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Evaluation of survival benefits by platinum and taxane for an unfavourable subset of carcinoma of unknown primary: a systematic review and meta-analysis

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Background: Although chemotherapeutic regimens containing a taxane or platinum agent have been widely recommended for unfavourable carcinoma of unknown primary (CUP), no evidence exists for the superiority of any administered regimens. To date, the efficacy has been mostly assessed in the limited setting of phase II trials, and few attempts have been made to synthesise all available data for survival outcomes.

Methods: Electronic databases were searched from 1980 to 2011. Survival results were combined for each pre-specified category of regimens using a random-effects model, and meta-regression models were used to adjust for heterogeneity in some known prognostic factors.

Results: A total of 32 studies were included for meta-analysis. Tendency towards better survival outcome by platinum or taxane was indicated. After adjustment for important prognostic factors, however, the difference between the platinum-based and non-platinum regimens became no longer significant. Survival benefits by the taxane-based regimens remained significant, with a prolonged median survival time of 1.52 months ($P=0.03$) and a higher 1-year survival rate of 6.25% ($P=0.05$), but the benefit did not sustain for 2 years.

Conclusion: Although no effective therapies have been established, this meta-analysis helps to fill an important gap of evidence. However, caution should still be taken because of the potential unmeasured confounding.

The European Society for Medical Oncology guideline lists recommended chemotherapeutic approaches containing platinum and taxane as commonly used low-toxicity chemotherapy regimens, but stated that no evidence exists for superior efficacy

of any of the administered regimens for unfavourable carcinoma of unknown primary (CUP) patients (Fizazi *et al*, 2011). Although CUP is a relatively common metastatic cancer constituting 3–5% of all human malignancies (McCredie *et al*, 1991; Levi *et al*, 2002), the

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prognosis of CUP is very poor, with a median survival of 3–4 months, and its 5-year survival rate is <10% (Abbruzzese *et al*, 1994). Certain CUPs, which are defined as favourable prognostic subsets, show a biology and behaviour similar to known metastatic primary carcinoma (Greco and Pavlidis, 2009). However, patients in the favourable subset constitute only a minority of CUP cases and most patients remain unresponsive to any treatment modalities and have a worse prognosis. As no effective therapies have been established through randomised controlled trials (RCTs) for such patients, empirical systemic treatments considering the performance status of patients are widely accepted.

Some efforts for quantitative integration of survival outcomes of previous clinical studies have been made. A descriptive summary estimated a 9-month median survival time for platinum-based chemotherapy by simply taking the median of the included studies' quoted median survival times (Pentheroudakis *et al*, 2009). A recent multiple treatment comparison meta-analysis based on a small number of RCTs on chemotherapy regimens in unfavourable CUP patients reported great uncertainty about the survival benefit of any particular regimen (Golfinopoulos *et al*, 2009). As the efficacy of chemotherapy for patients with unfavourable CUP has mostly been assessed in the setting of small phase II trials without controls, data from limited numbers of RCTs are insufficient evidence for the superiority of any particular regimen in terms of survival prolongation. Another meta-analysis that combined the response rate from 29 phase II trials showed that the chemotherapeutic regimens were not the sole significant parameters and recommended multivariate analysis to consider the heterogeneity of prognostic factors of previous CUP trials (Adenis *et al*, 2010).

Thus, this study critically evaluated all relevant studies on chemotherapeutic approaches to unfavourable CUPs, which were mostly from single-arm phase II trials, and evaluated whether platinum- or taxane-based regimens could improve survival in patients from the unfavourable subset of CUPs adjusting for some known prognostic factors.

METHODS

Search strategy. Ovid MEDLINE, EMBASE, and the Cochrane Central Register of Controlled Trials were searched comprehensively without language restrictions, with hand searches for further references. The search period was from 1980, when platinum was introduced into the market, to 2011.

We first searched a combination of keywords: (cancer OR carcinoma OR malignancy OR tumour OR neoplasm) AND (unknown primary OR occult primary OR unknown origin) AND (therapy OR therapeutic OR intervention). A team of reviewers in pairs screened the title and abstract of each article. Any discrepancies in the screening were resolved by re-evaluation and discussion. Each full text of the articles selected by the screening was then reviewed to determine its final inclusion.

Inclusion and exclusion criteria. Studies on chemotherapy for the unfavourable subset of CUP were included. Literature on the favourable subsets, such as papillary adenocarcinoma of the peritoneal cavity in women, poorly differentiated carcinoma with a midline distribution, poorly differentiated neuroendocrine carcinomas, adenocarcinoma involving only the axillary lymph nodes in women, squamous cell carcinoma involving cervical lymph nodes, blastic bone metastases and elevated prostate-specific antigen in men, and a single, small, potentially resectable tumour, were excluded. For the diagnosis of CUP, included studies had to evaluate the origin of metastasis at least by abdominal computerised tomography (CT) with either chest CT or plain chest radiography. Phase II and phase III clinical trials were included and consecutive case series performed with the prospective

enrolment of participants following clear inclusion criteria were considered for inclusion. Studies evaluating alternating regimens before disease progression or on second-line chemotherapy were excluded. Duplicate publications were excluded by retaining the one with the longest follow-up or the one that reported the data more comprehensively.

Quality assessment. The Methodological Index for Non-Randomised Studies (MINORS) (Slim *et al*, 2003) was adopted for quality assessment and applied primarily to all the included studies. Assessments for the Risk of Bias were also conducted for the RCTs using the Cochrane guideline (Higgins, 2008).

Data extraction. A pair of reviewers independently extracted relevant data from the studies using a standardised data extraction form. Information on the type of study design, patient follow-up, prognostic factors, chemotherapeutic regimens, and survival outcomes was abstracted. The survival outcomes were extracted using the methods of Tierney *et al* (2007), and confidence intervals were calculated based on the method of Simon (1986).

Statistical analysis. We first explored a combined estimate of survival outcome for each pre-specified regimen group: platinum *vs* no-platinum; taxane *vs* no-taxane; platinum and taxane (P_1T_1) *vs* platinum and no-taxane (P_1T_0) *vs* no-platinum and taxane (P_0T_1) *vs* no-platinum and no-taxane (P_0T_0). The outcomes considered were the median survival time, and the 1- and 2-year survival rates.

We pooled the results of each outcome for each category by the inverse variance method to generate a crude overall summary with a 95% confidence interval (95% CI) and descriptively contrasted the point estimates between categories. A multiple meta-regression model was then built up for testing the significance of the differences between platinum and no-platinum; taxane and no-taxane; and each of P_1T_1 , P_1T_0 , P_0T_1 , and P_0T_0 after adjustment for covariates. Being male, a histology of moderate- to well-differentiated adenocarcinoma, poor Eastern Cooperative Oncology Group (ECOG) performance status (≥ 2), liver metastasis, and multiple metastatic sites were considered potentially important prognostic factors (Pavlidis and Fizazi, 2005). The year of the study could also influence survival outcomes and was considered a potential source of heterogeneity.

A random-effects model was applied to combine the outcomes. The statistical heterogeneity among the studies was evaluated qualitatively with a forest plot as well as quantitatively with the I^2 measure and χ^2 test. The funnel plots asymmetry was assessed for small study effects.

RESULTS

Description of the included studies. A total of 1389 potentially relevant studies, 1385 from the electronic search and four from the hand searches, were identified, and 1281 were excluded by title/abstract screening. Full texts were retrieved for the remaining 108 studies, and 32 of them met all the criteria for inclusion in the analysis (Figure 1).

There were seven RCTs (Shildt *et al*, 1983; Dowell *et al*, 2001; Assersohn *et al*, 2003; Culine *et al*, 2003; Palmeri *et al*, 2006; Huebner *et al*, 2009; Hainsworth *et al*, 2010) and one non-randomised trial (Saghatchian *et al*, 2001) where two treatment arms were evaluated. Six of the included clinical trials primarily explored or estimated the treatment outcome separately; only one phase III trial had a primary objective of directly comparing the two treatment regimens (Hainsworth *et al*, 2010). Another phase III trial identified by electronic search was excluded because only the abstract was accessible and contacting the author yielded no further information (Gross-Goupil *et al*, 2009). Both the phase III

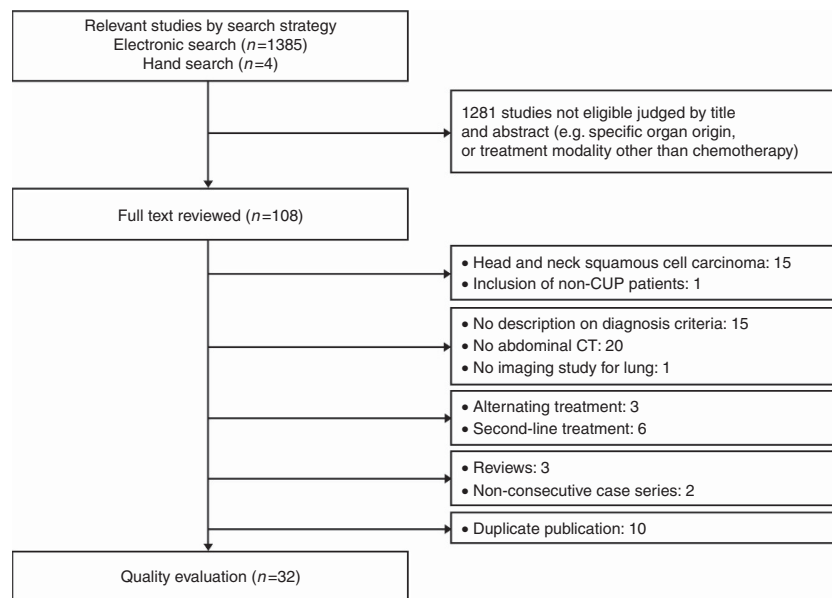


Figure 1. Flow chart of literature selection.

trials were stopped before enrolment of the planned sample size because of slow accrual.

There was one consecutive case series study (Gill *et al*, 1991) and the other 23 studies were single arm clinical trials (Raber *et al*, 1991; Briasoulis *et al*, 1998, 2000, 2008; Parnis *et al*, 2000; Greco *et al*, 2001, 2002; Guardiola *et al*, 2001; Macdonald *et al*, 2002; Balana *et al*, 2003; Mukai *et al*, 2003, 2010; Park *et al*, 2004; Piga *et al*, 2004; Pouessel *et al*, 2004; El-Rayes *et al*, 2005; Pittman *et al*, 2006; Schneider *et al*, 2007; Pentheroudakis *et al*, 2008; Hainsworth *et al*, 2009; Schuette *et al*, 2009; Yonemori *et al*, 2009; Moller *et al*, 2010). Three trials (Briasoulis *et al*, 1998, 2000; Pentheroudakis *et al*, 2008) also involved a small proportion of favourable CUP patients in their studies, and we extracted the data only from the unfavourable subset.

In the quality assessment using MINORS, most studies scored ≥ 10 points; however, three studies conducted before 2000 scored < 10 (Shildt *et al*, 1983; Raber *et al*, 1991; Parnis *et al*, 2000). Two independent reviewers evaluated the responses, which were the primary outcome of the study in only two studies (Culine *et al*, 2003; Moller *et al*, 2010). In all, 14 studies reported prospective calculation of the study size (Greco *et al*, 2002; Macdonald *et al*, 2002; Pouessel *et al*, 2004; El-Rayes *et al*, 2005; Palmeri *et al*, 2006; Pittman *et al*, 2006; Schneider *et al*, 2007; Pentheroudakis *et al*, 2008; Hainsworth *et al*, 2009, 2010; Huebner *et al*, 2009; Schuette *et al*, 2009; Yonemori *et al*, 2009; Mukai *et al*, 2010). Among the RCTs, only one study had a clear description of the method of allocation concealment (Assersohn *et al*, 2003). No RCT was blinded; however, we did not consider it a major threat to the validity since the outcomes should not be influenced by blinding.

Among the 42 regimens investigated, 35 were evaluated in studies conducted since 2000. In all, 26 regimens were tested based on a sample of 30 to 60 patients. The proportions of male patients in most of the studies were 50–75%. The efficacy of 14 regimens was tested with patients among whom 50–75% had moderate- to well-differentiated adenocarcinoma, and 13 regimens with patients among whom $< 50\%$ had the histological attribute. The proportion of patients with liver metastasis was $< 50\%$ in 17 regimens. A total of 24 regimens were assessed with patients among whom $< 20\%$ had an ECOG performance status ≥ 2 . The proportion of patients with multiple metastases was 50–75% in 21 regimens. Only two regimens out of 42 were monotherapies; 34 regimens contained

platinums and 17 included taxanes. Only one regimen had a follow-up duration shorter than 12 months, and 29 regimens were followed for longer than 24 months. The examined regimens and prognostic profiles are summarised in Table 1.

Survival estimation by regimen groups. The combined results by regimen group, 95% CIs, and statistics for heterogeneity are shown in Figure 2. The overall median survival time was 9.0 months (95% CI: 8.1–9.8), 1-year survival rate 35.6% (95% CI: 32.0–39.3), and 2-year survival rate 18.6% (95% CI: 15.4–21.7) across all chemotherapy regimens.

Most regimens (34 out of 42) contained a platinum component, demonstrating that using platinum in CUP treatment has become common. The majority of platinums were cisplatin or carboplatin, and oxaliplatin was used in only 2 out of the 34 regimens. For the meta-analysis, 26 regimens were included in the platinum group and 6 regimens in the non-platinum group. The platinum-based regimens showed a tendency to have better outcomes in the survival: median survival time of 9.4 *vs* 7.2 months; 1-year survival rate of 36.9% *vs* 29.6%; 2-year survival rate of 19.7% *vs* 11.9% for the platinum regimens *vs* non-platinum regimens.

When 16 taxane regimens and 16 non-taxane regimens were each combined, a slightly longer median survival (9.6 months) was estimated for the taxane regimens overall than the non-taxane regimens (8.3 months). A greater 1-year survival rate (41.3% *vs* 30.8%) was observed for the taxane group, and there was no overlapping of 95% CIs between the taxane group and the non-taxane group. However, the tendency towards better results in the taxane group was no longer definite but was sustained for the 2-year survival rate (21.2% *vs* 16.4%).

Among 14 P₁T₁, 12 P₁T₀, 2 P₀T₁, and 4 P₀T₀ regimens, we found P₁T₁ was the best regimen group and P₀T₀ was the worst for all survival outcomes. While a small number of studies were included, P₀T₁ tended to be better than P₁T₀ for median survival time (9.5 *vs* 8.9 months) and 1-year survival rate (36.6% *vs* 32.1%).

Investigation of heterogeneity. No significant association was found between the study year and outcomes shown by meta-regression analyses (Table 2). In the univariable meta-regression, the median survival was related to the histology and ECOG performance status. The 1- and 2-year survival probabilities were

Table 1. The chemotherapeutic regimens and major prognostic factors of the studies included

Author and year of study	Regimens	N	Major prognostic factors, N (%)					Follow-up (months)
			Male sex	MWD	ECOG ≥ 2	Liver meta	Multiple meta	
Shildt <i>et al</i> (1983)	5-Fluorouracil	20	—	—	—	—	—	10.2
Shildt <i>et al</i> (1983)	Cyclophosphamide, 5-fluorouracil, doxorubicin	16	—	—	—	—	—	24.5
Raber <i>et al</i> (1991)	Cisplatin, 5-fluorouracil, etoposide	46	—	26 (56.5)	2 (4.3)	—	—	—
Gill <i>et al</i> (1991)	Cisplatin, etoposide	16	9 (56.3)	6 (37.5)	—	5 (31.3)	10 (62.5)	28.75
Briasoulis <i>et al</i> (1998)	Carboplatin, etoposide, epirubicin	45	—	—	—	—	—	40
Pamis <i>et al</i> (2000)	Cisplatin, 5-fluorouracil, epirubicin	43	27 (62.8)	—	7 (16.3)	16 (37.2)	31 (72.1)	54
Briasoulis <i>et al</i> (2000)	Carboplatin, paclitaxel	33	18 (54.5)	22 (66.7)	—	—	—	47
Dowell <i>et al</i> (2001)	Carboplatin, etoposide	17	10 (58.8)	14 (82.4)	6 (35.3)	11 (64.7)	11 (64.7)	17.3
Dowell <i>et al</i> (2001)	Paclitaxel, 5-fluorouracil, leucovorin	17	10 (58.8)	11 (64.7)	3 (17.6)	10 (58.8)	10 (58.8)	16
Guardiola <i>et al</i> (2001)	Cisplatin, doxorubicin, cyclophosphamide	22	13 (59.1)	8 (36.4)	—	6 (27.3)	12 (54.5)	56.9
Greco <i>et al</i> (2001)	Carboplatin, paclitaxel, etoposide	71	35 (49.3)	34 (47.9)	12 (16.9)	—	43 (60.6)	50
Greco <i>et al</i> (2001)	Cisplatin, docetaxel	26	13 (50.0)	13 (50.0)	6 (23.1)	—	19 (73.1)	33
Greco <i>et al</i> (2001)	Carboplatin, docetaxel	47	25 (53.2)	18 (38.3)	12 (25.5)	—	32 (68.1)	24
Saghatchian <i>et al</i> (2001)	Cisplatin, etoposide, ifosfamide, bleomycin	30	15 (50.0)	0 (0)	7 (23.3)	3 (10)	17 (56.7)	32
Saghatchian <i>et al</i> (2001)	Cisplatin, 5-fluorouracil, interferon	18	6 (33.3)	18 (100)	1 (5.6)	6 (33.3)	8 (44.4)	32
Macdonald <i>et al</i> (2002)	Cisplatin, 5-fluorouracil, mitomycin	31	18 (58.1)	22 (71.0)	5 (16.1)	7 (22.6)	16 (51.6)	53
Greco <i>et al</i> (2002)	Carboplatin, paclitaxel, gemcitabine	120	64 (53.3)	63 (52.5)	16 (13.3)	51 (42.5)	78 (65)	27
Culine <i>et al</i> (2003)	Cisplatin, gemcitabine	39	26 (66.7)	26 (66.7)	6 (15.4)	21 (53.8)	29 (74.4)	22
Culine <i>et al</i> (2003)	Cisplatin, irinotecan	40	18 (45.0)	24 (60.0)	7 (17.5)	21 (52.5)	30 (75.0)	22
Assersohn <i>et al</i> (2003)	5-Fluorouracil	45	28 (62.2)	15 (33.3)	12 (26.7)	30 (66.7)	22 (48.9)	24
Assersohn <i>et al</i> (2003)	5-Fluorouracil, mitomycin	43	20 (46.5)	18 (41.9)	15 (34.9)	21 (48.8)	17 (39.5)	24
Balana <i>et al</i> (2003)	Cisplatin, gemcitabine, etoposide	40	25 (62.5)	15 (37.5)	4 (10.0)	6 (15.0)	27 (67.5)	29
Mukai <i>et al</i> (2003)	Cisplatin, docetaxel	5	1 (20.0)	4 (80.0)	1 (20.0)	—	—	—
Piga <i>et al</i> (2004)	Carboplatin, etoposide, doxorubicin	102	54 (52.9)	38 (37.3)	8 (7.8)	27 (26.5)	89 (87.3)	60
Pouessel <i>et al</i> (2004)	Docetaxel, gemcitabine	36	23 (63.9)	16 (44.4)	3 (8.3)	15 (41.7)	26 (72.2)	32
Park <i>et al</i> (2004)	Cisplatin, paclitaxel	37	20 (54.1)	31 (83.8)	8 (21.6)	12 (32.4)	35 (94.6)	31.5
El-Rayes <i>et al</i> (2005)	Carboplatin, paclitaxel	22	13 (59.1)	15 (68.2)	2 (9.1)	17 (77.3)	—	23.2
Palmeri <i>et al</i> (2006)	Cisplatin, paclitaxel, gemcitabine	33	25 (75.8)	25 (75.8)	1 (3.0)	8 (24.2)	14 (42.4)	20.5
Palmeri <i>et al</i> (2006)	Cisplatin, gemcitabine, vinorelbine	33	27 (81.8)	23 (69.7)	1 (3.0)	10 (30.3)	15 (45.5)	21.5
Pittman <i>et al</i> (2006)	Carboplatin, gemcitabine	50	23 (46.0)	—	12 (24.0)	38 (76.0)	—	35
Schneider <i>et al</i> (2007)	Carboplatin, gemcitabine, capecitabine	33	19 (57.6)	22 (66.7)	9 (27.3)	20 (60.6)	10 (30.3)	54.1
Pentheroudakis <i>et al</i> (2008)	Carboplatin, docetaxel	23	12 (52.2)	—	3 (13.0)	9 (39.1)	13 (56.5)	48
Briasoulis <i>et al</i> (2008)	Oxaliplatin, irinotecan	47	33 (70.2)	19 (40.4)	11 (23.4)	27 (57.4)	27 (57.4)	35.5
Huebner <i>et al</i> (2009)	Carboplatin, paclitaxel	42	23 (54.8)	—	5 (11.9)	23 (54.8)	36 (85.7)	26
Huebner <i>et al</i> (2009)	Gemcitabine, vinorelbine	45	27 (60.0)	—	10 (22.2)	29 (64.4)	37 (82.2)	22
Schuette <i>et al</i> (2009)	Oxaliplatin, capecitabine	51	35 (68.6)	7 (13.7)	5 (9.8)	20 (39.2)	—	24
Yonemori <i>et al</i> (2009)	Carboplatin, irinotecan	45	23 (51.1)	21 (46.7)	4 (8.9)	8 (17.8)	32 (71.1)	48
Hainsworth <i>et al</i> (2009)	Carboplatin, paclitaxel, bevacizumab, erlotinib	60	29 (48.3)	31 (51.7)	0	—	30 (50.0)	25
Mukai <i>et al</i> (2010)	Cisplatin, docetaxel	45	23 (51.1)	19 (42.2)	6 (13.3)	—	10 (22.2)	57.1
Moller <i>et al</i> (2010)	Cisplatin, paclitaxel, gemcitabine	98	51 (52.0)	—	0	49 (50.0)	53 (54.1)	80
Hainsworth <i>et al</i> (2010)	Carboplatin, paclitaxel, etoposide, gefitinib	93	44 (47.3)	49 (52.7)	9 (9.7)	50 (53.8)	75 (80.6)	60
Hainsworth <i>et al</i> (2010)	Gemcitabine, irinotecan, gefitinib	105	63 (60.0)	56 (53.3)	8 (7.6)	54 (51.4)	73 (69.5)	60

Abbreviations: MWD = moderate- to well-differentiated adenocarcinoma; ECOG = Eastern Cooperative Oncology Group performance status; meta = metastasis.

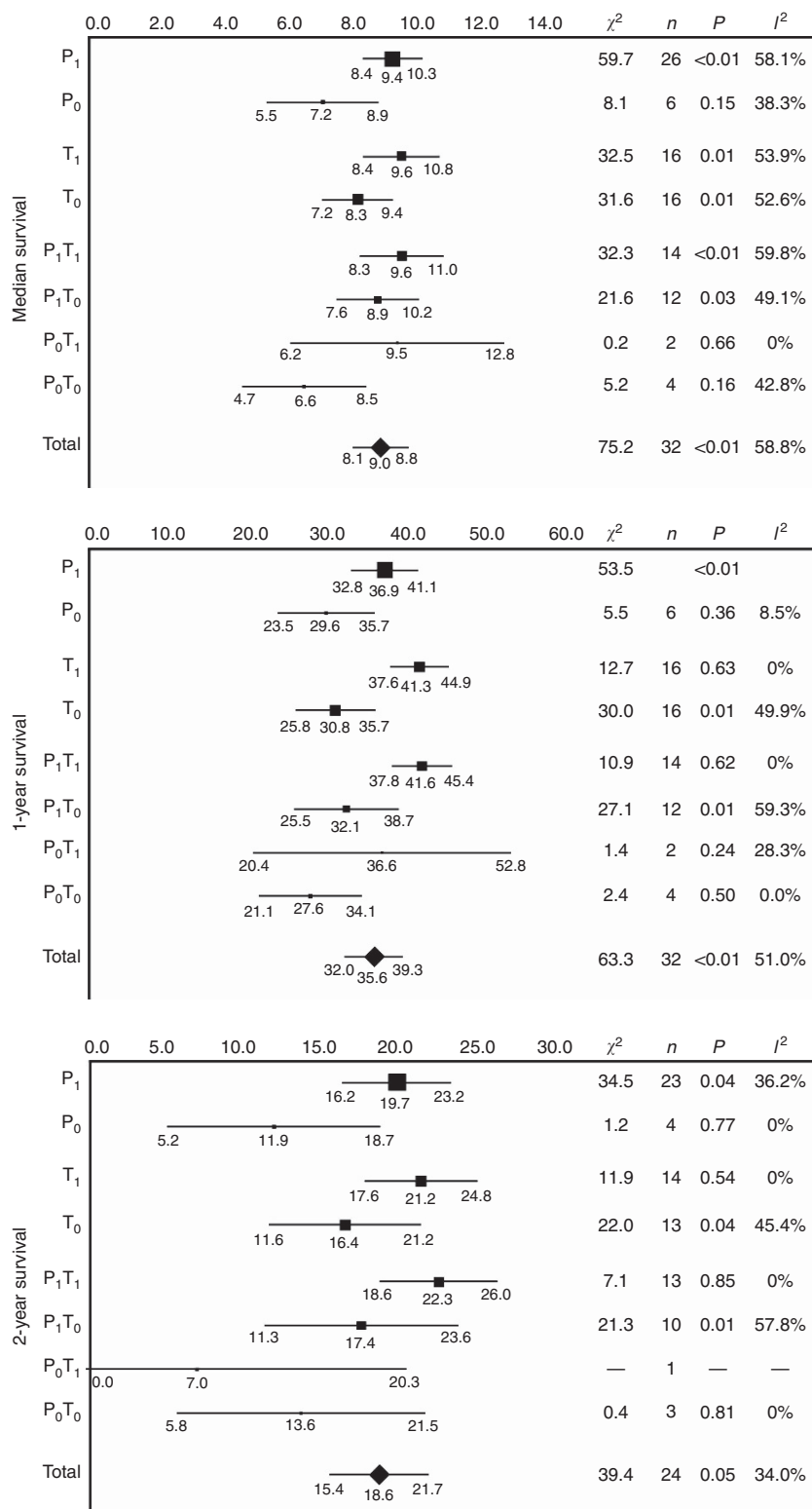


Figure 2. Combined estimates for the survival by treatment strategy. χ^2 , chi-squared statistics; n, the number of studies that were included to the regimen groups, I^2 , I-squared statistics; P₁, platinum; P₀, no platinum; T₁, taxane; T₀, no taxane; P₁T₁, platinum and taxane; P₁T₀, platinum and no taxane; P₀T₁, no platinum and taxane; P₀T₀, no platinum and no taxane; P-value is for the χ^2 test.

related to the gender, ECOG performance status, and whether liver metastasis occurred.

Assessment of treatment differences by regimen. The differences we observed in the descriptive comparison of the point estimates of crude overall summaries by platinum became less noticeable and

tended to be null after adjustment of the prognostic factors (Table 3). However, the difference in the median survival time by the taxane-containing combination increased to 1.52 months after adjustment for covariates and was statistically significant ($P=0.03$). The obvious superiority of the 1-year survival rate with taxane-based regimens remained significant ($P=0.05$), and a

Table 2. Relationship between prognostic factors and survival outcomes

Characteristics	Median survival		1-Year survival		2-Year survival	
	Coefficient ^a	P-value	Coefficient ^a	P-value	Coefficient ^a	P-value
Year of the study	-0.01	0.93	-0.11	0.85	-0.40	0.43
Gender	5.29	0.21	30.23	0.07 ^b	31.71	0.02 ^b
Histology	5.90	0.07 ^b	21.19	0.18	10.24	0.44
ECOG performance	-9.77	0.02 ^b	-62.17	<0.01 ^b	33.62	0.05 ^b
Liver metastasis	-3.78	0.15	-27.1	0.01 ^b	-26.61	<0.01 ^b
Multiple metastasis	-1.67	0.49	10.98	0.29	-7.34	0.40

Abbreviation: ECOG = Eastern Cooperative Oncology Group performance status.
^aCoefficient represents a linear relationship between each prognostic factor and the survival outcome by a univariable meta-regression model. The sign of the coefficient indicates the positivity or negativity of the relationship. The P-value of each coefficient indicates whether the linear correlation was significant or not.
^bP-value <0.1.

Table 3. Difference between subgroups adjusting prognostic factors as appropriate

Outcomes	Regimens	Coefficient (95% confidence interval) ^a	P-value
Median survival (months) ^b	Platinum	0.76 (-1.14 to 2.67)	0.43
	Taxane	1.52 (0.12 to 2.92)	0.03
	No-platinum, no-Taxane	Reference	
	Platinum, no-Taxane	0.78 (-1.16 to 2.72)	0.43
	No-platinum, taxane	2.58 (-1.09 to 6.26)	0.17
	Platinum and taxane	2.02 (-0.05 to 4.09)	0.06
1-Year survival (%) ^c	Platinum	-2.68 (-10.95 to 5.58)	0.52
	Taxane	6.25 (-0.05 to 12.55)	0.05
	No-platinum, no-taxane	Reference	
	Platinum, no-taxane	-2.93 (-11.76 to 5.89)	0.52
	No-platinum, taxane	8.28 (-6.93 to 23.48)	0.29
	Platinum and taxane	3.14 (-6.47 to 12.76)	0.52
2-Year survival (%) ^c	Platinum	2.21 (-5.93 to 10.35)	0.60
	Taxane	0.98 (-5.58 to 7.53)	0.77
	No-platinum, no-taxane	Reference	
	Platinum, no-taxane	-2.30 (-12.04 to 7.43)	0.64
	No-platinum, taxane	-11.63 (-28.04 to 5.63)	0.19
	Platinum and taxane	0.01 (-11.34 to 11.36)	0.99

Abbreviation: ECOG = Eastern Cooperative Oncology Group performance status.
^aThe coefficient from a multiple meta-regression model after adjustment for prognostic factors represents changes of the survival outcome by adding the particular agent(s) to a regimen in comparison to one without them or to the reference.
^bAdjusted for proportion of moderate- to well-differentiated adenocarcinoma, and proportion of performance status (in ECOG \geq 2).
^cAdjusted for proportion of male patients, proportion of performance status (in ECOG \geq 2), and liver metastasis.

possible treatment difference of 6.25% remained. No statistically significant survival benefit was observed from the comparison of the regimens containing platinum, taxane, or both to the chemotherapies with agents other than platinum or taxanes, except for the marginal significance of regimens containing both platinum and taxanes for a median survival time of 2.02 months ($P = 0.06$). The results also suggest a decrease in the 2-year survival rate by taxane regimens without a platinum agent, although it was not statistically significant. The only regimens to have positive results, although not reaching statistical significance, at all end points are the regimens containing both platinum and taxane in comparison to regimens containing neither of them.

Comparison of survival rates from randomised controlled clinical trials. None of the five RCTs providing data for calculating a comparative effect measure showed a hazard ratio (HR) significantly different from one for the overall survival rate (Figure 3). Apart from one out of the five studies, all studies involved at least one regimen containing taxane or platinum. Among the four studies comparing regimens containing taxane with others without taxane, two of the studies showed a direction of effect contrasting with the other two studies. The pooled HR was 0.95 (95% CI: 0.65–1.26), suggesting a nonsignificant difference. In examining the regimens for comparison of platinum against non-platinum, there was also heterogeneity in the direction of treatment

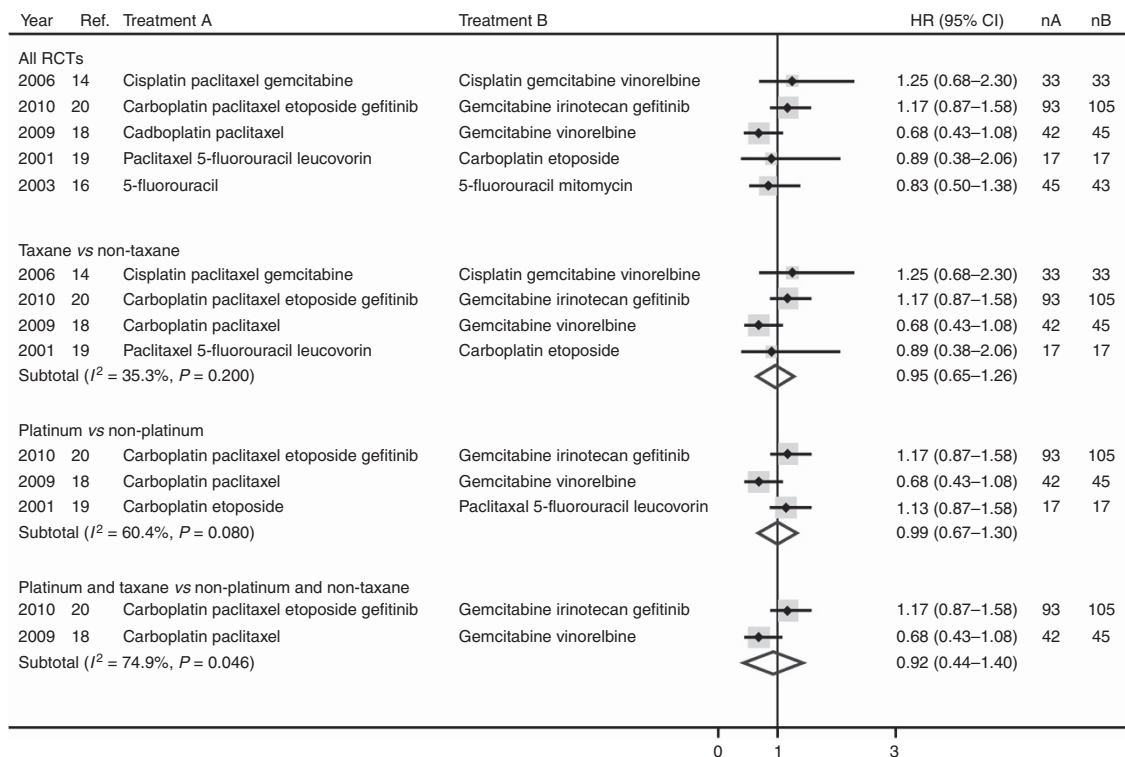


Figure 3. Comparison of survival rates from the randomised controlled clinical trials. The hazard ratio above 1 means that the risk of death by the first regimen is higher than that by the second regimen. nA, number of patients in the treatment A; nB, number of patients in the treatment B.

effect, which produced a pooled HR of 0.99 (95% CI: 0.67–1.30). For the comparison of regimens containing both taxane and platinum against neither of them, there was a significant heterogeneity between the two included results and they produced a pooled HR of 0.92 (95% CI: 0.44–1.40).

Assessment for funnel plot asymmetry. We observed an asymmetric funnel plot in only the median survival time ($P = 0.05$), but no apparent trend in the 1- or 2-year survival rate. The former indicated an association between the treatment results and study size (Figure 4). We performed the Egger test for the median survival time stratified by taxanes and platinum. The small study effect was more salient in the platinum regimens ($P = 0.04$) than non-platinum regimens ($P = 0.33$) and in non-taxane regimens ($P = 0.10$) than taxane regimens ($P = 0.96$).

DISCUSSION

An important target of treating cancer patients should be prolonging survival. However, for the unfavourable subset of CUP patients, there has been no clear understanding of the survival benefits by any chemotherapeutic treatment so far, although platinum- or taxane-based regimens have been commonly and empirically used in current clinical practice. The aim of our study was to estimate the survival efficacy of those regimens by systematically evaluating all relevant data in the literature.

In a recently issued guidance by the National Institute for Health and Clinical Excellence on metastasis with unknown primary, a combination of carboplatin and paclitaxel was suggested to have the highest total expected quality-adjusted life years among the therapeutic approaches, but the probability that the combination was cost-effective was $< 50\%$ at a given willingness to pay threshold in current UK clinical practice (Fowell, 2010; MacReady, 2010). The guidance also indicated a lack of information

supporting any particular chemotherapy in terms of survival prolongation in patients with CUP not belonging to a specific favourable group in the currently available literature. Our study suggested that the addition of a taxane to the regimen extended the median survival time by 1.52 months (95% CI: 0.12–2.92) and increased the 1-year survival rate by 6.25% (95% CI: 0.05–12.55). Although the survival benefit did not remain for 2 years, the suggested improvement in the median survival time and the 1-year survival probability, accounting for the heterogeneous settings, can provide clinically meaningful evidence of the survival benefit by including taxane in combinatorial chemotherapy.

The majority of platinum in the included regimens were cisplatin or carboplatin, but oxaliplatin was also used in 2 out of the 34 platinum regimens. Oxaliplatin has a particular indication (Cvitkovic, 1998; Graham and Cassidy, 2012) and may better be considered separately from other platinum agents in clinical use, whereas cisplatin and carboplatin can, for the most part, substitute for each other. For this reason, we conducted a reanalysis by removing the two oxaliplatin regimens from the platinum group, which had little impact on the results but a slight shift towards increasing survival: median survival time from 9.4 months (95% CI: 8.4–10.3) to 9.5 months (95% CI: 8.5–10.5); 1-year survival rate from 36.9% (95% CI: 32.8–41.1) to 38.9% (95% CI: 35.9–42.0); and 2-year survival rate from 19.7% (95% CI: 16.2–23.2) to 20.7% (95% CI: 17.9–23.6). Consequently, the differences between subgroups presented in Table 3 were not influenced much by exclusion of oxaliplatin from the platinum group.

Because RCTs provide another level of evidence, we also analysed them separately by categorising them to assess the effects of containing taxane, platinum, or both. Five RCTs were available with data for calculation of HR. The categorised pooled results suggested slightly favourable results toward taxane, platinum, or both in comparison to regimens without any of them, but failed to confirm the effect, mainly due to heterogeneity between the small number of small-sized studies. A previous meta-analysis of RCTs

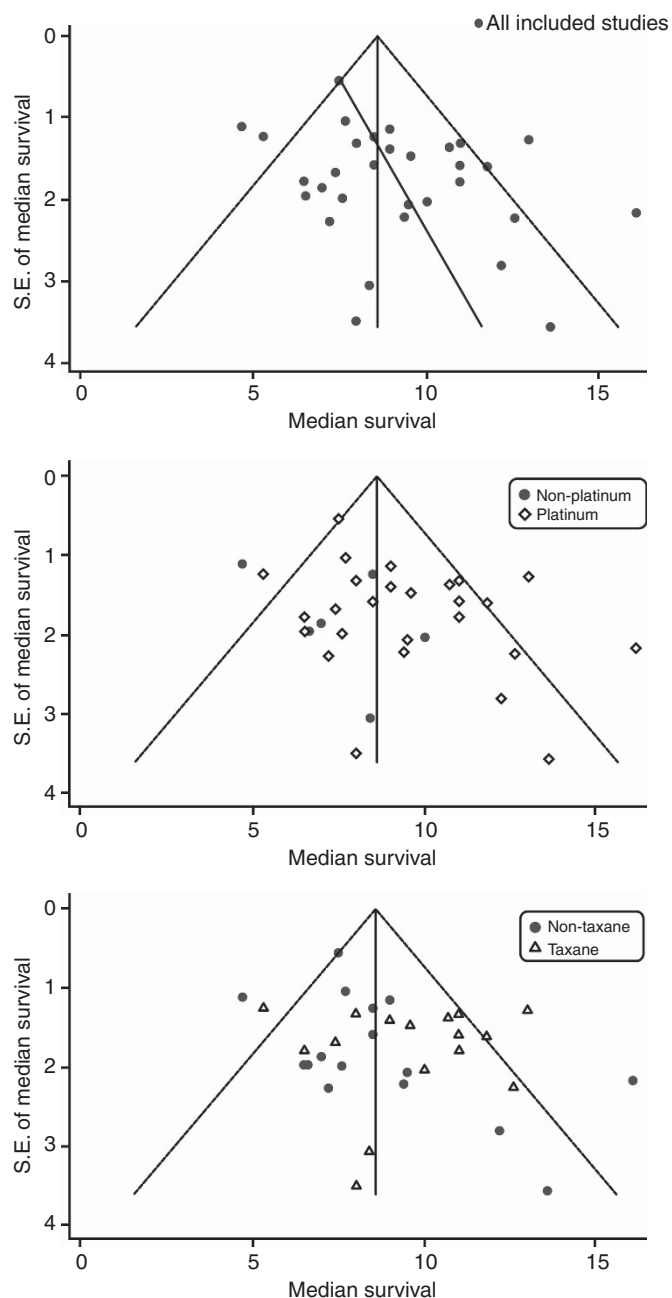


Figure 4. Funnel plot for evaluation of the small trial effect for survival. The fitted line corresponds to the regression test for funnel plot asymmetry proposed by Egger *et al* (1997).

conducted a multiple treatment comparison, a kind of indirect comparison among different treatments using a Bayesian statistical method, which also failed to demonstrate a survival benefit by any chemotherapeutic regimen for CUP (Golfopoulos *et al*, 2009). The analysis also suggested that the hazard ratios favoured platinum (0.69, 95% CI: 0.39–1.28); taxane (0.66, 95% CI: 0.22–2.08); or both (0.81, 95% CI: 0.34–1.89), in comparison with a monotherapy regimen with an agent other than platinum or taxane. Their analysis included a larger number of RCTs because of less restrictive inclusion criteria applied, in terms of method of diagnosis for CUP, than those used in our study. Their results also suggested a great uncertainty with a wider Bayesian version of CIs. Lack of a large enough number of RCTs, especially for phase III, could be a possible explanation of the failure to demonstrate

significance, although the possible efficacy of taxane or platinum was suggested consistently. From the five trials included in our meta-analysis of RCTs, only one trial was a study conducted to evaluate a comparative efficacy in the setting of phase III, but the study also stopped early because of slow accrual of patients. The other trials were randomised studies, but still were phase II studies, which aimed to assess efficacy in each treatment arm rather than for investigating a comparative efficacy. In such a circumstance, this study is of importance by filling the lack of certainty in the current evidence by utilising all available information with a maximum control of incomparability.

The funnel plot of the median survival showed an asymmetry indicating a small study effect. The scatter plot of survival outcome against its standard error is supposed to have a symmetric shape of an inverse funnel, representing a wide variability of estimates from small studies and narrow range of those from large studies. Thus, emptiness in the left base of the funnel suggests that small studies have a tendency to report favourable outcomes, which is referred as the small study effect and is often interpreted as an indication of publication bias. However, it is also a common phenomenon in the early development of a new treatment (Lau *et al*, 2006). The observed small study effect in the current analysis may also be due to such a true heterogeneity (Higgins, 2008), as a short-term benefit from chemotherapy is more likely in patients with high-performance status for tolerating toxicity of the treatment and such patients are more likely to be included in small, early phase clinical trials in the development of new treatment regimens. The stratified funnel plots and Egger test results suggest that the small study effect was more noticeable in the non-taxane or platinum regimens, which may imply that the survival benefits by taxanes were estimated rather conservatively, while those by platinum were still nonsignificant even with possible overestimation.

Our study results should be interpreted with caution because of several limitations. In this setting, a selection of studies with similar populations is particularly difficult, because of the heterogeneous nature of the CUP population, the availability of mostly single institute-based small studies, and the long interval from which the studies were selected. Although we tried to adjust for the differences in known prognostic factors among studies using a meta-regression in the evaluation of the treatment effect, the limitation of this study is that comparisons between those studies can barely substitute for RCTs even after correction for some characteristics across trials. This meta-analysis could still be susceptible to some bias from unmeasured confounders potentially influencing the overall survival, for example, palliative care offered or second-line treatment. Since the included studies have been conducted over a couple of decades, the identification of treatable subsets has changed during the time, which may also cause such heterogeneity as an unmeasurable confounder. Many of the trials, particularly in the first decade, contained some patients who would currently be identified and treated with regimens other than empiric therapy. On the other hand, a significant improvement in the supportive care and the health-care system can also influence the outcomes of the treatments. Most of the included studies were, in fact, found to be carried out within the past decade, and such an evolution effect might therefore have been minimised. We indeed observed lack of a relationship between the survival outcomes and the year of the study (Table 2).

The potential limitations notwithstanding, our study showed that combinations containing a taxane agent demonstrated a prolongation of the median survival time and higher 1-year survival rate in the first-line treatment of unfavourable CUP compared with those without taxanes. However, such a benefit did not seem to be sustained for 2 years. Summaries of available RCTs suggested addition of platinum, taxane, or platinum and taxane to a regimen tends to extend survival of CUP patients; however, those findings based on RCTs were said to be very uncertain owing to a

so small number of RCTs. This systematic review and meta-analysis utilising all available clinical trials fills an important gap of certainty in the current evidence by demonstrating a beneficial effect in the use of taxane for unfavourable CUP patients.

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