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# Oestrogen receptors $\beta$ 1 and $\beta$ cx have divergent roles in breast cancer survival and lymph node metastasis

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**Background:** The expression of oestrogen receptor (ER)  $\alpha$  characterises a subset of breast cancers associated with good response to endocrine therapy. However, the clinical significance of the second ER, ER $\beta$ 1, and its splice variant ER $\beta$ cx is still unclear.

**Methods:** We here report an assessment of  $ER\alpha$ ,  $ER\beta1$  and  $ER\betacx$  by immunohistochemistry using quantitative digital image analysis of 340 primary tumours and corresponding sentinel lymph nodes.

**Results:** No differences were seen in ER levels in primary tumours vs lymph node metastases. ER $\beta$ 1 and ER $\beta$ cx were equally distributed among age groups and tumour histological grades. Loss of ER $\beta$ 1 in the primary tumour was strongly associated with poor survival. Its prognostic impact was particularly evident in young patients and in high-grade tumours. The worst outcome was seen in the tumours lacking both ER $\alpha$  and ER $\beta$ 1. ER $\beta$ cx expression in the primary tumour correlated with a higher risk of lymph node metastasis, and with poor survival when expressed in sentinel node lymphocytes.

**Conclusions:** Our study reveals highly significant although antagonising roles of  $ER\beta1$  and  $ER\betacx$  in breast cancer. Consequently, we suggest that the histopathological assessment of  $ER\beta1$  is of value as a prognostic and potentially predictive biomarker.

Oestrogen receptor (ER)-mediated signalling has a fundamental role in breast cancer biology. In the majority of breast cancers, generally classified as the luminal subtypes, ER alpha (ER $\alpha$ ), one of the members of the ER family, works as a central hub governing tumour cell proliferation and tumour progression (Sorlie *et al*, 2006; Ross-Innes *et al*, 2012).

Clinical trials have confirmed the predictive role of ER $\alpha$  as a biomarker in response to adjuvant endocrine therapy and thereby its association with favourable outcomes. Hence, tamoxifen has been shown to reduce the risk of any recurrence by 39% over a 10-year treatment period (Early Breast Cancer Trialists' Collaborative Group (EBCTCG) *et al*, 2011); however, the prognostic value of

ER $\alpha$  seems to decrease after 5 years (Bentzon *et al*, 2008). This indicates that our understanding of ER signalling in breast cancer is still very limited. Since the discovery of a second ER, ER beta (ER $\beta$ ) in 1996 (Kuiper *et al*, 1996), the general focus has been to decipher its role in biology and its corresponding clinical importance. ER $\alpha$  and ER $\beta$  are located on different chromosomes but show considerable homology. Both ERs bind tamoxifen and their main ligand, oestradiol, with similar affinity, but the differences within the ligand-binding pockets are significant enough to permit the synthesis of ER $\alpha$ - and ER $\beta$ -selective ligands (Kuiper *et al*, 1998). In terms of gene regulation, experiments performed on breast cancer cell lines in the presence of oestradiol

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indicate partly overlapping transcriptomes induced by the two ER subtypes (Chang et al, 2006). However, experimental studies in vitro and during tumour xenograft growth have shown opposing roles of the two receptors in terms of proliferation (Ström et al, 2004; Hartman et al, 2006). Contributing to the complex picture of  $ER\beta$ signalling is the presence of several splice variants. The full-length  $\text{ER}\beta$  is known as  $\text{ER}\beta$ 1 whereas the best characterised splice variant in breast cancer is ER $\beta$ cx (also known as ER $\beta$ 2; Ogawa *et al*, 1998). We now know that  $ER\beta$  is expressed in the epithelium and stroma of normal as well as malignant mammary gland and mediates oestrogen response. Immunohistochemistry (IHC) studies indicate an association of ER $\beta$  expression with ER $\alpha$ -positive tumours and/or progesterone receptor (PR) status (Omoto et al, 2002; Fuqua et al, 2003). However, these studies were in part performed with pan-specific antibodies detecting all ER $\beta$  splice variants, including ER $\beta$ cx. Later on, the roles of different  $ER\beta$  splice variants have been dissected by the use of monoclonal C-terminus-targeted  $\mathrm{ER}\beta$  antibodies. It now appears that  $ER\beta$  expression is associated with tamoxifen response, particularly within ERa-negative tumours (Gruvberger-Saal et al, 2007; Honma et al, 2008). Other researchers have been unable to confirm this association and instead report an association between  $ER\beta cx$  and  $ER\alpha$  expression and a strong correlation of cytoplasmic  $ER\beta cx$  with poor survival (Shaaban *et al*, 2008). In a large populationbased study, Marotti et al (2009) could confirm a positive correlation of ER $\beta$ 1 expression with ER $\alpha$ , but not with survival; similar results were described by Borgquist et al (2008). One study has even shown an association of ER $\beta$ 1 with poor prognosis, but only in lymph nodepositive patients (Novelli et al, 2008). The majority of studies on larger patient populations have been performed by IHC on tissue microarrays (TMAs), a suboptimal platform for investigating heterogeneously expressed proteins. In the present study, we characterised ER $\beta$ 1 and ER $\beta$ cx and re-evaluated ER $\alpha$  expression by IHC of whole tumour sections from 340 patients with archived breast tumours and corresponding sentinel lymph nodes (SLNs).

# MATERIALS AND METHODS

Study population and follow-up. The study cohort was identified from the patient registry at the Department of Pathology, Karolinska University Hospital, Stockholm, Sweden. Only patients who had undergone sentinel node biopsy (SNB) from 2001 to 2006 were included. All patients had a preoperative diagnosis of breast cancer and a clinically negative axilla. A subset of the cohort originated from a prospective study evaluating the oncological safety of SNB, the results of which have been published elsewhere (Andersson et al, 2012). The surgery was performed at either Karolinska University Hospital, South General Hospital or at Sofiahemmet Hospital, all in the Stockholm area. Routinely, patients were followed up annually for 5-10 years, and then reintroduced into the national mammography-screening programme. All recurrences within the Stockholm area are routinely referred back to the Department of Oncology at Karolinska University Hospital where the study has been performed. Patients who had moved away from the Stockholm County during followup were censored at the time of their deregistration.

Clinicopathological parameters and data on received adjuvant therapy were extracted from patient medical records. As some of the routine assessments such as PR and proliferation markers changed during the period, cut-offs employed at the time of diagnosis were used. Oestrogen receptor  $\alpha$  was because if its central role in this study, however, re-evaluated throughout using IHC. Depending on tumour characteristics and stage of the disease, patients were treated with radiotherapy, chemotherapy and/or endocrine therapy. The occurrence of local, regional or distant relapse, death, breast cancer-specific death and the dates of last Specimen selection and immunohistochemistry. From each patient, one formalin-fixed paraffin tissue block of the primary tumour and one block of the matching SLN were identified. The sections were cut at  $4-\mu m$  thickness and mounted. Immunohistochemistry was performed either on an Autostainer (Dako, Glostrup, Denmark;  $ER\beta1$ ) or on IntelliPath FLX (BioCare medical, Concord, CA, USA; ER $\alpha$  and ER $\beta$ cx) according to the protocols and reagents provided by the manufacturers (BioCare medical and Dako), together with negative and positive controls. Heat-induced antigen retrieval in high pH solution was performed using a PT-linker (Dako) at 97°C for 20 min. The slides were incubated for 30 min at room temperature with the primary monoclonal antibodies. Anti-ER $\beta$ 1 (clone PPG5/10; Dako) 1:50, anti-ERßcx (clone 57/3; AbD Serotec, Oxford, UK) 1:200 and anti-ERa (clone NCL-L-ER-6F11; Novocastra, Wetzlar, Germany) 1:200 antibodies were used. 3,3'-diaminobenzidine (DAB) was used to detect primary antibody binding and haematoxylin as counterstaining.

Digitalisation of slides and image analysis. To set appropriate intensity cut-offs for the digital scoring, a subset of tumours and SLNs were first scored manually by two independent researchers for ER $\alpha$ , ER $\beta$ 1 (GR and JH) and ER $\beta$ cx (GR and GMK) in the following compartments: primary tumour, SLN metastasis, if present, and lymphocytes residing in the SLN. The Allred scoring system was used for the manual scoring. The method has been adapted to evaluate and quantify different proteins using IHC (Fuqua et al, 2003; Rosin et al, 2012). When the two researchers did not agree on the score, they re-evaluated the section together until an agreement could be reached. All slides were digitally scanned using a Pannoramic MIDI or Pannoramic 250 Flash (3DHistech, Budapest, Hungary). The Pannoramic viewer 1.15.2 software (3DHistech) was used for viewing the scanned images together with the built-in image analysis application, Nuclear-Quant (3DHistech), which has been validated and shown to be reproducible in the detection of ER in breast cancer (Krecsák et al, 2011). The software quantifies both the frequency and intensity of nuclear staining with DAB. Detection thresholds for the size, circularity and differences in contrast of the nuclei can be adjusted to distinguish cancer cells from other cell types such as fibroblasts or lymphocytes. The frequency score was based on the Allred system (Harvey et al, 1999), where a score between 0 and 5 is given depending on the frequency of positive cells (0% = 0, < 1% = 1, 1%to <10% = 2, 10% to <33% = 3, 33% to <66% = 4 and  $\geq$  66% = 5). The cut-off for the intensity (0 to 255) can also be changed to adjust the threshold to divide the nuclei stained into four different scoring categories ('no positive nuclei' = 0, 'low intensity' = 1, 'moderate intensity' = 2 and 'high intensity' = 3). The cut-offs were adjusted in a stepwise manner until the manual scoring and digital scoring matched on most occasions. The specific intensity threshold settings were for ER $\beta$ 1: 0,  $\geq$ 175; 1,  $< 175 \text{ and } \ge 125; 2, < 125 \text{ and } \ge 80; 3, < 80; \text{ for ER}\alpha: 0, \ge 170; 1,$ <170 and  $\geq$ 120; 2, <120 and  $\geq$ 70; 3, <70 and for ER $\beta$ cx: 0,  $\geq$ 177; 1, <177 and  $\geq$ 130; 2, <130 and  $\geq$ 70; 3, <70. The frequency and intensity scores were then combined into a final score ranging from 0 to 8 (excluding 1). For all sections, three representative areas of invasive tumour were annotated and then subjected to image analysis. Only nuclear ER staining was analysed. When possible, each of the three areas contained at least 1000 nuclei. An average score of 4 or higher was considered as positive ER expression, corresponding to 10% positive cells with weak intensity, a cut-off that has been used in several reports (Mann et al, 2001; Honma et al, 2008).

**Statistics.** For descriptive statistics, continuous variables are presented as median (range), while categorical variables are presented as numbers of cases and corresponding percentages. The Pearson's Chi-square test was used to test the hypothesis of equal distribution of ER expression in categorical variables (positive *vs* negative). For the evaluation of ER status in node-negative *vs* node-positive patients, the Pearson's Chi-square test was supplemented by additional logistic regression in order to estimate the odds ratio for the presence of metastasis in different ER status groups. For the testing of distribution of ER expression (positive *vs* negative) in paired samples, such as primary tumours and their corresponding SLNs, the McNemar test was applied for categorical variables.

Estimation of 10-year survival rates was performed using Kaplan-Meier survival analysis. For the analysis of overall survival, follow-up time was calculated from the date of primary surgery until death of any cause or the date of medical record review, as medical records are directly linked to the national death registry. For the analysis of breast cancer-specific survival, follow-up time was from the date of primary surgery until death caused by breast cancer or the last recorded follow-up visit as documented in medical records at the department of oncology. All patients who died with metastasised breast cancer were considered to have died of the disease. For the analysis of disease-free survival, follow-up time was recorded from the date of primary surgery until the date of any relapse or until the last recorded follow-up visit. The influence of ER status on survival was tested using the log-rank test within the Kaplan-Meier model. As endocrine treatment was assumed to strongly affect survival analysis regarding ER expression, analyses were also adjusted for any endocrine treatment by adding this information as strata into the Kaplan-Meier model. For the comparative analysis of the impact of known risk factors on survival rates and their comparison with the impact of ER receptor status, both uni- and multivariable Cox proportional hazard analyses were performed and results are presented as hazard ratios (HRs) with their 95% confidence intervals (CIs). All statistical computations were performed using IBM SPSS Statistics, Version 21. A *P*-value of  $\leq 0.05$  was considered statistically significant.

## RESULTS

Overall, 340 breast cancer patients operated between January 2001 and December 2006 were included. Of these, 322 tumours were stained and analysed for ER $\alpha$ , 316 for ER $\beta$ 1 and 315 for ER $\beta$ cx (see Supplementary Figure 1 for representative IHC stainings). Only nuclear staining of ER $\alpha$ , ER $\beta$ 1 and ER $\beta$ cx was analysed. Overall, 11 patients had no available tumour results of all three ERs, however, they were still included in the cohort since they had ER data from SLNs. Three patients had missing tumour data on ER $\alpha$  and ER $\beta$ 1, three patients on ER $\alpha$  and ER $\beta$ cx, and two patients on ER $\beta$ 1 and ER $\beta$ cx. Patient and tumour characteristics are presented in Table 1. Median follow-up for disease-free and breast cancer-specific survival was 81 months (range 0-148). Median follow-up for overall survival was 115 months (range 2-152). Thirty-six patients had died during the follow-up period, 16 of whom had died of breast cancer. Recurrences were found in 35 patients, sometimes multiple. In total, 10 local, 10 regional and 22 distant relapses were recorded. Ten patients developed contralateral breast cancer during the follow-up period, which was not considered a relapse. The Kaplan-Meier estimate of overall survival for the entire cohort was 83.9%, breast cancer-specific survival 92.1% and disease-free survival 81.6%. Adjuvant endocrine treatment had been given to 268 patients (78.8%) and chemotherapy to 132 patients (39.8%).

Oestrogen receptor status in different age groups and tumour histological grades. A cut-off of 10% was used to distinguish ERpositive from ER-negative tumours. As shown in Figure 1A, the percentage of ERa-positive tumours increased significantly with lower Elston-Ellis histological grade (P < 0.001). ER $\beta$ cx positivity followed a similar pattern but without reaching statistical significance (P = 0.23). ER $\beta$ 1, however, had an equal distribution in all tumour grades (P = 0.771). Patients were divided into four age groups: <40 years (N=12), 40-54 years (N=104), 55-64 years (N=141) and  $\geq 65$  years (N=81), to study age-related changes in ER positivity. As shown in Figure 1B, there was a significant increase in ERa-positive tumours with increasing age (P=0.011). There was a similar but non-significant trend for ER $\beta$ cx positivity (P=0.062) with increased age category. Again  $ER\beta 1$  showed a different expression pattern, where the positive tumours were equally distributed among all age groups (P = 0.22).

**Oestrogen receptor status in primary tumour and synchronous lymph node metastasis.** The ER status of primary tumours did not differ significantly compared to their paired synchronous SLN metastases for any of the ERs: ER $\alpha$  (P = 0.33), ER $\beta$ 1 (P = 1.0) and ER $\beta$ cx (P = 0.13). However, the proportion of ER $\beta$ cx-positive tumours was higher in node-positive than node-negative patients (P = 0.021; Table 2). This was confirmed by logistic regression, which resulted in an odds ratio of 2.54 (95% CI 1.15–5.61) for synchronous SLN metastasis in ER $\beta$ cx-positive compared to negative cases. No difference in risk of SLN metastasis was seen in patients with ER $\alpha$ - or ER $\beta$ 1-positive primary tumours.

 $ER\beta 1$  expression in the primary tumour strongly affects survival. Ten-year overall survival was significantly lower in women with ER $\beta$ 1-negative tumours (79.7% vs 91.1%, log rank P = 0.009, Figure 2A) with a HR of 2.48 (95% CI 1.23–5.01). The corresponding figures for patients receiving adjuvant endocrine treatment were 88.7% vs 92.0% and for untreated patients 48.6% vs 92.4% (endocrine treatment-adjusted log rank P = 0.005). No differences were seen when controlling for different types of endocrine treatment, such as tamoxifen or aromatase inhibitors. When primary tumours were stratified according to histological grades, 10-year overall survival was significantly worse for women with  $ER\beta$ 1-negative tumours of high grade compared to patients with ER $\beta$ 1-positive high-grade tumours (60% vs 87.8%, log rank P = 0.008). There was no significant survival difference with regard to  $\text{ER}\beta$ 1 expression within grade 1 and grade 2 tumours. The prognostic potential of  $ER\beta 1$  was also compared within the four age groups described above. Although the youngest age group included too few patients for subgroup analysis, 10-year overall survival was significantly lower in women with  $ER\beta$ 1-negative tumours aged 40–54 and 55–64 years than in their ER $\beta$ 1-positive counterparts (80% vs 98.7%, log rank P=0.001 and 84.7% vs 91.5%, log rank P = 0.042). No significant survival difference in regard to ER $\beta$ 1 status was seen in elderly women (>64 years).

We also examined 10-year breast cancer-specific survival. This was significantly lower in women with  $ER\beta$ 1-negative tumours (85.4% vs 97.7%, log rank P = 0.011; Figure 2B) with a HR of 3.44 (95% CI 1.24-9.49). The corresponding figures for patients given adjuvant endocrine treatment were 92.6% vs 93.9% and for untreated patients 59.9% vs 93.4% (endocrine treatment-adjusted log rank P = 0.020). Also for breast cancer-specific survival, no differences were seen when controlling for different types of endocrine treatment. When stratifying by grade, 10-year breast cancer-specific survival was significantly lower in high-grade  $ER\beta$ 1-negative tumours compared with high-grade  $ER\beta$ 1-positive tumours (58.2% vs 88.2%, log rank P = 0.036). Again, no difference was seen within lower histological grades. Similarly to overall survival, a lower 10-year breast cancer-specific survival was seen in women aged 40-54 years with  $ER\beta$ 1-negative tumours than in those with ER $\beta$ 1-positive tumours (76.4% vs 98.1%, log rank

Table 1. Patient and tumour characteristics of 340 breast cancer samples divided according to  $ER\beta1$  status and for all patients

| Clinicopathological<br>variables  | All<br>patients          | ERβ1<br>positive         | ERβ1<br>negative       |
|-----------------------------------|--------------------------|--------------------------|------------------------|
| Characteristics                   |                          |                          |                        |
| Number                            | 340                      | 241 (76.3) <sup>a</sup>  | 75 (23.7) <sup>a</sup> |
| Age (years) <sup>b</sup>          | 58 (23–86)               | 58 (23–86)               | 58 (32–82)             |
| Tumour diameter (mm) <sup>b</sup> | 15 (1–73)                | 15 (1–73)                | 15 (4–50)              |
| T stage (mm)                      |                          |                          |                        |
| T1 (≼20)                          | 224 (66.5)               | 155 (65.1)               | 54 (72.0)              |
| T2 (21–50)                        | 103 (30.6)               | 75 (31.5)                | 19 (25.3)              |
| T3 (>50)                          | 10 (3.0)                 | 8 (3.4)                  | 2 (2.7)                |
| Missing                           | 3                        | 3                        | 0                      |
| Histological type                 | 1                        | 1                        |                        |
| Ductal                            | 264 (78.6)               | 181 (76.4)               | 62 (82.6)              |
| Lobular<br>Mius d                 | 57 (16.9)                | 45 (19.0)                | 9 (12.0)               |
| Mixed<br>Other                    | 9 (2.7)<br>6 (1.8)       | 6 (2.5)<br>5 (2.1)       | 3 (4.0)                |
| Other<br>Missing                  | 6 (1.8)                  | 5 (2.1)<br>4             | 1 (1.3)<br>0           |
|                                   | 4                        | +                        | 0                      |
|                                   | 250 (20 1)               | 107 (00.1)               | F0 (74 0)              |
| Positive                          | 258 (80.1)<br>64 (19.1)  | 197 (82.4)<br>42 (17.6)  | 52 (71.2)<br>21 (28.8) |
| Negative<br>Missing               | 18                       | 42 (17.0)                | 21 (20.0)              |
| -                                 | 10                       | 5                        | 2                      |
| $ER\betacx$ staus                 |                          |                          |                        |
| Positive                          | 279 (88.6)               | 210 (90.1)               | 58 (81.7)              |
| Negative<br>Missing               | 36 (11.4)<br>25          | 23 (9.9)<br>8            | 13 (18.3)<br>4         |
| PR status                         | 23                       | 0                        | +                      |
|                                   |                          |                          |                        |
| Positive                          | 249 (73.9)               | 182 (76.2)               | 49 (66.2)              |
| Negative<br>Missing               | 88 (26.1)<br>3           | 57 (23.8)<br>3           | 25 (33.8)<br>1         |
|                                   |                          | 5                        | 1                      |
| Elston histological grading       | 1                        | T                        | Γ                      |
| 1                                 | 89 (26.4)                | 60 (25.2)                | 22 (29.3)              |
| 2                                 | 171 (50.7)               | 126 (52.9)               | 37 (49.3)              |
| 3<br>Missing                      | 77 (22.9)<br>3           | 52 (21.8)<br>3           | 16 (21.3)<br>0         |
| Proliferation                     |                          | 5                        |                        |
|                                   | 001 (11)                 | 450 117 1                |                        |
| Low                               | 221 (66.4)               | 159 (67.4)               | 47 (64.4)              |
| High<br>Missing                   | 112 (33.6)<br>7          | 77 (32.6)<br>236         | 26 (35.6)<br>2         |
| Sentinel lymph node statu         |                          | 230                      | <u> </u>               |
|                                   | 1                        | 134 /55 /                | 16 (41 2)              |
| N0<br>N1                          | 195 (57.4)<br>145 (42.6) | 134 (55.6)<br>107 (44.4) | 46 (61.3)<br>29 (38.7) |
| Adjuvant treatment                | (12.0)                   |                          | 2, (33.7)              |
| -                                 | 132 /20 01               | 05 /20 /\                | 30 (40 0)              |
| Chemotherapy<br>Radiotherapy      | 132 (39.8)<br>250 (75.3) | 95 (39.4)<br>182 (77.4)  | 30 (40.0)<br>52 (71.2) |
| All endocrine treatment           | 268 (80.7)               | 195 (83.0)               | 52 (71.2)              |
| Oestrogen receptor antagonist     | 152 (44.7)               | 110 (45.6)               | 29 (38.7)              |
| Aromatase inhibitor               | 63 (18.5)                | 40 (16.6)                | 20 (26.7)              |
| Combined endocrine treatment      | 53 (15.6)                | 45 (18.6)                | 6 (8.0)                |
|                                   |                          |                          |                        |

| Table 1. (Continued)             |                              |                       |                         |
|----------------------------------|------------------------------|-----------------------|-------------------------|
| Clinicopathological<br>variables | All<br>patients              | ERβ1<br>positive      | ERβ1<br>negative        |
| ERα/PR status                    |                              |                       |                         |
| Positive/positive                | 207 (67.0)                   | 165 (69.6)            | 42 (58.3)               |
| Positive/negative                | 40 (12.9)                    | 31 (13.1)             | 9 (12.5)                |
| Negative/positive                | 21 (6.8)                     | 15 (6.3)              | 6 (8.3)                 |
| Negative/negative                | 41 (13.3)                    | 26 (11.0)             | 15 (20.8)               |
| Missing                          | 31                           | 4                     | 3                       |
| ERa/PR status <sup>c</sup>       | All<br>patients <sup>c</sup> | $ER\beta cx$ positive | $ER\betacx$<br>negative |
| Positive/positive                | 208 (67.5)                   | 186 (68.1)            | 22 (62.9)               |
| Positive/negative                | 39 (12.7)                    | 37 (13.6)             | 2 (5.7)                 |
| Negative/positive                | 22 (7.1)                     | 19 (7.0)              | 3 (8.6)                 |
| Negative/negative                | 39 (12.7)                    | 31 (11.4)             | 8 (22.9)                |
| Missing                          | 32                           | 6                     | 1                       |

Abbreviations: ER = oestrogen receptor; PR = progesterone receptor. Data were collected from patient medical records, except for ER $\alpha$ , ER $\beta$ 1 and ER $\beta$ cx, which were evaluated within this study using immunohistochemistry. All numbers are cases (%) if not stated otherwise. <sup>a</sup>Twenty-four patients had missing ER $\beta$ 1 classification.

<sup>b</sup>Median (range).

<sup>c</sup>Note that data presented below is on ER $\beta$ cx expression.

P = 0.001). There were no differences in breast cancer-specific survival within the two higher age groups (55–64 and >64 years).

ER $\alpha$ -positive tumours were associated with better breast cancerspecific survival (log rank P = 0.048, Figure 2C) but not overall survival (log rank P = 0.20). There was no significant association between ER $\beta$ cx status in the primary tumour and overall or breast cancer-specific survival (Figure 2D and E), and no significant patterns were observed when stratifying for histological grades and age groups. None of the analysed ERs affected disease-free survival.

All potential prognostic variables were tested using univariable Cox regression analysis. Significant univariable factors were entered into a multivariable Cox regression model (Supplementary Table 1). As there were few patients in the youngest age group, no 95% CIs could be calculated for this specific group. In the multivariable model, loss of ER $\beta$ 1 was associated with worse prognosis with a HR of 2.40 (95% CI 1.16–4.94). Also the loss of PR (HR 3.38, 95% CI 1.76–6.50), older age (>65 years) and advanced nodal stage (HR 3.77, 95% CI 1.54–9.19) remained independent of prognostic factors for overall survival. For breast cancer-specific survival, only ER $\beta$ 1 (HR = 3.38, 95% CI 1.09–10.45) and advanced nodal stage (HR = 11.64, 95% CI 2.97–45.63) remained independent prognostic factors.

ER marker combinations affect breast cancer-specific and overall survival. Four combinations of intratumoural ER $\alpha$  and  $\text{ER}\beta 1$  expression were created:  $\text{ER}\alpha + /\text{ER}\beta 1 + (N = 197)$ ,  $\text{ER}\alpha + /$  $ER\beta 1 - (N=52)$ ,  $ER\alpha - /ER\beta 1 + (N=42)$  and  $ER\alpha - /ER\beta 1 -$ (N=21). Differential 10-year overall survival differed significantly between these four groups with 91.4%, 85.4%, 90.0% and 75.6%, respectively (log rank P = 0.050; Table 3). The same pattern was seen for 10-year breast cancer-specific survival rates of 93.9%, 91.0%, 93.6% and 72.1%, respectively (log rank P = 0.009). Interestingly, these differences lost their significance when adjusting for any endocrine treatment, however, significance was retained when adjusted only for ER antagonist-containing postoperative therapy (e.g., tamoxifen; log rank P = 0.032 and 0.031). Corresponding groups were created for ER $\beta$ 1 and ER $\beta$ cx (N = 210 $(\text{ER}\beta 1 + /\text{ER}\beta\text{cx} +), \quad 23 \quad (\text{ER}\beta 1 + /\text{ER}\beta\text{cx} -), \quad 58 \quad (\text{ER}\beta 1 - /$  $ER\beta cx + )$  and 13  $(ER\beta 1 - /ER\beta cx - ))$ ; in these groups, 10-year overall survival was 92.0%, 95.5%, 78.1% and 100%, respectively

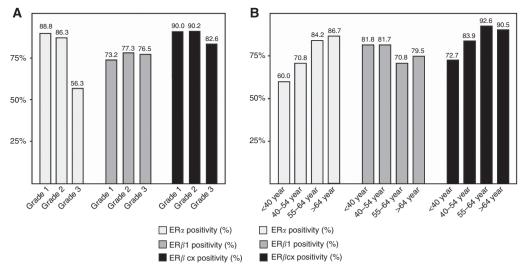


Figure 1. ER expression among different age groups and histological grades. (A) ER $\alpha$  positivity decreased with higher Elston–Ellis histological grade (white; P < 0.001, N = 80, 168, and 71). ER $\beta$ cx showed a similar trend, however not significant (black; P = 0.23, N = 80, 163, and 69). ER $\beta$ 1 was equally distributed among all three grades (grey; P = 0.771, N = 82, 163, and 68). (B) Patients were divided into four different age groups according to age at diagnosis, <40, 40–54, 55–64, and ≥65 years. ER $\alpha$  (white) positivity was lowest in the youngest age group and increased with age (P = 0.011, N = 10, 96, 139, and 75). A similar trend was seen with ER $\beta$ cx (black), however not reaching significance (P = 0.062, N = 11, 93, 135, and 73). ER $\beta$ 1 was equally distributed along all age groups (grey; P = 0.221, N = 11, 93, 137, and 74). Numbers above bars reflect percentage of positive tumours.

|          | Primary tumour<br>node negative | Primary tumour<br>node positive |  |
|----------|---------------------------------|---------------------------------|--|
| ERα      |                                 |                                 |  |
| Negative | 35 (19.3)                       | 29 (20.6)                       |  |
| Positive | 146 (80.7)                      | 112 (79.4)                      |  |
| P-value  | 0.7                             | 0.784                           |  |
| ERβ1     |                                 |                                 |  |
| Negative | 46 (25.6)                       | 29 (21.3)                       |  |
| Positive | 134 (74.4)                      | 107 (78.7)                      |  |
| P-value  | 0.3                             | 0.382                           |  |
| ERβcx    |                                 |                                 |  |
| Negative | 27 (15.2)                       | 9 (6.6)                         |  |
| Positive | 151 (84.8)                      | 128 (93.4)                      |  |
| P-value  | 0.0                             | 0.018*                          |  |

ER $\beta$ . SLN negative N=195, SLN positive N=145. Pearson's Chi-square test. \*P<0.05.

(log rank P = 0.011). Rates for 10-year breast cancer-specific survival were 95.9%, 92.3%, 80.0% and 100%, respectively (log rank P = 0.001). Adjusting for any endocrine treatment or for ER antagonists retained similar results (log rank P = 0.006 and 0.007 for overall survival and P = 0.002 and 0.003 for breast cancer-specific survival). Similar analyses of combinations of ER $\alpha$  and ER $\beta$ cx expression did not render any significant associations with survival.

ER $\beta$ cx expression in SLN lymphocytes is more common in node positivity and affects overall survival. ER $\beta$ 1 and ER $\beta$ cx positivity of lymphocytes residing in the SLN was seen in 202 out of 285 (70.9%) and 116 out of 292 (39.7%) patients, respectively. In

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contrast, ER $\alpha$  positivity in these cells was an extremely rare event with only one positive out of 248 cases (0.4%). ER $\beta$ cx positivity in SLN lymphocytes was more common in node-positive than nodenegative patients (57 out of 125 (45.6%) *vs* 59 out of 167 (35.3%)), even though this did not reach statistical significance (P = 0.076). This trend was supported by the fact that patients with ER $\beta$ cx positivity in their SLN had a higher mean number of axillary lymph node metastases (1.4) than those with ER $\beta$ cx negativity (0.9; P = 0.055). Ten-year overall survival was 93.0% for patients with ER $\beta$ cx negativity in their SLN lymphocytes, as compared with 84.8% in those with ER $\beta$ cx SLN lymphocyte positivity (log rank P = 0.053, Figure 2F). Interestingly, this finding turned significant when adjusting for adjuvant endocrine therapy (log rank P = 0.039). There was no effect of ER $\beta$ cx SLN lymphocyte status on breast cancer-specific survival or on disease-free survival.

# DISCUSSION

Our analysis revealed that  $ER\beta 1$  positivity within the primary tumour is an independent marker for good outcome, more powerful than ERa, which is classically associated with increased survival after adjuvant endocrine therapy. ER $\beta$ 1 expression remained an independent prognostic marker for both overall and breast cancer survival in a multivariable Cox regression model, which strengthens the prognostic value of the receptor. The number of events in our cohort during follow-up period was small due to the generally good prognosis of breast cancer today and considering that included patients were clinically node negative. It has been shown, however, that survival analysis can be reliable with even as little as five events per variable within the Cox regression model. This is especially evident when the association is plausible and hypothesised a priori (Vittinghoff and McCulloch, 2007). When performing survival analysis in breast cancer while studying ERs it is important with sufficient follow-up due to the risk of late recurrences. The follow-up for breast cancer-specific survival was shorter than for overall survival due to their definitions specified in the Statistics section. However, since any recurrent breast cancer

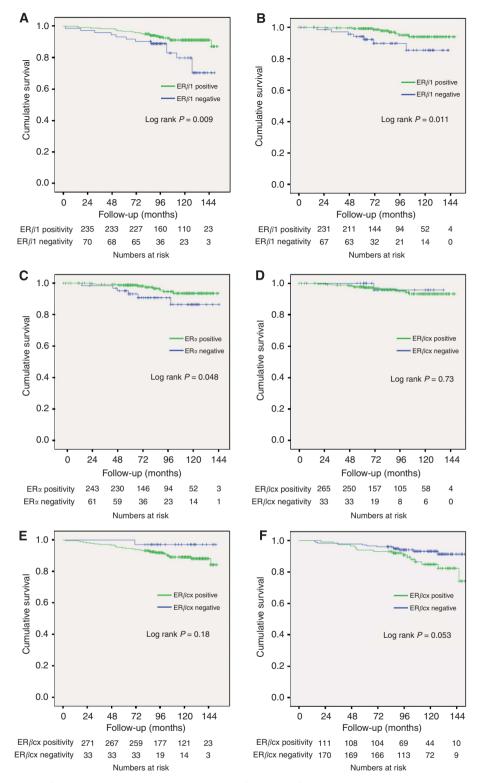


Figure 2. Kaplan–Meier plots of overall survival and breast cancer-specific survival for ER expression in primary tumour and sentinel node lymphocytes. (A) There was a higher overall survival in patients with  $ER\beta1$ -positive tumours (green) compared to  $ER\beta1$ -negative (blue; log rank P=0.009). (B) The same was seen for breast cancer-specific survival for  $ER\beta1$ -positive tumours (log rank P=0.011). (C) Breast cancer-specific survival was higher in patients with  $ER\alpha$ -positive tumours (log rank P=0.048). (D) Breast cancer-specific survival did not differ between patients with  $ER\betacx$ -positive or -negative tumours (log rank P=0.73). (E) There was a trend towards better overall survival in patients with  $ER\betacx$ -negative tumours; however, this did not reach statistical significance (log rank P=0.018). (F) Patients with  $ER\betacx$ -negative lymphocytes in the sentinel node showed a non-significant trend towards better overall survival (log rank P=0.053). Numbers at risk at each time point are given below each subfigure.

cases were referred back to the department where this study was performed, it is unlikely that such cases were missed. Thus, followup for breast cancer-specific survival and disease-free survival is probably underestimated. When we stratified patients for endocrine treatment, the survival was similarly high in patients with  $\text{ER}\beta$ 1-positive tumours. This is somewhat contradicting when

|                                 | ER $\beta$ 1 positive | ERβ1 negative         |  |
|---------------------------------|-----------------------|-----------------------|--|
| 10-year OS                      |                       |                       |  |
| ERα positive (%)                | 91.4                  | 85.4                  |  |
| ERα negative (%)                | 90.0                  | 75.6                  |  |
| P-value                         | 0.050*                |                       |  |
| 10-year BCSS                    |                       |                       |  |
| ERα positive (%)                | 93.9                  | 91.0                  |  |
| ERα negative (%)                | 93.6                  | 72.1                  |  |
| P-value                         | 0.009*                |                       |  |
|                                 | $ER\beta cx$ positive | $ER\beta cx$ negative |  |
| 10-year OS                      |                       |                       |  |
| ER $\beta$ 1 positive (%)       | 92.0                  | 95.5                  |  |
| $ER\beta1$ negative (%)         | 78.1                  | 100                   |  |
| P-value                         | 0.011*                |                       |  |
| 10-year BCSS                    |                       |                       |  |
| $ER\beta1$ positive (%)         | 95.9                  | 92.3                  |  |
| $\text{ER}\beta$ 1 negative (%) | 80.0                  | 100                   |  |
| P-value                         | 0.001*                |                       |  |
| 10-year OS                      |                       |                       |  |
| ERα positive (%)                | 90.4                  | 100                   |  |
| ERα negative (%)                | 84.6                  | 90.0                  |  |
| P-value                         | 0.283                 |                       |  |
| 10-year BCSS                    |                       |                       |  |
| ERα positive (%)                | 94.7                  | 100                   |  |
| ERα negative (%)                | 86.0                  | 85.7                  |  |
| P-value                         | 0.093                 |                       |  |

Abbreviations: BCSS = breast cancer-specific survival; ER = oestrogen receptor; OS = overall survival. The highest 10-year OS/BCSS was seen in ER $\alpha$  + /ER $\beta$ 1 + tumours. Either ER $\alpha$  + / ER $\beta$ 1 - or ER $\alpha$  - /ER $\beta$ 1 + was also associated with better prognosis, compared to double negative tumours (ER $\alpha$  - /ER $\beta$ 1 - that had the worst prognosis for both OS and BCSS. For ER $\beta$  and ER $\beta$ cx co-expression, the worst 10-year OS/BCSS was seen in patients with ER $\beta$ 1 - /ER $\beta$ cx + tumours, compared to ER $\beta$ 1 + /ER $\beta$ cx - , Re $\beta$ 1 - / ER $\beta$ cx - , No significant differences in 10-year OS BCSS were seen when comparing the co-expression of ER $\alpha$  and ER $\beta$ cx. \*P<0.05.

compared with the results of Honma et al (2008), where the increase in survival was only seen in  $ER\beta$ 1-positive patients treated with tamoxifen for >2 years. Our group of endocrine-treated women, however, received tamoxifen and/or aromatase inhibitors usually for 5 years, which may perhaps explain the differences in survival. When we examined survival regarding the co-expression of ER $\alpha$  and ER $\beta$ 1, patients who were either double or single positive had an equally good prognosis, meaning that  $ER\beta 1$  is a potential biomarker for distinguishing between patients with good or bad prognosis in ERa-negative tumours, generally considered as a group of patients with poor prognosis. ER $\beta$ 1 and ER $\beta$ cx coexpression was associated with survival in a similar manner, again the lowest survival was seen in the ER $\beta$ 1-negative tumours. Intriguingly, these tumours were also  $ER\beta cx$ -positive.  $ER\beta cx$  is thought to affect ER function through negative regulation by heterodimerisation, mainly with ERa, causing degradation of the receptor complex and therefore a decrease of ERa level (Zhao et al, 2007). This mechanism may result in a completely ER-negative tumour (ER $\alpha$  – and ER $\beta$ 1 – ). Therefore, this tumour could be less sensitive to endocrine therapy and perhaps more responsive to chemotherapy. As described earlier it is important to remember that the stratification and subgroup analysis of the co-expression data is performed on a small cohort and with few events, which merits caution when interpreting these results. However, we believe that the findings are both clinically and biologically relevant, but need further validation.

 $ER\beta1$  seems to be present in tumours in patients from all age groups, whereas ERa expression is usually less common in younger women. Most interestingly, the negative prognostic significance of the lack of ER $\beta$ 1 expression in the primary tumour is particularly strong in younger age groups and in tumours of high histological grades. Therefore,  $ER\beta$  should be considered a valuable prognostic biomarker and perhaps a therapeutical target in younger women in whom triple-negative breast cancer (absence of ERa, PR and HER2) is more common. Tamoxifen and other current endocrine therapies such as aromatase inhibitors, however, are designed to treat ERaexpressing cancers and might not be the optimal therapy for targeting ER $\beta$ 1. Our data suggest that the examination of ER $\beta$ 1 status should have an additional prognostic value during routine pathological examination of breast cancer, but not as a biomarker for endocrine responsiveness within ERa-negative cases as described elsewhere (Gruvberger-Saal et al, 2007; Honma et al, 2008). Further research is needed to identify the ideal  $ER\beta$ 1-targeting therapy and to validate our findings.

Although several prospective investigations on  $ER\beta$  isoforms in breast cancer have been performed, the potential association with clinicopathological parameters and survival remains unclear. The expression of the different  $ER\beta$  isoforms is a result of alternative splicing at the C-terminus; consequently, antibodies raised against the N-terminus will inevitably quantitate the total  $ER\beta$  level. This may cause a false view of the isoform-specific expression. In our study, we analysed ER $\beta$ 1 expression with a widely cited, wellvalidated C-terminal monoclonal antibody (Skliris et al, 2002; Carder et al, 2005; Weitsman et al, 2006; Novelli et al, 2008). As mentioned,  $ER\beta$  has been described as an anti-proliferative, tumour-suppressive receptor within breast cancer cells in vitro (Hartman et al, 2009). Consequently, it has been hypothesised that  $ER\beta$  should be downregulated during breast cancer progression (Roger et al, 2001). However, paired primary tumours and corresponding metastatic lesions are extremely scarce materials and the hypothesis is thus hard to prove. Often, the presence of locoregional lymph node metastases are used as a surrogate parameter, as they are one of the strongest risk factors for breast cancer death and distant metastases (Fisher et al, 1983; Andersson et al, 2010). In a study on 50 patients, Gschwantler-Kaulich et al (2011) observed that compared to the primary tumour, there was a reduction in both ER $\alpha$  and ER $\beta$  levels in axillary lymph node metastases. In our analysis, ER status differed between primary tumour and corresponding lymph node metastases in several patients, even though we could not identify any changes in the overall pattern of ER status. This discrepancy is probably explained by the fact that Gschwantler-Kaulich et al used TMAs with 1 mm cores while our analysis was based on whole sections. It is evident that intratumoural ER levels are heterogeneously expressed, hence, IHC on TMAs may not be an optimal method for ER assessment of breast cancer specimens. Furthermore, to reduce variation and bias in our study, all IHC stainings were assessed by computer-assisted image analysis (see Materials and Methods section) in a blinded manner. This method, if correctly performed, is able to reduce both intra- and interobserver variation in biomarker assessments (Krecsák et al, 2011). Our analysis further showed that patients with  $ER\beta cx$ positive primary tumours had an increased risk of lymph node metastasis. Nonetheless, the ER status of the corresponding lymph node metastasis did not correlate to outcome.

Oestrogen receptors, expressed in lymphocytes, are important in the maturation of B cells and play a crucial role in the peripheral immune system; this effect is mediated by both ER $\alpha$  and ER $\beta$ (Shim *et al*, 2006; Hill *et al*, 2011). Within the stroma of mammary gland and tumour adjacent tissue, ER $\beta$  is the predominating ER (Speirs *et al*, 2002). In the majority of our patients, we observed that lymphocytes within the SLNs express ER $\beta$ 1 (70.9%) while a minority expressed ER $\beta$ cx (39.4%) and <1% expressed ER $\alpha$ . We found that lymphocytes within lymph nodes containing metastatic breast cancer cells express higher levels of  $\text{ER}\beta$ cx. In patients treated with endocrine therapy,  $\text{ER}\beta$ cx-positive SLN lymphocytes indicated a poor prognosis and were furthermore associated with shorter 10-year overall survival. Our data imply that  $\text{ER}\beta$ cx within SLN lymphocytes may govern a yet unknown mechanism of breast cancer progression. Since the lymphocytes expressed little  $\text{ER}\alpha$ , the tumourigenic function could perhaps be mediated through heterodimerisation and inhibition of  $\text{ER}\beta$ 1.

In summary, this study indicates that  $ER\beta 1$  and the  $ER\beta$  splice variant  $ER\beta cx$  have several important roles during breast tumourigenesis. In the primary tumour,  $ER\beta 1$  was associated with good outcome and probably has a tumour-suppressive function.  $ER\beta cx$ , however, seems to play the most important role in regional lymph nodes where its presence in lymphocytes correlated to overall survival in breast cancer patients through an as yet unknown mechanism. The analysis of  $ER\beta 1$  and  $ER\beta cx$  by IHC provides useful clinical information, especially for younger women and tumours of high histological grade. Further research is needed to understand how to pharmaceutically target the individual ER subtypes in breast cancer patients.

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#### CONFLICT OF INTEREST

The authors declare no conflict of interest.

## AUTHOR CONTRIBUTIONS

JH, JdB, JB and JF designed research; GR, JH, JdB and GMK performed research; JdB and GR analysed data; JdB, GR and JH wrote the paper; JH led the project.

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