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# Clinical nodal staging scores for prostate cancer: a proposal for preoperative risk assessment

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**Background:** Pelvic lymph node dissection in patients undergoing radical prostatectomy for clinically localised prostate cancer is not without morbidity and its therapeutical benefit is still a matter of debate. The objective of this study was to develop a model that allows preoperative determination of the minimum number of lymph nodes needed to be removed at radical prostatectomy to ensure true nodal status.

**Methods:** We analysed data from 4770 patients treated with radical prostatectomy and pelvic lymph node dissection between 2000 and 2011 from eight academic centres. For external validation of our model, we used data from a cohort of 3595 patients who underwent an anatomically defined extended pelvic lymph node dissection. We estimated the sensitivity of pathological nodal staging using a beta-binomial model and developed a novel clinical (preoperative) nodal staging score (cNSS), which represents the probability that a patient has lymph node metastasis as a function of the number of examined nodes.

**Results:** In the development and validation cohorts, the probability of missing a positive lymph node decreases with increase in the number of nodes examined. A 90% cNSS can be achieved in the development and validation cohorts by examining 1–6 nodes in cT1 and 6–8 nodes in cT2 tumours. With 11 nodes examined, patients in the development and validation cohorts achieved a cNSS of 90% and 80% with cT3 tumours, respectively.

**Conclusions:** Pelvic lymph node dissection is the only reliable technique to ensure accurate nodal staging in patients treated with radical prostatectomy for clinically localised prostate cancer. The minimum number of examined lymph nodes needed for accurate nodal staging may be predictable, being strongly dependent on prostate cancer characteristics at diagnosis.

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Lymph node (LN) metastasis is the strongest prognostic factor for patients treated with radical prostatectomy (RP) for clinically localised prostate cancer (PCa) (Briganti *et al*, 2009; Heidenreich *et al*, 2011). Although extended pelvic LN dissection (ePLND) is the most accurate LN staging in patients undergoing RP for localised PCa (Abdollah *et al*, 2010; Heidenreich *et al*, 2011), PLND is not without morbidity and its therapeutical role is still a matter of debate (DiMarco *et al*, 2005; Briganti *et al*, 2006; Joslyn and Konety, 2006).

Previous tools have been developed to estimate the risk of LN metastasis and thereby guide the need and extent of PLND in PCa (Godoy *et al*, 2011; Briganti *et al*, 2012). It has been suggested that at least 10 LNs need to be removed and examined to achieve a 'standard' PLND (Touijer *et al*, 2007; Godoy *et al*, 2011), while at least 20 LNs removed and examined provides an ePLND (Weingartner *et al*, 1996). However, whether such numbers may vary according to individual patient characteristics is still unknown. This is key for planning the optimal treatment approach, given the strict correlation between PCa features and the risk of LN invasion (Heidenreich *et al*, 2002; Briganti *et al*, 2007; Abdollah *et al*, 2010).

We have recently developed a methodology that allows preoperative determination of the minimum number of LNs needed to take to ensure that a bladder cancer patient is correctly staged as LN negative (Shariat *et al*, 2012). The aim of this study was to develop a similar NSS for patients with PCa. We hypothesised that the true nodal status (no false-negative LN status) could be accurately predicted based on the number of LNs examined according to the clinical PCa features. Therefore, we used a large multicentre (i.e., heterogeneous) cohort of patients treated with RP and a variable extent of PLND to develop the novel NSS. Subsequently, we validated the novel model in a large single-centre cohort of RP patients who underwent an anatomically defined ePLND.

## PATIENTS AND METHODS

**Patients.** The development cohort (DC) included data from 4770 patients treated with RP and PLND between 2000 and 2011 for PCa at eight high-volume centres. In this cohort, the extent of PLND was at the discretion of each treating physician, thus mainly consisting of an anatomically defined limited PLND, including removal of all lymphatic tissue in the obturator fossa and along the external iliac vessels. However, ePLND was performed too. The validation cohort (VC) included 3595 patients treated between 1989 and 2012 at a single centre with RP and anatomically defined ePLND for PCa. Here, the ePLND consisted of excision of fibrofatty tissue along the external iliac vein, with the distal limit being the deep circumflex vein and the femoral canal. Lymph nodes along the internal iliac artery were also removed. Proximally, ePLND included the bifurcation of the common iliac artery (Briganti *et al* 2012). Preoperative staging was performed with prostate-specific antigen (PSA), Gleason score at biopsy, digital rectal examination, and imaging study results. No patient received preoperative radiotherapy, hormonal treatment, or chemotherapy. No patient had clinically evident distant metastases at the time of RP. This was an institutional review board-approved study, with all participating sites providing the necessary institutional data-sharing agreements prior to initiation of the study.

**Pathological evaluation.** Genitourinary pathologists assigned the pathological stage, which was reassigned according to the 2007 American Joint Committee on Cancer (AJCC) Tumor, Node, Metastasis (TNM) staging system when necessary. All lymphoid tissue removed was submitted for histological examination.

**Statistical analysis.** As in a previously described methodology for colorectal, bladder, and PCa (Gonen *et al*, 2009; Shariat *et al*, 2012; Kluth *et al*, 2013), we tested the probability of incorrect nodal

staging as a function of the number of examined LNs. The true LN status is practically unascertainable as there is no way to ascertain the false-negative rate. However, information from LN-positive patients can be used to determine whether the number of LNs examined and the number of LNs that are negative are sufficient to classify a patient as LN negative. For example, one can consider a patient with a large number of examined LNs and small number of positive LNs (called  $k$ ): if fewer LNs were examined, there would be a chance that this patient would be incorrectly deemed LN negative. In contrast, for a patient with a small number of examined LNs and large  $k$ , it is not likely that LN disease would have been missed, even though fewer LNs were examined. Thus, the data from LN-positive patients were used to interpret the data for the LN-negative patients. The probability that a LN-negative patient has LN disease can be computed using the following algorithm: computing the probability of missing a positive LN, computing the prevalence of pN1 status, and computing the NSS from sensitivity and prevalence.

**Probability of missing a positive node.** The probability of missing a positive LN (one minus the sensitivity) is inherent to the process of pathological detection and, as such, depends on the number of examined LNs – however, not on patient characteristics. For this purpose, we used a beta-binomial model, which allowed for heterogeneity in the intensity of LN spread across patients.

Three key assumptions underlie this step: (1) there are no false positives (if the specimen contains a negative LN, it will be correctly identified by the pathologist); (2) all LNs are exchangeable; that is, they all have an equal probability of being involved; and (3) sensitivity is the same for LN-positive and LN-negative patients.

These assumptions may not be completely tenable, but we consider them to be sufficient approximations to our biological understanding of LN spread and clinical practice of nodal staging.

**Estimation of prevalence of nodal disease.** The observed prevalence is an underestimate and needed to be adjusted for false negatives, which was done in two steps. The first step comprises assumption 1 and estimates  $\#FN_k$  as a function of  $k$ , which is the number of positive LNs from patients with LN involvement:

$$\#FN_k = \frac{[1 - P(FN_k)] \times \#TP_k}{P(FN_k)}$$

In this equation,  $\#TP_k$  is the number of true positives for a given  $k$ . Since prevalence is not a function of  $k$ , the second step receives the adjusted prevalence by averaging over  $k$ :

$$\text{Prev} = \frac{\sum_k (FN_k + TP_k)}{\sum_k (FN_k + TP_k + TN_k)}$$

Estimation of prevalence was stratified by clinical T stage, preoperative PSA, and Gleason score sum at biopsy for the cNSSs (preoperative). Furthermore, for a better clinical utilization of the cNSSs, prevalence was also estimated by a previously published clinical risk stratification (Briganti *et al*, 2006).

**Nodal staging score.** Adequate staging was assessed by computing the cNSS, the probability that a pathologically LN-negative patient is truly free from nodal disease:

$$\text{NSS} = \frac{1 - \text{Prev}}{1 - \text{Prev} + [\text{Prev} \times P(FN_k)]}$$

**Confidence intervals.** Precision of the reported estimates was assessed by creating 2000 bootstrap samples from the entire data set and replicating the estimation process. The 2.5th and 97.5th quartiles were used as the lower and upper 95% confidence intervals for the corresponding estimates, respectively.

**RESULTS**

Table 1 shows the clinical and pathological features of the 4770 patients in the DC and the 3595 patients in the VC. Among the 4770 patients in the DC, the median number of examined LNs was 6 (interquartile range (IQR): 7); 94.6% of the patients were deemed LN negative. Lymph node metastases were present in 16 of 3234 (0.5%) pT2 patients, 85 of 1105 (7.7%) pT3a patients, and 156 of 444 (35.1%) pT3b/T4 patients. Among the 3595 patients in the VC, the median number of examined LNs was 16 (IQR: 11); 87.1% were free from LN metastasis. Lymph node metastases were present in 34 of 2381 (1.4%) pT2 patients, 86 of 587 (14.7%) pT3a patients, and 344 of 627 (54.9%) pT3b/T4 patients.

In the DC, the resulting probability of missing nodal disease (one minus the sensitivity) as a function of the number of examined LNs is plotted in Figure 1A. As expected, the probability of missing LN disease decreased as the number of LNs examined increased: if only a single LN was examined in all patients, 84% of LN disease would be missed. The probability of missing a positive LN is 60% if 3 LNs are examined, 45% if 5 are examined, and drops to 22% when 10 are examined. The sensitivity of nodal staging exceeds 80% only when at least 11 LNs are examined. In the VC, as expected, the probability of missing LN metastasis decreases as the number of LNs examined increases (Figure 1B): if only a single LN is examined in all patients, 84% of LN disease would be missed. The probability of missing a positive LN decreases with an increasing number of LNs examined and is 62% if 3 LNs are examined, 47% if 5 are examined, and 30% if 10 are examined.

Table 1. Characteristics of 8365 patients who underwent radical prostatectomy and bilateral pelvic lymph node dissection for localised prostate cancer

	Development cohort (n = 4770)						Validation cohort (n = 3595)																																																																			
	Patients		Lymph nodes examined				Patients		Lymph nodes examined																																																																	
	No.	%	Mean	s.d.	Median	IQR	No.	%	Mean	s.d.	Median	IQR																																																														
<b>Clinical T stage</b>																																																																										
cT1	3298	69.1	6.9	6.0	5.0	6	1883	52.4	17.2	8.1	16.0	11																																																														
cT2	1367	28.7	8.2	6.5	6.0	8	1231	34.2	17.5	8.4	16.0	11																																																														
cT3	105	2.2	9.2	6.2	7.0	7	481	13.4	19.5	10.0	18.0	12																																																														
<b>Bx Gleason score</b>																																																																										
6	2121	44.5	5.94	5.0	5.0	5	2230	62.0	17.1	7.9	16.0	11																																																														
7	2063	43.2	8.0	6.6	6.0	8	1020	28.4	17.9	8.6	17.0	11																																																														
8–10	586	12.3	9.7	7.6	9.0	8	345	9.6	20.1	11.1	18.0	13																																																														
<b>Pathological T stage</b>																																																																										
pT2	3234	67.6	6.8	5.7	5.0	6	2381	66.2	17.9	8.0	16.0	11																																																														
pT3a	1103	23.1	8.2	6.6	6.0	7	587	16.3	18.2	8.6	17.0	11																																																														
pT3b/T4	443	9.3	9.7	7.4	8.0	9	627	17.4	19.1	9.9	17.0	10																																																														
<b>RP Gleason score</b>																																																																										
6	1135	23.8	5.6	4.4	5.0	4	1387	38.6	17.0	7.9	16.0	11																																																														
7	3170	66.4	7.7	6.4	6.0	7	1685	46.9	17.4	8.3	16.0	11																																																														
8–10	467	9.8	9.1	7.3	7.0	9	523	14.5	20.0	10.1	18.0	11																																																														
<b>Surgical margin</b>																																																																										
Negative	4029	84.5	7.2	6.1	5.0	6	2588	72.0	17.2	8.1	16.0	11																																																														
Positive	741	15.5	8.3	6.5	6.0	8	1007	28.0	18.7	9.3	17.0	11																																																														
<b>Pathological N stage</b>																																																																										
N0	4514	94.6	7.2	6.0	6.0	6	3131	87.1	17.1	7.9	16.0	11																																																														
N+	256	5.4	10.4	8.1	10.0	6	464	12.9	21.3	11.0	19.0	11																																																														
<b>Preoperative PSA (ng ml<sup>-1</sup>)</b>																																																																										
< 10	3830	80.3	7.1	6.0	5.0	6	2492	69.3	17.4	8.3	16.0	11																																																														
≥ 10	940	19.7	8.6	6.7	6.0	7	1103	30.7	18.1	9.0	17.0	10																																																														
<table border="1" style="width:100%; border-collapse: collapse;"> <thead> <tr> <th rowspan="2"></th> <th colspan="4">Development cohort (n = 4770)</th> <th colspan="4">Validation cohort (n = 3595)</th> </tr> <tr> <th>Median</th> <th>Mean</th> <th>Range</th> <th>IQR</th> <th>Median</th> <th>Mean</th> <th>Range</th> <th>IQR</th> </tr> </thead> <tbody> <tr> <td>Age (years)</td> <td>62.0</td> <td>61.3</td> <td>45–82</td> <td>10</td> <td>66.0</td> <td>65.5</td> <td>40–89</td> <td>10</td> </tr> <tr> <td>No. of examined LNs in all patients</td> <td>6.0</td> <td>7.4</td> <td>1–55</td> <td>7</td> <td>16.0</td> <td>17.6</td> <td>5–73</td> <td>11</td> </tr> <tr> <td>No. of examined LNs in patients with negative LNs</td> <td>6.0</td> <td>7.2</td> <td>1–55</td> <td>6</td> <td>16.0</td> <td>17.8</td> <td>5–66</td> <td>11</td> </tr> <tr> <td>No. of examined LNs in patients with positive LNs</td> <td>10.0</td> <td>10.4</td> <td>1–38</td> <td>10</td> <td>19.0</td> <td>21.3</td> <td>5–73</td> <td>11</td> </tr> <tr> <td>No. of positive LNs in patients with positive LNs</td> <td>1.0</td> <td>1.8</td> <td>1–23</td> <td>1</td> <td>2.0</td> <td>3.81</td> <td>1.6</td> <td>2</td> </tr> </tbody> </table>														Development cohort (n = 4770)				Validation cohort (n = 3595)				Median	Mean	Range	IQR	Median	Mean	Range	IQR	Age (years)	62.0	61.3	45–82	10	66.0	65.5	40–89	10	No. of examined LNs in all patients	6.0	7.4	1–55	7	16.0	17.6	5–73	11	No. of examined LNs in patients with negative LNs	6.0	7.2	1–55	6	16.0	17.8	5–66	11	No. of examined LNs in patients with positive LNs	10.0	10.4	1–38	10	19.0	21.3	5–73	11	No. of positive LNs in patients with positive LNs	1.0	1.8	1–23	1	2.0	3.81	1.6	2
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Abbreviations: IQR = interquartile range; LNs = lymph nodes; PSA = prostate-specific antigen; RP = radical prostatectomy; s.d. = standard deviation.																																																																										

Sensitivity of nodal staging exceeds 80% only when at least 15 LNs are examined.

The results for cNSS in the DC are shown in Figure 2 and Table 2. Although patients with a cT1 tumour have a 90% probability to be free of LN metastasis with one single LN examined, patients with cT2 and cT3 tumours will need 6 and 11 LNs examined to maintain similar levels of cNSS, respectively (Figure 2A). Although patients in the low- and intermediate-risk group achieve a cNSS of 90% with any examined LN, patients in the high-risk group will require eight examined LNs to achieve a similar probability (Figure 2D).

Figure 3 and Table 3 show the results for cNSS in the VC. For cT1 tumours, six examined LNs provide 90% confidence that the patient was indeed LN negative. For the same level of confidence, one would need 8 for cT2 tumours. Patients with a cT3 tumour will reach a cNSS of only 80% with 11 LNs examined, and 90% with 20 LNs examined (Figure 3A). Similar to the DC, in patients in the low- and intermediate-risk group, any LNs examined will result in a cNSS of >90%, whereas patients in the high-risk group will need 10 LNs examined to achieve a probability of 80% to be truly LN negative (Figure 3D).

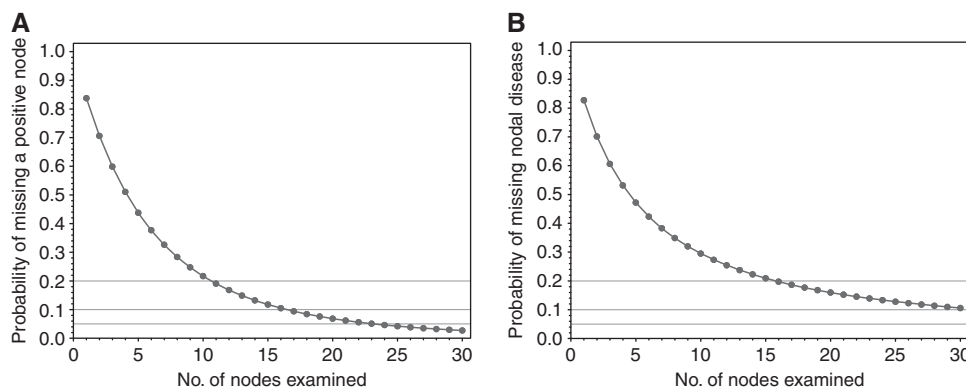


Figure 1. (A) Probability of missing nodal disease as a function of nodes examined in the development cohort of 4770 patients who were treated with radical prostatectomy and bilateral pelvic lymphadenectomy for localised prostate cancer. (B) Probability of missing nodal disease as a function of nodes examined in the validation cohort of 3595 patients who were treated with radical prostatectomy and bilateral pelvic lymphadenectomy for localised prostate cancer.

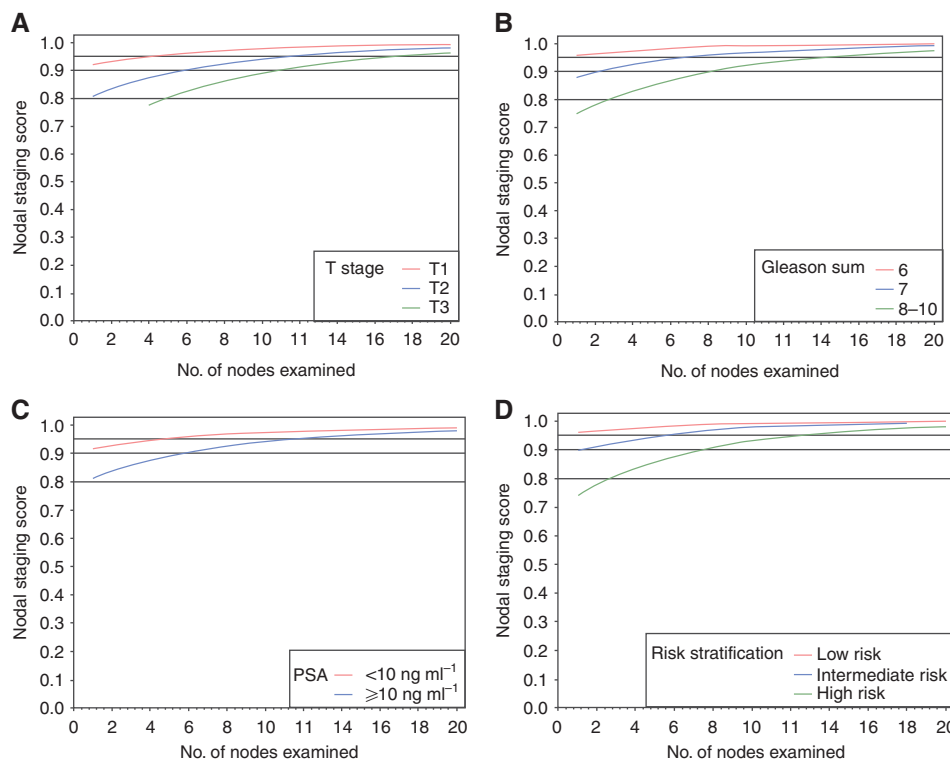


Figure 2. Clinical nodal staging scores in the development cohort of 4770 patients who were treated with radical prostatectomy and bilateral pelvic lymphadenectomy for localised prostate cancer. The vertical axis is the probability of missing nodal disease (one minus sensitivity); the horizontal axis is the number of examined nodes. (A) Sensitivity of the pathological evaluation of nodal disease stratified by clinical tumour stage. (B) Sensitivity of the pathological evaluation of nodal disease stratified by biopsy Gleason score. (C) Sensitivity of the pathological evaluation of nodal disease stratified by preoperative prostate-specific antigen. (D) Sensitivity of the pathological evaluation of nodal disease stratified by preoperative risk grouping.

**Table 2.** Clinical nodal staging scores for selected values of the number of LNs examined in the development cohort of 4770 patients who underwent radical prostatectomy and bilateral pelvic lymph node dissection for localised prostate cancer

No. of examined LNs	1	3	5	8	10	12	15	20
<b>Clinical T stage</b>								
T1	92.4	94.4	95.9	97.3	97.9	98.4	98.9	99.3
T2	80.7	85.4	88.9	92.5	94.2	96.4	96.8	98.1
T3	—	—	80.3	86.3	89.2	91.4	93.8	96.3
<b>Bx Gleason score</b>								
6	95.5	96.8	97.6	98.5	98.8	99.0	99.4	99.6
7	88.0	91.2	93.4	95.6	96.6	97.4	98.1	98.9
8–10	74.8	80.6	85.0	89.8	92.0	93.7	95.5	97.3
<b>Preoperative PSA (ng ml<sup>-1</sup>)</b>								
<10	91.9	94.1	95.6	97.1	97.8	98.3	98.8	99.3
≥10	81.5	86.1	89.4	92.9	94.5	95.7	96.9	98.2
<b>Risk stratification</b>								
Low risk	96.1	97.3	98.1	98.8	99.1	99.4	99.6	99.8
Intermediate risk	89.6	92.6	94.7	96.7	97.6	98.2	98.8	99.2
High risk	74.3	80.8	85.7	90.8	93.1	94.7	96.5	98.1
Abbreviations: LNs = lymph nodes; PSA = prostate-specific antigen.								

**Table 3.** Clinical nodal staging scores for selected values of the number of LNs examined in the validation cohort of 3595 patients who underwent radical prostatectomy and bilateral pelvic lymph node dissection for localised prostate cancer

No. of examined LNs	1	3	5	8	10	12	15	20
<b>Clinical T stage</b>								
T1	—	—	94.2	95.1	95.8	96.9	97.3	97.7
T2	—	—	—	90.0	91.9	92.6	94.3	95.2
T3	—	—	—	75.1	78.8	80.9	84.9	88.3
<b>Bx Gleason score</b>								
6	—	—	95.0	95.8	97.2	97.5	97.7	98.2
7	—	—	84.0	87.5	89.1	90.2	92.5	94.6
8–10	—	—	—	67.1	70.0	71.9	76.6	81.9
<b>Preoperative PSA (ng ml<sup>-1</sup>)</b>								
<10	—	—	93.6	95.2	95.9	96.4	97.0	97.7
≥10	—	—	78.0	82.7	85.0	86.8	88.9	91.3
<b>Risk stratification</b>								
Low risk	—	—	—	95.6	96.2	96.7	97.1	97.9
Intermediate risk	—	—	—	92.3	94.2	95.0	95.7	96.4
High risk	—	—	—	70.3	76.2	79.1	81.5	87.5
Abbreviations: LNs = lymph nodes; PSA = prostate-specific antigen.								

**DISCUSSION**

In patients undergoing RP for PCa, PLND is an essential procedure for accurate LN staging (Abdollah *et al*, 2010; Heidenreich *et al*, 2011). Lymph node metastasis is a powerful

prognostic factor in these patients (Briganti *et al*, 2009; Heidenreich *et al*, 2011) guiding clinical decision-making regarding patient counselling, follow-up scheduling and adjuvant therapies (Steuber *et al*, 2006; Boorjian *et al*, 2007; Abdollah *et al*, 2012). Many efforts have been made to estimate the number of LNs needed to be removed and examined in order to determine the true LN status (Heidenreich *et al*, 2002; Bader *et al*, 2003; Briganti *et al*, 2007; Abdollah *et al*, 2012). Anywhere between 10 and 28 LNs removed and examined has been recommended for accurate LN staging (Briganti *et al*, 2007; Touijer *et al*, 2007). The discrepancy between studies in the number of LNs needed to be removed may be due to the extent of dissection, pathological factors, variability in patients’ pelvic anatomy, and variability of disease severity (Bochner *et al*, 2001; Mattei *et al*, 2008; Joniau *et al*, 2013). Moreover, all these studies share one significant limitation: the lack of adjustment for the effect of clinical factors. This is crucial, given the strict association between cancer aggressiveness and rate of LN metastasis (Carlsson *et al*, 2013). Therefore, in this study we aimed to assess the association between the number of examined LNs and the true LN status according to preoperative patient characteristics.

We found that the number of LNs needed to be removed varies largely among patients according to clinical PCa features. The proportion of having a positive LN increased indeed proportionally with advancing clinical T stage, higher biopsy Gleason score, and higher PSA. Moreover, we found that all preoperative characteristics are highly correlated with the number of LNs needed to be examined to ascertain the true LN status. One might argue that such associations, where more aggressive cancers need more extensive PLND to detect a single nodal metastasis, seem to be counter-intuitive. In fact, while only few LNs are needed to be removed, ascertaining that a low-risk patient is LN negative is mainly driven by the extremely low rate of LNI in this patient group. Indeed, despite extension of the PLND to all anatomical nodal pelvic areas, the rate of LNI in this patient group remains <2% (Weingartner *et al*, 1996). Conversely, the notion that more LNs are required to achieve accurate LN-negative staging in men at higher risk of LNI is mainly due to the highly heterogenous and complex anatomical pathway of LN metastasis in PCa (Shariat *et al*, 2003a,b). Nodal dissemination in PCa does not indeed follow a pre-defined pathway of metastatic spread, but rather different lymphatic routes tributary to several primary lymphatic landing sites (Mattei *et al*, 2008). However, one has to realise that even with extensive PLND a 90% probability to be free from nodal metastasis will not be achieved. Therefore, even a well-performed, meticulous ePLND does not ensure 100% accuracy with regard to nodal status. Moreover, the risk of nodal invasion is not driven by a single cancer characteristic, but is highly dependent on the interaction of several clinical variables.

Unfortunately, the extent of PLND is too often based on the surgeon’s intuitive experience integrating his/her beliefs, patient characteristics such as performance status, comorbidities and tumour characteristics. Our model proves the concept that the needed extent of PLND is dependent on and can be predicted by the risk of LNI based on individual patient features, similar to a model we previously described for PCa (Kluth *et al*, 2013). Indeed, the one-size-fits-all approach is too inaccurate and one needs to estimate the individual risk of missing a positive LN. Performing an ePLND in all patients would certainly result in overtreatment of many patients, with resulting side effects and cost. Conversely, removing too few LNs may result in inaccurate staging and possibly inferior survival (Bader *et al*, 2003; Joslyn and Konety, 2006). Our tool may enable the physician and the patient to join in the decision-making by identifying risks through weighting the probability of missing a positive LN taking into account the risk of PLND.



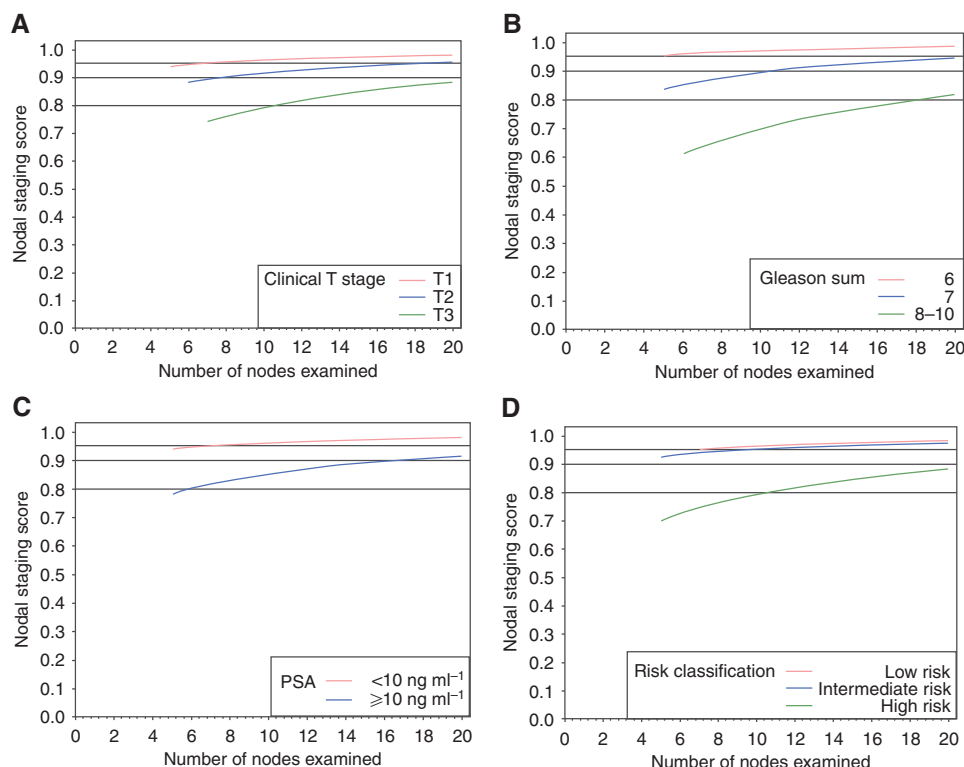


Figure 3. Clinical nodal staging scores in the validation cohort of 3595 patients who were treated with radical prostatectomy and bilateral pelvic lymphadenectomy for localised prostate cancer. The vertical axis is the probability of missing nodal disease (one minus sensitivity); the horizontal axis is the number of examined nodes. (A) Sensitivity of the pathological evaluation of nodal disease stratified by clinical tumour stage. (B) Sensitivity of the pathological evaluation of nodal disease stratified by biopsy Gleason score. (C) Sensitivity of the pathological evaluation of nodal disease stratified by preoperative prostate-specific antigen. (D) Sensitivity of the pathological evaluation of nodal disease stratified by preoperative risk grouping.

Although we were able to control for numerous potential confounders, we could not control for the surgeon's and pathologist's experience or treatment decisions. In the DC, we could not adjust for the anatomic template of the preferred PLND and the location of each removed LN, which has been shown to be important (Weingartner *et al*, 1996). Removing LNs from an area of high likelihood of malignancy may be more valuable than removing a LN that is less likely to be involved with cancer (Joniau and Van den Bergh, 2013). However, this limitation was not applicable to the VC, where all patients underwent an anatomically defined ePLND regardless of the risk of LNI. Nevertheless, one might argue that the number of LNs examined is not an exact surrogate for the extent or template of PLND. In addition, the number of LNs examined is also dependent on pathological evaluation and inherent differences among patients (Bochner *et al*, 2004). Specific characteristics of the DC and VC may increase the generalisability of this tool as data were derived from two different patient cohorts: a DC, in which the extent of PLND was based on the surgeon's discretion, and a VC, which was based on a large single-centre experience with a protocol for ePLND. Thus, our study examined the differential impact of the extent of PLND; it is noteworthy that the probability of freedom from missed LN metastasis was similar in both cohorts, thus supporting its generalisability and thereby proving its robustness. Finally, our model is based on assumptions. Although these assumptions might seem debatable, every single mathematical model and every theory is built on assumptions tested in an experimental setting, such as that used in this study. Prospective validation of our model can test whether the assumptions were realistic and robust (Kattan, 2011).

## CONCLUSIONS

For patients treated with RP for PCa, PLND is the only reliable technique to ensure accurate nodal staging. The minimum number of examined LN staging may be predictable, being strongly dependent on preoperative PCa characteristics. We developed a model to allow an informed decision-making based on these characteristics.

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