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Treatment and overall survival in renal cell carcinoma: a Swedish population-based study (2000–2008)

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Background: This retrospective register study assessed overall survival (OS) and influential factors on OS in Swedish renal cell carcinoma (RCC) patients.

Methods: Using three merged national health registers, Cox proportional-hazards analysis was conducted and, in three models, it was used to assess the impact of cytokine (interferon- α and tyrosine kinase inhibitor (TKI; sunitinib or sorafenib) treatment on OS in metastatic (m)RCC.

Results: From 2000 to 2008, 8009 patients were diagnosed with RCC and 2753 with mRCC (2002–2008). Median OS in RCC patients diagnosed from 2006 to 2008 compared with 2000–2005 was not reached vs 47.9 months ($P < 0.001$), and in mRCC patients diagnosed from 2006 to 2008 compared with 2002–2005, was 12.4 vs 9.6 months, respectively ($P = 0.004$). Factors associated with significantly improved OS in RCC were female gender, lower age, and previous nephrectomy, and, in mRCC female gender, previous nephrectomy, and any TKI prescription (Model 1: median-adjusted OS, 19.4 months (TKI patients) vs 9.7 months (non-TKI patients); hazard ratio, 0.621; $P < 0.001$).

Conclusion: OS was improved in Swedish patients diagnosed with RCC and mRCC in the period 2006–2008 compared with 2000–2005 (RCC) and 2002–2005 (mRCC). Although multifactorial in origin, results suggest that increased nephrectomy rates and the use of TKIs contributed to the improvement seen in mRCC patients.

Introduction of modern targeted therapies has greatly improved the prognosis of patients with metastatic renal cell carcinoma (mRCC). The multitargeted tyrosine kinase inhibitors (TKIs) sunitinib and sorafenib were the first new therapies approved for advanced RCC and have been available in the European Union since 2006. In a clinical trial of treatment-naïve mRCC patients, sunitinib significantly improved progression-free survival (PFS) vs interferon- α (IFN- α ; median PFS, 11 vs 5 months, $P < 0.001$; Motzer *et al*, 2007) and was associated with an improved median

overall survival (OS) of 26.4 and 21.9 months ($P = 0.051$), respectively (Motzer *et al*, 2009). In another trial, sorafenib significantly improved PFS vs placebo in patients after failure of one systemic therapy (cytokine in $> 80\%$ of cases); median PFS was 5.5 vs 2.8 months, respectively ($P < 0.01$; Escudier *et al*, 2007). OS was similar in the sorafenib and placebo groups (17.8 vs 15.2 months, respectively; $P = 0.146$), although the results were confounded by extensive cross-over (Escudier *et al*, 2009a). Despite impressive clinical trial results, there is a paucity of

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survival data from population-based studies reflecting these advances in clinical practice.

Sweden has a long heritage in maintaining public health registries and is therefore uniquely placed as a source of comprehensive and high-quality epidemiological data. The Swedish Cancer Register, for example, has existed since 1958 and holds information on almost all patients with a cancer diagnosis (Barlow *et al*, 2009).

The RENal COMPARison (RENCOMP) study was a retrospective, non-interventional study using data from the Swedish Cancer Register and two other national registries. The objectives were to assess temporal and regional trends in treatment patterns and survival of Swedish patients with RCC and to describe the impact of resource use.

The objectives of the RENCOMP analyses reported here were to assess changes in OS in Swedish patients with RCC and mRCC by comparing these nationwide patient cohorts diagnosed before and after the introduction of targeted therapies. We also investigated factors influencing OS in both populations and assessed the impact of first-line treatment (sunitinib, sorafenib, or IFN) and treatment sequences in mRCC.

MATERIALS AND METHODS

Study design and patients. This retrospective study used data from Swedish patients compiled in three registries by the National Board of Health and Welfare, Stockholm, Sweden (Table 1). Patients aged ≥ 16 years with a diagnosis of RCC (malignant neoplasm of the kidney, except renal pelvis) between 1958 and 2008 were identified by review of the Swedish Cancer Register for the International Classification of Diseases and Related Health Problems (ICD)-7 diagnosis codes I800 or I809 (excluding I801 (cancer of the renal pelvis)); these codes are equivalent to ICD-10 codes C64.0 and C64.9. The RCC study population, including mRCC patients, was limited to patients diagnosed from 2000 to 2008 to reflect modern RCC treatment. The mRCC population comprised patients diagnosed with metastatic disease (see below) from 2002 to 2008, including a number of patients with an RCC diagnosis before the year 2000. Cohorts for comparison purposes were defined as patients (with mRCC) diagnosed from 2000 (2002) to 2005 *vs* 2006–2008. For these patients, data from the Swedish Cancer Register were merged with data from the National Patient Register (December 2009) and Swedish Prescribed Drug Register (July 2005–October 2010) using personal identification numbers. Dates of death were updated as of October 2010. All data were made anonymous before use in the analysis.

Table 1. Summary of national Swedish registries used in this retrospective, non-interventional study

Registry	Year founded	Data	Completeness (%) ^a
Swedish Cancer Register	1958	Diagnosis and death records of all patients with a cancer diagnosis	100
National Patient Register	1987	Information on inpatient visits (since 1987) and outpatient visits (since 2001)	> 90
Swedish Prescribed Drug Register	2005	Dates and amounts of prescribed and dispensed drugs for individual patients	100

^aBarlow *et al*, 2009; Swedish National Board of Health and Welfare, 2011; Wettermark *et al*, 2007.

As the Cancer Register did not include full information on the date for mRCC diagnosis, the mRCC population was defined based on a derived date for diagnosis of metastatic disease. The date was derived as when one of three events (whichever occurred first) was recorded in the registry database. These events comprised (1) diagnosis of primary metastasis (M1 disease); (2) diagnosis of secondary (malignant) tumour; and (3) the first visit to an oncology clinic. These criteria could be properly applied from 2002 onwards, which explains why the RCC and mRCC cohorts differed in terms of years included.

TKI medication was defined as any medication dispensed after the date of diagnosis of mRCC with an Anatomical Therapeutic Classification (ATC) code L01XE04 (sunitinib) or L01XE05 (sorafenib); cytokine medication was defined as any medication dispensed after the date of mRCC diagnosis with an ATC code L03AB (includes different types of alpha-IFNs approved for RCC, and well as IFN- β and IFN- γ) or L03AC (includes IL-2 approved for RCC, as well as IL-11). Importantly, there is a large fraction of non-prescription TKI use in Sweden (on average, >40% of prescriptions made according to sales figures, with regional differences, and slowly increasing over time), in which TKIs are dispensed at clinics and not recorded in the Prescribed Drug Register.

Statistical methods. All data were presented using descriptive statistics, that is, frequency and relative frequency for categorical variables, and mean and s.d. and/or median and minimum and maximum for continuous variables. Median OS (reported in months, in which 1 month = 30 days) was estimated by Kaplan–Meier method and compared between patients with RCC and, separately, between patients with mRCC, by year and period of diagnosis using the log-rank test for univariate analysis. Multi-variate regression analysis was performed using the Cox proportional-hazards model. The regression model for OS in the overall RCC population, as well as models for OS in those with mRCC, included the following covariates: period of diagnosis, age, gender, institution size, nephrectomy status, and geographic region.

Three different Cox proportional-hazards models were used for analysis of OS in the cohort of patients with mRCC. mRCC Model 1 assessed the effect (hazard ratio (HR)) of at least one TKI prescription (compared with no TKI prescription) on OS; Model 2 examined the impact of first-line treatments (sunitinib, sorafenib, or IFN) as separate categories in the analysis of OS, which could include a subsequent treatment (Figure 1); and Model 3 investigated the effect of various treatment sequences on OS in mRCC patients (Figure 1). The reference category, that is, comparator, for the HR estimation of first-line treatment and treatment sequences in Models 2 and 3 included patients with no RCC drug treatment or other treatments (non-cytokine/non-TKI). From the results of the multivariate regression analyses, median OS was estimated to describe the differences of interest in Models 1–3.

All tests were two-sided and $P < 0.05$ was regarded as a statistically significant result. PASW Statistics v18 (IBM SPSS Statistics; IBM Corp., Armonk, NY, USA) was used for statistical analysis.

RESULTS

Patients. Between 2000 and 2008, 8009 patients were diagnosed with RCC, comprising the RCC study population, of whom 533 (7%) were diagnosed at autopsy. Between 2002 and 2008, 2753 patients, representing 44% of patients diagnosed with RCC during the same period, met the necessary criteria for a derived diagnosis of mRCC and comprised the mRCC study population. These included patients who presented with metastatic disease at initial diagnosis ($n = 872$; 31% of patients with a derived mRCC

Figure 1. Patients who received first- and second-line mRCC treatment included in the Cox proportional-hazards models.

diagnosis); patients diagnosed with metastatic disease by means of a secondary malignant tumour, subsequent to diagnosis of localised RCC ($n=1311$; 48%); and patients diagnosed based on their first visit to an oncologist ($n=570$; 21%).

Mean (s.d.) age at diagnosis of RCC and mRCC was 70 (12) and 69 (11) years, respectively (with 16% of mRCC patients aged ≥ 80 years). In all, 60% and 61% were male, respectively, and 39% of both RCC and mRCC patients were diagnosed in university clinics. Seventy-five percent of RCC patients who did not develop metastases during the observation period (80% in the period 2006–2008) and 60% of patients with M1 disease (by 2008) had a record of nephrectomy.

Patient characteristics for the RCC and mRCC study populations and the cohorts for period of diagnosis are listed in Table 2.

Drug treatment patterns in mRCC. Between July 2005 (the earliest year that prescription data were available) and October 2010, among patients diagnosed with mRCC from 2002 to 2008 ($n=2753$), 759 (28%) were dispensed a prescription for this condition, and 417 patients (15%) were dispensed at least one TKI either as first- or second-line treatment after another TKI or IFN. Of those patients diagnosed with mRCC from 2006 to 2008 ($n=1217$), 513 patients (42%) were dispensed a prescription and 360 (30%) were dispensed a TKI. The number of patients treated in first- or second-line with sunitinib or sorafenib, as well as IFN, is shown in Figure 1. Of the total 2002–2008 mRCC population, based on prescription, 8% of patients with mRCC received sunitinib as first-line treatment, 10% received first-line IFN, and 4%, first-line sorafenib; however, most patients (78%) received other or no treatment. These seemingly small shares depend on the availability of prescription data (July 2005–October 2010), introduction of TKIs in late 2006, and non-prescription TKI use as described above. Thirty-five percent of patients treated in

first line received a second-line treatment, and only 28% received another TKI after failing one previous TKI. Other treatments included tamoxifen, oral chemotherapeutic drugs, and other hormonal therapy.

Overall survival

RCC population. Between 2000 and 2008, 3- and 5-year OS rates for the total RCC population were 58% and 49%, respectively. Among all RCC patients diagnosed during this period ($N=8009$), 4181 (52%) died on or before October 2010, and median (unadjusted) OS, estimated by the Kaplan–Meier method, was 1816 days (i.e., 60.5 months (95% CI: 56.1–64.9)).

There was a statistically significant incremental improvement in median OS over time, as estimated by Kaplan–Meier method, in RCC patients diagnosed between 2000 and 2008 (Figure 2A), and in RCC patients diagnosed in the period 2006–2008 compared with 2000–2005, in which median OS was not reached vs 47.9 months (95% CI: 43.2–52.7), respectively ($P<0.001$; Figure 2B).

When controlling for confounding factors in the multivariate analysis, RCC patients diagnosed from 2006 to 2008 showed 28.6% reduction in the risk of death compared with those diagnosed from 2000 to 2005 (HR, 0.714, 95% CI: 0.664–0.767; $P<0.001$; Table 3), as reflected in median OS (not reached vs 46.7 months, respectively). Other factors associated with a significantly reduced risk of death in the total RCC population included female gender, lower age, previous nephrectomy, and treatment in the Stockholm, southern, and eastern regions, as compared with the western region of Sweden (Table 3).

mRCC population. For patients diagnosed with mRCC between 2002 and 2008, 3- and 5-year survival rates for the total population were 21% and 13%, respectively, and the median (unadjusted)

Table 2. Characteristics of Swedish patients diagnosed with RCC and mRCC by period of diagnosis

	RCC population		mRCC population	
Characteristic	2000–2005 (n = 5197)	2006–2008 (n = 2812)	2002–2005 (n = 1536)	2006–2008 (n = 1217)
Mean (s.d.) age, years	71 (11)	69 (12)	69 (10)	70 (11)
Male/female, %	60/40	60/40	61/39	62/38
T stage, % reported cases				
T0	<1	<1	—	—
T1	44	49	—	—
T2	22	19	—	—
T3	26	26	—	—
T4	5	3	—	—
Ta	0	<1	—	—
Tx	3	3	—	—
Previous nephrectomy, %	72	78	54	66
Deaths, %	60	37	88	76
Institution size, %				
Percentage of Large	37	35	42	35
Small	63	65	58	65
Region of Sweden, %				
South	18	19	17	18
Mid/central	23	23	21	22
Stockholm	18	17	21	18
East	12	11	12	13
North	11	11	10	11
West	18	19	20	19

Abbreviation: mRCC = metastatic renal cell carcinoma.

OS, estimated by the Kaplan–Meier method, was 10.9 months (95% CI: 10.0–11.8). There was a statistically significant incremental improvement in median OS over time, as estimated by Kaplan–Meier method, in mRCC patients diagnosed between 2002 and 2008 (Figure 3A). For mRCC patients diagnosed in the periods 2006–2008 and 2002–2005, median OS was 12.4 months (95% CI: 11.0–13.8) vs 9.6 months (95% CI: 8.5–10.7), respectively ($P = 0.004$; Figure 3B).

When controlling for confounding factors in the multivariate analysis, the risk of death was not statistically significantly different between mRCC patients diagnosed from 2006 to 2008 compared with those diagnosed from 2002 to 2005 (HR, 1.029, 95% CI: 0.941–1.125; $P = 0.536$; Table 4; mRCC Model 1), as reflected in the median-adjusted OS (10.6 vs 11.1 months, respectively). This finding was mainly explained by nephrectomy rates and use of TKIs (Supplementary Table S1). Accordingly, prescription of any TKI was associated with a significantly reduced risk of death (HR, 0.621, 95% CI: 0.547–0.705; $P < 0.001$; Table 4) for which the median-adjusted OS was 19.4 months for TKI patients vs 9.7 months for non-TKI patients (Figure 4).

According to mRCC Model 2, which evaluated the first-line treatment with sorafenib, IFN, and sunitinib, irrespective of subsequent treatment, all first-line treatment options were associated with a statistically significantly reduced risk of death (HR, 0.682–0.596; $P \leq 0.001$; Table 4) corresponding to a median-adjusted OS of 16.1 months with sorafenib to 19.6 months with sunitinib vs 9.2 months with other/no treatment (Figure 5).

mRCC Model 3, which investigated the effect of various treatment sequences on OS, confirmed a similar reduction in risk of death for patients receiving these sequences as compared with patients who received other or no treatment (Table 4). However, for patients who received one treatment solely, only sunitinib and IFN were associated with a statistically significantly reduced risk of death as compared with other or no treatment (Table 4).

Other factors associated with significantly reduced risk of death, and common to all mRCC models, included female gender, previous nephrectomy, and treatment in the Stockholm region, as compared with the western region of Sweden (Table 4).

DISCUSSION

By analysing three different merged national registries in Sweden, we had a unique opportunity to evaluate survival in patients diagnosed with RCC and mRCC before and after the introduction of targeted therapies in 2006. Specifically, we were able to demonstrate the improvements in survival that occurred in Swedish patients diagnosed with mRCC after the introduction of TKIs. The two RCC cohorts used for comparison had similar characteristics at diagnosis (Table 2), with only the following trends noted: an increase in the rate of nephrectomy, a decline in the mean age at diagnosis, and a slight stage migration, mainly from stage T2 to T1. Although these trends indicate a more active approach to the disease, these small differences alone cannot explain the differences in the survival between the two cohorts. We also showed that TKI treatment together with an increased nephrectomy rate were two important factors behind the variation in observed OS in the mRCC cohorts. This was demonstrated by an almost 10-month survival benefit among TKI-treated patients compared with patients who received non-TKI or no therapy.

The influence of other factors associated with OS was also analysed. For example, female gender was associated with lower risk of death (longer OS), an effect previously documented that may be related to less advanced disease at presentation in female patients (Aron *et al*, 2008; Woldrich *et al*, 2008). As expected, previous nephrectomy, the standard of care for patients with localised disease, was also associated with lower risk of death (longer survival). The impact of cytoreductive nephrectomy is well established in patients with RCC treated with IFN (Flanigan *et al*, 2004) and also retrospectively demonstrated with targeted therapy (Choueiri *et al*, 2011), although it tends to be used only in patients with better prognosis. In our study, the impact of nephrectomy, in terms of HR, was more marked in the overall RCC population than in mRCC patients, perhaps reflecting the potentially curative nature of surgery in early stage RCC.

Median unadjusted OS of the RCC population in our study, as estimated by Kaplan–Meier method, was 60.5 months, with a 5-year survival rate of 49%, both of which are less than the figures reported by the US SEER database (median OS 110 and 130 months for males and females, respectively, with 5-year survival rates of 65% and 69%, respectively; Aron *et al*, 2008). These differences may be explained by the younger age at diagnosis in the SEER population compared with the age at diagnosis of patients in our study (mean, 62 vs 70 years, respectively) and a slightly smaller proportion of patients presenting with M1 disease at diagnosis (9.9% vs 27%, respectively). More recent SEER data for the early targeted therapy era (2004–2007) showed a crude 3-year survival rate of 75% compared with 68% for the cytokine era (1998–2003), again with younger patients and fewer presenting with distant metastasis, which may account for the observed differences (Shek *et al*, 2012); in our study, 3-year survival was estimated at 58%. However, it should be noted that patients diagnosed at autopsy were included in our analysis (7%), shifting the survival curve downward.

Figure 2. Kaplan–Meier estimates of OS in Swedish patients diagnosed with RCC (**A**) by year of diagnosis and (**B**) between the periods 2000–2005 and 2006–2008. NA, not applicable. CI, confidence interval; NA, not available; NR, not reached.

Median unadjusted OS of the total mRCC population in our study was 10.9 months, as estimated by Kaplan–Meier method, 9.6 months in the 2002–2005 cohort and 12.4 months in the 2006–2008 cohort. These figures might seem low in comparison to published series on treated patients with cytokines and TKIs (Coppin *et al*, 2005; Gore *et al*, 2009; Stadler *et al*, 2010) but represent a population-based analysis including a large number of non-treated and poorly performing patients. It is also likely that the derived date of mRCC diagnosis in patients was set later than the actual date, as it was based on an estimate (algorithm) rather than an observed date. This suggests caution when interpreting and indirectly comparing median survival estimates across studies, while HRs likely provide a more accurate assessment of treatment benefit.

Nearly all of the three treatments analysed, alone or used in sequences, were associated with improved OS compared with other

or no treatment. The only exception was sorafenib used as first-line treatment followed by best supportive care, although patient numbers were low. Whether this implies inferior efficacy with sorafenib as first-line treatment compared with other TKIs, as previously shown in a randomised phase II trial (Escudier *et al*, 2009b), cannot be concluded with this study; however, the data suggest that sorafenib efficacy, in terms of OS, improves when used in sequence. By contrast, our findings related to sunitinib treatment appear to justify its recommendation as first-line treatment for mRCC in both European and US guidelines (Ljungberg *et al*, 2010; Motzer *et al*, 2011). mRCC Model 3 seemed to suggest an equal improvement in OS regardless of sequence, which is consistent with a recent Czech register study and German phase II study (Herrmann *et al*, 2011; Buchler *et al*, 2012). However, an Italian retrospective study suggested a benefit for the sorafenib–sunitinib sequence compared

Table 3. Results of multivariate Cox proportional-hazards regression analysis of factors correlated with OS in Swedish patients diagnosed with RCC between 2000 and 2008

RCC model (n = 8009)		
	HR (95% CI)	P value
Gender (female vs male)	0.897 (0.842–0.954)	0.001
Age	1.004 (1.001–1.006)	0.008
Nephrectomy (yes vs no)	0.152 (0.142–0.162)	<0.001
Institution size (large ^a vs small)	1.047 (0.981–1.117)	0.165
Year of diagnosis (2006–2008 vs 2000–2005)	0.714 (0.664–0.767)	<0.001
Region of Sweden (region of interest vs West)		
South	0.893 (0.808–0.988)	0.028
Mid/central	0.989 (0.899–1.088)	0.825
Stockholm	0.871 (0.786–0.965)	0.009
East	0.895 (0.802–0.999)	0.048
North	1.049 (0.936–1.176)	0.409

Abbreviations: CI = confidence interval; HR = hazard ratio; OS = overall survival; RCC = renal cell carcinoma. Note: for continuous variables (e.g., age), an HR > 1 equates to risk reduction when the value decreases and an HR < 1 equates to risk reduction when the value increases; for binary variables, an HR < 1 equates to risk reduction for the first category and an HR > 1 equates to risk reduction for the second category.
^aA large institution indicates a university urology clinic, where patients were diagnosed, and does not include university oncology clinics.

with sunitinib–sorafenib (Porta *et al*, 2011). A definitive answer regarding the most clinically beneficial sequence most likely requires a phase III trial, specifically designed to address this question; such a trial is currently ongoing in Germany, with results due in 2013 (NCT00732914). Regardless, consistent with findings by Heng *et al* (2010), a low fraction of patients in this study (28%) received second-line TKI treatment after failure of another TKI. Even if these data represent an early treatment era, this rate is substantially lower than might be expected by clinicians and clearly emphasises the importance of carefully selecting first-line treatment.

Regional differences were present in the analysis, consistent with findings in other therapeutic areas in Sweden (Swedish National Board of Health and Welfare, 2011). However, it is possible that other factors than those corrected for in the analysis could explain these regional differences.

There were several limitations to our study. One was that the multivariate analyses did not include important RCC prognostic factors, such as performance status, risk group stratification based on published Memorial Sloan–Kettering Cancer Center data (Motzer *et al*, 2002), or histology, as these data are not recorded in the registers. However, the lack of these data was not expected to affect the OS estimates in the Kaplan–Meier analysis as there is no reason to expect that the risk profile distribution changed over time.

Most importantly, however, this was a retrospective study, indicating a substantial risk of bias. Although selection bias is certainly a problem in the nephrectomy decision in mRCC, a clear selection strategy had not been established for TKI treatment in the

Figure 3. Kaplan–Meier estimates of OS in Swedish patients diagnosed with mRCC (A) by year of diagnosis and (B) between the periods 2002–2005 and 2006–2008. CI, confidence interval; NA, not available; NR, not reached.

Table 4. Results of a multivariate Cox proportional-hazards regression analysis of factors predictive for OS in Swedish patients diagnosed with mRCC between 2002 and 2008						
	mRCC Model 1 (n = 2753)		mRCC Model 2 (n = 2753)		mRCC Model 3 (n = 2753)	
	HR (95% CI)	P value	HR (95% CI)	P value	HR (95% CI)	P value
Gender (female vs male)	0.892 (0.819–0.971)	0.008	0.898 (0.824–0.977)	0.013	0.913 (0.836–0.998)	0.045
Age	0.999 (0.995–1.003)	0.736	0.998 (0.994–1.002)	0.414	0.999 (0.994–1.003)	0.489
Nephrectomy (yes vs no)	0.572 (0.525–0.624)	<0.001	0.578 (0.530–0.630)	<0.001	0.571 (0.522–0.625)	<0.001
Institution size (large ^a vs small)	0.992 (0.909–1.083)	0.863	0.997 (0.914–1.089)	0.955	1.010 (0.922–1.106)	0.838
Year of diagnosis (2006–2008 vs 2002–2005)	1.029 (0.941–1.125)	0.536	1.026 (0.937–1.123)	0.581	0.995 (0.905–1.093)	0.910
Region of Sweden (region of interest vs West)						
South	0.962 (0.842–1.100)	0.574	0.940 (0.821–1.075)	0.367	0.923 (0.803–1.061)	0.260
Mid/central	0.893 (0.785–1.016)	0.086	0.882 (0.775–1.003)	0.056	0.875 (0.766–1.000)	0.050
Stockholm	0.686 (0.599–0.784)	<0.001	0.675 (0.589–0.772)	<0.001	0.657 (0.572–0.756)	<0.001
East	0.898 (0.775–1.042)	0.157	0.906 (0.781–1.051)	0.193	0.869 (0.746–1.011)	0.068
North	1.076 (0.921–1.256)	0.356	1.040 (0.890–1.214)	0.623	1.045 (0.883–1.237)	0.606
Any TKI prescription (yes vs no)	0.621 (0.547–0.705)	<0.001	—	—	—	—
Treatment (regimen of interest vs other/no treatment)						
Sunitinib first-line (sole treatment in Model 3)	—	—	0.596 (0.503–0.706)	<0.001	0.599 (0.489–0.733)	<0.001
Sorafenib first-line (sole treatment in Model 3)	—	—	0.682 (0.549–0.847)	0.001	0.810 (0.616–1.066)	0.133
IFN first-line (sole treatment in Model 3)	—	—	0.636 (0.549–0.735)	<0.001	0.696 (0.586–0.827)	<0.001
Sunitinib → sorafenib	—	—	—	—	0.577 (0.417–0.798)	0.001
Sorafenib → sunitinib	—	—	—	—	0.536 (0.367–0.781)	0.001
IFN → sunitinib	—	—	—	—	0.509 (0.349–0.742)	<0.001
IFN → sorafenib	—	—	—	—	0.490 (0.334–0.719)	<0.001

Abbreviations: CI = confidence interval; HR = hazard ratio; IFN = interferon-α; mRCC = metastatic renal cell carcinoma; OS = overall survival; TKI = tyrosine kinase inhibitor. Note: for continuous variables (e.g., age), an HR > 1 equates to risk reduction when the value decreases and an HR < 1 equates to risk reduction when the value increases; for binary variables, an HR < 1 equates to risk reduction for the first category and an HR > 1 equates to risk reduction for the second category.
^aA large institution indicates a university urology clinic, where patients were diagnosed, and does not include university oncology clinics.

Figure 4. Cumulative OS adjusted by multivariate Cox proportional-hazards regression analysis (Model 1) in TKI- and non-TKI-treated Swedish patients diagnosed with mRCC. HR, hazard ratio; TKI, tyrosine kinase inhibitor.

Figure 5. Cumulative OS adjusted by multivariate Cox proportional-hazards regression analysis (Model 2) by first-line treatment in Swedish patients diagnosed with mRCC. HR, hazard ratio.

early years captured in this study, as physicians were learning the effects of these drugs. There is a reasonable possibility of some selection bias with regard to performance status that could overestimate the effects of TKI treatment. On the other hand, information bias resulting from the fact that the study population included a large fraction of patients with oral drugs dispensed at clinics or patients on intravenous treatment, such as temsirolimus or bevacizumab (also dispensed at clinics), as well as patients treated with IFN before 2005, introduced misclassification of patients as having received other or no treatment. Although these omissions did not affect OS estimates in the Kaplan–Meier

analyses (including the temporal comparisons), sensitivity analysis showed that the omissions affected the multivariate analysis by underestimating the importance of TKI use and blurring the impact of IFN; specifically, selection bias with regard to performance status and combinations with bevacizumab in the later period may have overestimated treatment effects, especially when used in sequence, while IFN prescribed before 2005 is categorised as other/no treatment and likely underestimates the effects of treatment. This implies that some graphs, such as Figures 4 and 5, should be interpreted with caution; however, biases to some extent antagonise each other, making a true

difference in terms of OS likely. Finally, the algorithm used to determine which patients had metastatic disease and the date of mRCC diagnosis has its flaws and we cannot exclude the possibility that (1) non-metastatic patients were included in the analysis and/or (2) some metastatic patients were excluded, despite a thorough optimisation process. Initially, a fourth criterion was used to define the derived date of mRCC diagnosis, namely if, secondary to an RCC diagnosis, a patient visited a urology department ≥ 4 times within a 12-month period. (In Sweden, it is not uncommon for patients living far from oncology clinics to be cared for in a surgical clinic.) Use of this criterion captured patients with a markedly better survival in later years and only a minority of confirmed mRCC patients. Sensitivity Kaplan–Meier and multivariate analyses (Supplementary Figure S1 and Supplementary Table S1), which incorporated this fourth criterion, clearly represented an overestimation of the treatment effects due to a high number of non-metastatic patients. This criterion was therefore omitted from the primary analysis. Furthermore, because the M1 criterion was not available in the Cancer Register from 2000 to 2001 (no registration) and no outpatient visit registration was made during 2000, it was decided to refrain from using patients from these years, as they did not appear to truly reflect a similar composition of mRCC patients, as compared with the other years. This likely means that the results reported here are conservative but consistent and regarded as the best achievable. The true crude OS estimates most probably lie somewhere in between the two models (Figure 3 and Supplementary Figure S1).

In conclusion, the results from our retrospective study are encouraging, consistent with findings from previously reported real-life register studies (Heng *et al*, 2009; Shek *et al*, 2012), and would appear to justify the optimism stemming from use of targeted agents. In addition, this provides verification of treatment benefit in a general population beyond clinical trials – evidence that is sparse in the mRCC literature. Specifically, we showed that the risk of death was lowered, that is, OS was improved, in Swedish patients diagnosed with RCC and mRCC, between 2006 and 2008 compared with before 2006, findings that are consistent with recent developments in diagnostic and therapeutic approaches for RCC. Although the observed increase in OS is multifactorial in origin, including improved palliative care, a contribution of targeted therapies is highly probable, as evidenced by the results in which prescription of any TKI was associated with improved OS in mRCC patients. However, the data represent the early years of TKI treatment in Sweden, primarily with sunitinib and sorafenib. More widespread use of targeted agents, the existence of additional treatments not available in this retrospective study (e.g., everolimus and pazopanib), as well as increased clinician experience, may result in further improved patient outcomes. Although the availability of new mRCC agents has not, as far as we know, provided a cure for patients, it has offered the potential to qualitatively extend life two to three times longer than was possible with cytokine-based regimens (Coppin *et al*, 2005). Furthermore, the added value of these agents' ability to shrink growing tumours and relieve tumour-related symptoms should not be underestimated from a quality-of-life perspective. Finally, our study highlights the value of health registers, as well as the limitations of real-life data analysis, underscoring the importance of prospective data collection.

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CONFLICT OF INTEREST

UH has reported receiving consultant/advisory fees from GlaxoSmithKline, Pfizer, Bayer, and Novartis, and research funding from Novartis. PS has reported receiving consultant/advisory fees from GlaxoSmithKline, Roche, Pfizer, and Novartis, honoraria from Bayer Schering Pharma, GlaxoSmithKline, Roche, and Pfizer, and research funding from Bayer Schering Pharma and Pfizer. SL has reported receiving honoraria from Roche, GlaxoSmithKline, and Pfizer, and research funding from Wyeth and Pfizer. BL has reported receiving consultant/advisory honoraria from Pfizer, Novartis, GlaxoSmithKline, and Bayer. TW, MJ, and RS are full-time employees of Pfizer. TW and RS hold Pfizer stock. JK is an employee of JK Biostatistics AB (Stockholm, Sweden) and was a paid contractor of Pfizer in the development of this manuscript and the analysis and interpretation of the data.

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