Estrogen Receptor 1 Gene (ESR1) is Associated with Restrictive Anorexia Nervosa

Audrey Versini[1,6], Nicolas Ramoz*[1,6], Yann Le Strat[1], Susann Scherag[2], Stefan Ehrlich[3], Claudette Boni[1], Anke Hinney[2], Johannes Hebebrand[2], Lucia Romo[4], Julien-Daniel Guelfi[5] and Philip Gorwood[1,5]

[1]INSERM U894-Team 1, Center of Psychiatry and Neurosciences, University Paris Descartes, 2ter rue d'Alésia, Paris, France; [2]Department of Child and Adolescent Psychiatry, University of Duisburg-Essen, Virchowstr 174, Essen, Germany; [3]Department of Child and Adolescent Psychiatry, Psychosomatic Medicine and Psychotherapy, Charité—Universitätsmedizin Berlin, Campus Virchow-Klinikum, Berlin, Germany; [4]Laboratoire Evaclispy, Université Paris 10 Ouest Nanterre La Défense, Nanterre, France; [5]Sainte-Anne Hospital Center (CHSA), Clinic for Mental and Brain Diseases (CMME), 100 rue de la Santé, Paris, France

Anorexia nervosa (AN) is a highly heritable young-onset psychiatric illness the etiology of which remains unknown. Estrogen alpha and beta receptors, encoded by ESR1 and ESR2 genes, are involved in food intake regulation and eating behavior, and may have a potential role in AN. We performed a family-based association study of 17 single-nucleotide polymorphisms (SNPs) encompassing ESR1 and ESR2 genes in a cohort of 321 French AN families. We attempted to replicate this finding in a cohort of 41 restrictive AN (RAN) families and in a population-based study of 693 young women. Using the transmission disequilibrium test, a significant over-transmission was detected between AN and ESR1 rs726281 and rs2295193. These SNPs and another among ESR1 were more specifically associated with the RAN subtype (rs726281, $p = 0.005$, odds ratio (OR) = 2.1, 95% confidence interval (95% CI) = 1.2–3.6; rs3798577, $p = 0.021$, OR = 1.6, 95% CI = 1.1–2.3; and rs2295193, $p = 0.007$, OR = 1.7, 95% CI = 1.2–2.5). A large eight-SNPs haplotype of ESR1 gene was also associated with AN ($p < 0.0001$, OR = 3.1, 95% CI = 1.8–5.1). Association of ESR1 SNPs and RAN was driven by paternal over-transmissions ($p < 0.0001$, OR = 3.7, 95% CI = 1.9–7.3). Furthermore, we confirmed the preferential paternal over-transmission of the ESR1 rs726281 on the independent German sample of 41 RAN trios ($p = 0.025$, OR = 3, 95% CI = 1.1–8.3). Finally, rs3798577 was associated with eating disorders in a population-based sample of 693 women ($p < 0.01$). Our findings are strongly in favor of an association between ESR1 polymorphisms and AN. In particular, ESR1 gene confers a high risk of vulnerability to the restrictive subtype of AN, and suggests that the estrogen pathway has to be further analyzed in AN.

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INTRODUCTION

Anorexia nervosa (AN, MIM 606788) is an eating disorder characterized by weight loss (body mass index (BMI) $<17.5 \, \text{kg/m}^2$) or failure to make expected weight gain during period of growth, intense fear of gaining weight or becoming fat, body image disturbance and amenorrhea (American Psychiatric Association, 1994; Gorwood et al, 2003; Ramoz et al, 2007). The prevalence of AN is between 0.3 and 0.6% (Hudson et al, 2007; Ramoz et al, 2007), with one of the highest mortality rate (approximately 10% per decade) mainly because of cachexia and suicide (Harris and

Barraclough, 1994; Hoek, 2006). Furthermore, the heritability of AN is high (70%), thus increasing the potential interests of genetic approaches (Gorwood et al, 2003; Ramoz et al, 2007). Studies have found evidence of a high genetic risk for AN, especially in the RAN subgroup probably because of heterogeneity reduction of the phenotype, with, therefore, an increased power to detect an association (Grice et al, 2002; Gershon et al, 1984).

According to the DSM-IV, AN patients are subdivided in two categories: restrictive type (RAN), characterizing patients with restricted food intake without binge eating or purging episodes, and binge eating/purging type (BPAN), characterizing patients with binge eating/purging episodes during anorexia and bulimia phases (American Psychiatric Association, 1994). RAN and BPAN patients represent two subtypes with specific clinical features, including significant lower BMI in RAN than in BPAN, and an increased impulsivity and a higher rate of self-harm and suicide in BPAN than RAN (Eddy et al, 2002; Foulon et al, 2007; Milos

et al, 2004). Hormone concentrations may also help to distinguish these two subtypes, because leptin is decreased in RAN compared with BPAN (Eddy et al, 2002).

Biological pathways of sexual hormones were poorly analyzed in AN, although the involvement of the estrogenic pathway in AN is supported by several cues. First a female predominance is observed in patients with a ratio of 9 women to 1–3 men (Hudson et al, 2007; Ramoz et al, 2007). Second, the onset of AN around puberty is correlated to the presence of estrogenic peaks). Third, recent studies suggested that sex hormones have a role in the risk of eating disorders (Culbert et al, 2008; Procopio and Marriott, 2007) but this observation was not replicated by other authors (Raevuori et al, 2008; Baker et al, 2009). Fourth, the anorexic effect of high estrogen levels was found in animal models (Couse and Korach, 1999; Wade and Gray, 1979). Moreover, estradiol levels were lower in anorexia patients than in controls (Ohwada et al, 2007; Brambilla et al, 2003). Fifth, the ESR1 and ESR2 genes, which code for estrogen alpha and beta receptors and are expressed in non-overlapping brain regions (Osterlund et al, 2000; Osterlund and Hurd, 2001), colocalize with corticotrophin-releasing factor and modulate its expression (Dagnault and Richard, 1997; Bao et al, 2005). They participate in the regulation of the hypothalamic pituitary adrenal (HPA) axis (Licinio et al, 1996). Thus, the disruption of the HPA axis reported in AN might be due to an alteration in the estrogen pathway in patient (Van de Stolpe et al, 2004). Finally, the involvement of the estrogen alpha receptor in the regulation of food intake and eating behavior was confirmed in mouse models (Musatov et al, 2006, 2007). Two case–control studies of AN with ESR1 and ESR2 genes were previously published (Eastwood et al, 2002; Rosenkranz et al, 1998). They reported no association of the ESR1 gene with AN, and an inconsistent association of AN with different variants within ESR2 gene.

Thus, our hypothesis is to identify an association between estrogen receptor and AN, especially in RAN subgroup, because a diminution of estradiol level was more important in the restrictive compared with the binging–purging subtype (Ohwada et al, 2007). We more precisely expect to observe an association between AN and ESR1 gene because ovariectomized rat showed an implication of ESR1 but not of ESR2 receptor in food intake, body weight and meal size (Santollo et al, 2007). Finally, we have been also analyzing the role of parental imprinting of this estrogen receptor in AN, because ESR1 belongs to the 6q25 region, which could be subjected to parental imprinting according to database (Morison et al, 2005). Furthermore, the absence of ESR1 transcription could be a result of aberrant methylation of promoter CpG islands, as shown, for example, in breast cancers (Wilson et al, 2008).

In this study, we performed the first family-based association study on a large number of AN families of French origin looking for an association between ESR1 gene and AN, and we took advantage of the narrowing definition of the disorder to show that this association is driven by the restrictive subtype of AN. We attempted to replicate the family-based association in an independent sample of German families (replication of the TDT approach), as well as in an independent population-based sample recruited from the general population.

MATERIALS AND METHODS

Subjects and Phenotype

The study was approved by each national ethics committee. All participants (and if underage, their parents) gave written informed consent.

The first family-based cohort was composed of 321 families of French origin, including 210 complete trios, of whom 102 were described previously (Gorwood et al, 2003). All participants were assessed by clinicians using the face-to-face semi-structured Diagnostic Interview for Genetic Studies (DIGS) (Nurnberger et al, 1994). Diagnosis of AN were made according to DSM-IV criteria (American Psychiatric Association, 1994), and included RAN with no lifetime binge-eating/purging episode ($N = 153$ families), and BPAN with at least one lifetime bingeing/purging episode, during anorexia or bulimia phases ($N = 154$ families). Information for subtype was incomplete for 14 probands.

Furthermore, the eating disorder inventory-2 (Garner et al, 1983) and the Eating Attitude Test-26 (Garner et al, 1982) were used to measure the psychological and behavioral dimensions and the broad range of symptoms of AN respectively.

A second family-based cohort was composed of 41 RAN trios of German origin, the probands being assessed with the Composite International Diagnostic Interview (CIDI) (Robins et al, 1988). The clinical features of the two family-based cohorts and subgroups are described in Table 1.

A third sample was composed of young women (mean age 20.30 ± 0.05 years) with Caucasian origin recruited for a genetic analysis of addictive behaviors, including eating disorders (SAGE study) (Le Strat et al, 2009). Eating disorders were assessed using the SCOFF scale, a brief and reliable questionnaire developed as a screening test of eating disorders in the general population with a sensitivity of 84.6% and a specificity of 89.6% (Hill et al, 2009). Participants were asked the following questions: 1. Do you make yourself sick because you feel uncomfortably full? 2. Do you worry you have lost control over how much you eat? 3. Have you recently lost more than 6 kg in a 3-month period? 4. Do you believe yourself to be fat when others say you are too thin? 5. Would you say that food dominates your life?

Participants were considered as having an eating disorder if they had a SCOFF score equal or above 2 ($n = 126$), and were considered as control if they had a SCOFF score <2 ($n = 567$).

SNPs and Genotyping

Genomic DNA from probands and parents was extracted from peripheral blood leukocytes. Thirteen single-nucleotide polymorphisms (SNPs) encompassing ESR1 gene and four SNPs within ESR2 gene were screened using TaqMan SNP genotyping assays (Applied Biosystems, Les Ullis, France). The selection of SNPs is based on tagged SNPs that depict the haplotype blocks across the genes, according to available databases (Perlegen and Hapmap) and positive associations with diseases.

Table 1 Demographic and Clinical Data for the French and German Family-Based Cohorts and the French RAN and BPAN Subgroups

Population disorder	French			German
	AN	**RAN**	**BPAN**	**RAN**
Families number (N)	321	153	154	41
Female (%)	97	97	97	95
Age (years)	22.3 ± 7.4	19.4 ± 6.5	25.2 ± 7.2[a]	16.0 ± 1.9[b]
Age of onset of anorexia (years)	16.2 ± 4.5	15.6 ± 4.4	16.8 ± 4.4[a]	14.6 ± 1.2
Duration of anorexia (years)	6.1 ± 6.0	3.9 ± 4.8	8.3 ± 6.2[a]	1.5 ± 1.3[b]
BMI minimum (kg/m^2)	13.6 ± 2.2	12.9 ± 1.7	14.4 ± 2.2[a]	13.5 ± 0.2[b]
Major depressive episode (%)	68	68	69	n.a.
Suicide attempt (%)	28	9	47[a]	n.a.
Alcohol and drug abuse/dependence (%)	6	2	11[a]	n.a.
EAT[c] dieting	16.4 ± 10.6	13.4 ± 10.2	19.5 ± 10.3[a]	n.a.
EAT bulimia	7.6 ± 6.0	4.6 ± 4.3	10.6 ± 5.8[a]	n.a.
EAT oral control	7.4 ± 5.3	7.7 ± 5.2	7.2 ± 5.3	n.a.

Abbreviation: n.a., not available.
[a]Significant difference between RAN and BPAN from the French cohort.
[b]Significant difference between French RAN subset and German RAN cohort.
[c]Eating attitude test.
Significant difference indicates a *p*-value inferior at 0.05 in one-way analysis of variance or *t*-test.
Data are indicated as *N*, %, or mean ± SD.

Table 2 Transmission Disequilibrium Test of *ESR1* SNPs with AN in the French Family-Based Cohort

| SNP#	Markers	Position	MAF	TDT			
				OT	**T:U**	**p**	**OR (CI$_{95}$)**
1	rs488133	152167137	T (0.39)	C	102:102	1	1 (0.8–1.3)
2	rs11155819	152241052	C (0.31)	C	105:100	0.727	1.1 (0.8–1.4)
3	rs12199722	152276593	G (0.32)	G	99:88	0.421	1.1 (0.8–1.5)
4	rs1884051	152324972	G (0.3)	A	81:71	0.417	1.1 (0.8–1.6)
5	rs726281	152344271	G (0.26)	A	85:58	**0.024**[a]	1.5 (1.0–2.0)
6	rs3020407	152348954	G (0.29)	A	95:76	0.146	1.3 (0.9–1.7)
7	rs17081994	152358440	C (0.09)	T	36:23	0.091	1.6 (0.9–2.6)
8	rs2982712	152399872	C (0.39)	T	106:95	0.438	1.1 (0.8–1.5)
9	rs3020371	152425513	T (0.31)	C	93:87	0.655	1.1 (0.8–1.4)
10	rs2228480	152461788	A (0.17)	G	55:50	0.626	1.1 (0.8–1.6)
11	rs3798577	152462823	C (0.46)	T	116:91	0.082	1.3 (1.0–1.7)
12	rs2295193	152494787	G (0.47)	G	117:85	**0.024**	1.4 (1.0–1.8)
13	rs2252837	152510513	T (0.35)	C	106:87	0.171	1.2 (0.9–1.6)

Abbreviations: MAF, minor allele and its frequency in the parentheses; OR (CI$_{95}$), odds ratio and the 95% confidence interval; OT, over-transmitted allele; T:U, transmitted *vs* untransmitted allele.
[a]Significant *p*-values are indicated in bold.

Details of SNPs are indicated in Table 2 and Supplementary Table 2 and their position within genes are shown on Figure 1 and Supplementary Figure 1 for *ESR1* and *ESR2* genes, respectively. A total of 362 AN patients and 613 parents were genotyped. Furthermore, 30 samples were genotyped in duplicate to assess the accuracy of the allelic call. SNP rs3798577 was genotyped using SNPlex technology (Applied Biosystems) for 693 women from the population-based sample (Tobler *et al*, 2005).

Statistical Analysis

Comparisons of clinical features between cohorts and subsets were performed using Student's *t*-test and χ^2 test. Hardy–Weinberg equilibrium tests, linkage disequilibrium D' values and minor allele frequencies were computed using Haploview 4.1 software (Barrett *et al*, 2005). Transmission disequilibrium tests (TDTs) for SNPs and haplotypes were carried out with FBAT and PLINK programs (Horvath *et al*,
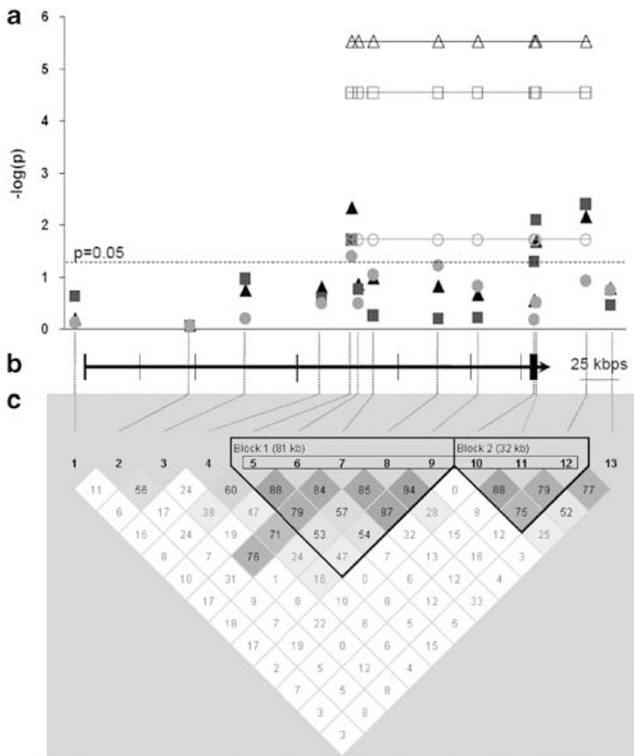
Figure 1 ESR1 SNPs, mapping, and association with restrictive AN. (a) Association between restrictive AN and SNPs (filled symbol) or haplotypes (open symbol) for combined parental (triangle), maternal (circle), and paternal (square) transmissions. The strongest association is observed for the eight-SNPs haplotype ($p = 6.10^{-6}$). (b) Genomic organization of ESR1 gene (bars correspond to exons and arrowhead indicates the orientation of transcription) and position of the 13 encompassing SNPs (interrupted bars) at scale. (c) Pairwise linkage disequilibrium of SNPs. Haplotype blocks 1 and 2 are indicated. The haplotype of eight-SNPs is boxed.

2001; Purcell et al, 2007). Parental origin of the association was calculated using ASPEX 2.5 package (Hinds and Risch, 1996). Gene–gene interaction was computed using the gene-based test of the PLINK program. The power of sample size for association tests was calculated using the Genetic Power Calculator program (http://pngu.mgh.harvard.edu/~purcell/gpc/) (Purcell et al, 2003). To take account the linkage disequilibrium between each SNP, we computed correction for multiple testing for SNPs using Single-Nucleotide Polymorphism Spectral Decomposition (SNPSpD) (http://gump.qimr.edu.au/general/daleN/SNPSpD/) (Nyholt, 2004).

RESULTS

Association of ESR1 Gene with AN and RAN Subset

We first genotyped 17 SNPs, encompassing ESR1 and ESR2 genes, in a French cohort of 321 AN families. All SNPs analyzed in this study were in Hardy–Weinberg equilibrium in the patients and parents. No discrepancy was observed in the genotyping of 30 duplicated samples. Two families were excluded because of Mendelian inheritance errors in the paternal transmissions.

Using TDTs, we found a significant over-transmission in AN with ESR1 SNPs rs726281 ($p = 0.024$, odds ratio (OR) $= 1.5$, 95% confidence interval (95% CI) $= 1–2$) and

rs2295193 ($p = 0.024$, OR $= 1.4$, 95% CI $= 1–1.8$), (Table 2). We showed that RAN and BPAN subjects differed significantly for several clinical variables, particularly regarding the age of onset of anorexia and the minimum lifetime BMI (Table 1). Therefore, we performed TDT analyses according to the RAN and BPAN subsets. Three ESR1 SNPs were significantly over-transmitted in the RAN subset, rs726281 ($p = 0.005$, OR $= 2.1$, 95% CI $= 1.2–3.6$), rs3798577 ($p = 0.021$, OR $= 1.6$, 95% CI $= 1.1–2.3$), and rs2295193 ($p = 0.007$, OR $= 1.7$, 95% CI $= 1.2–2.5$) (Table 3). In contrast, no SNP was associated with the BPAN subgroup (Supplementary Table 1).

No preferential transmission was found for the ESR2 SNPs in any of the AN subgroups (Supplementary Table 2). Furthermore, only one haplotype block encompasses the 112 kb of the ESR2 gene that renders unlikely the association between ESR2 gene with AN, or with the RAN or BPAN subsets (Supplementary Figure 1).

No significant epistatic association was found for combinations of ESR1 and ESR2 SNPs and the AN cohort, neither for the RAN nor for the BPAN subgroups (data not shown).

Haplotype Association in the RAN Subset

Linkage disequilibrium analysis of pairwise SNPs revealed several haplotype blocks across the 296 kb of ESR1 gene (Figure 1c). Over-transmissions in the RAN subset were found for common haplotypes based on five SNPs (block 1: rs726281*A-rs3020407*A-rs17080994*T-rs2982712*T-rs3020371*C, 0.503 frequency, $p = 0.06$), three SNPs (block 2: rs2228480*G-rs3798577*T-rs2295193*G, 0.410 frequency, $p = 0.002$), and the eight SNPs of the two blocks merged (rs726281*A-rs3020407*A-rs17080994*T-rs2982712*T-rs3020371*C-rs2228480*G-rs3798577*T-rs2295193*G, 0.230 frequency, $p = 6.10^{-6}$). This eight-SNPs haplotype presents an OR of 3.1 (95% CI $= 1.8–5.1$) for RAN.

Parental Origin of Association

In addition, we analyzed the parental origin of the association with RAN subset. Considering the combined TDT of the 13 ESR1 SNPs, we detected a highly significant excess of paternal transmission ($p = 9.10^{-7}$) and a significant excess of maternal transmission ($p = 0.0004$). More precisely in RAN, a paternal over-transmission of four SNPs (rs726281 $p = 0.024$, rs2228480 $p = 0.05$, rs3798577 $p = 0.008$ and rs2295193 $p = 0.004$) and a maternal over-transmission of rs726281 only ($p = 0.041$) were observed (Table 3). Interestingly, the rs726281 over-transmitted by both parents is located in haplotype block 1, while the three SNPs solely transmitted in excess by the father constitute block 2 (Figure 1c). Regarding the eight-SNPs haplotype in RAN, in comparison with the maternal over-transmission ($p = 0.02$, OR $= 2.1$, CI95 $= 1.1–3.8$), we found a higher statistical significance for the paternal over-transmission ($p = 3.10^{-5}$, OR $= 3.7$, 95% CI $= 1.9–7.3$) (Figure 1a).

ESR1 Association with an Independent RAN German Cohort

As a second step, we attempted to replicate the association by genotyping ESR1 SNPs in an independent German

Table 3 Transmission Disequilibrium Test of *ESR1* SNPs with Restrictive AN Subset of French Family-Based Cohort

SNP#	Markers	TDT				Maternal TDT		Paternal TDT	
		OT	T:U	*p*	OR (CI$_{95}$)	*p*	OR (CI$_{95}$)	*p*	OR (CI$_{95}$)
1	rs488133	C	56:51	0.629	1.1 (0.8–1.6)	0.758	0.9 (0.5–1.7)	0.237	1.5 (0.8–2.9)
2	rs11155819	C	51:49	0.841	1 (0.7–1.5)	0.866	1.1 (0.5–2.1)	0.873	1.1 (0.6–2.0)
3	rs12199722	G	53:40	0.178	1.3 (0.9–2.0)	0.631	1.2 (0.6–2.2)	0.114	1.7 (0.9–3.2)
4	rs1884051	A	42:30	0.157	2.0 (1–3.8)	0.330	1.4 (0.7–2.6)	0.257	1.5 (0.7–3.3)
5	rs726281	A	48:24	**0.005**[a]	2.1 (1.2–3.6)	**0.041**	2.2 (1.0–4.9)	**0.024**	2.3 (1.1–4.8)
6	rs3020407	A	52:38	0.140	1.4 (0.9–2.1)	0.317	1.4 (0.7–2.7)	0.170	1.6 (0.8–3.2)
7	rs17081994	T	20:11	0.106	1.8 (0.9–3.8)	0.090	2.4 (0.8–6.8)	0.564	1.4 (0.4–4.4)
8	rs2982712	T	61:46	0.147	1.3 (0.9–1.9)	0.064	1.8 (1.0–3.4)	0.647	1.2 (0.6–2.1)
9	rs3020371	C	53:41	0.216	1.3 (0.9–1.9)	0.150	1.6 (0.8–3.0)	0.622	1.2 (0.6–2.2)
10	rs2228480	G	31:23	0.276	1.3 (0.8–2.3)	0.670	0.8 (0.4–1.9)	**0.050**	2.3 (1.0–5.2)
11	rs3798577	T	66:42	**0.021**	1.6 (1.1–2.3)	0.307	1.4 (0.8–2.4)	**0.008**	2.4 (1.2–4.7)
12	rs2295193	G	68:40	**0.007**	1.7 (1.2–2.5)	0.123	1.6 (0.9–3.0)	**0.004**	2.6 (1.3–5.3)
13	rs2252837	C	57:43	0.162	1.3 (0.9–2.0)	0.182	1.6 (0.8–3.1)	0.355	1.3 (0.7–2.5)

Abbreviations: OR (CI$_{95}$), odds ratio and the 95% confidence interval; OT, over-transmitted allele; T:U, transmitted *vs* untransmitted allele.
[a]Significant *p*-values are indicated in bold.

Table 4 Transmission Disequilibrium Test of *ESR1* SNPs with Restrictive AN of German Family-Based Cohort

SNP#	Markers	MAF	Parental TDT				Maternal TDT			Paternal TDT		
			OT	T:U	*p*	OR (CI$_{95}$)	T:U	*p*	OR (CI$_{95}$)	T:U	*p*	OR (CI$_{95}$)
4	rs1884051	G (0.30)	G	20:15	0.398	1.3 (0.7–2.6)	09:09	1	1 (0.4–2.5)	11:06	0.225	1.8 (0.7–5.0)
5	rs726281	G (0.27)	G	23:14	0.139	1.6 (0.8–3.2)	08:09	0.808	0.9 (0.3–2.3)	15:05	**0.025**[a]	3 (1.1–8.3)
6	rs3020407	G (0.33)	A	17:16	0.862	1.1 (0.5–2.1)	11:06	0.225	1.8 (0.7–5.0)	06:10	0.317	0.6 (0.2–1.7)
7	rs17081994	C (0.09)	C	08:05	0.405	1.6 (0.5–4.9)	01:02	0.564	0.5 (0–5.5)	06:02	0.157	3 (0.6–14.9)
9	rs3020371	T (0.32)	T	22:16	0.330	1.4 (0.7–2.6)	09:08	0.808	1.1 (0.4–2.9)	12:07	0.251	1.7 (0.7–4.4)
10	rs2228480	A (0.19)	G	15:13	0.705	1.2 (0.5–2.4)	08:06	0.593	1.3 (0.5–3.8)	08:07	0.796	1.1 (0.4–3.2)
11	rs3798577	C (0.40)	T	23:13	*0.096*[b]	1.8 (0.9–3.5)	10:08	0.637	1.3 (0.5–3.2)	13:05	0.059	2.6 (0.9–7.3)
12	rs2295193	A (0.45)	–	21:21	1	1 (0.5–1.8)	11:11	1	1 (0.4–2.3)	10:10	1	1 (0.4–2.4)

Abbreviations: MAF, minor allele and its frequency in the parentheses; OR (CI$_{95}$), odds ratio and the 95% confidence interval; OT, over-transmitted allele; T:U, transmitted *vs* untransmitted allele.
[a]Significant *p*-values are indicated in bold.
[b]Trend of significance for *p*-values are indicated in italic.

cohort of 41 RAN trios (Table 4). Although no allele was transmitted in excess in the probands of this replication cohort, we found a preferential paternal over-transmission of the same SNP, which was initially associated in our sample, namely the rs726281 ($p = 0.025$), and a trend of significance for rs3798577 ($p = 0.059$), conferring a higher risk for RAN in this sample (OR = 3.0, 95% CI = 1.1–8.3, and OR = 2.6, 95% CI = 0.9–7.3, respectively).

ESR1 Association with Eating Disorders in an Independent Population-Based Sample

In a population-based sample of 693 young French Caucasian women (mean age 20.30 ± 0.05 years), tagged-SNP rs3798577 was found associated with the presence of eating disorders (additive model: $p = 0.008$) (Table 5).

Table 5 Association of *ESR1* rs3798577 Tag-SNP with AN in a Population-Based Sample of 693 French Caucasian Women

	Genotype						Total	
	CC		CT		TT			
	N	%	N	%	N	%	N	%
SCOFF < 2	128	22.6	284	50.1	155	27.3	567	100
SCOFF ⩾ 2	19	15.1	59	46.8	48	38.1	126	100

p-value (additive model) = 0.008.

DISCUSSION

In this study, we found an association between *ESR1* gene and AN, mainly observed in the restrictive subtype of AN. We showed no association of *ESR2* with AN, nor with RAN

or BPAN subgroups. Furthermore, we did not identified ESR1–ESR2 gene–gene interaction with either AN or restrictive subtype.

Only two works of AN with ESR1 and ESR2 genes were previously published (Rosenkranz et al, 1998; Eastwood et al, 2002). The initial study has screened for mutation in ESR2 gene (Rosenkranz et al, 1998) and found different distribution of the variant G1082A (rs1256049) in AN compared with BN or to obese–underweight subjects. In the second study, no association was reported between AN and three markers, or haplotypes, of ESR1 gene, but a significant association was reported with the rs1256049 SNP of ESR2 (Eastwood et al, 2002). Furthermore, an association was also reported with rs4986938 and rs928554 (but not with rs1256049) in bulimia nervosa and eating disorders not otherwise specified (Nilsson et al, 2004). In this study, no association was found between ESR2 SNPs, which are in linkage disequilibrium (D' > 0.9) with rs1256049 according to databases, and AN, RAN, or BPAN subsets. It is noteworthy that our screening of 321 AN families using the TDT was more powerful $(1-\beta = 89\%)$ than the two previous case–control studies $(1-\beta = 18$ and 63%, respectively) and distinguished RAN and BAN.

We reported a strong association of specific SNPs and haplotypes of the ESR1 gene and AN, and RAN subgroup, with a preferential paternal over-transmission. Although we genotyped a large cohort of families and performed TDTs that reduce the sex and ethnic stratification bias, the probability of a false-positive finding in this study cannot be excluded because of the number of tests performed. After a Bonferroni correction taking into account the 13 SNPs within ESR1 gene for the RAN subgroup, we found only trends of association for two SNPs, rs726281 $p_{corrected} = 0.065$ and rs2295193 $p_{corrected} = 0.091$, but still observed a strong statistical significance for the eight-SNPs haplotype $(p_{corrected} = 0.00008)$. It must be kept in mind that the Bonferroni correction is relevant for independent tests. Given that linkage disequilibrium between the 13 SNPs are mainly above zero in our sample, this correction is probably over-conservative. To assess a more appropriate p-value corrected for multiple testing for single SNP association, we carried out the SNPSpD method (Nyholt, 2004). Significance threshold required is a $p < 0.0051$ (rs726281) and effective number of independent SNP is 10 instead of 13. Thus, we found a significant association for rs726281 $p_{corrected} = 0.049$ and a trend for rs2295193 $p_{corrected} = 0.07$.

The three SNPs of ESR1 associated with RAN in our sample (rs726281, rs3798577, and rs2295193) are located in introns and 3' UTR region. To date, they have no known functional consequences. However, the associated haplotype block of eight SNPs covers the C terminal E/F region of ESR1 protein, which corresponds to the ligand fixation domain of nuclear receptor. So, it is tempting to speculate that the association of these SNPs with RAN could suggest changes of ligand fixation in RAN patients.

We also attempted to replicate our results in an independent German cohort of 41 RAN trios, but failed to confirm significant associations between ESR1 SNPs and RAN. Surprisingly, participants in the German RAN cohort showed a significant higher mean of minimal lifetime BMI compared with the French RAN cohort, while they were significantly younger and had shorter disease duration. As up to half of the RAN patients cross-over to binge eating/purging subtype in a 7 years follow-up study (Eddy et al, 2008), this difference may explain the absence of replication in the German sample, because in the latter, young RAN patients could still switch to the BPAN phenotype. Although this difference between the two cohorts decreases any chances of replication, we were able to observe the excess of paternal transmission of the main SNP associated with AN in the first sample. Owing to a small sample size, the RAN German cohort provided a reduce power (power = 33%) compared with the RAN French cohort (power = 59%) that can explain the low significance of association.

Finally, we tested a tagged-SNP of the ESR1 gene in a large population-based sample, and found an association with eating disorders, as defined by a SCOFF score equal or above 2. However, it should be noted that the SCOFF scale is a screening tool for eating disorders rather that a diagnostic instrument for AN, and therefore no direct comparison should be made between the two first clinical samples and this population-based sample. At most, this epidemiological sample gives indirect evidence for replication, which is usually regarded as needed for genetic association studies in complex disorders.

There are several limitations of this study. First, the instruments assessing the eating disorders are different in the three cohorts. French and German trios families were assessed with semi-structured interview, respectively, DIGS and CIDI, while the self-evaluation SCOFF scale was used to screen epidemiological eating disorders in a large woman population sample. Second, the sample size of the German trios is small with a shorter disease evolution period, although a cohort with a larger sample size and a disease duration equivalent to the French cohort is recommended to replicate our results. Third, the three SNPs of ESR1 (rs726281, rs3798577, and rs2295193), associated with RAN, have no known functional consequences up to now and are located in non-coding regions. These SNPs required a functional analysis regarding their potential effects of associated alleles in ESR1 receptor proprieties.

Although there is a lack of report on the genetic aspects of hormonal status as a risk factor for AN, recent investigations suggest that some hormonal environment might be associated with eating disorders (Procopio and Marriott, 2007; Young, 1991). Observations between same-sex and opposite-sex twins allowed to find a high AN prevalence in adult life in female compares with male, suggesting an influence of intrauterine exposure level to sex hormone (Procopio and Marriott, 2007). Interestingly, significant deficiency of serum levels for estradiol, the ligand of estrogen receptor, has been described in AN patients compared with controls, and a lower concentration in restrictive subtype rather than the binge eating/purging subtype has been reported (Ohwada et al, 2007). It is tempting to speculate that the association between ESR1 gene and RAN may lead to specific biological effects, such as a modified expression and/or an altered function of the estrogen receptor alpha. Furthermore, the paternal transmission of the association that we observed also supports a genetic imprinting, which might be influenced by sex hormones. In addition, it is important to consider DNA methylation in psychiatric disorder because it could help to the understanding of mechanism of the pathology and could

constitute potential therapeutic targets. Acute administration of a selective ER alpha agonist leads to ovariectomized rats that showed a decrease in daily food intake and body weight (Santollo *et al*, 2007). Furthermore, in female mice, agonists of ER receptors produce different effects on social learning of food preferences (Clipperton *et al*, 2008). Then, ER alpha antagonist treatment failed to reduce food intake decrease in chemically ovariectomized rats (Santollo and Eckel, 2009). Finally, in female rat, the level of ESR1 protein expression and estrogen sensitivity are modulated by oxytocin (Perry *et al*, 2009). Thus, it is difficult up to now to suggest a pharmacological treatment for AN. Further functional studies are needed to decipher the biological effect of the association identified in our study, which may influence the number or affinity of ESR1 receptor in RAN patients.

Our study suggests evidences that common variants in *ESR1* gene are associated with eating disorders, and more specifically RAN, highlighting the possible involvement of estrogenic hormonal pathway in AN that might open novel avenues of neuroendocrinopharmacological approaches in this neuropsychiatric disorder.

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DISCLOSURE

The authors declare no conflict of interest.

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