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Combined oral contraceptive use before the first birth and epithelial ovarian cancer risk

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Background: Combined oral contraceptive (COC) use reduces epithelial ovarian cancer (EOC) risk. However, little is known about risk with COC use before the first full-term pregnancy (FFTP).

Methods: This Canadian population-based case-control study (2001–2012) included 854 invasive cases/2139 controls aged \geq 40 years who were parous and had information on COC use. We estimated odds ratios (aORs) and 95% confidence intervals (CI) adjusted for study site, age, parity, breastfeeding, age at FFTP, familial breast/ovarian cancer, tubal ligation, and body mass.

Results: Among parous women, per year of COC use exclusively before the FFTP was associated with a 9% risk reduction (95% CI = 0.86-0.96). Results were similar for high-grade serous and endometrioid/clear cell EOC. In contrast, per year of use exclusively after the FFTP was not associated with risk (aOR = 0.98, 95% CI = 0.95-1.02).

Conclusions: Combined oral contraceptive use before the FFTP may provide a risk reduction that remains for many years, informing possible prevention strategies.

Combined oral contraceptive (COC) use is an established factor that consistently reduces the risk for epithelial ovarian cancer (EOC; Beral *et al*, 2008). Less is known about the association between EOC risk and COC use with respect to the timing of full-term births. Increasing parity reduces EOC risk (Hankinson and Danforth, 2006), but it is difficult to tease apart the independent

effects of COC use and parity. The total number of ovulatory years between menarche and menopause has been used, but this does not address the timing of COC use with respect to full-term births. Studies of breast cancer (Schlesselman, 1989; Romieu *et al*, 1990; Kahlenborn *et al*, 2006) and endometrial cancer (Cook *et al*, 2014) have reported a long-term effect with the use of COCs before the

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Table 1. Characteristics of invasive, epithelial ovariancancer cases and controls, parous women only, OVAL-BC,2002–2012								
	Cas N=		Controls, N=2139					
Characteristics	N	%	Ν	%				
Age (years)								
40–49	113	13	423	20				
50–59 60–69	272 278	32	773 670	36 31				
≥70	191	33 22	273	13				
Race								
White	711	83	1877	88				
Chinese/Japanese	35	4	57	3				
Other Asian	24	3	44	2				
Others Unknown	54 30	6 4	96 64	4 3				
Education	50		04	5				
High school or less	357	42	755	35				
Vocational school	216	25	573	27				
University	280	33	809	38				
Unknown	1	<1	2	<1				
$BMI\;(kgm^{-2})$								
<25 25–29.9	406 274	48	1015	47				
25–29.9 30–34.9	274 101	32 12	682 278	32 13				
≥ 35	73	9	161	8				
Unknown	0	0	3	<1				
Smoking								
Never	409	48	1046	49				
Current	95	11	156	7				
Former	350	41	937	44				
Family history breast and/or ovarian	673	79	1767	83				
Yes	163	19	328	03 15				
Unknown	18	2	44	2				
Menopausal status and HT								
Pre-menopausal	170	20	584	27				
Peri, post-menopausal								
No HT	399	47	933	44				
Oestrogen only Oestrogen plus progesterone only	131 101	15 12	252 252	12 12				
Other HT	53	6	114	5				
Unknown	0	0	4	<1				
сос								
No (never or <6 months)	452	39	506	20				
Yes	692	61	2007	80				
Yes, duration (years)								
<5	367	32	918	37				
5–10 ≥10	185 128	16 11	564 504	22 20				
Unknown	120	1	21	1				
	12	1	21	1				
Parity	389	46	1080	50				
2	288	34	736	34				
≥3	177	21	323	15				
Age at FFTP (years)	· · · · · · · · · · · · · · · · · · ·							
≤24	561	65	1417	66				
25–29	91	11	299	14				
≥30 Unknown	196 6	23 1	402 21	19 1				
Ever breastfed	0	I	21	I				
No	229	27	424	20				
Yes	625	73	424 1714	80				
Unknown	0	0	1	<1				

Table 1. (Continued)										
	Cas N=	ies, 854	Controls, N = 2139							
Characteristics	N	%	Ν	%						
Duration breastfeeding (months)										
Never <10 ≥10 Unknown	229 386 237 2	27 45 28 <1	424 872 838 5	20 41 39 <1						
Hysterectomy										
No Yes Unknown	631 221 2	74 26 <1	1683 454 2	79 21 <1						
Tubal ligation										
No Yes Unknown	582 272 0	68 32 0	1335 802 2	62 37 <1						
Abbreviations: $BMI = body$ mass index; $COC = combined$ oral contraceptives; $FFTP = first$ full-term pregnancy; $HT = hormone$ therapy.										

first full-term pregnancy (FFTP) among parous women. We therefore investigated the EOC risk associated with COC use, focusing on COC use before the FFTP.

MATERIALS AND METHODS

This Canadian population-based case-control study has been previously described (Cook *et al*, 2016) including ethics approvals (Conjoint Health Research Ethics Board, Calgary, Alberta (AB) and Research Ethics Board, British Columbia (BC) Cancer Agency, Vancouver, BC) and written informed consent. Briefly, cases were identified from the population-based BC and AB cancer registries who were: age 20–79 years (40–79 in AB); diagnosed with first primary, incident, histologically confirmed EOC (invasive EOC in AB); and able to complete study in English. A total of 1505 cases (60% of 2522 eligible) completed the study. Eligible controls identified from provincial health rosters and a mammography screening program (Eheman *et al*, 2014) were: aged 20–79 years (40–79 in AB); able to complete study in English; and, had at least one ovary. A total of 2564 (53% of 4838 eligible) completed the study.

Risk factor information was ascertained through the diagnosis date (month/year) for cases and an assigned reference date (month/year) for controls based on an age-frequency match with cases. Respondents completed a self-administered questionnaire (BC before 2005) or a telephone interview (AB and BC after 2005). In additional to demographic, lifestyle, and medical/reproductive factors, women provided information on COC use, including dates or ages of use. Specific COC names were not ascertained. Histotypes were determined by re-review of haematoxylin and eosin slides according to contemporary criteria (Köbel *et al*, 2014) for 979 women (85.6%).

The analysis was restricted to those ≥ 40 years of age at diagnosis/reference date (1144 invasive cases and 2513 controls). Combined oral contraceptive use was evaluated as: non-use (never or <0.5 years) *vs* ever use (≥ 0.5 years); continuous duration (years, ever users only) and, as categorical duration (non-use, <5 years, 5–10, ≥ 10 years, and unknown). We used logistic regression to estimate adjusted odds ratios (aORs) and 95% confidence intervals (CIs) in R software (R Development Team, 2015). All variables in Table 1 were evaluated as potential confounders. Final aORs included matching variables (Alberta, BC before 2005, BC

after 2005, and 40–49, 50–59, 60–69, \geq 70 years of age), parity (0, 1, 2, \geq 3 or 1, 2, \geq 3 when restricted to parous women), age at FTTP (\leq 24, 25–29, \geq 30 years), breastfeeding (never, ever), first degree family female breast or ovarian cancer (no, yes), tubal ligation (no, yes), and BMI (<25, 25–29.9, 30–34.9, \geq 35 kg m⁻²). Other variables did not alter the estimated ORs by more than 10%. Histotype-specific analyses were restricted to high-grade serous and combined endometrioid/clear cell, due to few cases of other histotypes. Because COC use exclusively before and exclusively after the FFTP were mutually exlcusive, they were modelled simultaneously, allowing direct comparisons of the two risk estimates using contrasts (Montgomery, 2012).

RESULTS

Characteristics of parous cases and controls are described in Table 1. Combined oral contraceptive use was common among parous women, reported by 61% of cases and 80% of controls. With respect to the timing of COC use (Table 2), use of COCs before and after the FFTP (aOR = 0.45, 95% CI = 0.34-0.59; per year of use: aOR = 0.94, 95% CI = 0.91-0.98) as well as exclusive use before the FFTP (aOR = 0.56, 95% CI = 0.42-0.75; per year of use: aOR = 0.91, 95% CI = 0.86-0.96) was associated with a reduced risk for EOC. Similarly, both before and after the FFTP as well as use exclusively before the FFTP was associated with a

reduced risk for both high-grade serous (aOR = 0.50, 95% CI = 0.35–0.72 and aOR = 0.49, 95% CI = 0.35–0.70, respectively) and endometrioid/clear cell (aOR = 0.52, 95% CI = 0.29–0.92 and aOR = 0.47, 95% CI = 0.27–0.83, respectively) EOC. In contrast, COC use exclusively after the first birth was associated with a smaller reduction in EOC risk (aOR = 0.78, 95% CI = 0.61–1.01) that was suggestive but there was no association with increasing duration (per year of use aOR = 0.98, 95% CI = 0.95–1.02). When the risk estimate for exclusive use before the FFTP was compared (via contrasts) with the risk estimate for exclusive use after, the aORs were found to be significantly different (*P*-value, < 0.01) (Table 2).

When we stratified by age at FFTP, COC use before and after as well as exclusively before the FFTP was consistently associated with a reduction in EOC risk regardless of age at first birth, a consistency that was not seen with COC use exclusively after the FFTP (Figure 1), although some results were unstable. Similar results were noted when stratified by parity, although risk estimates were more similar for parity ≥ 3 (Figure 1).

The association of COC use and EOC risk for our entire study population (both parous and non-parous women combined) was consistent with the reported literature (Supplementary Tables 1–4). Any COC use was associated with a reduction in risk (aOR = 0.58, 95% CI = 0.49, 0.69). Among COC users, risk was most strongly reduced with longer durations of use overall, within more recent time since last use, and for younger ages at first use.

Table 2. Risk for epithelial ovarian cancer among parous women associated with exclusive use of combination oral contraceptives (COCs) before and after the first full-term pregnancy (FFTP), overall and by histotype

			r				r							
			Α	ll epi	ithelial	cancer	Histotype-specific							
	Controls N=1574						Serous (high grade)				Endometrioid/clear cell			
			Cases N=720		OR ^a 95	95% CI	Cases N=375		ORª	95% CI	Cases N=113		ORª	95% CI
	N	%	N	%			N	%			N	%		
COC use														
No	427	27	323	45	1.00	referent	179	48	1.00	referent	46	41	1.00	referent
Yes, exclusive use														
Before and after FFTP Before FFTP After FFTP	535 703 444	25 45 28	119 186 211	14 26 29	0.45 0.56 ^b 0.78 ^b	(0.34, 0.59) (0.42, 0.75) (0.61, 1.01)		15 23 30	0.50 0.49 ^b 0.76 ^b	(0.35, 0.72) (0.35, 0.70) (0.56, 1.04)	28 38 29	20 34 26	0.52 0.47 ^b 0.73 ^b	(0.29, 0.92) (0.27, 0.83) (0.41, 1.30)
Yes, duration of use (ye	ears)	1						1				1		<u> </u>
Before and after FFTP <5 5-<10 ≥10 Unknown per year of use ^c	153 183 197 2	7 9 9 <1	46 36 37 0	5 4 4 0	0.62 0.39 0.36 0.94	(0.42, 0.92) (0.26, 0.59) (0.24, 0.54) (0.91, 0.98)	28 22 14 0	6 5 3 0	0.75 0.46 0.33 0.95	(0.46, 1.20) (0.27, 0.76) (0.17, 0.59) (0.90, 1.00)	11 7 10 0	8 5 7 0	0.79 0.29 0.56 0.91	(0.35, 1.64) (0.11, 0.67) (0.24, 1.18) (0.83, 0.99)
Exclusively before FFTP <5 5-<10 ≥10 Unknown per year of use ^c	397 193 112 1	25 12 7 <1	121 51 13 1	17 7 2 <1	0.61 0.57 0.22 0.91	(0.45, 0.83) (0.37, 0.85) (0.11, 0.42) (0.86, 0.96)	52 27 5 1	14 7 1 <1	0.50 0.62 0.18 0.92	(0.34, 0.74) (0.37, 1.03) (0.06, 0.44) (0.86, 0.99)	26 7 5 0	23 6 4 0	0.56 0.33 0.35 0.92	(0.30, 1.04) (0.12, 0.81) (0.10, 0.99) (0.83, 1.01)
Exclusively after FFTP <5 5-<10 ≥10 Unknown per year of use ^c	270 107 62 5	17 7 4 <1	128 48 32 3	18 7 4 <1	0.84 0.74 0.76 0.98	(0.63, 1.12) (0.49, 1.10) (0.47, 1.22) (0.95, 1.02)	67 24 19 1	18 6 5 <1	0.80 0.67 0.82 0.98	(0.56, 1.15) (0.39, 1.11) (0.45, 1.43) (0.94, 1.04)	17 7 4 1	15 6 4 1	0.70 0.84 0.68 0.97	(0.36, 1.35) (0.31, 2.01) (0.18, 1.97) (0.88, 1.07)

Abbreviations: OR = odds ratio; 95% CI = 95% confidence interval.

^aORs adjusted for study site (Alberta, BC before 2005, BC after 2005), age (40–49, 50–59, 60–69, ≥70 years), parity (1,2, ≥3), age at FFTP (≤24, 25–29, ≥30 years), breastfeeding (never, ever), first degree female family history of breast or ovarian cancer (no, yes), tubal ligation (no, yes), and BMI (<25, 25–29.9, 30–34.9, ≥35 kg m⁻²).

^b*P*-value for difference in ORs, <0.01.

^cAmong COC users only.

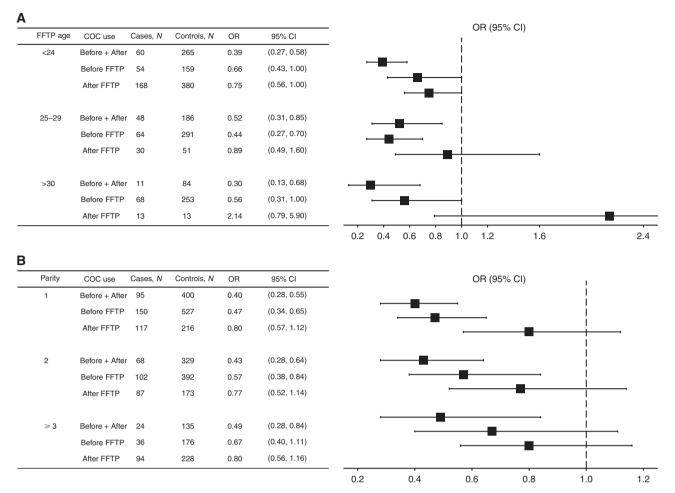


Figure 1. Combined oral contraceptive use with respect to the FFTP by age at first birth (**A**) and by number of births (parity) (**B**) among parous women. The aORs are adjusted for the following: study site (Alberta, BC before 2005, BC after 2005); age (40–49, 50–59, 60–69, \geq 70 years); parity in Panel **A** only (1, 2, \geq 3); age at FFTP in Panel **B** only (\leq 24, 25–29.9, 30–34.9, \geq 35 years); breastfeeding (never, ever); first degree female family history of breast or ovarian cancer (no, yes); tubal ligation (no, yes); BMI (<25, 25–29.9, 30–34.9, \geq 35 kg m⁻²).

CONCLUSION

When we assessed the timing of COC use exclusively before the FFTP among parous women, we found a strong reduction in risk (~40%), which was almost as strong as the ~50% risk reduction seen with COC before and after the FFTP. Even for fairly short-term COC use (<5 years) before the FFTP there was a significant and substantial reduction in risk years later in parous women. This result is surprising, given that these women all experienced the reduction in risk associated with being parous, and given that the literature (Beral *et al*, 2008) and our own results for parous and non-parous women indicating that last use of COCs in the more distant past is associated with weaker reductions in risk. In contrast, the effect of such use after the FTTP was of lesser magnitude, despite the assumption that the cessation of ovulation in these women should have equivalent effects regardless of the timing of COCs.

Consistent with our findings, other studies have reported that any use of COCs before age 20 years (Ness *et al*, 2000; Kumle *et al*, 2004; Beral *et al*, 2008; Lurie *et al*, 2008) or 25 years (Bosetti *et al*, 2002) is associated with a reduced EOC risk of 29–50% many years later. Ours is the first study to assess COC use exclusively before and after the FFTP to evaluate the timing of COC use with pregnancy.

Although the more immediate effects of COC use on biological end points such as hormone levels, gene expression, and ovulation are well documented, the long-term effects on EOC risk are largely attributed to fewer ovulations during reproductive life (Fathalla, 1971), with the assumption that the timing of ovulation reduction does not matter. Our results could be due, in part, to fewer ovulations because of COC use, but it is not clear why use before the FFTP would have such a strong, lasting impact on EOC risk. In breast cancer, the elevated risk noted with COC use before the FFTP has been hypothesised to be related to the carcinogenic susceptibility of undifferentiated breast tissue at this time (Romieu et al, 1990), and in endometrial cancer the reduction in risk with early COC use is unknown but may be related to a lasting effect on hormone levels that reduce cellular proliferation (Chan et al, 2007). Whether such mechanisms are also applicable to a long-lasting reduction in EOC risk is not clear. Regardless of the tissue of origin for EOC (fallopian tube, endometrium, ovary, etc.), our results suggest that the timing of ovulation reduction is important, and that there may be other long-term mechanisms for an EOC risk reduction beyond ovulation that manifest before the FFTP.

Study strengths include the population-based design; large sample size; restriction to first primary, histologically confirmed invasive EOC; detailed information on parity; assessment of contemporary histotypes (Köbel *et al*, 2014); high prevalence of COC use; and, restriction to parous women with adjustment for parity, thus minimising confounding by parity. Limitations include: no COC name/dosage information; cases recalling past COC use more fully than controls (but that would bias risk estimates to the null value); relatively low response percentage among the control women; and, possible residual confounding. In addition, COC use in this study represents formulations of COC available in the past, and current formulations may not have the same long-term effects.

In summary, the significant reduction in EOC risk observed with COC use before the FFTP among parous women is a novel and requires replication. Despite the consistently reported risk reduction in EOC with COCs, questions remain about the timing of use and the underlying biological mechanisms of long-term effects to guide future EOC risk prediction (Pearce *et al*, 2015) and directed chemoprevention strategies for high-risk women (Walker *et al*, 2015).

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CONFLICT OF INTEREST

The authors declare no conflict of interest.

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