## **Evaluation of diversity, specialization, and gene specificity in transcriptomes** Octavio Martínez\* and Humberto Reyes-Valdés\*\*

\*Laboratorio Nacional de Genómica para la Biodiversidad (Langebio) Cinvestav Irapuato, México. \*\* Department of Plant Breeding, Universidad Autónoma Agraria Antonio Narro, Saltillo, Coahuila, México

The transcriptom

The transcriptome is a set of genes transcribed in a given tissue under specific conditions and can be characterized by a list of genes with their corresponding frequencies of transcription. Transcriptome changes can be measured by counting gene tags from mRNA libraries or by measuring light signals in DNA microarrays.

## **Transcriptome Diversity**

Consider the division of an organism in tissues, then the transcriptomes of each tissue Can be simply described as the set of relative frequencies,  $p_{ij}$ , for the *i*-th gene (*i* = 1, 2, ..., *g*) in the *j*-th tissue (*j* = 1, 2, ..., *t*). Then the diversity of the transcriptome of each tissue can be quantified by an adaptation of Shannon's entropy formula (1),

$$H_j = -\sum_{i=1}^{g} p_{ij} \log_2(p_{ij})$$

 $H_i$  will vary from zero when only one gene is transcribed up to  $log_2(g)$  where all g genes are transcribed at the same frequency: 1/g. Gene specificity

Consider the average frequency of the i-th gene among tissues,  $p_i$ , then gene specificity is given by

$$S_i = \frac{1}{t} \left( \sum_{j=1}^r \frac{p_{ij}}{p_i} \log_2 \frac{p_{ij}}{p_i} \right)$$

S, will give a value of zero if the gene is transcribed at the same frequency in all tissues and a maximum value of  $log_2(t)$  if the gene is exclusively expressed in a single tissue.

**Relative Transcriptome Specialization** Tissue specialization can be measured by the average of the gene specificities, say,

$$\delta_j = \sum_{i=1}^{n} p_{ij} S_i$$

 $\delta_i$  varies from zero, if all the genes expressed in the tissue are completely  $\delta_j$  varies from zero, if all the genes expressed in the tissue are completely (Sepecific (S<sub>j</sub> = 0 for all *i*) up to a maximum of  $\log_2(i)$  when all the genes (S<sub>j</sub> = 0 for all *i*) up to a maximum of  $\log_2(i)$  when all the genes (S<sub>j</sub> = 1) presents an example of the point estimation of transcriptome diversity and specialization in 36 human tissues, derived from microarray data. Statistical properties of the estimators Asymming a multinomial distribution for the number of gene tags in the

3012.1

transcriptome it is possible to give approximate values for the expectation and  $\frac{1}{2}$  ince of the estimator of transcriptome diversity  $(H_i)$ . These expressions, found by Basharin (2), are given by

$$E[\hat{H}_{j}] = H_{j} - \frac{g-1}{2n_{j}} \log_{2}(e) + O(\frac{1}{n_{j}^{2}})$$
$$V[\hat{H}_{j}] = \frac{1}{n} \left[\sum_{i}^{s} p_{ij} (\log_{2}(p_{ij}))^{2} - H_{j}^{2}\right] + O(\frac{1}{n^{2}})$$

Gere E[•] and V[•] represent the expectation and variance operators represent the sequence of 1.442695 and results from the use of logarithms base 2 instead of the sequence The avalue of results and results from the use of togarithms base prime and the estimator is based on matrix and the standard of the standard

transcriptome specialization,  $\delta_j$ , are too complex to be obtained analytically. Bewever they can be assessed by re-sampling the original distribution to obtain approximate but robust inferences. Approximate confidence intervals for the information parameters

If is possible to obtain confidence intervals for the transcriptome diversity  $(H_i)$ by possible the asymptotic normal distribution of the estimator (2), and or some can be achieved for this parameter as well as for the transcriptome specialization ( $\delta_j$ ) and gene specificities (S) by re-sampling the gene tag defasets assuming the multimonial distribution (3).

The confidence interval methods tested were: 1) Asymptotic Interval or AI, assuming the normal distribution and using the approximation for the variance o[H]; 2) using the mean and variances obtained from the bootstrap procedure affit assuming normality. that we name here Bentreme to the second se assuming normality, that we name here Bootstrap Asymptotic Interval or **I** is a summing investigation of the second method proposed by Hall (4) based in re-sampling the error of the estimation and named here as Hall Bootstrap Percentiles or HBP. To test the methods for the estimation of confidence intervals we simulated three

datasets of sample sizes  $n_i = 100$ , 1000 and 10000, each one constituted by four transcriptomes (t=4) and 16 genes (g=16). The values of  $p_{ij}$  were selected to have values of  $H_i$  of approximately equal to 1, 2, 3 and 4 (the maximum) and also maximum diversity in specialization among the transcriptomes. A total of B=10,000 bootstrap replicates were obtained from each dataset.

From Fig 2 we can observe that for the maximum value of  $H_j = 4$  both, the BAI and BPI intervals fail to include the true value of the parameter, and the HBP fall to the right of the parametric value overestimating it out of its possible rank (obvious in sample size 100). In all the other cases, except for the value of  $H_j = 1$ with  $n_j = 100$ , the true value of  $H_j$  is included within the intervals. This means that when the sample size is small all methods could fail to include the true value of  $H_i$ . The differences between the lengths of the intervals obtained by the three methods (AI, BAI and BPI) for  $H_j$  for the same sample size are not large. However, given the lack of normality that could be present when the sample size is relatively small, it appears safer to use the bootstrap percentile (BPI or HBP) to obtain confidence intervals for this function

From Fig. 3 we can notice that the majority of the intervals, for all sample sizes, include the true value of the parameter, except for the case of the gene 4 for which all intervals for the sample size  $n_i=1000$  fails to include the true value. For all cases studied the specificity estimator is positively biased, i.e., it tends to overestimate the specificity of the gene. The Shapiro-Wilks test for normality of the bootstrap replicates of  $S_i$  demonstrated that in the majority of the cases its distribution is not normal, and thus for this case the BPI must be preferred over the other methods to obtain confidence intervals.

From Fig. 4 we can see that the true value of  $\delta_i$  is included in all the HBP intervals (limits shown as diamonds in the graph), however, for some cases the BAI and BPI intervals fail to include the true value of  $\delta_i$ . In the case of  $\delta_i$  the HBP intervals are much better centered on  $\delta_i$  than the BAI and BPI intervals. On these grounds the HBP method can be preferred for this function. We can also notice that the estimator of  $\delta_i$  is consistently biased to the right of the true parameter. The Shapiro-Wilk test applied to the bootstrap replicates of showed that normality can be assumed for the estimator of  $\delta_i$  only for large sample sizes.









Fig. 2 Approximated 95% confidence intervals for the values of  $H_{j}$ . Transcriptomes (i), are presented in the Y axis separated by grav discontinuous lines. The scale in the X axis is for the values of H. Intervals are colored by sample size; red for  $n_j$ =100, blue for  $n_j$ =1000 and green for  $n_j$ =10,000. Parametric values of  $H_j$  for each transcriptome are represented by black asterisks over the grey discontinuous lines. Values of the estimated  $H_j$  are represented by colored vertical lines, while the values of the means of the B=10,000 bootstrap replicates are shown as open circles. The limits for the BAI intervals are denoted by colored triangles pointing below (lower limit) or above (upper limits), AI limits are shown as black points while the BPI intervals are shown as continuous color lines and HBP limits are shown as red diamonds.









Fig. 5 Example of the estimated diversity and specialization in two human tissues from an analysis of MPSS data (1). Squares are the values originally estimated in the sample. Black lines are 95% approximate confidence intervals for the parameters using the BPI method. The intervals are positioned into the mean of the bootstrap replicates to show the bias in both functions (Hj and δ j). A red line joins the estimates in the original sample and the point where the means of the bootstrap replicates are.

Fig. 5 shows with real data, how the correction for bias and the estimation of confidence intervals are fundamental for the correct inference of the properties of the transcriptomes.

## Conclusion s and Perspectives

The methods presented here allow performing correct inferences about the diversity and specialization of the transcriptomes as well as about the specificities of the genes that are expressed in them. For transcriptome diversities  $(H_i)$  and gene specificities  $(S_i)$ , the bootstrap percentile method (BPI) appears to be a better choice. The methods shown are easily extended to perform test of hypothesis for the parameters of interest; this can be done by obtaining bootstrap estimates of the relevant difference and calculating the BPI. If the BPI contains zero (at a given confidence level) then the null hypothesis is not rejected. This permits to decide, for example, if two genes are equally specific for a set of transcriptomes or if two transcriptomes differ significantly in diversity or specialization. A lot of biologically relevant questions can be answered in this way.

The methods presented are only applicable to the counting of gene tags, as in SAGE, MPSS, large collections of ESTs, pirosequencing (454) or any other high throughput sequencing strategy. However, these methods are not directly applicable to continuous measurement of gene expression, as the ones performed in microarrays.

## References

1. Martínez, O. & Reyes-Valdés, H. (2008) Defining diversity, specialization, and gene specificity in transcriptomes through information theory. PNAS 105, 9709–9714. 2. Basharin, G. P. (1959) On a Statistical Estimate for the Entropy of a Sequ

Independent Random Variables. Theory of Probability and its Applications 4, 333-336. 3. Efron, B. & Tibshirani, R. J. (1993) An introduction to the bootstrap (Chapman & Hall, New York - London).

4. Hall, P. (1992) The Bootstrap and Edgeworth Expansion (Springer, New York). Acknowledg

We are thankful to Conacyt, Concyteg, Cinvestav and UAAAN for the support given to this project.

4th. EMBO Conference From Functional Genomics to Systems Biology, EMBL Heidelberg, 15 - 18 November, 2008.