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Evaluation of a methylenetetrahydrofolate-dehydrogenase 1958G>A polymorphism for neural tube defect risk

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Abstract Genetic variants of enzymes involved in the folate pathway might be expected to have an impact on neural tube defect (NTD) risk. Given its key role in folate metabolism, the methylenetetrahydrofolate dehydrogenase 1 (*MTHFD1*) gene could represent an attractive candidate in NTD aetiology. In this study, the impact of the *MTHFD1* 1958G>A polymorphism on NTD risk in the Italian population was examined both by hospital-based case-control and family-based studies. The *MTHFD1* 1958G>A polymorphism was genotyped in 142 NTD cases, 125 mothers, 108 fathers and 523 controls. An increased risk was found for the heterozygous 1958G/A (OR = 1.69; $P=0.04$) and homozygous 1958A/A (OR = 1.91; $P=0.02$) genotypes in the children. Significant association was also found when combined 1958G/A and 1958A/A genotypes of cases were compared with the 1958G/G genotype (OR = 1.76; $P=0.02$). The risk of an NTD-affected pregnancy of the mothers was increased 1.67-fold ($P=0.04$) only when a dominant effect (1958G/A or 1958A/A vs 1958G/G) of the 1958A allele was analysed. The combined TDT/1-TDT ($Z=2.11$; $P=0.03$) and FBAT ($Z=2.4$; $P=0.01$) demonstrated a significant excess of transmission of the 1958A allele to affected individuals. In summary, our results indicate that heterozygosity and homozygosity for the *MTHFD1* 1958G>A polymorphism are genetic determinants of NTD risk in the cases examined.

Keywords Folate metabolism · *MTHFD1* · Neural tube defects · Single nucleotide polymorphism

Introduction

Neural tube defects (NTDs) are one of the most common human congenital malformations in newborn children, affecting annually approximately 1 in 1,000 newborns in Italy (Mastroiacovo and Bianchi 1991) and 200,000 newborns or more worldwide (Moore et al. 1997). Myelomeningocele, commonly called Spina Bifida, is the most common form of NTD. Periconceptional supplementation with folic acid reduces a women's risk of an NTD-affected pregnancy by up to 75% (MRC Vitamin Study Research Group 1991; Czeizel and Dudas 1992). There have been numerous investigations into genes involved in folate metabolism, yielding at least one polymorphism—677C>T in the methylenetetrahydrofolate reductase (*MTHFR*) gene—that has been associated with an approximate doubling of the risk of NTDs (Botto and Yang 2000). However, these associations remain controversial, with several studies finding no association (Papapetrou et al. 1996; Mornet et al. 1997; Speer et al. 1997). Several other genes have been studied, including those encoding folate receptors, cystathionine- β -synthase, methionine synthase, methionine synthase reductase, and reduced folate carrier-1 (Lucock 2000; Melvin et al. 2000; Gelineau-van Waes and Finnell 2001), but many of these studies still await confirmation and warrant further investigation.

Given its central role in folate metabolism, the 5,10-methylenetetrahydrofolate-cyclohydrolase, 10-formyltetrahydrofolate synthetase (methylenetetrahydrofolate dehydrogenase: *MTHFD1*) could represent an attractive candidate gene in NTD aetiology. *MTHFD1* is a tri-functional protein catalysing three sequential reactions in the interconversion of 1-carbon derivatives of tetrahydrofolate (THF) that are substrates for methionine, thymidylate and de novo purine biosynthesis (Tan et al. 1997). Both NAD- and NADP-dependent *MTHFD1* have been identified (Mejia and Mackenzie 1985). *MTHFD1* is an NADP-dependent cytoplasmic enzyme (often referred to as C_1 -THF synthase) encoded by a

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gene on chromosome 14q24 (Rozen et al. 1989); the *MTHFD1* cDNA contains an open reading frame of 2,805 bp that encodes a protein of 935 amino acids (Hum et al. 1988). A mitochondrial trifunctional C₁-THF synthase has been recently identified by Prasanna et al. (2003). The gene spans 236 kb on chromosome 6 and encodes a protein of 978 amino acids, including an N-terminal mitochondrial signal sequence. The mitochondrial isoform is 61% identical to the human cytoplasmic isoform (Prasanna et al. 2003).

Hol et al. (1998) have analysed *MTHFD1* cDNA sequences of familial and sporadic cases of NTDs for the presence of mutations. Two amino acid substitutions were identified. The first (R293H) was detected in a patient with familial Spina Bifida and not in 300 control individuals. The second change, a 1958G>A substitution polymorphism resulting in replacement of the arginine residue at position 653 by glutamine (R653Q), turned out to be present in both patients and controls with similar frequencies. More recently, Brody et al. (2002) demonstrated that *MTHFD1* 1958G>A polymorphism is a maternal risk factor for NTD risk in the Irish population. Mothers who possess two copies of the *MTHFD1* 1958A allele have an increased risk of an NTD-affected pregnancy. There was also a suggestion that the *MTHFD1* 1958A allele decreases fetal viability (Brody et al. 2002).

In the present study, we have attempted to replicate this finding by studying the impact of the *MTHFD1* 1958G>A polymorphism on NTD risk in the Italian population in hospital-based case-control and family-based studies.

Materials and methods

Study population

The NTD study population consisted of children and their parents seeking treatment at the Spina Bifida Centre of the Gaslini Children's Hospital in Genoa, from May 2004 to the present. The Spina Bifida Centre of Gaslini Hospital has recruited children with NTDs from throughout Italy since 1976, and it sees about 150 new cases annually from both the north and the south of Italy. Of the 165 NTD families ascertained for the 1-year period, 142 (86%) agreed to participate and were enrolled. For the purposes of this study, 91 complete NTD triads and 51 incomplete triads, in which DNA was not available from all three family members (51 additional children with NTD, 34 mothers, 17 fathers) were eligible. Information about the type of defect was available for all the children with NTDs, and the cases group included 117 (82%) children (59% females; median age 12 years, range 1–26 years) affected by lumbar and sacral myelomeningocele (L1–S2 levels), 8 (6%) children with lipomyeloschisis, and 17 (12%) children with spina bifida occulta.

Of the 523 control subjects, 208 (40%) were randomly selected children who were admitted to the Gaslini Children's Hospital for miscellaneous illnesses (orthopaedic, otolaryngeal, dental disorders, acute surgical conditions, and trauma). The other 315 control subjects (60%) were healthy young adults who had contributed samples to the blood donor bank of the Gaslini Institute. None had a first-degree relative with an NTD. The samples from all control subjects were anonymous; information associated with these samples (children and adults alike) included sex, region of birth, and age. All participants were Italian with antecedents from all parts of the country. Individuals born in Sardinia were excluded.

The local ethical committee approved the protocol and written consent was obtained from all patients, parents, and control individuals.

Genetic analysis

Genomic DNA was extracted from fresh blood samples using the GenElute™ mammalian genomic DNA miniprep kit (Sigma, Taufkirchen, Germany). The presence of the *MTHFD1* 1958G>A polymorphism was investigated by restriction fragment length polymorphism (RFLP) assay as described by Hol et al. (1998). The digested PCR fragments were analysed by ethidium bromide-stained 4% MetaPhor agarose (FMC Bio-Products, Rockland, ME) gel electrophoresis. Repeat genotyping assays were performed on all samples that gave equivocal genotypes.

Statistical analysis

Allele frequencies in the NTD patients, mothers, fathers, and controls were determined by counting alleles and calculating proportions. Differences in genotype and allele frequency were evaluated by χ^2 analysis. The Hardy–Weinberg equilibrium (HWE) analysis was tested using χ^2 statistics for goodness of fit (1 degree of freedom). Logistic regression analysis was used to estimate the odds ratio (OR) for the presence of mutation *MTHFD1* 1958G>A in patients and controls. The OR was calculated with 95% confidence interval (CI) and the wild-type pattern was used as a reference. A codominant (the three genotypes studied separately), a dominant (1958A/A or 1958G/A genotypes compared with 1958G/G genotype) and a recessive (1958A/A genotype compared with 1958G/G or 1958G/A genotype) model for the effect of the 1958G>A variant was examined. Statistical analysis was performed using SPSS for Windows (SPSS, Chicago, IL).

The attributable fraction, representing the proportion of cases in the target population that is attributable to the presence of the mutation(s), was estimated using the equation proposed by Miettinen (1974): $(AF = [f_{Ca} (OR - 1)] / OR)$, where f_{Ca} is the fraction of cases with the

genotype(s) under study (1958G/A or 1958A/A) and OR is the associated OR. Power computations were performed employing the calculator available at <http://calculators.stat.ucla.edu/powercalc/>.

Three different statistical tests were applied for family-based studies: classical transmission disequilibrium test (TDT) (Spielman et al. 1993), combined TDT/1-TDT (Sun et al. 1999), and the family-based association test (FBAT) (Rabinowitz and Laird 2000). For TDT, we used 91 completed triads. Moreover, owing to 51 incomplete families, we applied the combined TDT/1-TDT test, which combines data from families in which both parental genotypes are available and families in which only one parental genotype is available. We also used the FBAT program (version 1.5.5; <http://www.biostat.harvard.edu/~fbat/fbat.htm>), which allows testing for association even if no parental genotype data are available. Based on the results of the unrelated case-control analysis, a biallelic test was performed using dominant inheritance model.

Results

Case-control study

Table 1 shows the *MTHFD1* genotype distributions, allele frequencies, and the results of HWE analysis. The genotype distribution for the *MTHFD1* 1958G > A mutation in the controls was as expected according to the HWE ($P=0.41$). The frequency of the A allele of the *MTHFD1* 1958G > A mutation was 0.487 among control individuals. No statistically significant heterogeneity was found in the allelic frequencies by type of controls (children or young adults) ($P=0.66$) or by sex ($P=0.54$). Moreover, we found that the genotype distribution obeyed the constraints of HWE also in NTD cases and their parents. Chi squared analysis revealed that allele frequencies were significantly different between cases and controls ($P=0.02$; data not shown).

To evaluate the impact of the *MTHFD1* 1958G > A variant on NTD risk, we calculated the OR and 95% CI (Table 2). We found a significant increased risk for both

Table 2 Odds ratios (OR) with 95% confidence intervals (CI) of the *MTHFD1* 1958G > A genotype in NTD cases, mothers and fathers

<i>MTHFD1</i> 1958G > A	OR (95% CI)	P-value
Patients-controls		
AA vs GG	1.91 (1.1–3.29)	0.02*
GA vs GG	1.69 (1–2.77)	0.04*
GA/AA vs GG	1.76 (1.1–2.82)	0.02*
AA vs. GG/GA	1.32 (0.88–1.99)	0.21
Mothers-controls		
AA vs GG	1.54 (0.86–2.77)	0.15
GA vs GG	1.73 (1.04–2.89)	0.04*
GA/AA vs GG	1.67 (1.02–2.73)	0.04*
AA vs GG/GA	1.05 (0.67–1.64)	0.91
Fathers-controls		
AA vs GG	1.15 (0.64–2.07)	0.64
GA vs GG	1.21 (0.72–2)	0.47
GA/AA vs GG	1.19 (0.73–1.92)	0.48
AA vs GG/GA	1.02 (0.63–1.64)	0.96

*Statistically significant

heterozygous (OR = 1.69; 95% CI: 1–2.77; $P=0.04$) and homozygous (OR = 1.91; 95% CI: 1.1–3.29; $P=0.02$) genotypes in the children. Moreover, when the dominant model was applied (1958G/A or 1958A/A vs 1958G/G), there was a nominally significant increase in NTD risk (OR = 1.76; 95% CI: 1.1–2.82; $P=0.02$). A recessive effect of the 1958A variant (1958A/A vs 1958G/G or 1958G/A) in cases was not significant (OR = 1.32; $P=0.21$). In cases homozygous for the 1958A allele, the power to detect an OR of 1.91 was 85% ($\alpha < 0.05$, frequency of the risk allele = 0.487, two-sided); thus, the likelihood of a type II error in our data set is low. We calculated that the proportion of NTD cases that could be attributed under a dominant effect model of the 1958A variant was about 35.6% (population attributable fraction).

A significantly elevated risk of NTD was observed in mothers who were heterozygous for the *MTHFD1* G1958 > A variant (OR = 1.73; 95% CI: 1.04–2.89; $P=0.04$). A maternal 1958A/A homozygous genotype was not a significant independent contributor to NTD risk (OR = 1.54; 95% CI: 0.86–2.77; $P=0.15$). Nevertheless, when a dominant effect (1958G/A or 1958A/A

Table 1 Genotype distribution and allele frequencies of the methylenetetrahydrofolate dehydrogenase 1 (*MTHFD1*) 1958G > A polymorphism in Italian neural tube defect (NTD) families

	Controls % ($n=523$)	NTD cases % ($n=142$)	Mothers % ($n=125$)	Fathers % ($n=108$)
Genotype distribution				
GG	27.3 ($n=143$)	17.6 ($n=25$)	18.4 ($n=23$)	24.1 ($n=26$)
GA	48 ($n=251$)	52.1 ($n=74$)	56 ($n=70$)	50.9 ($n=55$)
AA	24.7 ($n=129$)	30.3 ($n=43$)	25.6 ($n=32$)	25 ($n=27$)
Allele frequency				
G	0.513	0.437	0.464	0.495
A	0.487	0.563	0.536	0.505
Test for HWE ^a				
Chi-square	0.81	0.50	1.98	0.037
P-value	0.41	0.51	0.20	0.65

^aHardy-Weinberg equilibrium

vs 1958G/G) of the 1958A allele of the mothers was considered, we found a 1.67-fold increased risk ($P=0.04$). On the contrary, no significant association was seen when the recessive model of the 1958A variant of the mothers was applied (1958A/A vs 1958G/G or 1958G/A; OR = 1.05; 95% CI:0.67–1.64; $P = 0.91$).

No increased risk was found for the genotype of the fathers, regardless of the genetic model applied (codominant, dominant or recessive).

Family-based study

The results of family-based studies are presented in Table 3. The 142 NTD families included 91 for whom maternal and paternal DNA samples were available and 51 for whom only maternal or only paternal samples were obtained (incomplete families). We investigated whether there was any bias in the transmission of the A allele using a classical TDT test (Spielman et al. 1993), which tests allele transmission from heterozygous parents regardless of case genotype outcome. TDT indicated that, from 96 heterozygous parents, the A allele was transmitted 55 times (57%), whereas the G allele was transmitted 41 times (43%). This bias in A allele transmission is only suggestive of an effect ($P = 0.06$). Given the high number of incomplete families, we applied the combined TDT/1-TDT test, which combines data from families in which both parental genotypes are available and families in which only one parental genotype is available (Sun et al. 1999). In total, 51 families in which only one parent was available, and 91 triads were used for analysis by the combined TDT/1-TDT test. A significant excess of transmission of the A allele to affected individuals was found (Z score = 2.11; $P = 0.03$). More significant association was found by FBAT analysis ($Z = 2.4$; $P = 0.01$, for a dominant model).

Discussion

The aim of this study was to evaluate the role of a common polymorphism of the *MTHFD1* gene, the 1958G > A variant, on the risk of NTD in the Italian

population. The degree of risk was assessed both by hospital-based case-control and family-based studies.

The frequency of the 1958A allele in our controls was 0.487, which is comparable to the frequencies of 0.450 found in the control populations in Ireland and Germany (Hol et al. 1998; Brody et al. 2002). The frequency reported here is higher than that (0.39) found in a small sample of Italian DNAs from a CEPH diversity panel (Shi et al. 2003). This discrepancy may be due to the small size of the individual groups included in that panel compared to the current study.

The case-control study has indicated that cases carrying one (1958G/A) or two (1958A/A) alleles for the *MTHFD1* 1958G > A polymorphism have a significantly increased risk of NTD in comparison with individuals who have no copies of this polymorphism (1958G/G homozygotes). The magnitude of this association, i.e. the OR, was similar both when the genotypes were analysed separately (codominant model; 1958G/A = 1.69; 1958A/A = 1.91) and under a dominant model of action (1958G/A or 1958A/A = 1.76). However, the significance of the association studies is not impressive enough to allow definitive conclusions regarding whether this polymorphism is a risk factor for NTD. We found no association between any of the fathers' genotypes and NTDs. When we analysed maternal genotypes, a significant dominant maternal effect was observed (1.67-fold increased risk in mothers carrying 1958G/A or 1958A/A genotypes vs 1958G/G); no significant association was found when homozygosity of the mothers was considered separately (1958A/A vs 1958GG), or when a recessive model within mothers (1958A/A vs 1958G/A or 1958G/G) was analysed.

Furthermore, the results of the present case-control study are not in agreement with the report of Brody et al. (2002). In fact, in the Irish study the authors conclude that maternal *MTHFD1* 1958A/A homozygosity is a risk factor for NTDs, but they did not observe any difference in allele or genotype frequencies between the case and control groups.

Our finding of an association between the *MTHFD1* 1958G > A genotype of children and NTDs appears to be reasonably powered. Thus, a possible reason for the

Table 3 Results of family-based tests of the association between NTDs and the A allele of the G1958A polymorphism of the *MTHFD1* gene. TDT Transmission disequilibrium test, TDT/1-TDT combined TDT and TDT when only one parent is available (1-TDT), FBAT family-based association test

TDT				Combined TDT/1-TDT				FBAT			
N fam ^a	N inf fam ^b	T:NT ^c	P-value ^d	N fam ^a	N inf fam ^b	Z ^e	P-value ^d	N fam ^a	N inf fam ^b	Z ^e	P-value ^d
91	74	55:41	0.06	142	97	2.11	0.03	142	90	2.44	0.01

^aNumber of families available for the test. For TDT, each unit is a trio (an affected child and both parents), for TDT/1-TDT and FBAT each unit is a trio or an incomplete (only one parent available) family

^bThe number of informative families is the subset of families that are genotypically informative for the allele being tested

^cT:NT is the ratio of transmission to non-transmission of the test allele from heterozygous parents to affected children

^dAll P-values are uncorrected for multiple testing

^eTest statistic for association, distributed as a standard normal

discrepancy between our results and the Irish study may be due to population heterogeneity and, therefore, our results apply only to the Italian population. It is also possible that environmental differences can explain the discrepancies between the Irish and Italian case frequencies.

Three different statistical tests were applied to the family-based studies, only two of which, TDT/1-TDT and FBAT, yielded significant evidence for association between the *MTHFD1* G1958A polymorphism and NTDs. Despite their significance, these results should be interpreted with caution since the family-based studies were less powerful than the case-controls studies and need to be confirmed in further studies.

Although association studies demonstrated that the maternal *MTHFD1* 1958A/A genotype has no appreciable influence on NTD risk as an independent risk factor, it could be important when other factors are present, including the 1958A/A of the child. On the basis of the data presented here (association studies), the question of the extent, if any, to which maternal *MTHFD1* 1958A/A genotype contributes independently to NTD outcome, should be addressed by studies including much larger numbers of mother-child pairs, or by application of TDT to three-generation families that include heterozygous maternal grandparents, as suggested by Mitchell (1997).

The rationale for a possible association between *MTHFD1* genotype and NTD risk is based on the important role of the *MTHFD1* protein in folate pool maintenance. The *MTHFD1* 1958G>A substitution affects the synthetase domain of the enzyme, possibly altering nucleotide pools available for DNA synthesis and, thus, affect cell division. No specific biochemical phenotype has been associated with this genetic variant. The *MTHFD1* 1958G>A polymorphism is unlikely to have a direct effect on folate and homocysteine levels (Brody et al. 2002). Nevertheless, this polymorphism may produce an effect at the cellular level without causing perturbations in plasma metabolites.

In summary, our results indicate that the *MTHFD1* 1958G/A and 1958A/A genotypes in NTD cases are genetic determinants of NTD risk with dominant action. Further investigations are warranted to clarify if maternal genotype may confer risk, independently, or in the presence of a mutant *MTHFD1* genotype in the children. Clearly, our understanding of the role of the *MTHFD1* gene in NTD risk is incomplete, and studies of other variants in this gene that were not examined here are required.

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