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# QUANTITATIVE VARIATION WITHIN AND BETWEEN POPULATIONS OF THE WILD BARLEY, HORDEUM MURINUM

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#### SUMMARY

Variation in quantitative characters was studied in several wild populations and families of the weedy barley, *Hordeum murinum*. The hereditary, nonherediatary and interactive components of the variance were assessed for up to 13 metrical characters. Broad-sense heritabilities were calculated as were the relative contributions of populations, families and individuals to the overall variance. Correlation between characters was also considered. Significant genetic differences were observed between populations and between families within a population for developmental, morphological and reproductive characters. For most characters, the largest component of variance was due to population differences. Heritability was high for developmental characters and low for those associated with reproduction. The results have been discussed in light of experimental procedures used to assess the variation in wild populations for plant gene resource screening.

#### 1. INTRODUCTION

This paper describes a study in which genetic variation for quantitative characters was assessed within and between populations of a wild barley. The growing interest in screening of populations of wild and weedy relatives of crop plants for genetic resource purposes led to this study.

It is common practice to analyse quantitative variation using biometrical techniques in the trials of cultivated varieties and breeding lines. Various individuals whose genetic relationships are known, are subjected to a range of environments or treatments. With appropriate experimental designs, the Genotype, Environment and Genotype-Environment Interaction variance components can be estimated and the performance of the varieties under different environments can be predicted (*e.g.*, Yates and Cochran, 1938; Freeman and Dowker, 1973; Freeman, 1973; Ridgman, 1975; Moll and Stuber, 1974; *etc.*).

Assessment of the sources of quantitative variation in natural populations is more difficult (e.g. Hillel, Feldman and Simchen, 1973) due to problems associated with obtaining samples for which the genetic relationships between the individuals are known. Experiments must also be carried out under controlled conditions and there is a danger that inferences about the variation and its selection in the natural habitat may not be valid. However, measurement of quantitative variation using biometrical techniques in samples from wild populations grown in common environments can provide valuable information about the variation present in natural populations.

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As an example, the quantitative variation was analysed in populations of the cleistogamous wild barley, *Hordeum murinum*. Several wild-collected populations were grown in controlled environments to determine the heritable, non-heritable and interactive components of the quantitative variation between populations, families or individuals and to calculate genetic correlation between characters.

# 2. MATERIALS AND METHODS

# (i) Experimental design

Three experiments were carried out to investigate the variation in quantitative characters in *H. murinum*.

A randomised complete block design was used for the experiments. Equal numbers of plants from each family or population were allocated at random to each of four benches in a greenhouse. Each bench was treated as an experimental block; a random numbers table was used to allocate the plants within each block (table 1).

The three experiments consisted of the following:

Experiment I—16 individuals from each of 18 populations Experiment II—8 individuals from each of 10 populations Experiment III—8 seeds from each of the first heads of 47 individuals from 1 population

Material collected from the field as seed was used in Experiments I and II. Collection localities ranged from 50°N to 30°N latitude. In Experiment III each 8 seed set is a single full sib family obtained from a previous greenhouse experiment. Table 1 summarises the origins and experimental sample size of the material used in all three experiments.

The decimal code for Cereal Growth Stages was used as a standard by which to compare plants with different developmental times, morphology and reproductive behaviour (Zadoks *et al.*, 1975; Tootman *et al.*, 1979). Thirteen characters were analysed for this paper: Initial Seed Weight (SWt), Radicle Emergence (RAD), Coleoptile Emergence (COL), Leaf Through Coleoptile (LTC), First Leaf Unfolded (1st L), Second Leaf Unfolded (2nd L), Tiller (T)—number of days to attain a certain tiller number or numbers of tillers present on a certain day, Culm Length (CL), Flag Leaf Length (FLL), Flag Leaf Width (FLW), Spike Number (SpNo), Seed Number (SNo), and Fertile Seed Number (FSN).

### (ii) Statistical analysis

The data from all three experiments were analysed by analyses of variance assuming a random effects model. None of the block  $\times$  population/family mean squares were significantly different from the sample variances in Experiments I, II or III. Thus both sources of variance were pooled to give the appropriate estimate of the error variance.

A crude approximation of the relative contributions of population, family and individual to the overall variance was estimated from Experiments II and III in the following way:

Let Var 
$$(Y) = \hat{\sigma}_p^2 + \hat{\sigma}_f^2 + \hat{\sigma}_e^2$$

TABLE 1	LE 1
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Population	Latitude	Block	Population Family
Experiment I			
Hartlepool, Cleveland	54°40'N/1°15'W	4	16
York, Yorkshire	53°58'N/1°5'W	4	16
Shrewsbury, Shropshire	52°12'N/2°15'W	4	16
Lowestoft, Suffolk	52°29'N/1°45/E	4	16
Worcester, Worcestershire	52°12'N/2°15'W	4	16
Grantchester, Cambridgeshire	52°12′N/0°8′W	4	16
Cambridge, Cambridgeshire	52°12'N/0°8'W	4	16
Oxford, Oxfordshire	51°48'N/1°15'W	4	16
Porteynon, S. Glamorgan	51°34'N/3°15'W	4	16
Barry, S. Glamorgan	51°25'N/3°25'W	4	16
Guernsey, Channel Is.	49°31'N/3°25'W	4	16
Castel des Biches, France	49°15'N/3°10'W	4	16
Pisa, Italy	43°45'N/10°20'E	4	16
Aix, France	43°30'N/5°30'E	4	16
Paradis d'Ete, Italy	43°8'N/12°25'E	4	16
Golan Heights, Israel	33°5'N/35°35'E	4	16
Acco Plain, Israel	32°50'N/35°10'E	4	16
Western Negev, Israel	30°50'N/34°45'E	4	16
Experiment II			
Pisa, Italy	43°45'N/10°20'E	2	8
Aix, France	43°30′N/5°30′E	2	8
S. Peloponessos, Greece	36°30'N/22°30'E	2	8
W. Rhodes, Greece	36°15′N/27°50′E	2	8
Crete, Greece	35°20'N/23°30'E	2	8
Golan Heights, Israel	35°5′N/35°35′E	2	8
Acco Plain, Israel	32°50'N/35°10'E	2	8
Western Negev, Israel	30°50'N/34°45'E	2	8
Cairo, Egypt	30°3'N/31°15'E	2	8
Canterbury, N.Z.	44°45′S/171°35′E	2	8
Experiment III 47 Families from Pisa-grown in a previous experiment	43°45′N/10°20′E	2	8

Hordeum murinum populations and sample sizes used in statistical analysis of Experiments I, II and III

where  $\hat{\sigma}_p^2 =$  variance component due to populations from Experiment II,  $\hat{\sigma}_f^2 =$  variance component due to families from Experiment III,  $\hat{\sigma}_e^2 =$  variance component from the pooled error from Experiment III.

The percentage of total variation attributable to population, family or sampling error was calculated

$$R_p^2 = \frac{\hat{\sigma}_p^2}{\operatorname{Var}(Y)}.100, \qquad R_f^2 = \frac{\hat{\sigma}_f^2}{\operatorname{Var}(Y)}.100, \qquad R_e^2 = \frac{\hat{\sigma}_e^2}{\operatorname{Var}(Y)}.100$$

 $R_p^2$ ,  $R_f^2$  and  $R_e^2$  values were calculated for the nine characters which had been measured identically in Experiments II and III.

A broad-sense heritability (Falconer, 1960) was calculated for each character based on the variance components estimated from the analysis of variance (Experiment III).

To stabilise the variances, a logarithmic transformation was performed on the data for the characters SWt, RAD, COL, LTC, 1st L and 2nd L.

# 3. RESULTS

The results of the analyses of variance for 13 characters measured in Experiments I and II appear in table 2. No block effects were observed in Experiments I or II but significant block differences were obtained for the characters RAD, COL, LTC, 1st L, 2nd L, T2, TND134 and CL in Experiment III. Significant genetic differences were observed between populations for all characters measured in Experiments I and II and similarly between families in Experiment III for all but flag leaf and spike characters.

TABLE 2

Μ	lean :	squares j	from th	e Ana	lysis oj	f Var	iance, Ex	periments	I,	П,	Ш
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			df	SWt	RAD	COL	LTC	1st L
Experiment I							_	
block			3	0.0595	0.0400	0.0034		0.0033
between populations			17	0.3375	+ 0.1694+	0.1361		0.0953
pooled error			267	0.0036	0.0257	0.0254		0.0176
Experiment II								
block			3	0.0367	0.0103	0.0109		0.0032
between populastions			ğ	0.2305	1 0.05861	0.0153*		0.0037
pooled error			67	0.0269	0.0167	0.0072		0.0033
Experiment III								
block			3	0.0230	0.1180+	0.0707	0.0521+	0.0364+
between families			46	0.0626	+ 0.2385+	0.10851	0.0607	0.0555
pooled error			326	0.0209	0.0286	0.0133	0.0095	0.0084
* <i>P</i> < 0.05 † <i>P</i> < 0.01								
	T-B	T-C	CL	FLL	FLW	SpNo	SNo	FSN
Experiment I								
block	35.73		_		_			
between populations	160-32+				_			
pooled error	13.93						_	
Experiment II								
block	42.58	_	6108	88.10	1.683	15-11	5154	4778
between populations	370-03+		97229+	673.84+	4.283†	122.53+	55859+	35709
pooled error	44-48	-	4901	85.57	0.6261	8.42	2992	2243
Experiment III								
block	21.386	85-47	47972†	717.1	3-513	9.733	_	
between families	8.344+	41.62+	15740+	605 1*	2.798	6.938		
nooled error	2,402	14.36	01.55				-	

\* P < 0.05

÷ 0.01

The relative contributions by individuals, families and populations to the total variance were compared (table 3). In seven of the nine characters examined, the between population component of variance was generally the largest (40–90 per cent) while the individual component of variance accounted for only 4–6 per cent of the total. It is noteworthy that in the characters SWt, CL and SpNo more than 80 per cent of the total variance was contributed by the population component. In the flag leaf characters, however, the individual component of variance is large (23 per cent) and the family and population components are similar to each other ( $\simeq 40$  per cent).

The heritability of each character calculated from the analysis of variance for full sib families is shown in table 3. In general, heritabilities for the characters associated with reproduction are lower than those associated with early growth and development; e.g., the  $h^2$  calculated for SpNo or CL are lower than those for RAD or COL. This pattern is consistent with other studies in which low heritabilities are often found for characters closely associated with reproductive fitness (Falconer, 1960; Allard, 1960; Lewontin, 1974). In this study, three characters associated with reproduction (SpNo, CL, SWt) which have low  $h^2$ , also have  $R_p^2$  values in excess of 80 per cent. Very little of the total variance is contributed by the between-family component of variance. Flag leaf characters also have low  $h^2$  but the individual and family components of variance ( $R_f^2, R_e^2$ ) for these characters are sizeable. The sequence of developmental characters (RAD through 2nd L), have high  $h^2$  and relatively large  $R_f^2$ .

#### TABLE 3

Percentage of total variance contributed by populations  $(R_p^2)$ , families  $(R_f^2)$  and individual  $(R_e^2)$  and broad-sense heritability estimates with confidence intervals  $(t_{0.025})$  calculated from Experiment III by full-sib method (Falconer, 1960)

Character	R <sup>2</sup> <sub>p</sub>	R <sub>f</sub> <sup>2</sup>	R <sub>e</sub> <sup>2</sup>	Broad-dense heritability
Seed weight	80.17	14.87	4.96	$0.22 \pm 0.19$
Radicle emergence	38.81	54.65	6.54	$0.51 \pm 0.29$
Coleoptile emergence	52.78	42.05	5.17	$0.48 \pm 0.28$
Leaf through coleoptile				$0.45 \pm 0.27$
First leaf	59-96	34.85	5.26	$0.43 \pm 0.27$
Second leaf	67.96	28.37	3.67	$0.47 \pm 0.28$
Tiller 1				$0.22 \pm 0.19$
Tiller 2				$0.13 \pm 0.14$
Tiller number day 134			_	$0.20 \pm 0.13$
Culm length	80.27	13.00	6.73	$0.12 \pm 0.14$
Flag leaf length	40.71	36-55	22.74	$0.09 \pm 0.12$
Flag leaf width	46.01	30.06	23.93	$0.14 \pm 0.14$
Spike number	91.18	5.16	3.66	$0.07 \pm 0.07$

#### (ii) Correlation between characters

As a large number of biologically related characters have been chosen for study, the possibility of character correlation is high. The genotypic correlation matrix for Experiment III appears in table 4. Particular groups of characters were correlated. The groups were:

- (a) Sequential developmental characters: RAD, COL, LTC, 1st L, 2nd L; certain tiller characters.
- (b) Different measurements of the same morphological characters: FLL, FLW.
- (c) Different measurements of reproductively related characters: SpNo, SNo, FSN.

1. SWt	1.00		_				
2. RAD	-0·39	1.00			—		
3. COL	-0.31	0.95 1.00		_	_		
4. LTC	-0.29	0.95	0·98	1.00			
5. 1st L	-0.24	0· <b>92</b>	0·97	0 <b>·96</b>	1.00		
6. 2nd L	-0.48	0.92	0.93	0· <b>94</b>	0.91	1.00	_
7. T-1	0.48	-0.74	-0.75	-0.77	-0.75	-0.80	1.00
8. T-2	0.41	-0.65	-0.67	-0.68	-0.73	-0.73	0.72
9. TND134	0.15	0·19	0.15	0.19	0.23	0.18	-0.18
10. CL	-0.19	0.00	0.04	0.05	-0.05	0.02	-0.07
11. FLL	0.09	0.17	0.12	0.10	0.14	0.10	-0.04
12. FLW	-0·19	0.21	0.22	0.23	0.22	0.31	-0.24
13. SpNo	0.02	0.18	0.13	0.13	0.14	0.11	-0.20
	SWt	RAD	COL	LTC	1st L	2nd L	T-1
1. SWt					_		
2. RAD						_	_
3. COL		_			_		
4. LTC	_	_			—	_	
5. 1st L		_			_	_	
6. 2nd L	_					—	
7. T-1	_					-	
8. T-2	1.00			_	—	_	
9. TND134	-0.26	1.00				—	
10. CL	-0.04	-0.4	16	1.00		_	
11. FLL	-0.16	-0.0	)1	0.12 1.00			
12. FLW	-0.36	-0.1	13	0.26	0.54	1.00	
13. SpNo	-0.08	0.6	54	-0.35	0·19	0.14	1.00
-	T-2	TND	134	CL	FLL	FLW	SpNo

# TABLE 4 Genotypic correlation matrix for characters measured in Experiment III (n = 376)

The correlation coefficients for characters within the different character groups described above are highly significant and generally homogeneous (e.g., RAD, COL, LTC, 1st L, 2nd L). Correlation coefficients for characters which are not part of the same character group are either nonsignificant or opposite in sign (e.g., Tiller character characters vs. RAD, COL, LTC, 1st L, 2nd L). A more detailed treatment of the character correlations is not in keeping with the rest of this paper.

# 4. DISCUSSION

The experimental design used in these analyses was not very efficient since the family variance was estimated from one population only. Had a nested design of several families within each population being tested, more precise estimation of the variance components would have been possible. Nevertheless, the present analyses clearly indicate that:

(a) Ample genetic variation exists in populations and families of wild *Hordeum murinum* for a number of characters.

(b) The major contributing source of genetic variance differed with character.

For reproductive characters, most of the variance was found between populations. A large percentage of the total variance was contributed by families within populations in the developmental characters. Generally, the population component contributed the largest portion of the total variance.

(c) There was evidence for genetic correlation between characters, possibly suggesting that groups of characters might respond or may have responded to selection as a single unit.

It would be inappropriate to draw inferences from the results about the role of the phenotypic variation and selective forces in the habitats from which the plants were collected. However, the information gained from these types of experiments can be useful because such experiments create a common ground in which the relative genotypically or environmentally induced responses of plants obtained from diverse sources can be judged. In the experiments, the environments provided by the four blocks are micro-environments and are also not under the control of the investigator. Valuable information on the effects of environment on the range of character expression could be obtained by increasing the number of test environments. However, this study can make a contribution to the evaluation of gene resource material.

There is a growing realisation that little genetic variation upon which artificial selection can be practised remains in many cultivated plants. To counteract this situation, wild and weedy relatives and primitive varieties are collected and preserved to capture and make available to breeders as much genetic variation as possible. The material, however, must be evaluated for two reasons: (a) to determine the amount of genetic variation in agronomically useful characters in material already collected (*e.g.* IBPGR, 1978); (b) to establish the distribution pattern of the genetic variation within and between populations so that statistical recommendations for optimum sample sizes and number of collection sites can be made for maximal capture of genetic variation in future collections. Evaluation of resource material is an applied version of a problem well known to population geneticists; that is, how much genetic variation exists in natural populations and how is it distributed.

Of (a) and (b), (b) is the most difficult problem to address because the collection strategies recommended vary with the type of characters used to study the distribution of genetic variation in populations (Allard, 1970; Bennett, 1970; Qualset, 1975; Marshall and Brown, 1975, 1981; Tolbert et al., 1980; Witcombe and Gilani, 1979; Bogyo et al., 1980; Brown and Munday, 1982). The results of this paper suggest that collection strategies will vary with the character of interest. If the aim is to capture variation for reproductive characters (e.g., yield, flowering time) then clearly the best strategy is to collect material from different populations over a wide eco-geographic range because the genetic variance observed within a population for these characters is low. Alternatively, if the interest lies in characters such as germination time or rate of seedling development, then ample variation may be found within a population. Overall, the between population component of variance was large for all characters used in this study including the seedling developmental characters (tables 2 and 3). Thus if a general recommendation can be made for capture of maximum

genetic variation from populations of *Hordeum murinum*, the best strategy would maximise the number of collection sites in as many distinct ego-geographic areas as possible.

Ultimately, the answer to (b) requires a reliable guide to genetic structure of populations. We do not yet have the ability to reconstruct the overall pattern of genetic variation in a population from the patterns indicated by a small sample of characters. Thus to assess variation in arbitrarily chosen or easy-to-assess sets of characters, with the hope that the variation uncovered is typical of the genome as a whole, is little more than pious hope (see Giles, in press). The first step in examining a wild population for genetic variation is to have a clear idea of what characters should be examined; this being determined by the use to which the variation will be put. It is only through the use of suitable experimental designs that a reasonably true estimate of genetic variation in wild-collected material can be made. The methodology of biometrical genetics provides an excellent framework for further studies of the genetic variation in natural populations. A great deal of valuable information could be obtained by exploiting the potential of this methodology.

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