression between healthy controls and MDD or between healthy controls and sub-types of MDD?
- (2) Is there evidence of elevated immune ageing MDD compared to healthy controls?

## Methods

### Participants

This was a non-interventional study, conducted as part of the Wellcome Trust Consortium for Neuroimmunology of Mood Disorders and Alzheimer's disease (NIMA). There were five clinical study centres in the UK: Brighton, Cambridge, Glasgow, King's College London, and Oxford. All procedures were approved by an independent Research Ethics Committee (National Research Ethics Service East of England, Cambridge Central, UK; approval number 15/EE/0092) and the study was conducted according to the Declaration of Helsinki. All participants provided informed consent in writing and received £100 compensation for taking part.

### Sample and eligibility criteria

We recruited four groups of participants: treatment-resistant depression, treatment-responsive depression, untreated depression, and healthy volunteers.

Eligibility criteria can be viewed in full in Supplementary Information (Supplementary Table 1).

Patients were assigned to one of three subgroups or strata, per protocol:

- (i) treatment-resistant (DEP+MED+) patients who had total Hamilton Depression Rating Scale (HAM-D) score >13 and had been medicated with a monoaminergic drug at a therapeutic dose for at least 6 weeks;
- (ii) treatment-responsive (DEP-MED+) patients who had total HAM-D <7 and had been medicated with a monoaminergic drug at a therapeutic dose for at least 6 weeks; and
- (iii) untreated (DEP+MED-) patients who had HAM-D >17 and had not been medicated with an antidepressant drug for at least six weeks.

### Questionnaire assessments

Psychological symptoms and childhood adversity were assessed by administration of questionnaires as previously described<sup>9</sup> (see Supplementary Information). Baseline depression severity was rated using the 17-item HAM-D.

### Sampling and isolation of PBMCs

Whole blood was collected in K2EDTA tubes (BD, USA) by peripheral venepuncture and allowed to cool to room temperature for a minimum of 45 min. PBMCs were collected from the interphase following density gradient

centrifugation. RNA was extracted using the RNeasyMini Kit (Qiagen, Germany) as per the manufacturer's instructions. RNA was eluted in 50  $\mu$ l RNase-free H<sub>2</sub>O and stored at  $-80^{\circ}\text{C}$  before being sent for sequencing.

### RNA-sequencing and processing

PBMC samples were taken from four separate population groups as follows: 44 healthy controls, 94 MDD treatment-resistant, 47 MDD treatment-responsive and 46 MDD untreated patients. All PBMC samples had an RNA Integrity Number (RIN)  $\geq 8$  and were analysed for gene expression levels by RNA-Seq to an average depth of 54.5 million read pairs. Reads were trimmed using Cutadapt 1 (version cutadapt-1.9.dev2)<sup>10</sup>. The reference used for mapping was the Homo sapiens genome from Ensembl, assembly GRCh38, annotation version 84. Reads were aligned to the reference genome using STAR 2 (version 2.5.2b)<sup>11</sup>. Reads were assigned to features of type 'exon' in the input annotation grouped by gene\_id in the reference genome using featureCounts 3 (version 1.5.1). Strandedness was set to 'reverse' and a minimum alignment quality of 10 was specified. After filtering for only protein coding genes, we observed a median of 40 million exonic aligned reads per sample (>85%).

### RNA-sequencing differential expression analysis

Differential expression analysis was performed using DESeq2 (version 1.18.1)<sup>12</sup>. The count matrix was initially filtered to include only coding genes, with a mean of >1 read per sample. For the comparisons of binary clinical covariates (e.g. gender, tobacco) one group was compared to the other. For continuous clinical covariates (e.g. age, BMI) the patients in the lower quartile were compared to those in the upper quartile. No additional covariates were used in the DESeq2 model when comparing clinical covariates. For the comparisons between HC group and the MDD groups the 15 clinical covariates (Fig. 1b) identified as having >5 significant associated genes (adjusted  $p < 0.01$ ) and "batch" were included as covariates in the model. To control for extreme outlier values typical in large and heterogeneous datasets, a Cooks cut-off of 0.2 was used. All other parameters were left to default. Significance was set at an adjusted  $p$  of <0.01. For full details see the Supplementary Information.

### Deconvolution analysis

The per sample distribution of cell types was estimated by Cibersort<sup>13</sup>, using the DESeq2 normalised expression values (no additional covariates) as the mixture file, and the LM22 (22 immune cell types) signature gene file. Quantile normalisation was disabled. All other parameters were left to default.

### RNA-sequencing randomised cases and controls

The 231 samples were randomised using the *r* function "sample" (without replacement), and were then split into two random groups, one with 44 samples and one with 187 samples (in line with the real group distribution and *n*). These two groups were then differentially compared using DESeq2 as described above. For full details see the Supplementary Information (refs. <sup>14,15</sup>).

### Co-expression analysis

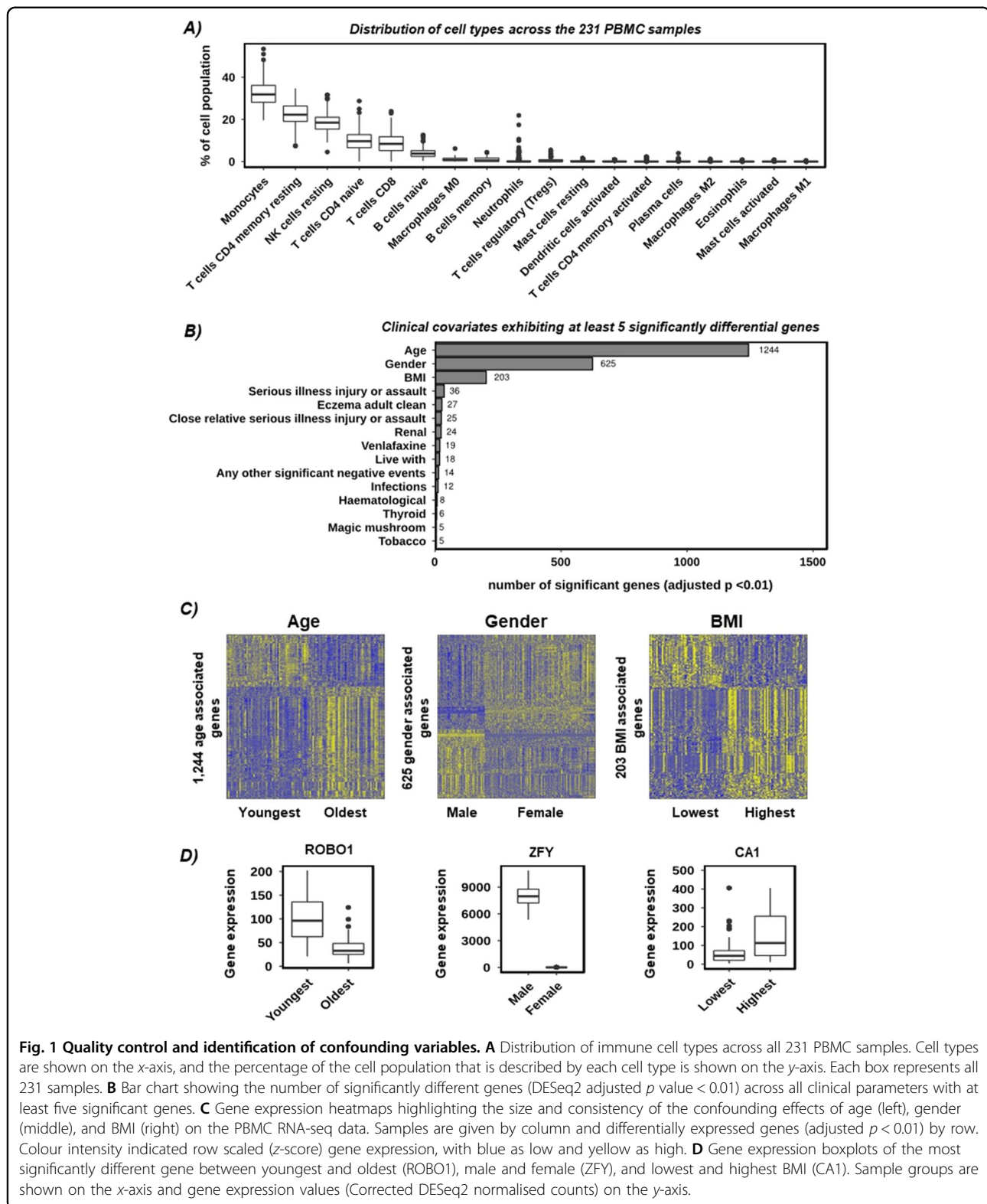
The co-expression network cluster analysis was based on the analysis performed by Le et al.<sup>16</sup> and used their code as a template. The method is detailed in full in Supplementary Information. Briefly, a correlation tree was generated from the expression matrix based on Pearson correlation coefficients and a topological overlap matrix. Clusters were identified by cutting the tree at a height of 0.95. To identify any clusters with significantly different gene expression between HC and MDD samples, a metagene for each cluster was generated using per gene Z-scores. For each cluster the mean expression z-score across all genes in that cluster was calculated, for each sample. The resultant scores for the HC samples were compared to that of the MDD samples using an unpaired, two-tailed *T*-test. *p* values were adjusted using the Benjamini–Hochberg procedure.

### Expression microarray analysis

The GSK-HiTDiP MDD<sup>17</sup> microarray data was downloaded from GEO (GSE98793) and the 22 samples that were reported to have failed QC were removed. The expression data was then quantile normalised using Limma<sup>18</sup>. Unannotated probe sets were removed. To control for genes represented by several different probe sets, Jetset<sup>19</sup> was used to select the probe set for each gene with the highest Jetset score. This resulted in 20,191 valid probe sets. Differential expression analysis was performed between the HC and MDD groups using Limma, and included batch, age, gender and anxiety as additional covariates. All other parameters were left to default. The quantile normalised expression values were corrected for batch using Limma's "removeBatchEffect" function.

### RNA-sequencing biological age meta-genes

A list of PBMC age associated genes was identified by using DESeq2 to compare the samples of lowest to highest quartile of age, as described above. Next the expression values (non-corrected but outlier capped) for the PBMC age related genes were scaled (per gene z-score), with the sign inversed for genes that were downregulated with age. Finally, the mean scaled value (across all sig genes) per sample was calculated. This value was considered as the samples biological age. The samples biological age was



**Fig. 1 Quality control and identification of confounding variables.** **A** Distribution of immune cell types across all 231 PBMC samples. Cell types are shown on the x-axis, and the percentage of the cell population that is described by each cell type is shown on the y-axis. Each box represents all 231 samples. **B** Bar chart showing the number of significantly different genes (DESeq2 adjusted  $p$  value  $< 0.01$ ) across all clinical parameters with at least five significant genes. **C** Gene expression heatmaps highlighting the size and consistency of the confounding effects of age (left), gender (middle), and BMI (right) on the PBMC RNA-seq data. Samples are given by column and differentially expressed genes (adjusted  $p < 0.01$ ) by row. Colour intensity indicated row scaled (z-score) gene expression, with blue as low and yellow as high. **D** Gene expression boxplots of the most significantly different gene between youngest and oldest (ROBO1), male and female (ZFY), and lowest and highest BMI (CA1). Sample groups are shown on the x-axis and gene expression values (Corrected DESeq2 normalised counts) on the y-axis.

then plotted against the samples chronological age, and the spearman correlation value determined. To optimise this metric, we repeated over a range of adjusted  $p$  and

$\log_2$ fold change cut-offs and selected the combination with the greatest correlation with patient age. For full details see the Supplementary Information.



## Results

### Quality control and identification of confounding variables

The NIMA samples were deeply sequenced and aligned to the human genome, exhibiting a high per sample alignment and feature counts rate (>85% alignment, median of 40 million exonic reads, per sample). Deconvolution analysis<sup>13</sup> showed the distributions of cell types to be consistent between the samples and typical of PBMCs (Fig. 1A). Deseq2 Differential expression analysis identified fifteen potentially confounding clinical covariates (each with >5 significant genes each at adjusted  $p < 0.01$ ) from a panel of 87 (Fig. 1B), with Age, Gender and BMI showing the strongest effects by an order of magnitude (1244, 625 and 203 significant genes respectively). The expression profiles for the Age, Gender and BMI associated genes were consistent across all samples (Fig. 1C) and the most differential genes (Fig. 1D) were consistent with the relevant biology (e.g. the most significant gender related genes were UTX and HYA which are X and Y linked<sup>20,21</sup>). We therefore concluded firstly that the data was of a high quality both technically and experimentally, and secondly that, given the size of the observed effect in the primary data, it was appropriate to control for the fifteen confounding clinical covariates in the downstream analysis.

### There is no robust evidence for a differential expression signature between HC and MDD in PBMCs

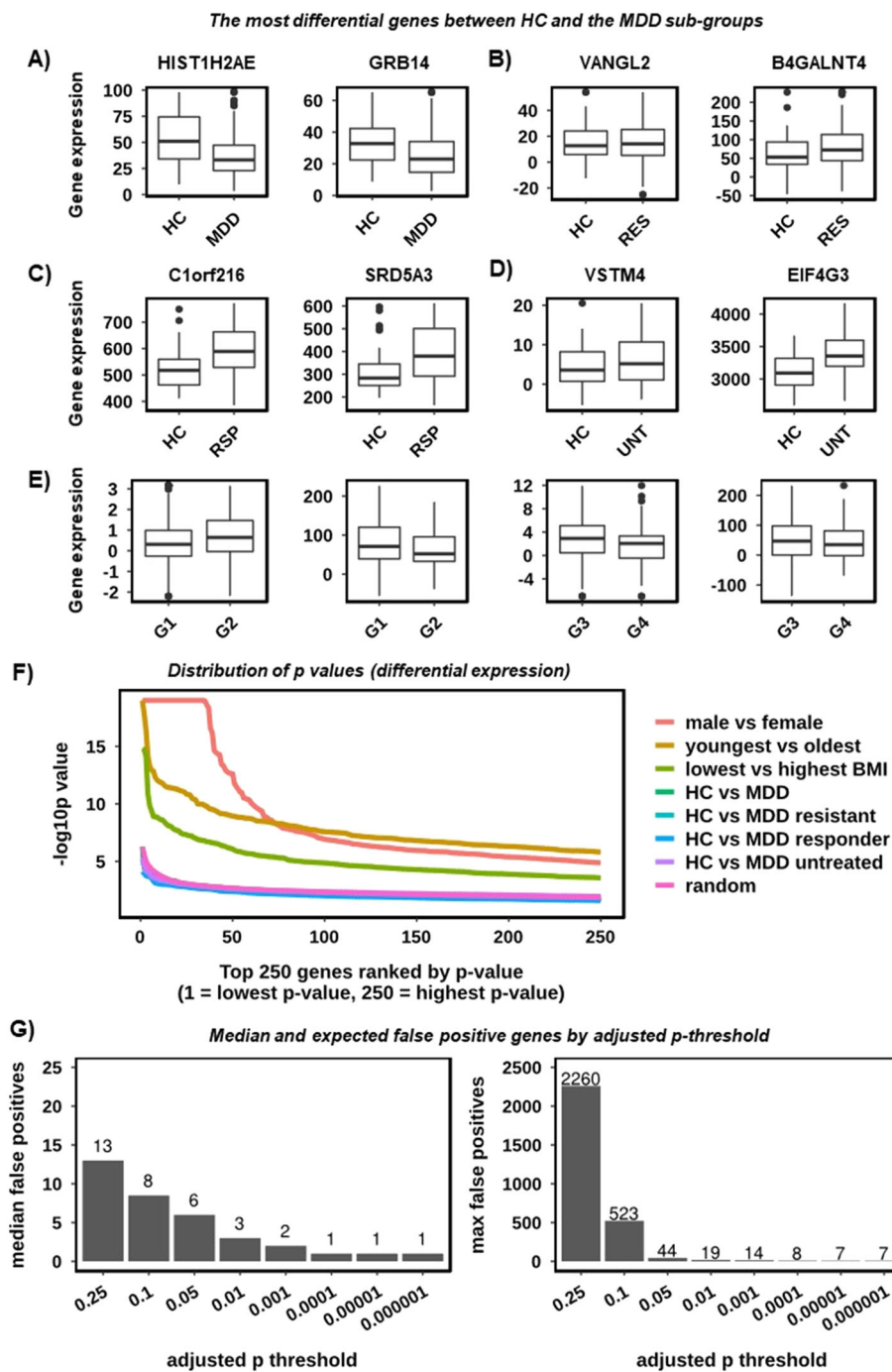
We used differential expression analysis to characterise any differences between HC and each of the MDD groups (MDD, treatment-resistant, treatment-responsive and untreated), using an adjusted  $p$  cut-off of  $<0.01$ , and including all 15 confounding clinical covariates plus batch as interaction terms. One significantly different gene was evident between HC and MDD (*HIST1H2AE*, adjusted  $p = 0.008$ ) and none between HC and MDD responders, MDD resistant or MDD untreated. We additionally tried reduced differential models – without BMI, with Age, BMI and Gender only and with Batch only, however it made no meaningful difference to these results. Observing only one significant gene suggested that either (1) the adjusted  $p$ -value threshold was too strict, or (2) the adjusted  $p$ -value threshold was reasonable, and we were observing type I error at *HIST1H2AE*. When we viewed the per sample expression at *HIST1H2AE* (Fig. 2A) it showed the difference in expression between HC and the MDD groups to be highly subtle. This was also true for the two genes of lowest  $p$ -value (non-significant) for each of the four comparisons (Fig. 2A–D).

Fig. 2E highlights the two most significant genes from each of two comparisons of randomised cases and controls. Randomised groups are labelled G1–G4. At the 250 most highly significant genes for each comparison the distributions of  $p$ -values were almost identical to that of

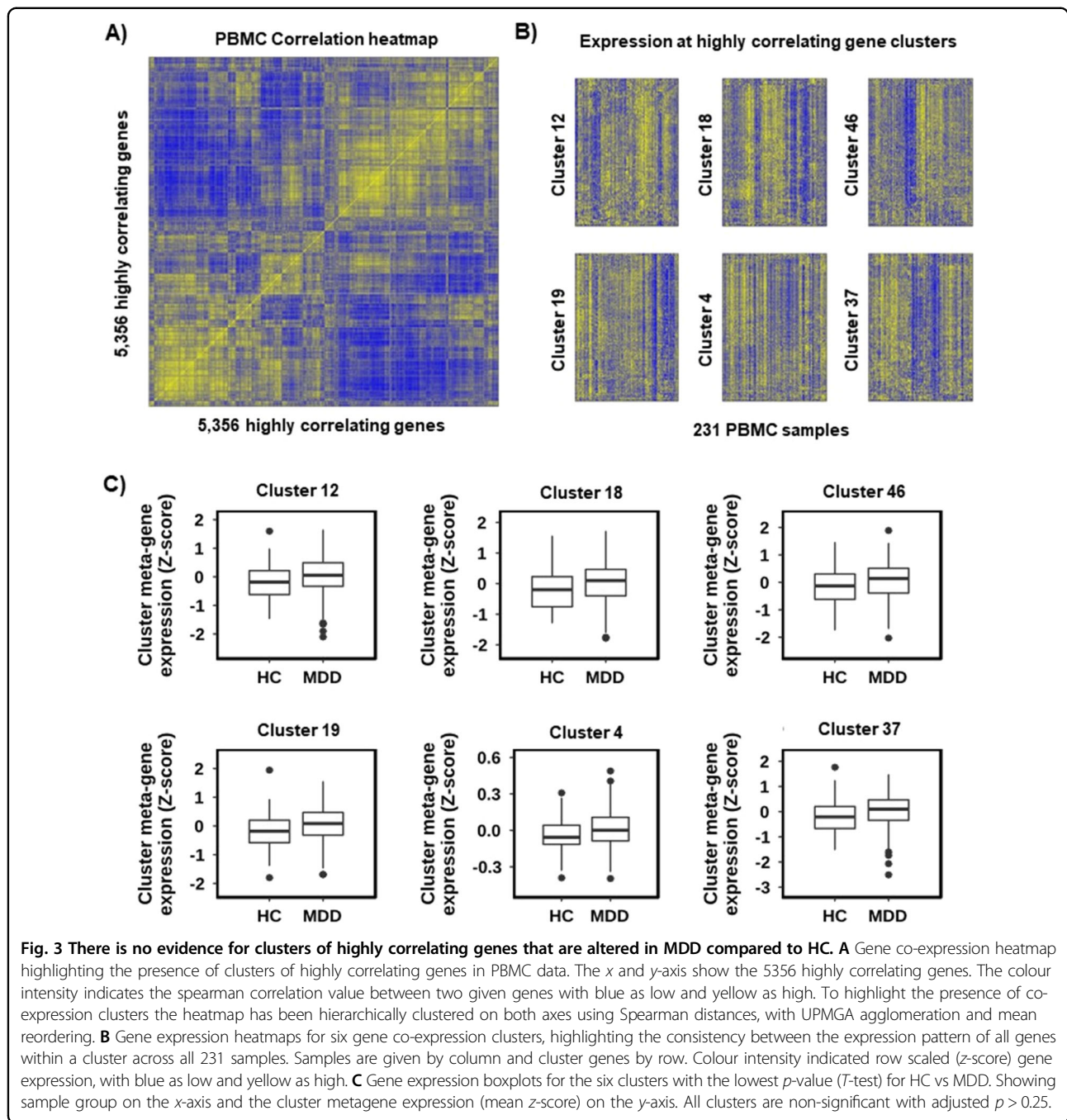
randomised cases and controls (Fig. 2F). This was in stark contrast to age, gender and BMI. These observations suggested that relaxing the adjusted  $p$ -threshold would not increase the number of true positives. We next estimated the number of false positives expected in this dataset at a range of adjusted  $p$  thresholds by generating 50 differential expression comparisons using randomised cases and controls and taking the median and maximum numbers of significant genes (Fig. 2G). The results showed that we would expect on average three false positives at adjusted  $p < 0.01$ , suggesting that it was not unlikely for *HIST1H2AE* to be false positive in this case. Though it is difficult to prove a negative outright, the balance of probabilities suggest that the data more strongly supported the absence of a HC vs MDD differential expression signature in PBMCs.

### There is no evidence for clusters of highly correlating genes that are altered in MDD compared to HC

We next considered the possibility that a HC vs MDD differential signature in PBMCs could be too subtle to detect using single gene interactions. This could occur for example if it originated from a subset of cells within the population. Several transcriptomic studies have shown<sup>22–25</sup> that subtle signatures can be reliably detected by collapsing clusters of highly correlating genes into representative metagenes for differential expression analysis. This acts to reduce noise and multi-sample correction stringency at the expense of single gene resolution. To do so we removed genes with low expression (mean  $>10$ , in the Combat corrected data) or with exceptionally high coefficient of variability (standard deviation/mean  $<0.15$ ), to reduce the chance that correlations could be driven by technical variability. Next, we generated a gene co-expression matrix from the remaining 5356 genes and plotted it as a hierarchically clustered heatmap (Fig. 3A). The heatmap showed clear structure and confirmed the existence of several clusters of highly correlating genes. To identify the correlation clusters, we used the method as described in Le et al.<sup>16</sup> (Supplementary Information). We identified 48 gene clusters with at least 50 genes in each. To validate these clusters, we plotted them as expression heatmaps (Fig. 3B), which confirmed the highly correlating nature of the genes in each. Next, we set out to determine whether the expression at cluster metagenes differed between HC and MDD. We generated per cluster metagenes and compared the metagene expression for HC samples to MDD samples. We observed no significant difference ( $p < 0.25$ , unpaired, two tailed  $t$ -test with Benjamini–Hochberg correction) between HC and MDD in any cluster (data not shown). Boxplots of the six clusters of lowest  $p$ -value (non-significant) highlighted the absence of any convincing biological differences at each cluster (Fig. 3C). We therefore concluded that there was no evidence for clusters of



**Fig. 2** There is no evidence for a classical differential expression signature between HC and MDD in PBMCs. **A** Gene expression boxplots highlighting the most significantly different genes between HC and MDD. Sample groups are shown on the x-axis and gene expression values (DESeq2 normalised counts) on the y-axis. **B** As **A** however for HC vs the MDD treatment-resistant group. **C** As **A** however for HC vs the MDD treatment-responsive group. **D** As **A** however for HC vs the MDD untreated group. **E** As **A** however for the two most significant genes from each of two comparisons of randomised cases and controls. Randomised groups are labelled G1–G4. **F** Distribution of differential expression *p*-values highlighting the consistency between HC vs MDD and randomised cases and controls. The 250 most significant genes for each comparison are shown on the x-axis (ranked from lowest to highest) and the *p* value (as  $-\log_{10}$ ) on the y-axis. Lines are given for the three confounding variables Gender ('male vs female'), Age ('youngest vs oldest'), BMI ('lowest vs highest'), HC vs the four MDD types (MDD, MDD treatment-resistant, MDD treatment-responsive and MDD untreated) and for the average of 50 comparisons of randomised cases and controls ('random'). **G** Bar charts highlighting the number of differentially expressed genes that were expected to be false positives by adjusted *p* threshold, based on 50 iterations of randomised cases and controls. The adjusted *p* threshold is given on the x-axis and the median (left) and maximum (right) number of expected false positives on the y-axis.



highly correlating genes that are altered in MDD compared to HC in this dataset.

**False positive genes were not random in PBMC data**

We observed in our 50 differential comparisons involving randomised cases and controls that the most significant genes included genes of immune function (such as TNF and IFIT2) more frequently than we expected. This raised the possibility that false positives genes might preferentially be immune genes when looking at PBMCs.

To test this hypothesis, for each gene we took the mean *p*-value across the fifty randomised comparisons, then selected the 50 most highly significant genes by mean *p*-value. We ran over representation analysis on the genes (using DAVID with GO biological processes and a background of the PBMC expressed coding genes) and found nine significantly enriched (<5% FDR) gene ontologies (Supplementary Table 2). All were immune related with the top three being “response to virus”, “type I interferon signaling pathway” and “cellular response to

interleukin-1” and included the genes IFIT1, IFIT2, IFIT3 and CCL8. As we used a background specific to PBMCs, this enrichment was relative to PBMCs and not other cell types. Thus, was suggestive that genes of these immune functions are more prone than other classes of genes to the type of stochastic noise that can result in a false positive. Therefore, indicating that false positives are not random in these data and show a significant bias towards certain immune functions. This further supported that it would not be reasonable to relax the adjusted  $p$  threshold when comparing HC to MDD, as it would likely introduce an erroneous immune signal that could be confused for bona-fide.

#### Relative to patient age biological age is significantly greater in MDD patients than HC

To explore whether MDD patients showed increased biological ageing compared to HC, we estimated the biological age of each sample by taking the mean expression value ( $z$ -score) across all the age-related genes (see Methods section and Supplementary Methods for full details) and plotted it against chronological age (Fig. 4A). As expected, we observed a strong positive and significant linear correlation between biological and chronological age (Spearman Correlation Coefficient (SCC) = 0.72,  $p < 0.01$ ). To determine whether MDD or HC patients showed altered biological ageing (relative to chronological age) we performed a linear regression using the model biological age  $\sim$  chronological age (Fig. 4A). Next, we counted the number of HC or MDD patients above or below the regression line and found a subtle (HC – 26 below (59%), 18 above (41%), MDD – 78 below (42%), 109 above (58%)) but significant difference ( $p < 0.05$ , Fisher’s exact test). To illustrate the difference in distribution, we used the residuals – i.e. the distance along the  $y$ -axis of each dot from the regression line (Fig. 4B). Finally, to validate the result we replicated the analysis using the GSK-HiTDiP MDD whole blood microarray data. The results were comparable to PBMCs (Fig. 4A, B), with the MDD patients showing significantly elevated biological ageing relative to chronological ageing (HC – 35 below (61%), 22 above (39%), MDD – 48 below (42%), 65 above (58%)),  $p < 0.05$ , Fisher’s exact test).

#### Discussion

In this large, well-controlled and deeply sequenced data-set, we find no evidence for a differential expression signature in PBMCs between HC and MDD – as a whole or in the subtypes described; nor is there evidence for clusters of highly correlating genes that are altered in MDD compared to HC. We also found that biological age relative to chronological age is significantly greater in MDD patients than in HC.

Our differential analysis showed only one gene to differ significantly (adjusted  $p < 0.01$ ) between HC and MDD and none between HC and MDD sub-groups. Further investigation concluded that, given the concurrence between the distribution of  $p$  values for random samples and the HC and MDD group comparisons, the very low difference in expression between HC and MDD groups at these genes, and the numbers of expected false positives at this adjusted  $p$ -threshold, this was most likely a false positive, unlikely to be biologically meaningful, and that there was no justification for relaxing the  $p$  value threshold in this data. To test whether any HC vs MDD signature might be too subtle to detect at the single gene level, we generated 48 gene co-expression clusters and compared the metagenes between HC and MDD. We observed no significant differences in any cluster, or any convincing biological differences. We therefore concluded that the data more strongly supported the absence of a HC vs MDD differential expression signature in PBMCs.

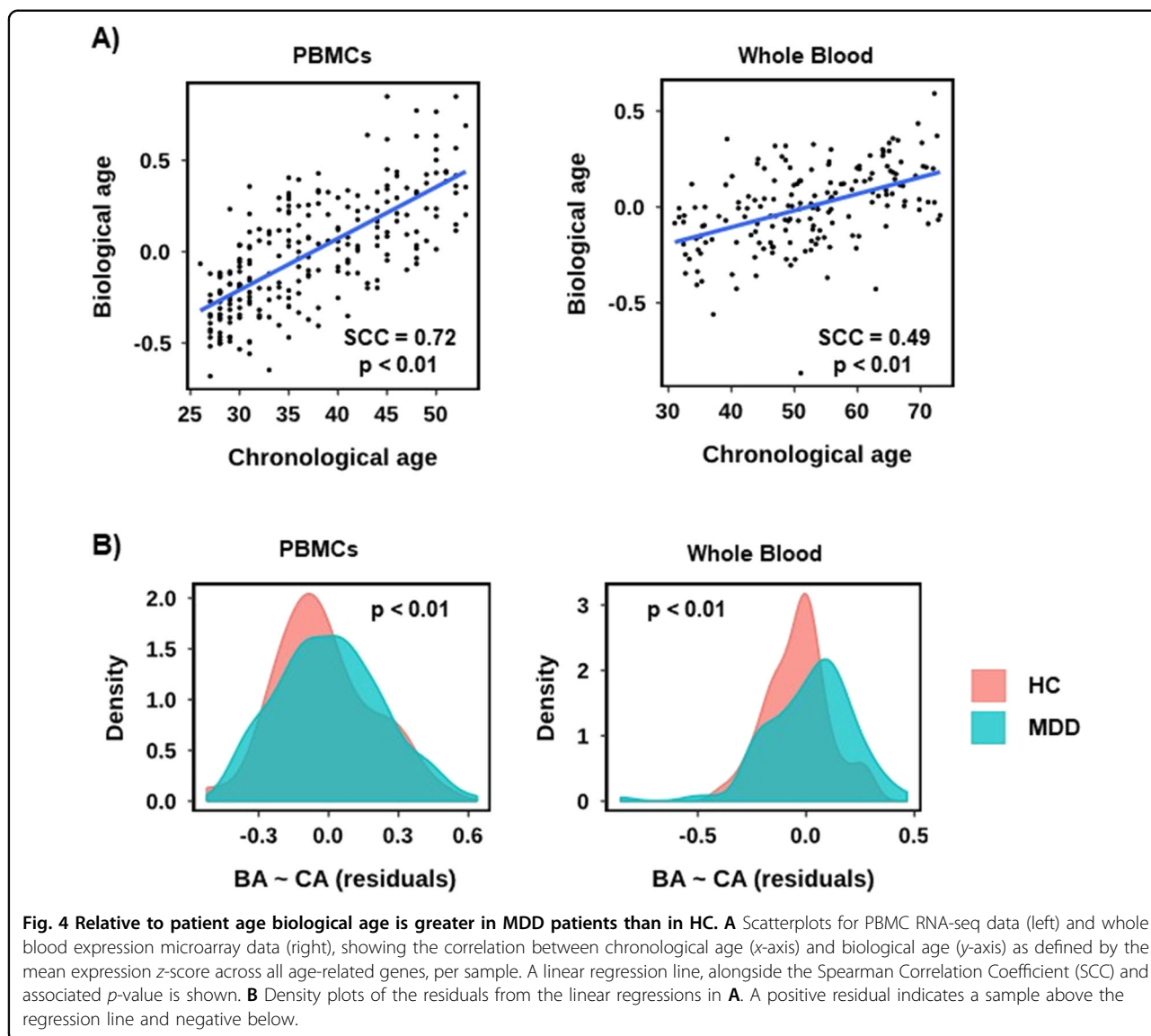
In addition, when we randomised cases and controls 50 times and performed over representation analysis, we found the most significant false positives to not be random but to have a significant immune phenotype, including “response to virus” and “type I interferon signalling pathway”. This further justified not relaxing the adjusted  $p$  threshold in this data, as doing so would likely introduce an erroneous immune signature that could be interpreted as bona fide.

These results are, in many ways, comparable to previous transcriptomic studies in whole blood which also found no signature at adjusted  $p < 0.05$  using larger sample numbers<sup>2</sup>. One strength of our approach is that we control for age, gender and BMI in our sample selection. In our opinion, we could not justify relaxing our adjusted  $p$  threshold. However, other studies identified signatures at adjusted  $p$  values ranging from  $p < 0.1$  to  $p < 0.25$ .

A further strength of our study is that we present the per sample expression values for all genes of interest. We would argue that as other data<sup>2,26</sup> presented signatures that were detectable only at adjusted  $p > 0.05$  using  $\sim 1000$  samples each, these signatures are likely to be subtle. However, the omission of per sample expression data at the genes of interest, makes it difficult to establish how subtle and so it is difficult to form a robust opinion of how biologically meaningful these expression differences are.

As mentioned in the introduction, evidence for an inflammatory protein signature in MDD is substantial. This is particularly the case for the proinflammatory cytokine IL-6, with several meta-analyses confirming this. There is also a longitudinal association between MDD and IL-6<sup>27</sup>, yet the tissue source of cytokines remains unclear. Our data strongly suggest that in this sample of MDD, the source of cytokines is unlikely to be PBMCs. Reflecting on other potential sources; *neutrophils* are increasingly seen as





important for fine regulation of the immune-inflammatory response, outnumbering PBMCs by one or two orders of magnitude<sup>28</sup>. Neutrophils produce a large variety of chemokines and cytokines upon stimulation and can differentially switch phenotypes, displaying distinct subpopulations in different microenvironments<sup>29</sup>. If neutrophils confer the cytokine signature, it would be expected that gene expression studies of whole blood would capture their contribution. Another potential cell source are *endothelial cells*. These are ubiquitous in both brain and periphery. Recently, Blank et al. demonstrated a specific role in relation to aspects of depression-relevant behaviour in mice by showing that downstream signalling of brain endothelial cells induces fatigue and cognitive impairment via impaired neurotransmission in the hippocampus<sup>30</sup>. However, assessing the individual contribution of endothelial cells in

humans would be technically very challenging. Nevertheless, considering findings presented in a recent GWAS of MDD, it is important to consider that peripheral tissues may have less of an overall contribution than the brain. Wray et al. integrated their GWAS data with functional genomic data, comparing their findings with bulk tissue RNAseq from genotype tissue expression (GTEx)<sup>1</sup>. Here only brain tissue showed enrichment, with the areas showing the most significant enrichment being cortical. This was in contradistinction to other tissue types including whole blood.

The issue of body mass in MDD is complex. Wray et al. found significant positive genetic correlations with body mass<sup>1</sup> and Mendelian randomization (MR) analysis was consistent with BMI being causal or correlated with causal risk factors for depression. Also, negative MR results

provide important evidence of no direct causal relationship between MDD and *subsequent* changes in BMI. Adipose tissue actively secretes cytokines and obesity is itself associated with changes in the secretome of adipocytes leading to increased production of proinflammatory cytokines<sup>31</sup>. This raises the possibility that *adipocytes* may be a potential source of inflammatory cytokines acting as a tissue “reservoir”. Careful consideration should be applied when deciding whether BMI should be treated as a confounding variable in MDD or incorporated as part of disease pathogenesis.

We demonstrated that MDD samples showed significantly elevated biological age compared to HC. Although significant, the effect was relatively subtle, comparable to that identified in CpG methylation data<sup>32</sup>. Diniz et al. found MDD exhibited greater molecular senescence in young and middle-aged adults by examining the impact of MDD on the senescence associated secretory phenotype (SASP), a dynamic secretory molecular pathway indicative of cellular senescence<sup>4</sup>. More severe episodes of depression present with higher SASP indices and a significant interaction between current MDD episode and overweight, thus comorbid current MDD plus being overweight had the highest SASP index. While we have not correlated with direct measures of senescence such as SASP indices or epigenetic markers, we would argue that our finding is consistent with the literature and points to a potentially interesting biology.

The strengths of this study lie in the high-quality RNA and large clinical dataset, sequenced to an average depth of >54.5 million reads, which aligned with >70% of the reads mapping to exons. Thus, a deeply sequenced, well-controlled clinical sample. The limitations of this study relate to heterogeneity inherent in MDD. Within our study, there was also some heterogeneity within the assessing of prior medications as this was done using retrospective self-reporting, albeit based on a comprehensive structured instrument completed by an interviewer. The lack of medical comorbidities was an a priori decision and can be seen as both a strength in removing the potential confound of comorbid inflammatory illness and a limitation in respect of the typicality of an MDD group.

The age range is relatively narrow and could be seen as a limitation in relation to interpretation of biological ageing. Lastly, we did not measure cytomegalovirus (CMV) serostatus, which has been associated with differences in immune cell profiles<sup>33</sup>.

## Conclusion

This study was a detailed and careful examination of the transcriptomic signal in PBMCs in MDD and HCs. The lack of a significant differentiating signal between MDD and HCs was confirmed by the randomisation of

the cases and controls. There was, however, evidence of elevated biological ageing relative to patient age in MDD vs HC. Future work should endeavour to expand clinical sample sizes, reduce MDD heterogeneity and account for confounds from the outset. Advances in RNA-seq at the level of the single cell may help uncover further, more subtle differences. However, the subtlety of any signature mitigates against convincing use as a diagnostic or predictive biomarker, and tissue enriched data is strongly indicative of brain tissue being the most informative in this regard.

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## Conflict of interest

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